# Appendix

# **Table of Contents**

A	Proof of Proposition 4	13	
B	Proof of Proposition 7	14	
С	2 Proof of Theorem 12		
D	Proof of Proposition 15		
Е	Proof of Theorem 19	17	
F	Proofs of Section 5 Results           F.1         Proof of Corollary 16	<b>19</b> 19 20 20 21	
G	Technical Lemmas	21	

# A Proof of Proposition 4

We recall Proposition 4

**Proposition 4.** Let f be  $(\mu, d)$ -convex. Suppose  $\hat{x}$  minimizes f over S and  $\tilde{x}$  minimizes f over star $(\hat{x}, S)$ . Then, for any  $x \in S$ ,

$$f(x) - f(\tilde{x}) \ge \mu \left(\frac{1}{3}d(x, \tilde{x})\right).$$

$$H_{2}$$

$$x \text{ (case 2)}$$

$$f_{\tilde{x}}$$

$$f_{\tilde{x}}$$

$$f_{\tilde{x}}$$

$$f_{\tilde{x}}$$

$$f_{\tilde{x}}$$

$$H_{1}$$

Figure 3: Illustration of Proposition 4.

x (case 1)

*Proof.* The proof relies on the observation that if  $\lambda \mapsto z_{\lambda}$  is a line segment such that  $z_0 \in \partial f_x$ (i.e.  $f(z_0) = f(x)$ ),  $z_1 \in \partial f_y$  (i.e.  $f(z_1) = f(y)$ ), and  $z_{\lambda} \in f_x \setminus f_y^{\circ}$  (i.e.  $f(x) \ge f(z_{\lambda}) \ge f(y)$ ) for  $\lambda \in [0, 1]$ , then

 $f(x) - f(y) = f(z_0) - f(z_1) \ge \mu(d(z_0, z_1)).$ 

This is seen by repeating the proof of Lemma 3: since  $z_{\lambda} \notin f_{y}^{\circ}$  we must have  $(f(z_{\lambda}) - f(z_{1}))/\lambda \ge 0$ . Plugging in  $z_{\lambda} = z_{0} + \lambda(z_{1} - z_{0})$  and taking the limit infimum as  $\lambda \downarrow 0$  gives  $0 \le \langle \nabla f | z_{1}, z_{0} - z_{1} \rangle$ . By  $(\mu, d)$ -convexity of f, then,

$$f(z_1) - f(z_0) \ge f(z_1) - f(z_0) - \langle \nabla f |_{z_1}, z_0 - z_1 \rangle \ge \mu(d(z_1, z_0)).$$

Finally, if k such segments  $z_{\lambda}$  form a path from x to y, then at least one of them must have  $d(z_1, z_0) \ge \frac{1}{k} d(x, y)$ . This is due to the triangle inequality for d.

We now restrict our attention to the plane P containing  $(x, \hat{x}, \tilde{x})$ . Let C be the minimal cone containing  $f_{\tilde{x}}$  with vertex at  $\hat{x}$ . We note that, by optimality of  $\tilde{x}$ , no point  $x \in S$  can lie in the interior of C.

Then  $C \cap P$  is complementary to the union of two half-planes  $H_1$  and  $H_2$  with boundary lines  $\ell_1$ and  $\ell_2$ , respectively. These lines intersect  $\partial f_x$  at two respective points,  $s_1$  and  $s_2$ , and are tangent to  $f_{\tilde{x}}$  at two respective points,  $t_1 \equiv \tilde{x}$  and  $t_2$ . Finally,  $f_x \supseteq f_{\hat{x}}$  by optimality of  $\hat{x}$ . This is depicted in Figure 3 above.

There are two cases to consider. In the first case,  $x \in H_1$ . Then the line segment connecting  $\tilde{x}$  and x is contained entirely in  $D = f_x \setminus f_{\tilde{x}}^{\circ}$ . By the preceding discussion,

 $f(x) - f(\tilde{x}) \ge \mu(d(x, \tilde{x})) \ge \mu\left(\frac{1}{3}d(x, \tilde{x})\right).$ 

In the second case  $x \in H_2 \setminus H_1$ . In this case, the segment from  $\tilde{x}$  to  $s_1$  along  $\ell_1$  lies entirely in D. Similarly, the segments from  $s_1$  to  $t_2$  and from  $t_2$  to x are line segments contained in D. Each of these three segments connects  $\partial f_{\tilde{x}}$  to  $\partial f_x$ , and they together form a path from  $\tilde{x}$  to x. By the preceding discussion,  $f(x) = f(\tilde{x}) > x(1 d(x, \tilde{x}))$ 

$$f(x) - f(\hat{x}) \ge \mu\left(\frac{1}{3}d(x,\hat{x})\right).$$

This completes the proof.

#### **B** Proof of Proposition 7

We recall Proposition 7:

**Proposition 7.** Let  $\mathcal{F}$  be a model class,  $\psi$  a  $(\mu, d)$ -convex loss, and  $f^*$  the population minimizer of the  $\psi$ -risk. Then, the star estimator  $\tilde{f}$  satisfies the excess risk bound

$$\mathbb{E}\Psi(\mathcal{E}(\tilde{f})) \le \mathbb{E}\Psi\left(\sup_{f \in \mathcal{F}'} \left\{ \frac{1}{n} \sum_{i=1}^{n} 2\varepsilon_i'(\psi_i(f^*) - \psi_i(f)) - (1 + \varepsilon_i')\mu(\frac{1}{3}d(f_i, f_i^*)) \right\} \right)$$
(10)

where the  $\varepsilon'_i$  are i.i.d. symmetric Rademacher random variables,  $\mathcal{F}' = \bigcup_{\lambda \in [0,1]} \lambda \mathcal{F} + (1-\lambda) \mathcal{F}$ , and  $\Psi : \mathbb{R} \to \mathbb{R}$  is any increasing, convex function.

Proof. We'll work forwards from (9), which says that

$$\mathcal{E}(\tilde{f}) = \mathbb{E}\psi(\tilde{f}_i) - \mathbb{E}\psi(f^*)_i \le \sup_{f \in \mathcal{F}'} \left\{ (\mathbb{E}_n - \mathbb{E})(\psi(f_i^*) - \psi(f_i)) - \mathbb{E}_n \mu(d(f_i, f_i^*)/3) \right\}$$

Using the notation  $\Delta_i(f) = \psi(f_i^*) - \psi(f)_i$ ,  $\nu_i(f) = \frac{1}{2}\mu(\frac{1}{3}d(f_i, f_i^*))$ , we can rewrite this as

$$\mathcal{E}(\tilde{f}) \le \sup_{f \in \mathcal{F}'} \left\{ \frac{1}{n} \sum_{i=1}^n (1 - \mathbb{E}) \Delta_i(f) - 2\nu_i(f) \right\}.$$

Adding and subtracting  $\mathbb{E}\nu_i(f)$  gives

$$= \sup_{f \in \mathcal{F}'} \left\{ \frac{1}{n} \sum_{i=1}^n (1 - \mathbb{E}) (\Delta_i(f) - \nu_i(f)) - (1 + \mathbb{E}) \nu_i(f) \right\}.$$

By applying  $\mathbb{E}\Psi$  to both sides and applying Lemma G1 (a symmetrization result along the lines of Liang et al. (2015)) with  $A = \Delta$ ,  $B = \nu$ , and  $T = \mathcal{F}$ , we get

$$\leq \mathbb{E}\Psi\left(2\sup_{f\in\mathcal{F}'}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}'(\Delta_{i}(f)-\nu_{i}(f))-\nu_{i}(f)\right\}\right)$$
$$=\mathbb{E}\Psi\left(\sup_{f\in\mathcal{F}'}\left\{\frac{1}{n}\sum_{i=1}^{n}2\varepsilon_{i}'(\psi(f^{*})_{i}-\psi(f)_{i}-\frac{1}{2}\mu(\frac{1}{3}d(f_{i},f_{i}^{*})))-\mu(\frac{1}{3}d(f_{i},f_{i}^{*}))\right\}\right),$$

where we plugged in the definition of  $\Delta_i$  and  $\nu_i$ . This completes the proof.

# C Proof of Theorem 12

We recall Theorem 12:

**Theorem 12.** Let  $\psi$  be an  $\eta$ -exp-concave loss function taking values in [0, m]. Then the star estimator in  $\mathcal{F}$  satisfies the excess risk bound

$$\mathbb{E}\Psi(\mathcal{E}(\tilde{f})) \le \mathbb{E}\Psi\left(\sup_{f,g\in\mathcal{F}'}\left\{\frac{1}{n}\sum_{i=1}^{n}4\varepsilon_i'(\psi_i(f)-\psi_i(g)) - \frac{\eta(\psi_i(f)-\psi(g_i))^2}{18m\eta\vee 36}\right\}\right).$$
 (15)

where  $\Psi$  is any increasing, convex function and  $\mathcal{F}' = \bigcup_{\lambda \in [0,1]} \lambda \mathcal{F} + (1-\lambda) \mathcal{F}$ . Alternatively, when  $\psi$  is *p*-uniformly convex with modulus  $\alpha$  and  $\|\psi\|_{\text{lip}}$ -Lipschitz, we have

$$\mathbb{E}\Psi(\mathcal{E}(\tilde{f})) \le \mathbb{E}\Psi\left(\sup_{f,g\in\mathcal{F}'}\left\{\frac{1}{n}\sum_{i=1}^{n}4\|\psi\|_{\mathrm{lip}}\left(f_{i}-g_{i}\right)\varepsilon_{i}'-\frac{\alpha|f_{i}-g_{i}|^{p}}{3^{p}}\right\}\right).$$
(16)

Proof. We'll work forwards from (10), which says that

$$\mathbb{E}\Psi(\mathcal{E}(\tilde{f})) \leq \mathbb{E}\Psi\left(\sup_{f\in\mathcal{F}'}\left\{\frac{1}{n}\sum_{i=1}^{n}2\varepsilon_{i}'(\psi(f^{*})_{i}-\psi(f)_{i}-\frac{1}{2}\mu(\frac{1}{3}d(f_{i},f_{i}^{*})))-\mu(\frac{1}{3}d(f_{i},f_{i}^{*}))\right\}\right)$$
$$\leq \mathbb{E}\Psi\left(\sup_{f,g\in\mathcal{F}'}\left\{\frac{1}{n}\sum_{i=1}^{n}2\varepsilon_{i}'(\psi(g)_{i}-\psi(f)_{i}-\frac{1}{2}\mu(\frac{1}{3}d(f_{i},g_{i})))-\mu(\frac{1}{3}d(f_{i},g_{i}))\right\}\right),$$

where the second step enlarges the domain in the supremum. For (15), we plug in the definition of the offset from Proposition 11

$$\mu(\frac{1}{3}d(x,y)) = \frac{|\psi(x) - \psi(y)|^2}{18m \vee 36/\eta}$$

Under the condition that  $\psi$  takes values in [0, m], we have that

$$|\psi(x) - \psi(y)|^2 \le 2m|\psi(x) - \psi(y)| \le (2m \lor 4/\eta)|\psi(x) - \psi(y)|.$$

It follows that we can apply our "contraction lemma" for offset processes, Lemma G2, with the contractions

$$\begin{split} |\psi(x) - \psi(y) - \frac{1}{2}\mu(\frac{1}{3}d(x,y))| \\ &\leq |\psi(x) - \psi(y)| + |\frac{1}{2}\mu(\frac{1}{3}d(x,y))| \\ &\leq |\psi(x) - \psi(y)| + \frac{(m \vee 2/\eta)|\psi(x) - \psi(y)|}{18m \vee 36/\eta} \leq \frac{19}{18}|\psi(x) - \psi(y)|. \end{split}$$

For (16), we first require the following lemma.

**Lemma C1.** Let  $\psi : \mathbb{R} \to \mathbb{R}$  be  $(\mu, d)$ -convex and  $\|\psi\|_{\text{lip}}$ -Lipschitz with respect to d(x, y) = |x - y|. Let  $r \ge 0$  be the largest constant such that  $\mu(cx) \le c^r \mu(x)$  for  $c \le 1$  which is non-negative by monotonicity of  $\mu$ . Then

$$\mu(\frac{1}{3}|x-y|) \le (\frac{2}{3})' \|\psi\|_{\text{lip}} |x-y|.$$

Applying the lemma with r = p, we simply apply the contractions

$$|\psi(x) - \psi(y) - \frac{1}{2}\mu(\frac{1}{3}d(x,y))| \le |\psi(x) - \psi(y)| + \frac{1}{2}\mu(\frac{1}{3}d(x,y)) \le (1 + 2^{p-1}/3^p) \|\psi\|_{\text{lip}} |x - y|$$

using Lemma G2, and then plug in

$$\mu(\frac{1}{3}d(x,y)) = \frac{\alpha|x-y|^p}{3^p},$$

from the definition of *p*-uniform convexity.

*Proof of Lemma C1.* Let z be the minimizer of  $\psi$  over [x, y]. By Lemma 3 we have  $\psi(x) - \psi(z) > \mu(|x - z|), \qquad \psi(y) - \psi(z) > \mu(|y - z|).$ 

Since 
$$|x - y| = |x - z| + |y - z|$$
 because  $z \in [x, y]$ , we have  
 $\mu(\frac{1}{3}|x - y|) \le \mu(\frac{1}{3}(|x - z| + |y - z|)).$ 

By monotonicity of  $\mu$ , this is

$$\leq \mu(\frac{2}{3}(|x-z| \lor |y-z|)) \\ \leq (\frac{2}{3})^r \{\mu(|x-z|) \lor \mu(|y-z|)\}.$$
Using (25), we have  

$$\leq (\frac{2}{3})^r |\psi(x) - \psi(z)| \lor (\frac{2}{3})^r |\psi(y) - \psi(z)| \\ \leq (\frac{2}{3})^r ||\psi||_{\text{lip}} |x-z| \lor (\frac{2}{3})^r ||\psi||_{\text{lip}} |y-z| \\ \leq (\frac{2}{3})^r ||\psi||_{\text{lip}} |x-y|,$$
where in the last step we again used that  

$$|x-y| = |x-z| + |y-z| \ge |x-z| \lor |y-z|$$

where in the last st

$$|x - y| = |x - z| + |y - z| \ge |x - z| \lor |y - z|$$

by our choice of z. This completes the proof.

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(25)

#### **Proof of Proposition 15** D

We recall Proposition 15.

**Proposition 15.** If  $\psi$  is a self-concordant loss and  $\hat{f}$  is the empirical risk minimizer in a convex class *F*, then

$$\mathbb{E}\Psi(\mathcal{E}(\tilde{f})) \leq \mathbb{E}\Psi\left(\sup_{f \in \mathcal{F}'} \left\{ \frac{1}{n} \sum_{i=1}^{n} 4(\psi_i(f) - \psi(f_i^*))\varepsilon_i' - \omega\left(\|f_i - f_i^*\|_{\psi, f_i^*}\right) \right\} \right), \quad (18)$$

for  $\omega(z) = z - \log(1+z)$ ,  $||z||_{\psi,w} \doteq \sqrt{z^2 \psi''(w)}$ , and  $(\varepsilon'_i)_{i=1}^n$  are independent, symmetric Rademacher random variables and  $\Psi$  is any increasing, convex function.

*Proof.* Combining the self-concordance inequality Lemma 14 with Lemmas 3 and 5 immediately gives us

$$\mathcal{E}(\hat{f}) = \mathbb{E}\psi(\hat{f}) - \mathbb{E}\psi(f^*) \ge \mathbb{E}\omega(\|\hat{f} - f^*\|_{\psi, f^*})$$
(26)

for the empirical risk minimizer  $\hat{f}$  in a convex class. Since  $\hat{f}$  is the risk minimizer, we also have  $\mathbb{E}_n \psi(f^*) - \mathbb{E}_n \psi(\hat{f}) \ge 0$ . Adding these two and rearranging, we have

$$\begin{split} \mathcal{E}(\hat{f}) &\leq 2\mathcal{E}(\hat{f}) - \mathbb{E}\omega(\|\hat{f} - f^*\|_{\psi, f^*}) \\ &\leq 2(\mathbb{E} - \mathbb{E}_n)\psi(\hat{f}) - 2(\mathbb{E} - \mathbb{E}_n)\psi(f^*) - \mathbb{E}\omega(\|\hat{f} - f^*\|_{\psi, f^*}) \\ &\leq 2\sup_{f \in \mathcal{F}} \left\{ (\mathbb{E} - \mathbb{E}_n)\psi(\hat{f}) - (\mathbb{E} - \mathbb{E}_n)\psi(f^*) - \frac{1}{2}\mathbb{E}\omega(\|f - f^*\|_{\psi, f^*}) \right\} \end{split}$$

Applying  $\mathbb{E}\Psi$  on both sides gives

$$\mathbb{E}\Psi(\mathcal{E}(\hat{f})) \leq \mathbb{E}\Psi\left(2\sup_{f\in\mathcal{F}}\left\{(\mathbb{E}-\mathbb{E}_n)\psi(\hat{f}) - (\mathbb{E}-\mathbb{E}_n)\psi(f^*) - \frac{1}{2}\mathbb{E}\omega(\|f-f^*\|_{\psi,f^*})\right\}\right)$$
$$= \mathbb{E}\Psi\left(2\sup_{f\in\mathcal{F}}\left\{(\mathbb{E}-\mathbb{E}_n)(\psi(\hat{f}) - \psi(f^*)) - \frac{1}{4}(\mathbb{E}+\mathbb{E})\omega(\|f-f^*\|_{\psi,f^*})\right\}\right)$$

By Jensen's inequality, this is

1

$$\leq \mathbb{E}\Psi\left(2\sup_{f\in\mathcal{F}}\left\{(\mathbb{E}-\mathbb{E}_n)(\psi(\hat{f})-\psi(f^*))-\frac{1}{4}(1+\mathbb{E})\omega(\|f-f^*\|_{\psi,f^*})\right\}\right).$$

The proof is then complete after applying Lemma G1 with  $A(f) = 2(\psi(f) - \psi(f^*)), T = \mathcal{F}$ , and  $B(f) = \frac{1}{2}\omega(\|f - f^*\|_{\psi, f^*}).$ П

### E Proof of Theorem 19

We recall Theorem 19.

**Theorem 19.** Let  $\psi$  be an  $\eta$ -exp-concave loss taking values in [0, m]. Then, with probability at least  $1 - 9e^{-z}$ , the star estimator  $\tilde{f}$  applied to  $(\psi, \mathcal{F})$  satisfies

$$\mathcal{E}(\tilde{f}) \le \inf_{0 \le \alpha \le \gamma} \left\{ 4\alpha + \frac{10}{\sqrt{n}} \int_{\alpha}^{\gamma} \sqrt{H_2(s)} \, ds + \frac{2H_2(\gamma)}{cn} + \frac{\gamma\sqrt{8\pi}}{\sqrt{n}} + \left(\frac{2}{cn} + \frac{\gamma\sqrt{8}}{\sqrt{n}}\right) z \right\}, \quad (23)$$

where  $H_2(s) \doteq H_2(s, \psi \circ \mathcal{F}')$  and  $c = 36^{-1}(1/m \wedge \eta/2)$ .

Proof. We work forwards from (15), which tells us

$$\mathbb{E}\Psi(\mathcal{E}(\tilde{f})) \le \mathbb{E}\Psi\left(\sup_{t\in T} \left\{Z_t - cZ_t^2\right\}\right),\,$$

where we define Z, t, T, and c according to

$$Z_t = \frac{1}{n} \sum_{i=1}^n 4\varepsilon_i' t(X_i, Y_i) = \frac{1}{n} \sum_{i=1}^n 4\varepsilon_i' (\psi(f(X_i), Y_i) - \psi(g(X_i), Y_i)),$$
  

$$t = \psi(f(X_i), Y_i) - \psi(g(X_i), Y_i),$$
  

$$T = \psi \circ \mathcal{F}' - \psi \circ \mathcal{F}',$$
  

$$c = \frac{1}{36} \left(\frac{1}{m} \lor \frac{\eta}{2}\right).$$

Let V be a covering of T at resolution  $\gamma$  in  $L^2(\mathbb{P}_n)$  that is chosen to include 0, so that  $\#V \leq \exp(2H_2(\gamma))$  almost surely by construction of T and definition of  $H_2(-)$ . Then we can choose  $\pi: T \to V$  with the properties that (1)  $\|t - \pi(t)\|_{2,\mathbb{P}} \leq \gamma$  uniformly over  $t \in T$ , and (2)  $\pi(t) = 0$  if  $\|t\|_{2,\mathbb{P}} < \gamma$ .

The proof will proceed in three lemmas which will be stated below and proved subsequently. The first lemma shows that  $\sup_{t \in T} \{Z_t - cZ_t^2\}$  can be controlled in terms of (i) the local complexity of  $(Z_t)_{t \in T}$  at scale  $\gamma$  and (ii) the offset complexity of a finite approximation to  $(Z_t)_{t \in T}$  at resolution  $\gamma$ . The second and third lemmas develop high-probability bounds for these two terms.

Lemma E1 (from Liang et al. (2015, Lemma 6)). It holds almost surely that

$$\sup_{t \in T} \left\{ Z_t - cZ_t^2 \right\} \le \sup_{t \in T} \left\{ Z_t - Z_{\pi(t)} \right\} + \sup_{v \in V} \left\{ Z_v - (c/4)Z_v^2 \right\}.$$
(27)

Lemma E2.

$$\mathbb{P}\left(\sup_{t\in T}\left\{Z_t - Z_{\pi(t)}\right\} \ge 4\alpha + \frac{10}{\sqrt{n}}\int_{\alpha}^{\gamma}\sqrt{2H_2(s)}\,ds + \gamma\sqrt{\frac{8\pi}{n}} + x\right) \le 2e^{-nx^2/(8\gamma^2)}$$
(28)

Lemma E3.

$$\mathbb{P}\left(\sup_{v\in V}\left\{Z_v - (c/4)Z_v^2\right\} > \frac{4H_2(\gamma) + 2x}{cn}\right) \le e^{-x}.$$
(29)

Applying a union bound to the event in (28) with  $x = z\gamma\sqrt{8}$ , the event in (29) with z = x, and the complement of the event (27), we obtain that with probability at least  $1 - 3e^{-z}$ 

$$\sup_{t\in T} \left\{ Z_t - Z_{\pi(t)} \right\} \le 4\alpha + \frac{10}{\sqrt{n}} \int_{\alpha}^{\gamma} \sqrt{H_2(s)} \, ds + \frac{2H_2(\gamma)}{cn} + \frac{\gamma\sqrt{8\pi}}{\sqrt{n}} + \left(\frac{2}{cn} + \frac{\gamma\sqrt{8}}{\sqrt{n}}\right) z$$

Finally, since  $\mathbb{E}\Psi(\mathcal{E}(\tilde{f})) \leq \mathbb{E}\Psi\left(\sup_{t \in T} \{Z_t - cZ_t^2\}\right)$  for all convex and increasing  $\Psi$ , we can apply the following.

**Lemma E4** (Panchenko (2003, Lemma 1)). If  $\mathbb{E}\Psi(X) \leq \mathbb{E}\Psi(Y)$  for all convex and increasing functions  $\Psi$ , then  $\mathbb{P}(Y \geq t) \leq Ae^{-at} \implies \mathbb{P}(X \geq t) \leq Ae^{1-at}.$  Thus, we have

$$\mathcal{E}(\tilde{f}) \le 4\alpha + \frac{10}{\sqrt{n}} \int_{\alpha}^{\gamma} \sqrt{H_2(s)} \, ds + \frac{2H_2(\gamma)}{cn} + \frac{\gamma\sqrt{8\pi}}{\sqrt{n}} + \left(\frac{2}{cn} + \frac{\gamma\sqrt{8}}{\sqrt{n}}\right) z$$

with probability at least  $1 - (3e)e^{-z}$ . After noting that  $0 \le \alpha \le \gamma$  is arbitrary and  $3e \le 9$ , the proof is complete.

Proof of Lemma E1. We have

$$\sup_{t \in T} \left\{ Z_t - cZ_t^2 \right\} = \sup_{t \in T} \left\{ (Z_t - Z_{\pi(t)}) + ((c/4)Z_{\pi(t)}^2 - cZ_t^2) - \left( cZ_t^2 + Z_{\pi(t)} - (c/4)Z_{\pi(t)}^2 \right) \right\}$$
$$\leq \sup_{t \in T} \left\{ Z_t - Z_{\pi(t)} \right\} + \sup_{v \in V} \left\{ Z_v - (c/4)Z_v^2 \right\},$$

provided we can show that the middle term  $(c/4)Z_{\pi(t)}^2 - cZ_t^2$  is a.s. non-positive. To see this, note that either  $||t||_{2,\mathbb{P}_n} < \gamma$ , in which case by construction  $\pi(t) = 0$  and  $Z_{\pi(t)}^2 = 0$ , so we are done, or else we have

$$\|\pi(t)\|_{2,\mathbb{P}_n} \le \|\pi(t) - t\|_{2,\mathbb{P}_n} + \|t\|_{2,\mathbb{P}_n} \le \|t\|_{2,\mathbb{P}_n} + \gamma \le 2 \|t\|_{2,\mathbb{P}_n},$$

so that  $\|\pi(t)\|_{2,\mathbb{P}_n}^2 \leq 4 \|t\|_{2,\mathbb{P}_n}^2$ . But, after plugging in the definition of  $Z_t$ , the middle term is precisely

$$\frac{16c}{n} \left( \frac{\|\pi(t)\|_{2,\mathbb{P}_n}^2}{4} - \|t\|_{2,\mathbb{P}_n}^2 \right),\,$$

so we are done.

*Proof of Lemma E2.* Keeping in mind that  $||t - \pi(t)||_{2,\mathbb{P}_n} \leq \gamma$  and applying the chaining result in Srebro et al. (2010, Lemma A.3) gives us

$$\mathbb{E}_{\varepsilon} \sup_{t \in T} \left\{ Z_t - Z_{\pi(t)} \right\} \le 4\alpha + \frac{10}{\sqrt{n}} \int_{\alpha}^{\gamma} \sqrt{2H_2(s)} \, ds \tag{30}$$

almost surely with respect to the data, where we used that

$$\ln N(s, T, L^2(\mathbb{P}_n)) \le 2\ln N(s, \psi \circ \mathcal{F}', L^2(\mathbb{P}_n)) \le 2H_2(s)$$

by definition of T and the fact that  $H_2(-)$  is an almost-sure bound on the logarithm of the  $L^2(\mathbb{P}_n)$  covering numbers. It follows by applying Ledoux and Talagrand (1991, Theorem 4.7) with  $\sigma^2(X) = \gamma^2/n$  that

$$\mathbb{P}_{\epsilon}\left(\sup_{t\in T}\left\{Z_t - Z_{\pi(t)}\right\} \ge M_{\epsilon} + x\right) \le 2e^{-nx^2/(8\gamma^2)},\tag{31}$$

where  $\mathbb{P}_{\varepsilon}$  denotes the probability with respect to the multipliers  $\varepsilon$  conditional upon the data and  $M_{\epsilon}$  is a conditional median of  $\sup_{t \in T} \{Z_t - Z_{\pi(t)}\}$ . Finally, we can deduce the upper bound

$$\mathbb{E}_{\varepsilon} \sup_{t \in T} \left\{ Z_t - Z_{\pi(t)} \right\} - M_{\epsilon} \\
\leq \mathbb{E}_{\epsilon} \left[ \left( \sup_{t \in T} \left\{ Z_t - Z_{\pi(t)} \right\} - M_{\epsilon} \right) \mathbb{1} \left\{ \sup_{t \in T} \left\{ Z_t - Z_{\pi(t)} \right\} > M_{\epsilon} \right\} \right] \\
= \int_0^\infty \mathbb{P}_{\epsilon} \left( \sup_{t \in T} \left\{ Z_t - Z_{\pi(t)} \right\} - M_{\epsilon} > t \right) dt \\
\leq \int_0^\infty 2e^{-nt^2/(8\gamma^2)} dt = \gamma \sqrt{\frac{8\pi}{n}}$$
(32)

Finally, putting together (30), (31) and (32) gives us

$$\mathbb{P}_{\epsilon}\left(\sup_{t\in T}\left\{Z_t - Z_{\pi(t)}\right\} \ge 4\alpha + \frac{10}{\sqrt{n}}\int_{\alpha}^{\gamma}\sqrt{2H_2(s)}\,ds + \gamma\sqrt{\frac{8\pi}{n}} + x\right) \le 2e^{-nx^2/(8\gamma^2)}.$$

Since this conditional bound holds almost surely with respect to the data, we immediately deduce (28).  $\hfill \square$ 

*Proof of Lemma E3.* Working conditionally upon the data, we can compute by applying Markov's inequality that

$$\mathbb{P}_{\varepsilon}\left(\sup_{v\in V}\left\{Z_{v}-(c/4)Z_{v}^{2}\right\}>t\right)=\mathbb{P}_{\varepsilon}\left(\exp\left(rn\sup_{v\in V}\left\{Z_{v}-(c/4)Z_{v}^{2}\right\}\right)>e^{rnt}\right)$$
$$\leq \mathbb{E}_{\varepsilon}\exp\left(rn\sup_{v\in V}\left\{Z_{v}-(c/4)Z_{v}^{2}\right\}\right)e^{-rnt}.$$

We can further compute that

$$\mathbb{E}_{\varepsilon} \exp\left(rn \sup_{v \in V} \left\{ Z_v - (c/4) Z_v^2 \right\} \right) = \mathbb{E}_{\varepsilon} \sup_{v \in V} \exp\left(\sum_{i=1}^n r\varepsilon_i' v_i - (c/4) r v_i^2 \right)$$
$$\leq \sum_{v \in V} \mathbb{E}_{\varepsilon} \exp\left(\sum_{i=1}^n \frac{r^2 v_i^2}{2} - \frac{cr v_i^2}{4}\right),$$

by applying Hoeffding's lemma to each expectation with respect to the variables  $\varepsilon_i$ . Taking r = c/2, this is precisely #V. Thus, we have that

$$\mathbb{P}_{\varepsilon}\left(\sup_{v\in V}\left\{Z_{v}-(c/4)Z_{v}^{2}\right\}>t\right)\leq \exp\left(\ln(\#V)-\frac{cnt}{2}\right)$$

Since  $\ln(\#V) \leq 2H_2(\gamma)$  almost surely, we can deduce the unconditional bound

$$\mathbb{P}\left(\sup_{v\in V}\left\{Z_v - (c/4)Z_v^2\right\} > t\right) \le \exp\left(2H_2(\gamma) - \frac{cnt}{2}\right)$$

Taking  $t = (4H_2(\gamma) + 2x)/cn$  gives (29).

## F Proofs of Section 5 Results

#### F.1 Proof of Corollary 16

We state Corollary 16.

**Corollary 16** (cf. Mendelson (2002, Theorem 5.1)). Let  $\psi(f, y) = |f - y|^p$  for p > 1 and let the class  $\mathcal{F}$  and response Y take values in [-B, B]. Then there exists a universal  $C(p, B) = O(p2^pB^p)$  such that the star estimator  $\tilde{f}$  has excess  $\psi$ -risk bounded as

$$\mathbb{P}\left(\mathcal{E}(\tilde{f},\mathcal{F}) \ge \epsilon + \frac{C(p,B)(H_2(\epsilon/C_{p,B},\mathcal{F}) + \ln(1/\epsilon) + \ln(1/\rho))}{n}\right) \le \rho,\tag{20}$$

Proof. This follows as a result of the more general bound (19), which says that

$$\mathbb{P}\left(\mathcal{E}(\tilde{f}) \ge \epsilon + \left(36m \vee \frac{72}{\eta}\right) \left(\frac{H_2(\epsilon, \psi \circ \mathcal{F}') + \ln(1/\rho)}{n}\right)\right) \le \rho.$$

In order to deduce (20), we need to bound the quantities m,  $1/\eta$ , and  $H_2(\psi \circ \mathcal{F}')$ . For m, since  $|f|, |y| \leq B$ , it must hold that  $|f - y|^p \leq 2^p B^p$ . For  $\eta$ , we can compute that

$$(\psi')^2/\psi'' = \frac{p^2 z^{2p-2}}{p(p-1)z^{p-2}} \le \frac{pz^p}{p-1} \le \frac{p2^p B^p}{p-1}$$

for  $z = |f - y| \le 2B$ . Finally, we have  $\|\psi\|_{\text{lip}} \le p2^p B^{p-1}$  by bounding the first derivative, so that we have the entropy estimates

$$H_2(\epsilon, \psi \circ \mathcal{F}') \le H_2\left(\frac{\epsilon}{p2^p B^{p-1}}, \mathcal{F}'\right) \le 2H_2\left(\frac{\epsilon}{p2^{p+1} B^{p-1}}, \mathcal{F}\right) + \ln\left(\frac{4B}{\epsilon}\right),$$

where the last step follows by applying Lemma G4 with  $R \leq \sup_{f,y} |f - y| \leq 2B$ .

#### F.2 Proof of Lemma 17

We recall Lemma 17.

**Lemma 17** (Foster et al. (2018)). For all f and  $\delta \in (0, 1/2]$ , the excess risk relative to  $\mathcal{L}_{\delta}$  satisfies  $\mathcal{E}(f; \mathcal{L}_{\delta}) \leq \mathcal{E}(f; \mathcal{L}) + 2\delta$ 

*Proof.* We can compute that

$$\ln(f) - \ln((1-\delta)f + \delta) = \ln\left(\frac{f}{(1-\delta)f + \delta}\right) \le \ln\left(\frac{1}{1-\delta}\right) \le 2\delta,$$
(33)

since  $0 \le -\ln(1-\delta) \le 2\delta$  for  $0 \le \delta \le 1/2$ . Consequently, for any g,

$$\begin{split} \mathcal{E}(g;\mathcal{L}_{\delta}) &= \mathbb{E}[-\ln g] - \inf_{f \in \mathcal{L}_{\delta}} \mathbb{E}[-\ln f] \\ &= \mathbb{E}[-\ln g] - \inf_{f \in \mathcal{L}} \mathbb{E}[-\ln f] + \inf_{f \in \mathcal{L}} \mathbb{E}[-\ln f] - \inf_{f \in \mathcal{L}_{\delta}} \mathbb{E}[-\ln f] \\ &= \mathcal{E}(g;\mathcal{L}) + \left(\inf_{f \in \mathcal{L}} \mathbb{E}[-\ln f] - \inf_{f \in \mathcal{L}_{\delta}} \mathbb{E}[-\ln f]\right) \end{split}$$

By separability of the two infima, this is the same as

$$= \mathcal{E}(g; \mathcal{L}) + \sup_{h \in \mathcal{L}_{\delta}} \inf_{f \in \mathcal{L}} \mathbb{E}[\ln h - \ln f].$$

By choosing  $h = (1 - \delta)f + \delta$ , the outer supremum may be bounded as

$$\geq \mathcal{E}(g; \mathcal{L}) + \inf_{f \in \mathcal{L}} \mathbb{E}[\ln((1-\delta)f + \delta) - \ln f]$$
  
$$\geq \mathcal{E}(g; \mathcal{L}) - 2\delta,$$

where the final step follows from negating (33).

#### F.3 Proof of Corollary 18

We recall Corollary 18.

**Corollary 18.** With probability at least 
$$1 - \rho$$
, the star estimator  $f_{\delta}$  in  $\mathcal{L}_{\delta}$  satisfies

$$\mathcal{E}(\tilde{f}_{\delta};\mathcal{L}) \le \epsilon + 2\delta + C\ln(1/\delta) \left(\frac{H_2(\delta\epsilon,\mathcal{L}) + \ln(1/\epsilon\delta) + \ln(1/\rho)}{n}\right)$$
(21)

Let  $\mathcal{L}$  be the generalized linear model corresponding to

$$\mathcal{F}_{B} = \left\{ x \mapsto Wx \, \big| \, W \in \mathbb{R}^{k \times q}, \left\| W \right\|_{2 \to \infty} \le B \right\}$$

with A-Lipschitz, surjective link  $\varphi$  and features  $X \in \mathbb{R}^q$  that satisfy  $||X||_2 \leq R\sqrt{q}$ . Then with probability at least  $1 - \rho$ 

$$\mathcal{E}(\varphi^{\dagger} \circ \tilde{f}_{\delta}; \mathcal{F}_B) \leq \frac{\ln(n)}{n} \left\{ Ckq \ln(ABRn\sqrt{k}) + \ln(1/\rho) \right\}.$$
(22)

Proof. By Lemma 17, it suffices for (21) to show instead that

$$\mathbb{P}\left(\mathcal{E}(\tilde{f}_{\delta};\mathcal{L}_{\delta}) > \epsilon + C\ln(1/\delta)\left(\frac{H_2(\delta\epsilon,\mathcal{L}) + \ln(1/\epsilon\delta) + \ln(1/\rho)}{n}\right)\right) \le \rho$$

This in turn follows from the general inequality (19), which says in this context that

$$\mathbb{P}\left(\mathcal{E}(\tilde{f}_{\delta},\mathcal{L}_{\delta}) \geq \epsilon + \left(36m \vee \frac{72}{\eta}\right) \left(\frac{H_2(\epsilon, -\ln\circ\mathcal{L}_{\delta}') + \ln(1/\rho)}{n}\right)\right) \leq \rho$$

Since  $\mathcal{L}'_{\delta}$  takes values in  $[\delta, 1]$ , the log loss takes values in  $[0, \ln(1/\delta)]$ , so we choose  $m = \ln(1/\delta)$ . Since the log loss is 1-exp-concave, we choose  $\eta = 1$ . Finally, the log loss in this domain is  $(1/\delta)$ -Lipschitz, so we have the estimates

$$H_2(\epsilon, -\ln\circ\mathcal{L}'_{\delta}) \le H_2(\delta\epsilon, \mathcal{L}'_{\delta}) \le 2H_2(\delta\epsilon/2, \mathcal{L}_{\delta}) + \ln(2\ln(1/\delta)/\delta\epsilon),$$

where the last step follows from Lemma G4 with  $R = \ln(1/\delta)$ . Finally, use that  $H_2(-, \mathcal{L}_{\delta}) \leq H_2(-, \mathcal{L})$  since  $\mathcal{L}_{\delta}$  is the image of  $\mathcal{L}$  under a pointwise contraction, and simplify (for example absorbing  $\ln \ln(1/\delta) \leq \ln(1/\delta)$ ) into the constant C). For (18), we plug in the covering estimates

$$H_2(\delta\epsilon, \mathcal{L}) \le H_2\left(\frac{\delta\epsilon}{A}, \mathcal{F}_B\right) \le \ln\left(\frac{ABR\sqrt{k}}{\epsilon\delta}\right)^{kd} = kd\ln\left(\frac{ABR\sqrt{k}}{\delta\epsilon}\right)$$

for  $\mathcal{F}_B$ , which are standard. Finally, we take  $\epsilon = \delta = 1/n$  and simplify.

#### F.4 Proof of Corollary 20

We recall Corollary 20.

**Corollary 20.** Consider a generalized linear model with A-Lipschitz loss  $f \mapsto -\ln \langle \varphi(f), y \rangle$ . Suppose the entropy numbers  $H_2(\epsilon; \mathcal{F})$  are of order  $\epsilon^{-q}$ . Then, the regularized star estimator  $\varphi^{\dagger} \circ \tilde{f}_{\delta}$  with  $\delta = 1/n$  satisfies the rates appearing on the left.

On the other hand, for an arbitrary class  $\mathcal{L}$  taking values in [0,1] subject to the log loss, the regularized star estimator  $\tilde{f}'_{\delta'}$ —for appropriately chosen  $\delta'$ —attains the rates appearing on the right. Here, the symbol  $\leq_{\rho}$  denotes an upper bound that holds with probability  $1 - \rho$ , hiding universal constants and a multiplicative factor  $\ln(1/\rho)$ .

$$\mathcal{E}(\varphi^{\dagger} \circ \tilde{f}_{\delta}) \lesssim_{\rho} \begin{cases} A^{q} n^{-2/(2+q)} \ln(n) & q < 2\\ A^{q} n^{-1/2} \ln(n) & q = 2\\ A^{q} n^{-1/q} & q > 2 \end{cases} \qquad \mathcal{E}(\tilde{f}_{\delta'}) \lesssim_{\rho} \begin{cases} n^{-1/(1+3q/2)} & q < 2\\ n^{-1/4} \ln(n) & q = 2\\ n^{-1/(2q)} & q > 2 \end{cases}$$
(24)

*Proof.* These bounds are all derived by applying Theorem 19 under different assumptions on the entropy function. In particular combining (23) with Lemma 17—using the fact that the log loss over  $\mathcal{L}_{\delta}$  takes values in  $[0, \ln(1/\delta)]$  and is 1-exp-concave—gives us that with probability  $1 - \rho$ ,

$$\mathcal{E}(\tilde{f}) \lesssim_{\rho} 2\delta + \inf_{0 \le \alpha \le \gamma} \left\{ 4\alpha + \frac{10}{\sqrt{n}} \int_{\alpha}^{\gamma} \sqrt{H_2(s)} \, ds + \frac{H_2(\gamma) \ln(1/\delta)}{n} + \frac{\gamma}{\sqrt{n}} \right\},$$

where the symbol  $\leq_{\rho}$  hides universal constants and a multiplicative factor  $\ln(1/\rho)$ .

For the left-hand side results, the entropy numbers scale as  $(A/\epsilon)^q$ ; choosing  $\delta = \frac{1}{n}$ , we get

$$\frac{2}{n} + 4\alpha + \frac{10A^{q/2}}{\sqrt{n}} \int_{\alpha}^{\gamma} s^{-q/2} \, ds + \frac{\gamma^{-q}A^q \ln n}{n} + \frac{\gamma}{\sqrt{n}}$$

For the q < 2 case we take  $\alpha = 0$  and  $\gamma = n^{-1/(2+q)}$ . For the case q = 2 we take  $\alpha = 1/n$  and  $\gamma = 1$ . For the case q > 2 we take  $\alpha = n^{-1/q}$  and  $\gamma = 1$ . For the right-hand side results, the entropy numbers scale as  $(1/\delta\epsilon)^q$ , giving us the bound

$$2\delta + 4\alpha + \frac{12\delta^{-q/2}}{\sqrt{n}} \int_{\alpha}^{\gamma} s^{-q/2} \, ds + \frac{\gamma^{-q}\delta^{-q}}{n} + \frac{\gamma}{\sqrt{n}}.$$

For q < 2, we take  $\alpha = 0$ ,  $\delta = n^{-1/(1+3q/2)}$ , and  $\gamma = n^{-1/(2+3q)}$ . For q = 2 we take  $\delta = n^{-1/4}$ ,  $\alpha = 1/n$ , and  $\gamma = 1$ . For q > 2 we take  $\delta = \alpha = n^{-1/2p}$  and  $\gamma = 1$ .

#### **G** Technical Lemmas

**Lemma G1** (Offset symmetrization). For every increasing and convex function  $\Psi$ ,

$$\mathbb{E}\Psi\left(\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\left(1-\mathbb{E}\right)A_{i}(t)-(1+\mathbb{E})B_{i}(t)\right\}\right)\leq\mathbb{E}\Psi\left(2\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}^{\prime}A_{i}(t)-B_{i}(t)\right\}\right).$$

*Proof.* Noting that  $\mathbb{E}A_i(t) = \mathbb{E}A'_i(t)$  and  $\mathbb{E}B_i(t) = \mathbb{E}B'_i(t)$ , where  $A'_i$  (respectively,  $B_i$ ) is an independent copy of  $A_i$  (resp.  $B'_i$ ), and finally moving the expectations outside by applying Jensen's inequality to the convex function  $\Psi(\sup_{t\in T}(-))$ , we have

$$\mathbb{E}\Psi\left(\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\left(1-\mathbb{E}\right)A_{i}(t)-(1+\mathbb{E})B_{i}(t)\right\}\right)$$
$$\leq \mathbb{E}\Psi\left(\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}A_{i}(t)-A_{i}'(t)-B_{i}(t)-B_{i}'(t)\right\}\right).$$

Since  $A_i - A'_i$  is symmetric, it is equal in distribution to  $\varepsilon'_i(A_i - A'_i)$ , where  $\varepsilon'_i$  is a symmetric Rademacher r.v. independent of  $(A, \hat{A}', B, B')$ , hence we can write

$$= \mathbb{E}\Psi\left(\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}'(A_{i}(t)-A_{i}'(t))-B_{i}(t)-B_{i}'(t)\right\}\right)$$
$$= \mathbb{E}\Psi\left(\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}'A_{i}(t)-B_{i}(t)+\frac{1}{n}\sum_{i=1}^{n}(-\varepsilon_{i}')A_{i}'(t)-B_{i}'(t)\right\}\right).$$
$$= \mathbb{E}\Psi\left(\sup_{t\in T}\left\{2\mathbb{E}_{\sigma}\left[\frac{\sigma}{n}\sum_{i=1}^{n}\varepsilon_{i}'A_{i}(t)-B_{i}(t)+\frac{1-\sigma}{n}\sum_{i=1}^{n}(-\varepsilon_{i}')A_{i}'(t)-B_{i}'(t)\right]\right\}\right),$$

where  $\sigma$  is an independent symmetric Bernoulli r.v. By a final application of Jensen's inequality and equality of the distributions of  $(\sigma \varepsilon_i')_{i=1}^n$  and  $((1 - \sigma)(-\varepsilon_i'))_{i=1}^n$ , this is

$$\leq \mathbb{E}\Psi\left(2\sup_{t\in T}\left\{\frac{\sigma}{n}\sum_{i=1}^{n}\varepsilon_{i}'A_{i}(t)-B_{i}(t)+\frac{1-\sigma}{n}\sum_{i=1}^{n}(-\varepsilon_{i}')A_{i}'(t)-B_{i}'(t)\right\}\right),\$$

$$=\mathbb{E}\Psi\left(2\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}'A_{i}(t)-B_{i}(t)\right\}\right),\tag{34}$$
so what we aimed to show.

which is what we aimed to show.

**Lemma G2** (Offset contraction). Suppose that  $|A_i(s) - A_i(t)| \le |C_i(s) - C_i(t)|$  for all  $s, t \in T$ . *Then, for all increasing and convex*  $\Psi$ *, we have* 

$$\mathbb{E}\Psi\left(2\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}^{\prime}A_{i}(t)-B_{i}(t)\right\}\right)\leq\mathbb{E}\Psi\left(2\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}^{\prime}C_{i}(t)-B_{i}(t)\right\}\right),$$
(35)

whenever the  $\varepsilon'_i$  are symmetric Rademacher variables that are independent of A, B and C.

*Proof.* To simplify notation, put

$$S_m(t) = \sum_{i=1}^m \varepsilon_i' A_i(t) - B_i(t).$$

Writing out the expectation with respect to  $\varepsilon'_n$  gives

$$\mathbb{E}\Psi\left(2\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}'A_{i}(t)-B_{i}(t)\right\}\right)$$
  
=  $\mathbb{E}\Psi\left(\frac{1}{n}\sum_{j=1}^{2}\sup_{t\in T}\left\{(-1)^{j}A_{n}(t)+S_{n-1}(t)-B_{n}(t)\right\}\right)$   
=  $\mathbb{E}\Psi\left(\frac{1}{n}\sup_{s,t\in T}\left\{A_{n}(s)-A_{n}(t)+(S_{n-1}(s)-B_{n}(s))-(S_{n-1}(t)-B_{n}(t))\right\}\right)$ 

Applying our assumption that  $|A_i(s) - A_i(t)| \le |C_i(s) - C_i(t)|$ , this is

$$\leq \mathbb{E}\Psi\left(\frac{1}{n}\sup_{s,t\in T}\left\{|C_n(s) - C_n(t)| + (S_{n-1}(s) - B_n(s)) - (S_{n-1}(t) - B_n(t))\right\}\right)$$

Since the argument of the supremum is symmetric in (s, t), we can remove the absolute value, yielding

$$\leq \mathbb{E}\Psi\left(\frac{1}{n}\sup_{s,t\in T}\left\{C_n(s) - C_n(t) + (S_{n-1}(s) - B_n(s)) + (S_{n-1}(t) - B_n(t))\right\}\right).$$

Since the supremum is now separable in (s, t), we further have

$$= \mathbb{E}\Psi\left(\frac{1}{n}\sum_{j=1}^{2}\sup_{t\in T}\left\{(-1)^{j}C_{n}(t) - B_{n}(t) + S_{n-1}(t)\right\}\right)$$
$$= \mathbb{E}\Psi\left(\frac{2}{n}\sup_{t\in T}\left\{\varepsilon_{n}C_{n}(t) - B_{n}(t) + S_{n-1}(t)\right\}\right).$$

Applying these manipulations to each summand r from n - 1 down to 1 gives us

$$\leq \mathbb{E}\Psi\left(\sup_{t\in T}\left\{\frac{2}{n}\sum_{i=r}^{n}\varepsilon_{i}C_{i}(t)-B_{i}(t)+S_{r-1}(t)\right\}\right)$$
$$\leq \mathbb{E}\Psi\left(2\sup_{t\in T}\left\{\frac{1}{n}\sum_{i=1}^{n}\varepsilon_{i}C_{i}(t)-B_{i}(t)\right\}\right),$$

which is what we aimed to show.

**Lemma G3** (Log margin computation). For  $|z| \leq c$ ,

$$e^{-z} + z - 1 \ge \frac{z^2}{2c \lor 4}.$$

*Proof.* Note that  $z - 1 \ge z/2 \ge z^2/(2c)$  for  $z \ge 2$ . So it suffices to check the inequality for z < 2. On the other hand, one can check by minimizing the left-hand side that

$$\frac{e^{-z}+z-1}{z^2} \geq \frac{1}{4}$$

for 0 < z < 2 (the derivative of the left-hand side is negative, and the inequality holds at z = 2). Finally, the inequality for  $z \le 0$  follows by noting that

$$e^{-z} - 1 + z = \frac{z^2}{2} + \sum_{k=3}^{\infty} \frac{(-z)^k}{k!},$$

by the series expansion for  $e^{-z}$  and the remainder term must be non-negative for  $z \leq 0$ .

**Lemma G4** (cf. Mendelson (2002, Lemma 4.5)). Put  $\mathcal{F}' = \bigcup_{\lambda \in [0,1]} \lambda \mathcal{F} + (1-\lambda) \mathcal{F}$  and  $R_{\mu} = \sup_{f \in \mathcal{F}} \|f\|_{L^{2}(\mu)}$ . Let  $N_{2}(\epsilon, S, \mu)$  denote the  $\epsilon$ -covering number of the set S in  $L^{2}(\mu)$ . Then

$$N_2(\epsilon, \mathcal{F}', \mu) \le \left(\frac{2R_\mu}{\epsilon}\right) N_2(\epsilon/2, \mathcal{F}, \mu)^2.$$
(36)

Consequently, if  $R = \sup_{\mu} R_{\mu}$  where the supremum is over probability measures,

$$H_2(\epsilon, \mathcal{F}') \le 2H_2(\epsilon/2, \mathcal{F}) + \ln\left(\frac{2R}{\epsilon}\right).$$
 (37)

*Proof.* Let S denote a minimal covering of  $\mathcal{F}$  in  $L^2(\mu)$  at resolution  $\epsilon/2$ . Given some  $(s,t) \in S^2$ , let T(s,t) denote an  $\epsilon/2$  covering of the line segment interpolating s and t. This line segment has length at most  $2R_{\mu}$  in  $L^2(\mu)$ , hence  $\#T(s,t) \leq \frac{2R_{\mu}}{\epsilon}$ . We are therefore done if we can show that

$$\bigcup_{(s,t)\in S^2} T(s,t)$$

is an  $\epsilon$  covering of  $\mathcal{F}'$ .

To this end, let  $f \in \mathcal{F}'$  be given. By definition, we may write  $f = \lambda f_1 + (1 - \lambda)f_2$  for  $f_1, f_2 \in \mathcal{F}$ , and we can choose  $s_1, s_2 \in S$  such that

$$||s_1 - f_1||_{L^2(\mu)}, ||s_2 - f_2||_{L^2(\mu)} \le \epsilon/2.$$

Due to convexity of the norm, we must have that

$$\|(\lambda s_1 + (1 - \lambda)s_2) - f\|_{L^2(\mu)} \le \epsilon/2.$$

By construction, there exists some  $h \in T(s_1, s_2)$  such that

$$\|(\lambda s_1 + (1 - \lambda)s_2) - h\|_{L^2(\mu)} \le \epsilon/2.$$

Using the triangle inequality, we deduce  $||f - h||_{L^2(\mu)} \le \epsilon$  and the proof of (36) is complete; (37) then follows by first taking logarithms, then taking the supremum over probability measures  $\mu$ .  $\Box$ 

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