# Appendix

**Outline** We provide detailed proofs for all of our theories in Secs. A to F. Sec. G provides multiple additional experiments demonstrating that pseudo-labeling improves transfer learning and that combining pseudo-labeling with adversarial training in the source further improves transferability. Sec. H provides additional details about our experiments.

Recall that in the main context, in Algorithm 1, we have  $\hat{W}_1 \leftarrow \text{top-}r \text{ SVD of } [\hat{\beta}_1, \hat{\beta}_2, \cdots, \hat{\beta}_T]$ . Specifically, we assign the columns of  $\hat{W}_1$  as the collection of the top-r left singular vectors of  $[\hat{\beta}_1, \hat{\beta}_2, \cdots, \hat{\beta}_T]$ .

The rest of proofs are based on the above methodology.

#### A Proof of Lemma 1

Let us define  $\hat{\mu}_t = \sum_{i=1}^{n_t} x_i^{(t)} y^{(t)} / n_t$  and  $\mu_t = Ba_t$  for all  $t \in [T+1]$ . Notice that

$$\hat{J} = (\hat{\mu}_1 / \|\hat{\mu}_1\|, \cdots, \hat{\mu}_T / \|\hat{\mu}_T\|) = (\hat{\mu}_1, \cdots, \hat{\mu}_T) \operatorname{diag}(\|\hat{\mu}_1\|^{-1}, \cdots, \|\hat{\mu}_T\|^{-1})$$

As a result, doing SVD for  $\hat{J}$  to obtain left singular vectors is equivalent to doing SVD for  $\hat{\Phi} = (\hat{\mu}_1, \dots, \hat{\mu}_T)$  to obtain left singular vectors (up to an orthogonal matrix, meaning rotation of the space spanned by the singular vectors) since multiplying a diagonal matrix on the right does not affect the collection of left singular vectors. It further means doing SVD for  $\hat{J}$  to obtain left singular vectors is equivalent to obtaining left singular vectors for  $\hat{\Phi} = (\hat{\mu}_1, \dots, \hat{\mu}_T) \text{diag}(\|\mu_1\|^{-1}, \dots, \|\mu_T\|^{-1})$  (up to an orthogonal matrix).

We mainly adopt the Davis-Kahan Theorem in [60]. We further denote  $\Phi = (\mu_1, \dots, \mu_T) \operatorname{diag}(\|\mu_1\|^{-1}, \dots, \|\mu_T\|^{-1}).$ 

**Lemma 2** (A variant of Davis–Kahan Theorem). Assume  $\min\{T, p\} > r$ . For simplicity, we denote  $\hat{\sigma}_1 \geq \hat{\sigma}_2 \geq \cdots \geq \hat{\sigma}_r$  as the top largest r singular value of  $\hat{\Phi}$  and  $\sigma_1 \geq \sigma_2 \geq \cdots \geq \sigma_r$  as the top largest r singular value of  $\Phi$ . Let  $V = (v_1, \cdots, v_r)$  be the orthonormal matrix consists of left singular vectors corresponding to  $\{\sigma_i\}_{i=1}^r$  and  $\hat{V} = (\hat{v}_1, \cdots, \hat{v}_r)$  be the orthonormal matrix consists of left singular vectors corresponding to  $\{\hat{\sigma}_i\}_{i=1}^r$ . Then,

$$\|\sin\Theta(\hat{V},V)\|_F \lesssim \frac{(2\sigma_1 + \|\hat{\Phi} - \Phi^*\|_{op})\min\{r^{0.5}\|\hat{\Phi} - \Phi^*\|_{op}, \|\hat{\Phi} - \Phi^*\|_F\}}{\sigma_r^2}.$$

Moreover, there exists an orthogonal matrix  $\hat{O} \in \mathbb{R}^{r \times r}$ , such that  $\|\hat{V}\hat{O} - V\|_F \leq \sqrt{2}\|\sin\Theta(\hat{V},V)\|_F$ , and

$$\|\hat{V}\hat{O} - V\|_F \lesssim \frac{(2\sigma_1 + \|\hat{\Phi} - \Phi^*\|_{op})\min\{r^{0.5}\|\hat{\Phi} - \Phi^*\|_{op}, \|\hat{\Phi} - \Phi^*\|_F\}}{\sigma_r^2}.$$

It is worth noticing that actually B plays the exact same role as V. Since B has orthonormal columns, for  $\phi$  we have

$$\Phi = B(a_1, \cdots, a_T) \operatorname{diag}(\|\mu_1\|^{-1}, \cdots, \|\mu_T\|^{-1})$$
  
=  $B(a_1, \cdots, a_T) \operatorname{diag}(\|a_1\|^{-1}, \cdots, \|a_T\|^{-1}).$ 

Thus, B is a solution of the SVD step in Algorithm 1.

**Lemma 3** (Restatement of Lemma 1). Under Assumption 1, if  $n > c_1 \max\{pr^2/T, r^2 \log(1/\delta)/T, r^2\}$  for some universal constant  $c_1 > 0$  and  $2r \le \min\{p, T\}$ , for all  $t \in [T]$ . For  $\hat{W}_1$  obtained in Algorithm 1, with probability at least  $1 - O(n^{-100})$ ,

$$\|\sin \Theta(\hat{W}_1, B)\|_F \lesssim r\left(\sqrt{\frac{1}{n}} + \sqrt{\frac{p}{nT}} + \sqrt{\frac{\log n}{nT}}\right).$$

Proof. By a direct application of Lemma 2, we can obtain

$$\|\sin\Theta(\hat{W}_1, B)\|_F \lesssim \frac{(2\sigma_1 + \|\hat{\Phi} - \Phi\|_{op})\min\{r^{0.5}\|\hat{\Phi} - \Phi\|_{op}, \|\hat{\Phi} - \Phi\|_F\}}{\sigma_r^2}$$

Besides, we know that the left singular vectors of  $\Phi$  are the same as the ones of  $M = [a_1, \dots, a_T]$ since  $\Phi = BM \operatorname{diag}(||a_1||^{-1}, \dots, ||a_T||^{-1})$ .

To estimate  $\|\hat{\Phi} - \Phi\|_{op} = \sup_{v \in \mathbb{S}^{p-1}} \|v^{\top}(\hat{\Phi} - \Phi)\|$ , for any fixed  $v \in \mathbb{S}^{p-1}$ , by standard chaining argument in Chapter 6 in [57], we know that

$$\mathbb{P}\left(\|v^{\top}(\hat{\Phi} - \Phi)\| \gtrsim \sqrt{\frac{T}{n}} + \sqrt{\frac{\log(1/\delta)}{n}}\right) \le \delta$$

Then, we use chaining again for the  $\psi_2$ -process  $\{v : \|v^{\top}(\hat{\Phi} - \Phi)\|\}$ , we obtain

$$\mathbb{P}\left(\sup_{v\in\mathbb{S}^{p-1}}\|v^{\top}(\hat{\Phi}-\Phi)\|\gtrsim\sqrt{\frac{p}{n}}+\sqrt{\frac{T}{n}}+\sqrt{\frac{\log(1/\delta)}{n}}\right)\leq\delta.$$

Besides, we know  $\sigma_r(M) = \sqrt{T/r}$  by assumption 1, and we also have  $\sum_{i=1}^r \sigma^2(M) = T$ , thus, we know that  $\sigma_1(M)$  and  $\sigma_r(M)$  are both of order  $\Theta(\sqrt{T/r})$ 

$$\|\sin\Theta(\hat{W}_1, B)\|_F \lesssim \frac{(\sqrt{\frac{p}{n}} + \sqrt{\frac{T}{n}} + \sqrt{\frac{\log(1/\delta)}{n}} + \sqrt{T/r})\sqrt{r}(\sqrt{\frac{p}{n}} + \sqrt{\frac{T}{n}} + \sqrt{\frac{\log(1/\delta)}{n}})}{T/r}$$

by simple calculation, we further have

$$\|\sin\Theta(\hat{W}_1, B)\|_F \lesssim r\sqrt{r}(\frac{1}{n} + \frac{p}{nT} + \frac{\log(1/\delta)}{nT}) + r(\sqrt{\frac{1}{n}} + \sqrt{\frac{p}{nT}} + \sqrt{\frac{\log(1/\delta)}{nT}}).$$

If we further have  $n > r \max\{p/T, \log(1/\delta)/T, 1\}$ , we further have

$$\|\sin\Theta(\hat{W}_1, B)\|_F \lesssim r(\sqrt{\frac{1}{n}} + \sqrt{\frac{p}{nT}} + \sqrt{\frac{\log(1/\delta)}{nT}}).$$

Plugging into  $\delta = n^{-100}$ , the proof is complete.

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### **B** Proof of Corollary 1

**Corollary 2** (Restatement of Corollary 1). Under Assumption 1, if  $n > c_1 \max\{pr^2/T, r^2 \log(1/\delta)/T, r^2, rn_{T+1}\}$  for some universal constant  $c_1 > 0$ ,  $2r \le \min\{p, T\}$ , then for  $\hat{W}_1$  obtained in Algorithm 1, with probability at least  $1 - O(n^{-100})$ ,

$$\mathcal{R}(\hat{W}_1, \hat{w}_2^{(T+1)}) \lesssim \sqrt{\frac{r + \log n}{n_{T+1}}} + \sqrt{\frac{r^2 p}{nT}}.$$

*Proof.* By DK-lemma, we know there exists a  $W_1^*$  such that  $W_1^* \in \operatorname{argmin}_{W \in \mathbb{O}_{p \times r}} \|W^\top \mu_{T+1}\|$  (the minimizer is not unique, so we use  $\in$  instead of = to indicate  $W_1^*$  belongs to the set consists of minimizers) and  $\|W_1^* - \hat{W}_1\|$  is small.

$$\begin{aligned} \mathcal{R}(\hat{W}_{1}, \hat{w}_{2}^{(T+1)}) &= L(\mathcal{P}_{x,y}^{(T+1)}, \hat{w}_{2}^{(T+1)}, \hat{W}_{1}) - \min_{\|w_{2}\| \leq 1, W_{1} \in \mathbb{O}_{p \times r}} L(\mathcal{P}_{x,y}^{(T+1)}, w_{2}, W_{1}) \\ &= -\langle \frac{\hat{W}_{1}^{\top} \hat{\mu}_{T+1}}{\|\hat{W}_{1}^{\top} \hat{\mu}_{T+1}\|}, \hat{W}_{1}^{\top} \mu_{T+1} \rangle + \|W_{1}^{*\top} \mu_{T+1}\| \\ &= -\langle \frac{\hat{W}_{1}^{\top} \hat{\mu}_{T+1}}{\|\hat{W}_{1}^{\top} \hat{\mu}_{T+1}\|}, \hat{W}_{1}^{\top} \mu_{T+1} \rangle + \langle \frac{W_{1}^{*\top} \hat{\mu}_{T+1}}{\|W_{1}^{*\top} \hat{\mu}_{T+1}\|}, W_{1}^{*\top} \mu_{T+1} \rangle \\ &- \langle \frac{W_{1}^{*\top} \hat{\mu}_{T+1}}{\|W_{1}^{*\top} \hat{\mu}_{T+1}\|}, W_{1}^{*\top} \mu_{T+1} \rangle + \|W_{1}^{*\top} \mu_{T+1}\| \\ &\lesssim \|\hat{W}_{1} - W_{1}^{*}\|\|\mu_{T+1}\| + \|W_{1}^{*\top} \mu_{T+1} - W_{1}^{*\top} \hat{\mu}_{T+1}\| \\ &\lesssim \|\hat{W}_{1} - W_{1}^{*}\|\|\mu_{T+1}\| + \|B^{\top} \mu_{T+1} - B^{\top} \hat{\mu}_{T+1}\| \end{aligned}$$

if  $n > r^2 \max\{p/T, \log(1/\delta)/T, 1\}$ . The last formula is due to the fact that  $W_1^*$  and B are different only up to an orthogonal matrix.

By standard chaining techniques, we have with probability  $1 - \delta$ 

$$\|B_1^{\top}\mu_{T+1} - B_1^{\top}\hat{\mu}_{T+1}\| \lesssim \sqrt{\frac{r}{n_{T+1}}} + \sqrt{\frac{\log(1/\delta)}{n_{T+1}}}.$$

Thus, we can further bound  $\|\hat{W}_1 - W_1^*\|$  by  $\sqrt{2}\|\sin\Theta(\hat{W}_1, B)\|_F$ , thus, by Lemma 1, we have

$$\mathcal{R}(\hat{W}_1, \hat{w}_2^{(T+1)}) \lesssim \sqrt{\frac{r + \log(1/\delta)}{n_{T+1}}} + r(\sqrt{\frac{1}{n}} + \sqrt{\frac{p}{nT}} + \sqrt{\frac{\log(1/\delta)}{nT}}).$$

Now, if we further have  $n > rn_{T+1}$ , we have

$$\mathcal{R}(\hat{W}_1, \hat{w}_2^{(T+1)}) \lesssim \sqrt{\frac{r + \log(1/\delta)}{n_{T+1}}} + \sqrt{\frac{r^2 p}{nT}}.$$

Plugging into  $\delta = n^{-100}$ , the proof is complete.

#### C Proof of Theorem 1

**Theorem 5** (Restatement of Theorem 1). Under Assumption 2 and 3, for  $||a_{T+1}|| = \alpha = \Omega(1)$ , if  $n > c_1 \max\{r^2, r/\alpha_T\} \cdot \max\{p \log T, \log n/T, 1\}$  and  $n > c_2(\alpha \alpha_T)^2 rn_{T+1}$  for universal constants  $c_1, c_2, 2r \le \min\{p, T\}$ . There exists a universal constant  $c_3$ , such that if we choose  $\varepsilon \in [\max_{t \in S_1} ||a_t|| + c_3 \sqrt{p \log T/n}, \min_{t \in S_2} ||a_t|| - c_3 \sqrt{p \log T/n}]$  (this set will not be empty if T, n are large enough), for  $\hat{W}_1^{adv}$ ,  $\hat{w}_2^{adv,(T+1)}$  obtained in Algorithm 2 with q = 2, with probability at least  $1 - O(n^{-100})$ ,

$$\|\sin\Theta(\hat{W}_1^{adv}, B)\|_F \lesssim (\alpha_T)^{-1} \left(\sqrt{\frac{r^2}{n}} + \sqrt{\frac{pr^2}{nT}} + \sqrt{\frac{r^2\log n}{nT}}\right),$$

and the excess risk

$$\mathcal{R}(\hat{W}_1^{adv}, \hat{w}_2^{adv, (T+1)}) \lesssim \alpha \sqrt{\frac{r + \log n}{n_{T+1}}} + (\alpha_T)^{-1} \left(\sqrt{\frac{r^2 p}{nT}}\right).$$

*Proof.* For  $\ell_2$ -adversarial training, we have

$$\begin{split} \hat{\beta}_t^{adv} &= \operatorname{argmin}_{\|\beta_t\| \le 1} \max_{\|\delta_i\|_p \le \varepsilon} \frac{1}{n_t} \sum_{i=1}^{n_t} -y_i^{(t)} \langle \beta_t, x_i^{(t)} + \delta_i \rangle \\ &= \operatorname{argmin}_{\|\beta_t\| \le 1} \max_{\|\delta_i\|_p \le \varepsilon} \frac{1}{n_t} \sum_{i=1}^{n_t} -y_i^{(t)} \langle \beta_t, x_i^{(t)} \rangle + \varepsilon \|\beta_t\| \end{split}$$

Recall  $\hat{\mu}_t = \frac{1}{n_t} \sum_{i=1}^{n_t} y_i^{(t)} x_i^{(t)}$ , if we have  $\|\hat{\mu}_t\| \ge \varepsilon$ , then  $\hat{\beta}_t^{adv} = \hat{\mu}_t / \|\hat{\mu}_t\|$ , otherwise,  $\hat{\beta}_t^{adv} = 0$ . We denote

$$\hat{G} = [\hat{\beta}_1^{adv}, \cdots, \hat{\beta}_T^{adv}].$$

Since  $|S_1| = \Theta(T)$ , there exists a universal constant  $c_3$  such that with probability  $1 - \delta$ , we have for all  $i \in S_1$ ,  $\hat{\mu}_i \leq ||a_i|| + c_3\sqrt{p\log T/n}$ . Thus, if T is large enough, the set  $[\max_{t \in S_1} ||a_t|| + c_3\sqrt{p\log T/n}, \min_{t \in S_2} ||a_t|| - c_3\sqrt{p\log T/n}]$  is non-empty. If we choose  $\varepsilon \in [\max_{t \in S_1} ||a_t|| + c_3\sqrt{p\log T/n}, \min_{t \in S_2} ||a_t|| - c_3\sqrt{p\log T/n}]$ , for all  $t \in S_2$ ,  $\hat{\beta}_t^{adv} = \hat{\mu}_t/||\hat{\mu}_t||$ . Meanwhile,  $\hat{G}_{S_1}$  is a zero matrix.

Notice that the left singular vectors obtained by applying SVD to  $\hat{G}$  for left singular vectors is equivalent to applying SVD for left singular vectors to  $\hat{G}_{S_2}$ , which is further equivalent to applying SVD for left singular vectors to  $\hat{\Phi}_{S_2}$ , given that  $\hat{G}_2$  is equal to  $\hat{\Phi}_{S_2}$  times a diagonal matrix on the right. Thus, we have

$$\|\sin\Theta(\hat{W}_1^{adv}, B)\|_F \lesssim \frac{(2\sigma_1(\Phi_{S_2}) + \|\hat{\Phi}_{S_2} - \Phi_{S_2}\|_{op})\min\{r^{0.5}\|\hat{\Phi}_{S_2} - \Phi_{S_2}\|_{op}, \|\hat{\Phi}_{S_2} - \Phi_{S_2}\|_F\}}{\sigma_r^2(\Phi_{S_2})}.$$

By our assumptions, we know that

$$\mathbb{P}\left(\sup_{v\in\mathbb{S}^{p-1}}\|v^{\top}(\hat{\Phi}_{S_2}-\Phi_{S_2})\|\gtrsim \alpha_T^{-1}(\sqrt{\frac{p}{n}}+\sqrt{\frac{T}{n}}+\sqrt{\frac{\log(1/\delta)}{n}})\right)\leq \delta.$$

As a result,

$$\|\sin\Theta(\hat{W}_1, B)\|_F \lesssim \alpha_T^{-2} r \sqrt{r} (\frac{1}{n} + \frac{p}{nT} + \frac{\log(1/\delta)}{nT}) + \alpha_T^{-1} r (\sqrt{\frac{1}{n}} + \sqrt{\frac{p}{nT}} + \sqrt{\frac{\log(1/\delta)}{nT}}).$$

If we further have  $n > \frac{r}{\alpha_T} \max\{p/T, \log(1/\delta)/T, 1\}$ , we further have

$$\|\sin\Theta(\hat{W}_1, B)\|_F \lesssim (\alpha_T)^{-1} r(\sqrt{\frac{1}{n}} + \sqrt{\frac{p}{nT}} + \sqrt{\frac{\log(1/\delta)}{nT}}).$$

Now, if we further have  $n > (\alpha \alpha_T)^2 r n_{T+1}$ , we have

$$\mathcal{R}(\hat{W}_1, \hat{w}_2^{(T+1)}) \lesssim \alpha \sqrt{\frac{r + \log(1/\delta)}{n_{T+1}}} + (\alpha_T)^{-1} \sqrt{\frac{r^2 p}{nT}}$$

Plugging into  $\delta = n^{-100}$ , the proof is complete.

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**Remark 6** ( $\ell_2$ -adversarial training v.s. standard training). The proof of the counterpart of Lemma 1 under the setting of Theorem 1 basically follows similar methods in the proof of Lemma 1. The only modification is that we need an extra step:

$$\mathbb{P}\left(\sup_{v\in\mathbb{S}^{p-1}}\|v^{\top}(\hat{\Phi}-\Phi)\|\gtrsim\sqrt{\frac{p}{n}}+\sqrt{\frac{T}{n}}+\sqrt{\frac{\log(1/\delta)}{n}}\right)\leq\mathbb{P}\left(\sup_{v\in\mathbb{S}^{p-1}}\|v^{\top}(\hat{\Phi}_{S_{1}}-\Phi_{S_{1}})\|\gtrsim\sqrt{\frac{p}{n}}+\sqrt{\frac{T}{n}}+\sqrt{\frac{\log(1/\delta)}{n}}\right)\\
+\mathbb{P}\left(\sup_{v\in\mathbb{S}^{p-1}}\|v^{\top}(\hat{\Phi}_{S_{2}}-\Phi_{S_{2}})\|\gtrsim\sqrt{\frac{p}{n}}+\sqrt{\frac{T}{n}}+\sqrt{\frac{\log(1/\delta)}{n}}\right)$$

and recall that both  $|S_1|$  and  $|S_2|$  are of order  $\Theta(T)$ .

#### **D Proof of Theorem 2**

**Theorem 6** (Restatement of Theorem 2). Under Assumptions 1 and 4, if  $n > c_1 \cdot r^2 \max\{s^2 \log^2 T/T, rn_{T+1}, 1\}$  for some universal constants  $c_1 > 0$ ,  $2r \le \min\{p, T\}$ . There exists a universal constant  $c_2$ , such that if we choose  $\varepsilon > c_2 \sqrt{\log p/n}$ , for and  $\hat{W}_1^{adv}$ ,  $\hat{w}_2^{adv,(T+1)}$  obtained in Algorithm 2 with  $q = \infty$ , with probability at least  $1 - O(n^{-100}) - O(T^{-100})$ ,

$$\|\sin\Theta(\hat{W}_1^{adv}, B)\|_F \lesssim r\left(\sqrt{\frac{1}{n}} + \sqrt{\frac{s^2}{nT}}\right) \cdot \log(T+p)$$

and the excess risk

$$\mathcal{R}(\hat{W}_1^{adv}, \hat{w}_2^{adv, (T+1)}) \lesssim \left(\sqrt{\frac{r+\log n}{n_{T+1}}} + r\sqrt{\frac{s^2}{nT}}\right) \cdot \log(T+p).$$
(7)

*Proof.* For  $\ell_{\infty}$ -adversarial training, we have

$$\begin{split} \hat{\beta}_t^{adv} &= \operatorname{argmin}_{\|\beta_t\| \le 1} \max_{\|\delta_i\|_{\infty} \le \varepsilon} \frac{1}{n_t} \sum_{i=1}^{n_t} -y_i^{(t)} \langle \beta_t, x_i^{(t)} + \delta_i \rangle \\ &= \operatorname{argmin}_{\|\beta_t\| \le 1} \frac{1}{n_t} \sum_{i=1}^{n_t} -y_i^{(t)} \langle \beta_t, x_i^{(t)} \rangle + \varepsilon \|\beta_t\|_1 \\ &= \operatorname{argmin}_{\|\beta_t\| \le 1} \langle \beta_t, \frac{1}{n_t} \sum_{i=1}^{n_t} -y_i^{(t)} x_i^{(t)} \rangle + \varepsilon \|\beta_t\|_1 \end{split}$$

Recall  $\hat{\mu}_t = \frac{1}{n_t} \sum_{i=1}^{n_t} y_i^{(t)} x_i^{(t)}$ . By observation, when reaching minimum, we have to have  $sgn(\beta_{tj}) = sgn(\hat{\mu}_{tj})$ , therefore

$$\begin{aligned} \operatorname*{argmax}_{\|\beta_t\|=1} &\sum_{j=1}^d \hat{\mu}_{tj} \beta_{tj} - \varepsilon |\beta_{tj}| \\ = \operatorname*{argmax}_{\|\beta_t\|=1} &\sum_{j=1}^d (\hat{\mu}_{tj} - \varepsilon \cdot sgn(\hat{\mu}_{tj})) \beta_{tj} \\ = &\frac{T_{\varepsilon}(\hat{\mu})}{\|T_{\varepsilon}(\hat{\mu})\|}, \end{aligned}$$

where  $T_{\varepsilon}(\hat{\mu})$  is the hard-thresholding operator with  $(T_{\varepsilon}(\hat{\mu}))_j = sgn(\hat{\mu}_j) \cdot \max\{|\hat{\mu}_j| - \varepsilon, 0\}$ . We denote

$$\hat{G} = [\hat{\beta}_1^{adv}, \cdots, \hat{\beta}_T^{adv}].$$

By the choice of  $\varepsilon$ ,  $\varepsilon \gtrsim C\sqrt{\frac{\log p}{n}}$  for sufficiently large C, we have that the column sparsities of  $\hat{G}$  is no larger than  $s \log T$ . As a result, the total number of non-zero elements in  $\hat{G}$  is less than  $O(Ts \log T)$  with probability at least  $1 - T^{-100}$ .

Now we divide the rows of  $\hat{G}$  by two parts:  $[p] = A_1 \cup A_2$ , where  $A_1$  consists of indices of rows whose sparsity smaller than or equal to s, and  $A_2$  consists of indices of rows whose sparsity larger than s.

Since the number of non-zero elements in  $\hat{G}$  is less than  $Ts \log T$ , we have  $|A_2| \leq T \log T$ . Using the similar analysis as in the proof of Lemma 1, we have

$$\|\hat{\Phi}_{A_2} - \Phi_{A_2}\| \le \sqrt{\frac{T\log T}{n}}$$

For the rows in  $A_1$ , all of them has sparsity  $\leq s$ , so the maximum  $\ell_1$  norm of these rows

$$\|\hat{\Phi}_{A_1} - \Phi_{A_1}\|_{\infty} = O_P(s\sqrt{\frac{\log T}{n}}).$$

Similarly, the maximum  $\ell_1$  norm of the columns in  $\hat{G}_{A_1}$  satisfies

$$\|\hat{\Phi}_{A_1} - \Phi_{A_1}\|_1 = O_P(s\sqrt{\frac{\log p}{n}}).$$

Therefore, we have

$$\|\hat{\Phi}_{A_1} - \Phi_{A_1}\| \le \sqrt{\|\hat{\Phi}_{A_1} - \Phi^*_{A_1}\|_{\infty}} \|\hat{\Phi}_{A_1} - \Phi_{A_1}\|_1 = O_P(s\sqrt{\frac{\log p + \log T}{n}}).$$

Consequently,

$$\|\hat{\Phi} - \Phi\| \le \|\hat{\Phi}_{A_1} - \Phi_{A_1}\| + \|\hat{\Phi}_{A_2} - \Phi_{A_2}\| = O_P(s\sqrt{\frac{\log p + \log T}{n}})$$

As a result, when  $s\sqrt{\frac{\log p + \log T}{n}} \lesssim T/r$ , applying Lemma 2, we obtain

$$\|\sin\Theta(\hat{W}_1, B)\|_F \lesssim \sin\theta(\hat{W}_1^{adv}, B) \lesssim (\sqrt{\frac{r}{n}} + \sqrt{\frac{rs^2}{nT}}) \cdot \log(T+p).$$

Now, if we further have  $n > (\alpha \alpha_T)^2 n_{T+1} / \nu$ , we have

$$\mathcal{R}(\hat{W}_1, \hat{w}_2^{(T+1)}) \lesssim \sqrt{\frac{r + \log(1/\delta)}{n_{T+1}}} + \sqrt{\frac{rs^2}{nT}} \cdot \log(T+p).$$

**Remark 7** ( $\ell_{\infty}$ -adversarial training v.s. standard training). The proof of the counterpart of Lemma 1 under the setting of Theorem 2 follows exact the same method in the proof of Lemma 1.

#### E Proof of the case with pseudo-labeling

**Theorem 7** (Restatement of Theorem 3). Denote  $\tilde{n} = \min_{t \in [T]} n_t^u$  and assume  $\tilde{n} > c_1 \max\{pr^2/T, r^2 \log(1/\delta)/T, r^2, n\}$  for some constant  $c_1 > 0$ . Assume  $\sigma_r(M^\top M/T) = \Omega(1/r)$  and  $n^{c_2} \gtrsim \tilde{n} \gtrsim n$  for some  $c_2 > 1$ , if  $n \gtrsim (T+d)$  and  $\min_{t \in [T]} ||a_t|| = \Theta(\log^2 n)$  and  $\eta_i^{(t)} \sim \mathcal{N}_p(0, \rho_t^2 I^2)$  for  $\rho_t = \Theta(1)$ . Let  $\hat{W}_{1,aug}$  obtained in Algorithm 3, with probability  $1 - O(n^{-100})$ ,

$$\|\sin\Theta(\hat{W}_{1,aug},B)\|_F \lesssim r\left(\sqrt{\frac{1}{\tilde{n}}} + \sqrt{\frac{p}{\tilde{n}T}} + \sqrt{\frac{\log n}{\tilde{n}T}}\right)$$

*Proof.* Let us first analyze the performance of pseudo-labeling algorithm in each individual task. In the following, we analyze the properties of  $y_i^{u,(t)}$  and  $\hat{\mu}_{final}^{(t)} = \frac{1}{n_t^u + n_t} \sum_{i=1}^{n_t^u + n_t} (\sum_{i=1}^{n_u^t} x_i^u y_i^u + \sum_{i=1}^{n_t} x_i^u y_i^u)$ . Since  $\tilde{n} \gtrsim n$  and we only care about the rate in the result. In the following, we derive the results for  $\hat{\mu}_{final}^{(t)} = \frac{1}{n_t^u} \sum_{i=1}^{n_t^u + n_t} (\sum_{i=1}^{n_u^t} x_i^u y_i^u)$ . Also, for the notational simplicity, we omit the index t in the following analysis.

We follow the similar analysis of Carmon et al. [11] to study the property of  $y_i^u$ . Let  $b_i$  be the indicator that the *i*-th pseudo-label is incorrect, so that  $x_i^u \sim N((1-2b_i)y_i^u\mu, I) := (1-2b_i)y_i^u\mu + \varepsilon_i^u$ . Then we can write

 $\hat{\mu}_{final} = \gamma \mu + \tilde{\delta},$ 

where  $\gamma = \frac{1}{n_u} \sum_{i=1}^{n_u} (1 - 2b_i)$  and  $\tilde{\delta} = \frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_i^u y_i^u$ .

Let's write  $y_i^u = sign(x_i^\top \hat{\mu})$ . Using the rotational invariance of Gaussian, without loss of generality, we choose the coordinate system where the first coordinate is in the direction of  $\hat{\mu}$ . Then  $y_i^u = sign(x_i^\top \hat{\mu}) = sign(x_{i1}) = sign(y_i^* \frac{\mu^\top \hat{\mu}}{\|\hat{\mu}\|} + \varepsilon_{i1}^u)$  and are independent with  $\varepsilon_{ij}^u$   $(j \ge 2)$ .

As a result,

$$\frac{1}{n_u}\sum_{i=1}^{n_u}\varepsilon_{ij}^u\cdot y_i^u\stackrel{d}{=}\frac{1}{n_u}\sum_{i=1}^{n_u}\varepsilon_{ij}^u,\quad \text{for }j\geq 2.$$

Now let's focus on  $\frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_{i1}^u \cdot y_i^u$ . Let  $y_i^* = (1 - 2b_i)y_i^u$ , we have

$$\frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_{i1}^u \cdot y_i^u = \frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_{i1}^u \cdot y_i^* + 2\frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_{i1}^u \cdot b_i \stackrel{d}{=} \frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_{i1}^u + 2\frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_{i1}^u \cdot b_i.$$

Since

$$\left(\frac{1}{n_u}\sum_{i=1}^{n_u}\varepsilon_{i1}^u\cdot b_i\right)^2 \le \left(\frac{1}{n_u}\sum_{i=1}^{n_u}(\varepsilon_{i1}^u)^2\right)\left(\frac{1}{n_u}\sum_{i=1}^{n_u}b_i^2\right) \lesssim \frac{1}{n_u}\sum_{i=1}^{n_u}b_i^2 = \frac{1}{n_u}\sum_{i=1}^{n_u}b_i \lesssim \mathbb{E}[b_i] + \frac{1}{\sqrt{n_u}} \lesssim +\frac{1}{n} + \frac{1}{\sqrt{n_u}},$$

where the last inequality is due to the fact that

$$\begin{split} \mathbb{E}[b_i] = \mathbb{P}(y_i^u \neq y_i^*) &= \mathbb{P}(sign(y_i^* \frac{\mu^\top \hat{\mu}}{\|\hat{\mu}\|} + \varepsilon_{i1}^u) \neq y_i^*) \\ \leq \mathbb{P}(sign(y_i^* \frac{\mu^\top \hat{\mu}}{\|\hat{\mu}\|} + \varepsilon_{i1}^u) \neq y_i^* \mid \frac{\mu^\top \hat{\mu}}{\|\hat{\mu}\|} > \frac{1}{2} \|\mu\|) + \mathbb{P}(\frac{\mu^\top \hat{\mu}}{\|\hat{\mu}\|} > \frac{1}{2} \|\mu\|) \\ \lesssim \exp^{-\|\mu\|/2} + \frac{1}{n^C} \end{split}$$

As a result, we have

$$\tilde{\delta} \stackrel{d}{=} \frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_i^u + e,$$

where  $||e||_2 \lesssim \frac{1}{\sqrt{n_u}} + \frac{1}{n^C}$ .

Additionally, we have  $\gamma = \frac{1}{n_u} \sum_{i=1}^{n_u} (1-2b_i) = 1 - \frac{2}{n_u} \sum_{i=1}^{n_u} b_i = 1 - O(\frac{1}{\sqrt{n_u}} + \frac{1}{n^c}).$ As a result, for each  $t \in [T]$ , we have

$$\hat{\mu}_t = \mu_t + \frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_i^u + e',$$

with  $||e'||_2 \lesssim \frac{1}{\sqrt{n_u}} + \frac{1}{n^{C'}}$  being a negligible term.

Since e' is negligible, we can then follow the same proof as those in Section A by considering  $\tilde{\mu}_t = \mu_t + \frac{1}{n_u} \sum_{i=1}^{n_u} \varepsilon_i^u$  and obtain the desired results.

Similarly, due to the negligibility of e', we can prove Theorem 4 by following the exact same techniques in Sections C and D.

#### F Lower bound proof

**Proposition 2** (Restatement of Proposition 1). Let us consider the parameter space  $\Xi = \{A \in \mathbb{R}^{p \times r}, B \in \mathbb{R}^{p \times r} : \sigma_r(A^\top A/T) \gtrsim 1, B^\top B = I_r\}$ . If  $nT \gtrsim rp$ , we then have

$$\inf_{\hat{W}_1} \sup_{\Xi} \mathbb{E} \| \sin \Theta(B, \hat{W}_1) \|_F \gtrsim \sqrt{\frac{rp}{nT}}.$$

We first invoke the Fano's lemma.

**Lemma 4** ([54]). Let  $M \ge 0$  and  $\mu_0, \mu_1, ..., \mu_M \in \Theta$ . For some constants  $\alpha \in (0, 1/8), \gamma > 0$ , and any classifier  $\hat{G}$ , if  $\operatorname{KL}(\mathbb{P}_{\mu_i}, \mathbb{P}_{\mu_0}) \le \alpha \log M$  for all  $1 \le i \le M$ , and  $L(\mu_i, \mu_j)$  for all  $0 \le i \ne j \le M$ , then

$$\inf_{\hat{\mu}} \sup_{i \in [M]} \mathbb{E}_{\mu_i}[L(\mu_i, \hat{\mu})] \gtrsim \gamma.$$

Now we take  $B_0, B_1, ..., B_M$  as the  $\eta$ -packing number of  $O^{p \times r}$  with the  $\sin \theta$  distance. Then according to [41, 52], we have

$$\log M \asymp rd\log(\frac{1}{\eta}).$$

For any  $i \in [M]$ , we have

$$\mathrm{KL}(\mathbb{P}_{B_i}, \mathbb{P}_{B_0}) = \sum_{t=1}^T n \| (B_i - B_0) a_t \|^2 \le nT\eta^2.$$

Let  $\eta = \sqrt{\frac{rd}{nT}}$ , we complete the proof.

### **G** Additional Empirical Results

We provide additional results on transfer performance with varied amounts of pseudo-labels in Table 2. Here, we train models with both adversarial (allowed maximum perturbations of  $\varepsilon = 1$  with respect to the  $\ell_2$  norm) and non-adversarial (standard) training on ImageNet. The observed trend is the same as on the CIFAR-10 and CIFAR-100 tasks from Table 1 – both using robust training and additional pseudo-labeled data improve performance.

Table 2: Additional results extending Table 1. Effect of amount of pseudo-labels on transfer task performance (measured with accuracy). At 0%, we just use 10% of data from the source task; at 900%, we use all remaining 90% of data with pseudo-labels (this is 9 times the train set size). Adversarial training corresponds to using  $\ell_2$ -adversarial training with  $\varepsilon = 1$  on the source task. As per Section 7 of [46], images in all datasets down-scaled to  $32 \times 32$  before scaling back to  $224 \times 224$ .

Source Task	Target Task	+0% Pseudo-labels	+200% Pseudo-labels	+500% Pseudo-labels	+900% Pseudo-labels
ImageNet	Aircraft [35]	17.3%	17.6%	17.9%	19.9%
ImageNet (w/adv.training)	Aircraft	21.2%	20.9%	24.0%	24.5%
ImageNet	Flowers [40]	60.7%	64.9%	65.4%	66.5%
ImageNet (w/adv.training)	Flowers	66.9%	68.1%	70.0%	70.1%
ImageNet	Food [8]	33.7%	36.0%	36.7%	37.2%
ImageNet (w/adv.training)	Food	35.8%	37.5%	39.4%	40.8%
ImageNet	Pets [42]	43.2%	44.9%	48.4%	49.0%
ImageNet (w/adv.training)	Pets	47.9%	53.1%	58.9%	59.6%

## **H** Experiment Details

#### H.1 Training Hyperparameters

All of our experiments use the ResNet-18 architecture. When transferring to the target task, we only update the final layer of the model. Our hyperparameter choices are identical to those used in [46]:

- 1. ImageNet (source task) models are trained with SGD for 90 epochs with a momentum of 0.9, weight decay of 1e 4, and a batch size of 512. The initial learning rate is set to 0.1 and is updated every 30 epochs by a factor of 0.1. The adversarial examples for adversarial training are generated using 3 steps with step size  $\frac{2\varepsilon}{3}$ .
- 2. Target task models are trained for 150 epochs with SGD with a momentum of 0.9, weight decay of 5e 4, and a batch size of 64. The initial learning rate is set to 0.01 and is updated every 50 epochs by a factor of 0.1.

Data augmentation is also identical to the methods used in [46]. As per Section 7 of [46], we scale all our target task images down to size  $32 \times 32$  before rescaling back to size  $224 \times 224$ .

Experiments were run on a GPU cluster. A variety of NVIDIA GPUs were used, as allocated by the cluster. Training time for each source task model was around 2 days (less when using subsampled data) using 4 GPUs. Training time for each target task model was typically between 1-5 hours (depending on the dataset) using 1 GPU.

#### H.2 Pseudo-label Generation

When subsampling ImageNet (our source task), the sampled 10% with ground truth labels preserves the class label distribution. This sample is fixed for all our experiments. All ImageNet pseudo-labels are generated by a model trained on this 10% without any adversarial training. This model has a source task test accuracy (top-1) of 44.0%.

When training models with pseudo-labels, we preserve the class label distribution of the original training set (i.e., we add less pseudo-labels for those classes that have fewer examples in the entire training set).