Nearly Horizon-Free Offline Reinforcement Learning

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Abstract

We revisit offline reinforcement learning on episodic time-homogeneous Markov Decision Processes (MDP). For tabular MDP with \( S \) states and \( A \) actions, or linear MDP with anchor points and feature dimension \( d \), given the collected \( K \) episodes data with minimum visiting probability of (anchor) state-action pairs \( d_m \), we obtain nearly horizon \( H \)-free sample complexity bounds for offline reinforcement learning when the total reward is upper bounded by \( 1 \). Specifically:

- For offline policy evaluation, we obtain an \( \tilde{O} \left( \sqrt{\frac{1}{Kd_m}} \right) \) error bound for the plug-in estimator, which matches the lower bound up to logarithmic factors and does not have additional dependency on \( \text{poly}(H, S, A, d) \) in higher-order term.

- For offline policy optimization, we obtain an \( \tilde{O} \left( \sqrt{\frac{1}{Kd_m}} + \frac{\min(S, d)}{Kd_m} \right) \) sub-optimality gap for the empirical optimal policy, which approaches the lower bound up to logarithmic factors and a high-order term, improving upon the best known result by [1] that has additional \( \text{poly}(H, S, d) \) factors in the main term.

To the best of our knowledge, these are the first set of nearly horizon-free bounds for episodic time-homogeneous offline tabular MDP and linear MDP with anchor points. Central to our analysis is a simple yet effective recursion based method to bound a “total variance” term in the offline scenarios, which could be of individual interest.

1 Introduction

Reinforcement Learning (RL) aims to learn to make sequential decisions to maximize the long-term reward in unknown environments, and has demonstrated success in game-playing [2, 3], robotics [4], and automatic algorithm design [5]. These successes rely on being able to deploy the algorithms in a way that directly interacts with the respective environment, allowing them to improve the policies in a trial-and-error way. However, such direct interactions with real environments can be expensive or even impossible in other real-world applications, e.g., education [6], health and medicine [7, 8], conversational AI [9] and recommendation systems [10]. Instead, we are often given a collection of logged experiences generated by potentially multiple and possibly unknown policies in the past.

This lack of access to real-time interactions with an environment led to the field of offline reinforcement learning [11]. Within this, offline policy evaluation (OPE) focuses on evaluating a policy, and offline policy optimization (OPO) focuses on improving policies: both rely only upon the given fixed past experiences without any further interactions. OPE and OPO are in general notoriously difficult, as unbiased estimators of the policy value can suffer from exponentially increasing variance in terms of horizon in the worst case [12, 13].

To overcome this “curse of horizon” in OPE, [14, 15] first introduced marginalized importance sampling (MIS) based estimators. They showed that if (1) all of the logged experiences are generated

from the same behavior policy, and (2) the behavior policy is known, then the exponential dependency on horizon can be improved to polynomial dependency. Subsequently, [16–18] showed that the polynomial dependency could be achieved even without assumptions (1) and (2). The basic idea for these MIS-based estimators is estimating the marginal state-action density ratio between the target policy and the empirical data, so as to adjust the distribution mismatch between them. On the algorithmic side, marginal density ratio estimation can be implemented by either plug-in estimators [15, 19], temporal-difference updates [20, 21], or solving a min-max optimization [16, 17, 22, 18]. These OPE estimators can also be used as one component for OPO, resulting in the algorithms in [23–25], which also inherit the polynomial dependency on horizon.

In a different but related line of work, [26, 27] recently showed that, for the online episodic time-homogeneous tabular Markov Decision Process (MDP) that allows for the interactions with environments, the sample complexity only has poly log dependency on the horizon. This motivates us to consider the following question:

Can offline reinforcement learning escape from the polynomial dependency on the horizon?

In this paper, we provide an affirmative answer to this question. Specifically, considering the episodic time-homogeneous tabular MDP with \( S \) states, \( A \) actions and horizon \( H \), or the linear MDP with anchor points and feature dimension \( d \), assuming the total reward for any episode is upper bounded by 1 almost surely, we obtain the following nearly \( H \)-free bounds:

- For offline policy evaluation (OPE), we show that the plug-in estimator has a finite-sample error of \( \tilde{O} \left( \sqrt{\frac{1}{Kd_m}} \right) \) (Theorem 1 and 5), where \( K \) is the number of episodes and \( d_m \) is the minimum visiting probability of (anchor) state-action pairs, that matches the lower bound up to logarithmic factors. We emphasize that the bound has no additional poly \((H, S, A, d)\) dependency in the higher order term, unlike the known results of [19, 24].

- For offline policy optimization (OPO), we show that the policy obtained by model-based planning on empirical MDP has a sub-optimality gap of \( \tilde{O} \left( \sqrt{\frac{1}{Kd_m}} + \frac{\min(S,d)}{Kd_m} \right) \) (Theorem 3 and 6), which matches the lower bound up to logarithmic factors and a high-order term. This also improves upon the best known result from [1] by removing additional poly \((H, S, d)\) factors in the main term.

To the best of our knowledge, these are the first set of nearly horizon-free bounds for both OPE and OPO on time-homogeneous tabular MDP and linear MDP with anchor points. To achieve such sharp bounds, we propose a novel recursion based method to bound a “total variance” term (introduced below) that is broadly emerged in the offline reinforcement learning, which could be of individual interest.

Technique Overview. With a sequence of fairly standard steps in the literature, we can bound the error of the plug-in estimator via terms related to the square root of the “total variance” (also known as the Cramer-Rao type lower bound illustrated in [13]):

\[
\sqrt{\sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \text{Var}_{P(s,a)} (V_{h+1}^\pi(s'))}
\]

where \( \xi^\pi \) is the reaching probability, \( P \) is the transition and \( V^\pi \) is the value function under policy \( \pi \). (For a more formal definition, see Section 3.). An improper bound of this term will introduce unnecessary dependency on the horizon, either in the main term or in the higher-order term. We instead bound this term with a recursive method, by observing that this “total variance” term can be approximately upper bounded by the square root of the “total variance of value square”

\[
\sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \text{Var}_{P(s,a)} (V_{h+1}^\pi(s'))^2
\]

and the error of plug-in estimator (see Lemma 1 for the detail). Applying this argument recursively, we can finally obtain a poly \( \log H \) upper bound on the total variance, which eventually gets rid of the polynomial dependency on \( H \).

The idea of higher order expansion has been investigated in [28, 27]. We notice that the recursion introduced in [28] was designed for the infinite horizon setting, and how to generalize their technique to finite horizon setting is still unclear. Our recursion is conceptually more similar to the recursion in [27]. However, [27] considered the online setting, where we only need to bound the error on the visited state-action pairs. For the offline setting, we need to bound the error on every state-action pair that can be touched by the policy \( \pi \). This introduces the reaching probability \( \xi^\pi \) in the “total variance” term, which we need to deal with using the MDP structure. As a result, our recursion is significantly different from their counterpart, especially in the case of linear MDP.
**Organizer.** Our paper is organized as follows: in Section 2, we review the related literature on offline reinforcement learning, and then briefly introduce the problem we consider in Section 3. In Section 4 and Section 5, we show our results of offline policy evaluation and offline policy optimization on tabular MDP correspondingly, and in Section 6, we show how to generalize our results to linear MDP with anchor points. We finally conclude and discuss our results in Section 7.

2 Related Work

In this section, we briefly discuss the related literature in three categories, i.e., offline policy evaluation, offline policy optimization, and horizon-free online reinforcement learning. Notice that, for the setting that assumes an additional generative model, typical model-based algorithms first query equal number of data from each state-action pair, then perform offline policy evaluation/optimization based on the queried data. Thus we view the reinforcement learning with generative model as a special instance of offline reinforcement learning. To make the comparison fair, for method and analysis that do not assume Assumption 1, we scale the error and sample complexity, by assuming per-step reward is upper bounded by $1 - \gamma$ and $1/H$ under infinite-horizon and finite-horizon setting correspondingly.

Table 1: A comparison of existing offline policy evaluation results. The sample complexity in infinite horizon setting is the number of queries of transitions we need while in episodic setting is the number of episodes we need. If Non-Uniform Reward, the MDP we consider satisfies Assumption 1; otherwise, we assume the per-step reward is upper bounded by $1 - \gamma$ and $1/H$ correspondingly.

<table>
<thead>
<tr>
<th>Analysis</th>
<th>Setting</th>
<th>Non-Uniform Reward</th>
<th>Sample Complexity</th>
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<tbody>
<tr>
<td>[28]</td>
<td>Infinite Horizon Tabular</td>
<td>Yes</td>
<td>$\tilde{O}(\frac{1}{\min(1 - \gamma \epsilon)})$</td>
</tr>
<tr>
<td>[29]</td>
<td>Infinite Horizon Tabular</td>
<td>Yes</td>
<td>$\Omega(\frac{1}{\min(1 - \gamma \epsilon)})$</td>
</tr>
<tr>
<td>[19]</td>
<td>Finite Horizon time-inhomogeneous Tabular</td>
<td>No</td>
<td>$\tilde{O}(\frac{1}{\min(1 - \gamma \epsilon)} + \frac{\sqrt{SA}}{H \epsilon})$</td>
</tr>
<tr>
<td>[13]</td>
<td>Finite-Horizon time-inhomogeneous Tabular</td>
<td>No</td>
<td>$\Omega(\frac{1}{\min(1 - \gamma \epsilon)})$</td>
</tr>
<tr>
<td>This work</td>
<td>Finite Horizon time-homogeneous Tabular/Linear</td>
<td>Yes</td>
<td>$\tilde{O}(\frac{1}{\min(1 - \gamma \epsilon)})$</td>
</tr>
<tr>
<td>Lower Bound</td>
<td>Finite Horizon time-homogeneous Tabular/Linear</td>
<td>Yes</td>
<td>$\Omega(\frac{1}{\min(1 - \gamma \epsilon)})$</td>
</tr>
</tbody>
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**Offline Policy Evaluation.** For OPE in infinite horizon tabular MDP, [28] showed that plug-in estimator can achieve the error of $\tilde{O}(\frac{1}{\min(1 - \gamma \epsilon)})$ under Assumption 1, which matches the lower bound in [29] up to logarithmic factors. For OPE in finite horizon time-inhomogeneous tabular MDP, [19, 24] provided an error bound of $\tilde{O}(\sqrt{\frac{1}{H \epsilon m} + \frac{\sqrt{SA}}{H \epsilon m}})$ under the uniform reward assumption, which matches the lower bound [13] up to logarithmic factors and an additional higher-order term. We here consider the time-homogeneous MDP, and obtain an error bound of $\tilde{O}(\sqrt{\frac{1}{K \epsilon m}})$, that does not have the additional $\sqrt{SA}$ in higher-order term, which is different from [19, 24].

Beyond the tabular setting, [30] considered the performance of plug-in estimator with linear function approximation under the assumption of linear MDP without anchor points, and [31, 32] provided more detailed analyses on the statistical properties of different kinds of estimators under different assumptions, which are not directly comparable to our work. Recently, there are also works [e.g. 33] focusing on the interval estimation of the policy for practical application.

**Offline Policy Optimization.** Offline policy optimization for infinite horizon MDP can date back to [35]. [28] recently showed that a perturbed version of model-based planning can find $\epsilon$-optimal policy within $\tilde{O}(\frac{1}{\min(1 - \gamma \epsilon) \epsilon})$ queries of transitions in infinite horizon tabular MDP when the total reward is upper bounded by 1, that matches the lower bound up to logarithmic factors. For the finite horizon time-inhomogeneous tabular MDP setting, [24] showed that model-based planning can identify $\epsilon$-optimal policy with $\tilde{O}(\frac{H}{\min(1 - \gamma \epsilon) \epsilon})$ episodes, that matches the lower bound for time-inhomogeneous MDP up to logarithmic factors. When it comes to finite horizon time-homogeneous tabular MDP and linear MDP with anchor points, [1] provided a $\tilde{O}(\frac{\min(H, S, d)}{d \epsilon^2})$ episode complexity for model-based
Table 2: A comparison of existing offline learning results. The sample complexity in infinite horizon setting is the number of queries of transitions we need while in episodic setting is the number of episodes we need. If Non-Uniform Reward, the MDP we consider satisfies Assumption 1; otherwise, we assume the per-step reward is upper bounded by $1 - \gamma$ and $1/H$ correspondingly.

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<tr>
<td>[34]</td>
<td>Infinite Horizon Tabular</td>
<td>No</td>
<td>$\tilde{O}\left(\frac{1}{d_m \epsilon^2}\right)$</td>
</tr>
<tr>
<td>[28]</td>
<td>Infinite Horizon Tabular</td>
<td>Yes</td>
<td>$\tilde{O}\left(\frac{1}{d_m \epsilon^2}\right)$</td>
</tr>
<tr>
<td>[24]</td>
<td>Finite Horizon time-inhomogeneous Tabular</td>
<td>No</td>
<td>$\tilde{O}\left(\frac{H}{d_m \epsilon^2}\right)$</td>
</tr>
<tr>
<td>[1]</td>
<td>Finite Horizon time-homogeneous Linear</td>
<td>No</td>
<td>$\tilde{O}\left(\min(H, S, d)\right)$</td>
</tr>
<tr>
<td>[27]</td>
<td>Finite Horizon time-homogeneous Tabular Online</td>
<td>Yes</td>
<td>$\tilde{O}\left(\frac{S A}{\epsilon} + \frac{S^2 A}{\epsilon}\right)$</td>
</tr>
<tr>
<td>This Work</td>
<td>Finite Horizon time-homogeneous Tabular/Linear</td>
<td>Yes</td>
<td>$\tilde{O}\left(\frac{1}{d_m \epsilon^2} + \frac{\min(S, d)}{d_m \epsilon}\right)$</td>
</tr>
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<td>$\Omega\left(\frac{1}{d_m \epsilon^2}\right)$</td>
</tr>
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</table>

planning, which is $\min\{H, S, d\}$ away from the lower bound. We provide a $\tilde{O}\left(\frac{1}{d_m \epsilon^2} + \frac{\min\{S, d\}}{d_m \epsilon}\right)$ episode complexity, that matches the lower bound up to logarithm factors and a higher order term.

A recent work [36] considered solving the offline policy optimization with model-free $Q$-learning with variance reduction. Although their algorithm can match the sample complexity lower bound $\Omega\left(\frac{H^2}{d_m \epsilon^2}\right)$ when per-step reward is upper bounded by 1, model-free algorithms generally need at least $\Omega\left(\frac{H^2}{d_m^2}\right)$ episodes of data to finish the algorithm (a.k.a the sample size barrier in [28]) and it is unclear how to translate their results to the setting with total reward upper bounded by 1.

**Horizon-Free Online Reinforcement Learning.** There are several works obtained nearly horizon-free sample complexity bounds for online reinforcement learning. In the time-homogeneous setting, whether the sample complexity needs to scale polynomially with $H$ was an open problem raised by [37]. The problem was first addressed by [26] who proposed an $\epsilon$-net over the optimal policies and a simulation-based algorithms to obtain a sample complexity that only scales logarithmically with $H$, though their dependency on $S, A$ and $\epsilon$ is not optimal and their algorithm is not computationally efficient. The sample complexity bound was later substantially improved by [27] who obtained an $\tilde{O}\left(\frac{S A}{\epsilon^2} + \frac{S^2 A}{\epsilon}\right)$ bound, which nearly matches the contextual bandits (tabular MDP with $H = 1$) lower bound $\Omega\left(\frac{S^2 A}{\epsilon}\right)$ up to an $\frac{S^2 A}{\epsilon}$ factor [38]. Their key ideas are (1) a new bonus function and (2) a recursion-based approach to obtain a tight bound on the sample complexity. Such kind of recursion-based approach cannot be directly applied to the offline setting, and we develop a novel recursion method to suit the offline scenarios.

### 3 Problem Setup

**Notation:** Throughout this paper, we use $[N]$ to denote the set $\{1, 2, \cdots, N\}$ for $N \in \mathbb{Z}^+$, $\Delta(E)$ to denote the set of the probability measure over the event set $E$. Moreover, for simplicity, we use $\epsilon$ to denote polylog $(S, A, H, d, 1/\delta)$ factors (that can be changed in the context), where $\delta$ is the failure probability. We use $\tilde{O}$ and $\Omega$ to denote the upper bound and lower bound up to logarithm factors.

#### 3.1 Markov Decision Process

Markov Decision Process (MDP) is one of the most standard models studied in the reinforcement learning, usually denoted as $\mathcal{M} = (\mathcal{S}, \mathcal{A}, R, P, \mu)$, where $\mathcal{S}$ is the state space, $\mathcal{A}$ is the action space, $R: \mathcal{S} \times \mathcal{A} \to \Delta(\mathbb{R}^+)$ is the reward, $P: \mathcal{S} \times \mathcal{A} \to \Delta(\mathcal{S})$ is the transition, and $\mu$ is the initial state distribution. We additionally define $r: \mathcal{S} \times \mathcal{A} \to \mathbb{R}^+$ to denote the expected reward.
We focus on the episodic MDP with the horizon\(^1\) \(H \in \mathbb{Z}^+\) and time-homogeneous setting\(^2\) that \(P\) and \(R\) do not depend on the level \(h \in [H]\). A (potentially non-stationary) policy \(\pi\) is defined as \(\pi = (\pi_1, \pi_2, \cdots, \pi_H)\), where \(\pi_h : S \rightarrow \Delta(A), \forall h \in [H]\). Following the standard definition, we define the value function \(V^\pi_H(s) := \mathbb{E}_{\mu,P}[\sum_{i=0}^{H} R(s_i, a_i)|s_0 = s]\) and the action-value function (i.e. the Q-function) \(Q^\pi_H(s, a) := \mathbb{E}_{\mu,P}[\sum_{i=0}^{H} R(s_i, a_i)|(s_0, a_0) = (s, a)]\), which are the expected cumulative rewards under the transition \(P\) and policy \(\pi\) starting from \(s_0 = s\) and \((s_0, a_0) = (s, a)\). Notice that, even though \(P\) and \(R\) keep invariant under the change of \(h\), \(V\) and \(Q\) always depend on \(h\) in the episodic setting, which introduces additional technical difficulties compared with the infinite horizon setting.

The expected cumulative reward under policy \(\pi\) is defined as: \(v^\pi := \mathbb{E}_\mu[V^\pi_H(s)]\), and our ultimate goal is finding the optimal policy \(\pi^*\) of \(\mathcal{M}\), which can be written as: \(\pi^* = \arg \max_\pi (v^\pi)\). We additionally define the reaching probabilities \(\xi^\pi_H(s) = \mathbb{P}_\mu,\mu,\pi(s_h = s), \forall h \in [H]\), and state-action pair \((s, a)\) at time step \(h\). Obviously we have \(\sum_s \xi^\pi_H(s) = 1\), and we also have the following relations between \(\xi^\pi_H(s, a)\): \(\xi^\pi_H(s, a) = \xi^\pi_H(s) \pi(a|s)\), \(\xi^\pi_{h+1}(s') = \sum_{s, a} \xi^\pi_h(s, a) P(s'|s, a)\).

With \(\xi^\pi_H\) at hand, we can write \(v^\pi\) in an equivalent way: \(v^\pi = \sum_{s,a} \left(\sum_{h \in [H]} \xi^\pi_h(s, a) \right) r(s, a)\), which provides a dual perspective on the policy evaluation\(^[18]\).

### 3.2 Offline Reinforcement Learning

Generally, \(P\) and \(r\) are not revealed to the learner, which means we can only learn about \(\mathcal{M}\) and identify the optimal policy \(\pi^*\) with data from different kinds of sources. In offline reinforcement learning, the learner can only have access to a collection of data \(D = \{(s_i, a_i, r_i, s_i')\}_{i \in [n]}\) where \(r_i \sim R(s_i, a_i)\) and \(s_i' \sim P(\cdot|s_i, a_i)\), that is collected in \(K\) episodes with (known or unknown) behavior policy (so that \(n = KH\)). For simplicity, define \(n(s, a)\) as the number of data that \((s_i, a_i) = (s, a)\), while \(n(s, a, s')\) is the number of data that \((s_i, a_i, s_i') = (s, a, s')\).

With \(D\), the learner is generally asked to do two kinds of tasks. The first one is the offline policy evaluation (a.k.a. off-policy evaluation), that aims to estimate \(v^\pi\) for the given \(\pi\). The second one is the offline policy optimization, that aims to find the \(\hat{\pi}^*\) that can perform well on \(\mathcal{M}\). We are interested in the statistical limit due to the limited number of data, and how to approach this statistical limit with simple and computationally efficient algorithms.

### 4 Offline Policy Evaluation

In this section, we first consider the offline policy evaluation for tabular MDP with number of state \(S = |S| < \infty\), number of action \(A = |A| < \infty\), which is the basis for the more general settings. We first introduce the plug-in estimator we consider, which is equivalent to different kinds of estimators that are widely used in practice. Then we describe the assumptions we use, show the error bound of the plug-in estimator, and provide the proof sketch of the error bound.

#### 4.1 The Plug-in Estimator

Here we introduce the plug-in estimator. We first build the empirical MDP \(\hat{\mathcal{M}}\) with the data: \(\hat{P}(s'|s, a) = \frac{n(s, a, s')}{n(s, a)}, \hat{r}(s, a) = \frac{\sum_{i \in [n]} \mathbb{1}_{(s_i, a_i) = (s, a)} r_i}{n(s, a)}\), where \(\mathbb{1}\) is the indicator function. Then we correspondingly define \(\hat{Q}^\pi_H, \hat{V}^\pi_H\) and finally the estimator \(\hat{v}^\pi, \forall h \in [H]\), by substituting the \(P\) and \(r\)

\(^1\)A common belief is that we can always reproduce the results between episodic time-homogeneous setting and infinite horizon setting via substitute the horizon \(H\) in episodic setting with the “effective horizon” \(\gamma^{-1}\) in infinite horizon setting. However, this argument does not always hold, for example the dependency decouple technique used in\(^[34, 28]\) cannot be directly applied in the episodic setting.

\(^2\)Some previous work consider time-inhomogeneous setting\([e.g. 39]\), where \(P\) and \(R\) can be varied for different \(h \in [H]\). It is noteworthy that we need an additional \(H\) factor in the sample complexity to identify \(\epsilon\)-optimal policy for time-inhomogeneous MDP compared with time-homogeneous MDP. Transforming the improvement analysis from time-homogeneous setting to time-inhomogeneous setting is trivial (we only need to replace \(S\) with \(HS\), but not vice-versa, as the analysis for time-inhomogeneous setting probably do not exploit the invariant transition sufficiently.)
in $Q_h^π, V_h^π$ and $v^π$ with $\hat{P}$ and $\hat{r}$. Such computation can be efficiently implemented with dynamic programming. We also introduce the reaching probabilities $\xi_h^π(s) = \mathbb{P}_{\mu, \hat{P}, \pi}(s_h = s), \xi_h^π(s, a) = \mathbb{P}_{\mu, \hat{P}, \pi}(s_h = s, a_h = a)$ in the empirical MDP $\hat{M}$, which will be helpful in our analysis.

The plug-in estimator has been studied in [30] under the assumption of linear transition. It’s known that the plug-in estimator is equivalent to the MIS estimator proposed in [19] and a certain version of DualDICE estimator with batch update is proposed in [16], due to the observation that $\hat{\nu} = \sum_{s,a} \left( \left( \sum_{h \in [H]} \xi_h^π(s, a) \right) \hat{r}(s,a) \right)$.

### 4.2 Theoretical Guarantee

Here we first summarize the assumptions we use for the tabular MDP.

**Assumption 1** (Bounded Total Reward). $\forall π$, we have $\sum_{h \in [H]} r_h \leq 1$ almost surely, where $s_1 \sim \mu, a_h \sim \pi(\cdot|s_h), r_h \sim R(s_h, a_h)$ and $s_{h+1} \sim P(\cdot|s_h, a_h), \forall h \in [H]$. This also means $\mathbb{P}(r \sim R(s,a)|r > 1) = 0, \forall (s,a)$.

This is the key assumption used in [26, 27] to escape the polynomial dependence of horizon in episodic setting. As discussed in [37, 26, 27], this assumption is also more general than the uniformly bounded reward assumption: $\forall (s,a), r(s,a) \leq 1/H$. Thus, all of the results in this paper can be generalized to the uniformly bounded reward with a proper scaling of the bounded total reward.

**Assumption 2** (Data Coverage). $\forall (s,a), n(s,a) \geq nd_m$.

This assumption has been used in [19, 24] and is similar to concentration coefficient assumption originated from [40]. Intuitively, the performance of the offline reinforcement learning should depend on $d_m$, since the state-action pair with less visitation will introduce more uncertainty.

Notice that, $d_m \in (0, (SA)^{-1}]$. Assuming access to the generative model, we can query equal number of samples from each state action pair, where $d_m = (SA)^{-1}$. For the offline data sampled with a fixed behavior policy $π_{BEH}$, we can view $d_m \approx \frac{1}{K} \min_{s,a} \sum_{h \in [H]} \xi_h^{π_{BEH}}(s, a)$, which measures the quality of exploration for $π_{BEH}$. When the number of episodes $K = \Omega(1/d_m)$, by standard concentration, we know $\min_{s,a} n(s,a) \approx nd_m = HKd_m$.

We remark that, some of the recent Liu et al. [41], Yin et al. [36] define the data coverage via the coverage on the visitation of the optimal policy when considering offline policy optimization. This kind of data coverage can be covered by our Assumption 2, by considering the sub-MDP which only consists of the state-action pair that can be visited by the optimal policy. As we already know that the optimal policy will not leave this sub-MDP, any near-optimal policy on this sub-MDP will be near-optimal on the original MDP, and hence our results can be directly applied under this alternative definition of data coverage.

With these assumptions at hand, we can present our main results.

**Theorem 1.** Under Assumption 1 and Assumption 2, suppose $K = \Omega(1/d_m)$, then

$$|v^π - \hat{v}^π| \leq \sqrt{\frac{1}{Kd_m^4}}$$

holds with probability at least $1 - \delta$, where $K$ is the number of episodes, $d_m$ is the minimum visiting probability and $\delta$ absorbs the poly log factors.

To demonstrate the tightness of our upper bound, we provide a minimax lower bound in Theorem 2.

**Theorem 2.** There exists a pair of MDPs $\mathcal{M}_1$ and $\mathcal{M}_2$, and offline data $\mathcal{D}$ with $|\mathcal{D}| = KH$ and minimum visiting probability $d_m$, such that for some absolute constant $c_0$, we have

$$\inf_{\hat{v}^π} \sup_{\mathcal{M}_i \in \{\mathcal{M}_1, \mathcal{M}_2\}} \mathbb{P}_{\mathcal{M}_i} \left( |\hat{v}^π(\mathcal{D}) - v^π| > \frac{c_0}{\sqrt{Kd_m^4}} \right) > 0.25,$$

where $\hat{v}^π$ is any estimator that takes $\mathcal{D}$ as input.

**Remark** Theorem 1 and 2 show that, even with the simplest plug-in estimator, we can match the minimax lower bound $\Omega \left( \sqrt{\frac{1}{Kd_m^4}} \right)$ for offline policy evaluation in time-homogeneous MDP up to

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To the best of our knowledge, no lower bound has been provided for finite horizon time-homogeneous setting, so here we provide a minimax lower bound, and the proof can be found in Appendix B.
logarithmic factors. As a result, we can conclude that time-homogeneous MDP is not harder than the bandits in offline policy optimization.

Remark The assumption that $K = \Omega\left(1/d_m\right)$ is a mild and necessary assumption, as with only $o\left(1/d_m\right)$ episodes, there can exist some under-explored state-action pair which unavoidably leads to constant error (this is also how we construct the hard instance for the minimax lower bound).

Remark Central to our analysis is a recursion based upper bound on the “total variance” term (see Lemma 1), which enables us to have a sharper bound for the plug-in estimator compared with previous work [e.g. 19, 24] that include the unnecessary $\text{poly}(S, A)$ factors in the higher-order term.

4.3 Proof Sketch

The detailed proof can be found in Appendix A, and here we sketch our proof in short. With the value difference lemma (see Lemma 3 in the Appendix A), we need to focus on bounding the term:

$$\sum_{h \in [H]} \sum_{s,a} \xi_h^s(a, s) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^\pi(s') \right],$$

which, by Bernstein’s inequality and Cauchy-Schwartz inequality associated with the Assumption 1 and Assumption 2, can be upper bounded by the following term with high probability:

$$\sqrt{\frac{1}{Kd_m}} \sqrt{\sum_{h \in [H]} \sum_{s,a} \xi_h^s(s, a) \text{Var}_{P(s,a)} \left[V_{h+1}^\pi(s') \right]}$$

To bound the “total variance” term $\sum_{h \in [H]} \sum_{s,a} \xi_h^s(s, a) \text{Var}_{P(s,a)} \left[V_{h+1}^\pi(s') \right]$, we use a novel recursion-based method based on the following observation:

Lemma 1 (Variance Recursion). Assume $V_h(s) \in [0, 1], \forall h \in [H], s \in S,$ and then we have that

$$\sum_{h \in [H]} \sum_{s,a} \xi_h^s(s, a) \text{Var}_{P(s,a)} \left[V_{h+1}^\pi(s') \right]$$

$$\leq \sum_{h \in [H]} \sum_{s,a} \xi_h^s(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^\pi(s') \right] + 2^{i+1} \left[ \sum_s \mu(s) V_i(s) + \sum_{h \in [H]} \sum_{s,a} \xi_h^s(s, a) \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^\pi(s') \right].$$

Remark Lemma 1 enables us to obtain a sharp bound with a simple recursion-based analysis. Notice that here we make no assumptions on $V$ except that $V_h(s) \in [0, 1]$, which makes Lemma 1 broadly applicable. Moreover, the term $\sum_s \mu(s) V_i(s)$ can incorporate the information of $V$ into the recursion, which will be helpful in offline policy optimization.

Here the term in (1) can be bounded with the total variance of higher-order value function $\sum_{h \in [H]} \sum_{s,a} \xi_h^s(s, a) \text{Var}_{P(s,a)} \left[V_{h+1}^\pi(s') \right]$, via Bernstein’s inequality. Thus Lemma 1 can be applied iteratively to obtain a tight bound for the “total variance” term. Define

$$\Delta_1(i) = \sum_{h \in [H]} \sum_{s,a} \xi_h^s(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^\pi(s') \right].$$

Applying Lemma 1 with $V_i^\pi(s)$, we can write the following recursion:

$$\Delta_1(i) \leq \sqrt{\frac{1}{Kd_m}} \left( \Delta_1(i+1) + 2^{i+1} (v + \Delta_1(0)) \right) + \frac{1}{Kd_m}.$$

Solve this recursion, and we can obtain the results in Theorem 1.

5 Offline Policy Optimization

In this section, we further consider the offline policy optimization for tabular MDP, which is the ultimate goal for offline reinforcement learning. We first introduce the model-based planning algorithm, which is probably the simplest algorithm for offline policy optimization. Then we analyze the performance gap between the policy obtained by model-based planning and the optimal policy.
5.1 Model-Based Planning

We consider the optimal policy on the empirical MDP $\hat{M}$, which can be defined as
\[ \hat{\pi}^* = \arg \max_\pi \hat{\pi}. \] (3)

Here $\hat{\pi}^*$ can be obtained by dynamic programming with the empirical MDP, which is also known as model-based planning. We remark that our analysis is independent to the algorithm used for solving (3). In other words, the result also applies to the optimization-based planning with the empirical MDPs [42, 43], as long as it solves (3).

5.2 Theoretical Guarantee

Theorem 3 provides an upper bound on the sub-optimality of $\hat{\pi}^*$.

**Theorem 3.** Under Assumption 1 and Assumption 2, suppose that $K = \tilde{\Omega}(1/d_m)$, and then
\[ |v^{\pi^*} - v^{\hat{\pi}^*}| \leq \sqrt{\frac{1}{Kd_m} + \frac{S_t}{Kd_m}}, \]
holds with probability at least $1 - \delta$, where $K$ is the number of episodes, $d_m$ is the minimum visiting probability and $t$ absorbs the poly log factors.

We also provide a minimax lower bound under the finite horizon time-homogeneous setting.

**Theorem 4.** There exists a pair of MDPs $\mathcal{M}_1$ and $\mathcal{M}_2$, and offline data $\mathcal{D}$ with $|\mathcal{D}| = KH$ and minimum state-action pair visiting frequency $d_m$, such that for some absolute constant $c_0$, we have
\[ \inf_{\hat{\pi}} \sup_{\mathcal{M}_1, \mathcal{M}_2} \mathbb{P}_{\mathcal{M}_1} \left( |v^{\hat{\pi}}(\mathcal{D}) - v^{\pi^*}| > \frac{c_0}{\sqrt{Kd_m}} \right) > 0.25, \]
where $\hat{\pi}$ is any planning algorithm that takes $\mathcal{D}$ as input.

**Remark** Theorem 3 provides a bound approaching the minimax lower bound in Theorem 4 up to logarithmic factors and a higher-order term, which shows that the error of offline policy optimization does not scale polynomially on the horizon. Notice that if $d_m = \Omega \left( \frac{1}{S^2A} \right)$, we can obtain an error bound of $\tilde{O} \left( \frac{SA}{K} + \frac{SA^2}{Kd_m} \right)$, which can be translated to sample complexity $\tilde{O} \left( \frac{SA}{v^{\pi^*}} + \frac{SA^2}{Kd_m} \right)$ that matches the best known result of sample complexity for online finite-horizon time-homogeneous setting [27]. We conjecture that the additional $S$ factor in the higher-order term is only an artifact (see Lemma 2) and can be eliminated with more delicate analysis. We leave this as an open problem.

**Remark** There are also works considering local policy optimization [e.g. 44–46] when the offline data are not sufficient exploratory. We want to emphasize that, as Theorem 4 suggests, to identify the global optimal policy, we need the offline data sufficient exploratory.

5.3 Proof Sketch

The detailed proof can be found in Appendix A, and here we sketch our proof in short. Notice that
\[ v^{\pi^*} - v^{\hat{\pi}^*} = v^{\pi^*} - \hat{\pi}^{\pi^*} + \hat{\pi}^{\pi^*} - \hat{\pi}^{\hat{\pi}^*} + \hat{\pi}^{\hat{\pi}^*} - v^{\hat{\pi}^*} \leq 0 \quad \text{ Error on Fixed Policy Error on Data-Dependent Policy} \] (4)

We can directly apply Theorem 1 to bound the error on fixed policy. For the error on data-dependent policy, since the policy $\hat{\pi}^*$ depends on data, we need to consider
\[
\sum_{h \in [H]} \sum_{s,a} \hat{\xi}_h^*(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^{\hat{\pi}^*}(s') \right] 
= \sum_{h \in [H]} \sum_{s,a} \hat{\xi}_h^*(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^{\pi^*}(s') \right] 
+ \sum_{h \in [H]} \sum_{s,a} \hat{\xi}_h^*(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) \left( V_{h+1}^{\pi^*}(s') - V_{h+1}^{\hat{\pi}^*}(s') \right) \right].
\] (5)
As \( \pi^* \) is independent of \( \hat{P} \), the term in (5) can be similarly handled with the techniques used in offline policy evaluation. However, due to the dependency of \( \hat{\pi}^* \) and \( \hat{P} \), we need to deal with the term in (6) more carefully. To decouple the dependency of \( \hat{\pi}^* \) and \( \hat{P} \), we first introduce the following lemma:

**Lemma 2.** \( \forall V_h(s) \in [0, 1], \forall h \in [H], s \in S \), then we have that with high probability,

\[
\left| \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_h(s') \right| \leq \sqrt{\frac{\text{Var}_{P(s,a)}[V_h(s')]}{n(s,a)}} + \frac{S_h}{n(s,a)}.
\]

**Remark** Lemma 2 has been widely used in the design and analysis of online reinforcement learning algorithms, [e.g., 27]. It holds even \( V_h(s) \) depends on \( \hat{P}(s'|s, a) \), however, at the cost of an additional \( S \) factor. This is the source of the additional \( S \) factor in the higher-order term, and we believe a more fine-grained analysis can help avoid this additional \( S \) factor.

Denote

\[
\Delta_2(i) = \sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) (V_{h+1}^\pi(s') - V_{h+1}^{\hat{\pi}}(s'))^{2^i} \right].
\]

Applying Lemma 2, and then Lemma 1 with \( V_h(s) = V_h^\pi(s) - V_h^{\hat{\pi}}(s) \), we have that

\[
\Delta_2(i) \leq \sqrt{\frac{S_h}{K_d m}} (\Delta_2(i+1) + 2(i+1)(v^\pi - v^{\hat{\pi}} + \Delta_2(0))) + \frac{S_h}{K_d m}.
\]

Also, with (4) and (5), we have that \( v^\pi - v^{\hat{\pi}} \leq \Delta_2(0) + O \left( \frac{1}{K_d m} \right) \). Solve this recursion, then we can obtain the results in Theorem 3.

**Remark** We want to emphasize that, as Lemma 1 does not make specific assumptions on \( V \), we can bound the term in (5) with almost identical analysis in offline policy evaluation. Moreover, the \( \sum_k \mu_k(s) (V_k^\pi(s) - V_k^{\hat{\pi}}(s)) = v^\pi - v^{\hat{\pi}} \) is the term of interest, which helps to obtain a tight bound and also further demonstrate the effectiveness of Lemma 1. Compared with the absorbing MDP technique [28, 1], our recursion-based analysis is much more simple without any need for constructing \( \epsilon \)-net on \( V \).

### 6 Extensions to Linear MDP with Anchor Points

In this section, we first introduce the definition of the linear MDP with anchor points [47, 1], and then generalize our results of offline policy evaluation and optimization to this setting.

**Definition 1** (Linear MDP with Anchor Points [47, 1]). For the MDP \( \mathcal{M} = (S, A, R, P, \mu) \), assume there is a feature map \( \phi : S \times A \rightarrow \mathbb{R}^d \), such that \( r \) and \( P \) admits a linear representation:

\[
r(s, a) = \langle \phi(s, a), \theta_r \rangle, \quad P(s'|s, a) = \langle \phi(s, a), \mu(s') \rangle,
\]

where \( \mu \) is an unknown (signed) measure of \( S \). Furthermore, assume there exists a set of anchor state-action pairs \( K \), such that \( \forall (s, a) \in S \times A \), \( \exists \{\lambda_k^{s,a}\}_{k \in K} \),

\[
\phi(s, a) = \sum_{k \in K} \lambda_k^{s,a} \phi(s_k, a_k), \quad \sum_{k \in K} \lambda_k^{s,a} = 1, \quad \lambda_k^{s,a} \geq 0, \forall k \in K.
\]

With the definition of Linear MDP as well as the anchor points assumption, we can find that \( r(s, a) = \sum_{k \in K} \lambda_k^{s,a} r(s_k, a_k), \quad P(s'|s, a) = \sum_{k \in K} \lambda_k^{s,a} P(s'|s_k, a_k) \), which can lead to an empirical estimation of \( P \) and \( r \) by replacing \( \{r(s_k, a_k)\}_{k \in K}, \{P(s'|s_k, a_k)\}_{k \in K} \) with the empirical counterpart \( \{\hat{r}(s_k, a_k)\}_{k \in K}, \{\hat{P}(s'|s_k, a_k)\}_{k \in K} \) estimated from the offline data. Following [47, 1], we additionally make the following assumption on the offline data:

**Assumption 3** (Anchor Point Data [47, 1]). Assume \( |K| = d \). For the offline data \( D = \{(s_i, a_i, r_i, s_i')\}_{i \in [n]}, (s_i, a_i) \in \{(s_k, a_k)\}_{k \in K}, \forall i \in [n] \). Furthermore, \( \forall k \in K, n(s_k, a_k) \geq nd_m \).

We now present the main theorem on the offline policy evaluation and offline policy improvement on the linear MDP with anchor points.
Theorem 5. Under Assumption 1 and Assumption 3, suppose \( n = \Omega(H/d_m) \), and then for a given policy \( \pi \), the plug-in estimator \( \hat{\pi}^* \) satisfies

\[
|v^\pi - \hat{v}^\pi| \leq \sqrt{\frac{H}{n d_m}}
\]

with probability at least \( 1 - \delta \), where \( n \) is the number of offline data, \( d_m \) is the minimum visiting probability of anchor points and \( \delta \) absorbs the polylog factors.

Theorem 6. Under Assumption 1 and Assumption 3, suppose that \( n = \Omega(H/d_m) \), and then for the \( \hat{\pi}^* \) obtained by model-based planning,

\[
|v^{\pi^*} - v^{\hat{\pi}^*}| \leq \sqrt{\frac{H}{n d_m}} + \frac{d H}{n d_m},
\]

holds with probability at least \( 1 - \delta \), where \( n \) is the number of offline data, \( d \) is the feature dimension, \( d_m \) is the minimum visiting probability of anchor points and \( \delta \) absorbs the polylog factors.

Proof of both theorems can be found in Appendix C. From a high-level perspective, we observe that for the unseen state-action pair, we still have Bernstein-type concentration bound and a lemma similar to Lemma 2. Hence we can apply our recursion introduced in Lemma 1 and use the similar techniques for tabular MDP to obtain the desired results. We want to remark that such results demonstrate the broad applicability of our recursion-based analysis in different kinds of offline scenarios.

Remark Compared with the results in [1], we remove the additional dependency of \( \min\{H, |S|, d\} \) in the main term and approach the optimal complexity shown in the [47] up to logarithmic factors, which shows that model-based planning is minimax optimal for Linear MDP with anchor points and demonstrates the effectiveness of our recursion-based analysis. There are still several open problems. For example, can we provide a sharper analysis for the policy obtained by model-based planning without any additional factors on higher-order term? Can we extend to MDP with more general assumption? We leave these problems as future work.

7 Conclusion

In this paper, we revisit the offline reinforcement learning on episodic time-homogeneous MDP. We show that, if the total reward is properly normalized, offline reinforcement learning is not harder than the offline bandits counterparts. Specifically, we provide performance guarantee for algorithms based on empirical MDPs that match the lower bound up to logarithmic factors for offline policy evaluation, and up to logarithmic factors and a higher-order term for offline policy optimization, and both do not have polynomial dependency on \( H \). There are still several open problems. For example, can we provide a sharper analysis for the policy obtained by model-based planning without any additional factors on higher-order term? Can we extend to MDP with more general assumption? We leave these problems as future work.
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References


Checklist

1. For all authors...
   (a) Do the main claims made in the abstract and introduction accurately reflect the paper’s contributions and scope? [Yes]
   (b) Did you describe the limitations of your work? [Yes]
   (c) Did you discuss any potential negative societal impacts of your work? [N/A]
   (d) Have you read the ethics review guidelines and ensured that your paper conforms to them? [Yes]

2. If you are including theoretical results...
   (a) Did you state the full set of assumptions of all theoretical results? [Yes]
   (b) Did you include complete proofs of all theoretical results? [Yes]

3. If you ran experiments...
   (a) Did you include the code, data, and instructions needed to reproduce the main experimental results (either in the supplemental material or as a URL)? [N/A]
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A Proof of the Main Theorems

A.1 Proof for Offline Policy Evaluation

The proof of all of the technical lemmas can be found in Appendix E. Our proof is organized as follows: we first decompose the estimation error to the errors introduced by reward estimation \( \hat{r} \) and transition estimation \( \hat{P} \) with Lemma 3, then Lemma 4 provides an upper bound for the error introduced by \( \hat{r} \). For error introduced by \( \hat{P} \), we first show it can be upper bounded by a “total variance” term (in (8)). A naive bound for this “total variance” term will introduce an additional \( H \) factors, thus we apply a recursion-based method to upper bound this “total variance” term (see Lemma 1). Solve the recursion in Lemma 5, and put everything together, we eventually obtain the bound in Theorem 1.

**Lemma 3** (Value Difference Lemma).

\[
v^{\pi} - \tilde{v}^{\pi} = \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \left[ r(s, a) - \hat{r}(s, a) + \sum_{s'} \left[ \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^\pi(s') \right] \right].
\]

(7)

The following lemma provides an upper bound on the estimation error introduced by \( \hat{r} \), i.e., the first term in (7).

**Lemma 4** (Error from Reward Estimation). Suppose Assumption 1 holds, then we have that

\[
\left| \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \left[ r(s, a) - \hat{r}(s, a) \right] \right| \leq \sqrt{\frac{\lambda_h}{K d_m}} + \frac{h}{\sqrt{K d_m}}.
\]

holds with probability at least \( 1 - \delta \).

We then use a recursive method to bound the error introduced by \( \hat{P} \), i.e., the second term in (7). Denote

\[
\Delta_1 := \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^\pi(s') \right].
\]

With Bernstein’s inequality, we know that with high probability

\[
\Delta_1 \leq \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \left[ \sqrt{\frac{\text{Var}_{P(s, a)}(V_{h+1}^\pi(s')) t}{n(s, a)}} + \frac{t}{n(s, a)} \right]
\]

\[
\leq \sqrt{\frac{t}{K d_m}} \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \text{Var}_{P(s, a)}(V_{h+1}^\pi(s')) + \frac{t}{K d_m},
\]

(8)

where the second inequality is due to Cauchy-Schwartz inequality associated with the Assumption 2.

We then upper bound \( \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \text{Var}_{P(s, a)}(V_{h+1}^\pi(s')) \) in (8) with Lemma 1. Again, by Bernstein’s inequality and Cauchy-Schwarz inequality, we have the following with high probability:

\[
\sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}^\pi(s')^{2^{i+1}} \right]
\]

\[
\leq \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \left[ \sqrt{\frac{\text{Var}_{P(s, a)}(V_{h+1}^\pi(s')^{2^{i+1}}) t}{n(s, a)}} + \frac{t}{n(s, a)} \right]
\]

\[
\leq \sqrt{\frac{t}{K d_m}} \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \text{Var}_{P(s, a)}(V_{h+1}^\pi(s')^{2^{i+1}}) + \frac{t}{K d_m}.
\]

(9)

Define

\[
V_1(i) := \sum_{h \in [H]} \sum_{s, a} \xi_h^\pi(s, a) \text{Var}_{P(s, a)}(V_{h+1}^\pi(s')^{2^{i+1}}).
\]
Apply Lemma 1 with (9), we have the recursion as

\[ V_1(i) \leq \sqrt{\frac{t}{Kd_m}}V_1(i+1) + \frac{t}{Kd_m} + 2^{i+1}(\Delta_1 + v^\pi). \]  

(10)

Notice that \( V_1(i) \leq H, \forall i \). Now we can solve the recursion with the following lemma:

**Lemma 5.** For the recursion formula:

\[ V(i) \leq \sqrt{\lambda_1 V(i+1) + \lambda_1 + 2^{i+1}\lambda_2}, \]

with \( \lambda_1, \lambda_2 > 0 \), if \( V(i) \leq H, \forall i \), then we have that

\[ V(0) \leq 6(\lambda_1 + \lambda_2), \]

and we need to do the recursion at most \( O(\log \frac{H\lambda_1}{\lambda_2}) \) times.

**Remark** We want to emphasize that, the recursion needs to be done at most \( O(\log \frac{H\lambda_1}{\lambda_2}) \) times. Thus, by union bound, such recursion only introduces an additional \( \log \log \) factor in the error when \( \lambda_1, \lambda_2 = \text{poly}(S, A, H) \), that can be absorbed by \( \iota \). For simplicity, we still use \( \iota \) to denote the \( \text{poly log} \) factors in the following derivation.

Apply Lemma 5 with \( \lambda_1 = \frac{\iota}{Kd_m}, \lambda_2 = \Delta_1 + v^\pi, \) we have that

\[ V_1(0) = O\left(\frac{t}{Kd_m} + \Delta_1 + v^\pi\right), \]

and also notice in (8) that

\[ \Delta_1 \leq \sqrt{\frac{t}{Kd_m}}V_1(0) + \frac{t}{Kd_m}. \]

Combine these two inequalities, we have that

\[ \Delta_1 \leq O\left(\sqrt{\frac{t}{Kd_m}}(\frac{t}{Kd_m} + \Delta_1 + v^\pi)\right) + \frac{t}{Kd_m}. \]

Suppose \( K = \tilde{\Omega} \left(\frac{1}{d_m}\right) \), then with Assumption 1 and discussion in Appendix D, we have that

\[ \Delta_1 \leq O\left(\sqrt{\frac{t}{Kd_m}} + \frac{t}{Kd_m}\right). \]

(11)

Combined (11) with Lemma 4 and \( K = \tilde{\Omega} \left(\frac{1}{d_m}\right) \), we conclude the proof of Theorem 1.

**A.2 Proof for Offline Policy Optimization**

We first make the following standard decomposition:

\[ v^\pi - v^{\hat{\pi}^*} = v^\pi - \hat{v}^\pi + \hat{v}^\pi - \hat{\hat{v}}^\pi + \hat{\hat{v}}^\pi - v^{\hat{\pi}^*} \]

\[ \leq \text{Error on Fixed Policy} + \text{Error on Data-Dependent Policy} \]

(12)

The first term characterizes the evaluation difference of optimal policy on original MDP and the empirical MDP, and the second term characterize the evaluation difference of the planning result \( \hat{\pi}^* \) from the empirical MDP on original MDP and the empirical MDP.

We can directly apply Theorem 1 to bound the first term in (12). However, as \( \hat{\pi}^* \) has complicated statistical dependency with \( \hat{P} \), we cannot apply Theorem 1 on the second term in (12) for the evaluation error on data-dependent policy. Notice that a direct application of the absorbing MDP techniques introduced in [34, 28] for the second term will introduce additional \( H \) or \( S \) factors in the main term as shown in [1]. Thus, we further generalize our recursion-based method to keep the main
We can further apply Lemma 2 and Cauchy-Schwartz inequality to the first term in (14), which

\[ \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \left[ \left( \hat{r}(s,a) - r(s,a) \right) + \sum_{s'} (\hat{P}(s'|s,a) - P(s'|s,a))V_{h+1}^{\pi}\right] \]

\[ = \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \left( \Delta_r(s,a) + \Delta_P(s,a) + \Delta_{PV}(s,a) \right), \]

where

\[ \Delta_r(s,a) := \hat{r}(s,a) - r(s,a), \]

\[ \Delta_P(s,a) := \sum_{s'} (\hat{P}(s'|s,a) - P(s'|s,a))V_{h+1}^{\pi}, \]

\[ \Delta_{PV}(s,a) := \sum_{s'} (\hat{P}(s'|s,a) - P(s'|s,a)) \left( V_{h+1}^{\pi}(s') - V_{h+1}^{\pi}(s') \right). \]

For the inner product of \( \xi_h^\pi(s,a) \) with \( \Delta_r \) and \( \Delta_P \) term, as \( r \) and \( \pi^* \) is independent of \( \hat{P} \), we can identically apply the result for offline evaluation, that leads to a \( \tilde{O} \left( \frac{1}{K\alpha_m} + \frac{1}{K\alpha_m} \right) \) error. We then consider the \( \Delta_{PV} \) term.

With Lemma 2, we can again use the recursion to bound the error introduced by \( \Delta_{PV} \), and finally obtain the bound in Theorem 3. By Lemma 2 and Cauchy-Schwartz inequality, we have that

\[ \Delta_2 := \left| \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \Delta_{PV}(s,a) \right| \]

\[ \leq \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \left[ \sqrt{\frac{S \cdot \text{Var}_{P(s,a)}(V_{h+1}^{\pi}(s') - V_{h+1}^{\pi}(s'))}{n(s,a)}} + \frac{S_t}{n(s,a)} \right] \]

\[ \leq \frac{S_t}{Kd_m} \sqrt{\sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \text{Var}_{P(s,a)}(V_{h+1}^{\pi}(s') - V_{h+1}^{\pi}(s'))} + \frac{S_t}{Kd_m}. \quad (13) \]

Now we turn to \( \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \text{Var}_{P(s,a)}(V_{h+1}^{\pi}(s') - V_{h+1}^{\pi}(s')) \) in (13). We still bound this term with the recursive methods. With Lemma 1, we have that

\[ \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \left( V_{h+1}^{\pi}(s') - V_{h+1}^{\pi}(s') \right)^2 \]

\[ \leq \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \left[ \sum_{s'} (P(s'|s,a) - \hat{P}(s'|s,a)) \left( V_{h+1}^{\pi}(s') - V_{h+1}^{\pi}(s') \right)^{2i+1} \right] + 2^{i+1} \left( \Delta_2 + (v^{\pi^*} - v^{\hat{\pi}^*}) \right). \quad (14) \]

We can further apply Lemma 2 and Cauchy-Schwartz inequality to the first term in (14), which eventually lead to the recursion formula. Specifically, denote

\[ \mathbb{V}_2(i) := \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \left( V_{h+1}^{\pi}(s') - V_{h+1}^{\pi}(s') \right)^2, \]

we have the recursion as

\[ \mathbb{V}_2(i) \leq \frac{S_t}{Kd_m} \mathbb{V}_2(i+1) + \frac{S_t}{Kd_m} + 2^{i+1} \left( \Delta_2 + (v^{\pi^*} - v^{\hat{\pi}^*}) \right). \]
This recursion can be solved similarly as (10) by applying Lemma 5 with \( \lambda_1 = \frac{S_t}{K d_m} \) and \( \lambda_2 = \Delta_2 + v^{\pi*} - v^{{\bar{\pi}}*} \), which leads to
\[
V_2(0) \leq O \left( \frac{S_t}{K d_m} + (v^{\pi*} - v^{{\bar{\pi}}*} + \Delta_2) \right).
\] (15)

Meanwhile, from (13) we have that
\[
\Delta_2 \leq \sqrt{\frac{S_t}{K d_m} V_2(0)} + \frac{S_t}{K d_m}.
\] (16)

Combine (15) and (16), we have
\[
\Delta_2 \leq O \left( \sqrt{\frac{S_t}{K d_m} \left( \frac{S_t}{K d_m} + v^{\pi*} - v^{{\bar{\pi}}*} + \Delta_2 \right)} + \frac{S_t}{K d_m} \right).
\]

Moreover, with (12) and Theorem 1, we have that
\[
v^{\pi*} - v^{{\bar{\pi}}*} \leq O \left( \sqrt{\frac{1}{K d_m} + \frac{1}{K d_m}} \right) + \Delta_2.
\]

Thus, with the discussion in Appendix D, we can conclude that \( v^{\pi*} - v^{{\bar{\pi}}*} \leq O \left( \sqrt{\frac{1}{K d_m} + \frac{S_t}{K d_m}} \right) \), which finishes the proof for Theorem 3.

### B Proof of the Lower Bounds

#### B.1 Lower Bound for Offline Policy Evaluation

Our lower bound instance is adapted from the instances in [35, 49, 29] for finite horizon time-homogeneous setting.

**Proof.** We consider a two-state MDP with state \( s_1, s_2 \), with an unique action \( a \). \( s_1 \) is an absorbing state, which means \( P(s_1|s_1, a) = 1 \), while \( P(s_2|s_2, a) = p \), \( P(s_1|s_2, a) = 1 - p \). We assume the reward is deterministic with \( r(s_1, a) = 0 \), \( r(s_2, a) = \frac{1}{H} \) that satisfies Assumption 1. Assume we want to have an accurate estimation of \( V_1(s_2) \), which is equivalent to have a sufficient accurate estimation of \( p \). With straightforward calculation, we have that
\[
V(p) := V_1(s_2) = \frac{p - p^{H+1}}{1 - p} \frac{1}{H}.
\]

Notice that
\[
\frac{\partial V(p)}{\partial p} = \frac{1 - (1 + (1 - p)H)p^H}{(1 - p)^2} \frac{1}{H}.
\]

Let \( p_1 = 1 - \frac{c_1}{H} \), where \( c_1 \) is an absolute constant, we know that
\[
\frac{\partial V(p_1)}{\partial p_1} = \frac{1 - c_1(1 - \frac{c_1}{H}H)}{(1 - p_1)^2} \frac{1}{H} \geq \frac{1 - c_1 e^{-c_1}}{(1 - p_1)^2} \frac{1}{H} = \frac{1 - c_1 e^{-c_1}}{c_1^2} \frac{H}{H},
\]
which is monotonically decreasing w.r.t \( c_1 \). Assume \( p_2 = 1 - \frac{c_2}{H} \) where \( c_2 < c_1 \) is another absolute constant, we have that
\[
V(p_2) - V(p_1) \geq \frac{1 - c_1 e^{-c_1}}{(1 - p_1)^2} \frac{1}{H} (p_2 - p_1) = \frac{1 - c_1 e^{-c_1}}{c_1^2} (c_1 - c_2).
\]

We now use Le Cam’s method to show that without sufficient number of data from Bern(p), we cannot identify \( p = p_1 \) or \( p = p_2 \) with high probability, and thus cannot have ideal estimation error on both of \( p_1 \) and \( p_2 \). We start from the following lemma:

**Lemma 6.**
\[
KL(\text{Bern}(p)||\text{Bern}(q)) \leq \frac{(p - q)^2}{q(1 - q)}.
\]
Proof.

\[
\begin{align*}
\text{KL}(\text{Bern}(p)||\text{Bern}(q)) &= p \log \frac{p}{q} + (1 - p) \log \frac{1 - p}{1 - q} \\
&\leq p \frac{|p - q|}{q} + (1 - p) \log \frac{1 - p}{1 - q} \\
&= \frac{(p - q)^2}{q(1 - q)},
\end{align*}
\]

where the inequality is due to the fact that \(\log(1 + x) \leq x\). \(\square\)

Assume \(\Psi : [0, 1]^{n(s, a)} \rightarrow \{p_1, p_2\}\) is a test with \(n\) i.i.d samples from \(\text{Bern}(p)\), and use \(P_1\) and \(P_2\) to denote the probability measure under \(p_1\) and \(p_2\), we have that

\[
\inf_{\Psi} \{P_1(\Psi(D) \neq p_1) + P_2(\Psi(D) \neq p_2)\}
\begin{align*}
&\geq 1 - \frac{\text{KL}(\text{Bern}(p_1)||\text{Bern}(p_2))}{2} \quad \text{(Le Cam’s inequality)} \\
&\geq 1 - \sqrt{\frac{n(s, a)\text{KL}(\text{Bern}(p_1)||\text{Bern}(p_2))}{2}} \quad \text{(Pinsker’s inequality)} \\
&\geq 1 - \sqrt{\frac{n(s, a)(p_1 - p_2)^2}{p_1(1 - p_1)}} \\
&= 1 - \sqrt{\frac{n(s, a)(c_1 - c_2)^2}{c_1(H - c_1)}}.
\end{align*}
\]

Take \(c_2 = c_1 - \frac{c_1(H - c_1)}{2n(s, a)}\), we know that with probability at least 0.5 we cannot identify \(p = p_1\) or \(p = p_2\). And notice that

\[
V(p_2) - V(p_1) \geq 1 - c_1 e^{-c_1} \sqrt{\frac{c_1(H - c_1)}{2n(s, a)}} \geq 2c_0 \sqrt{\frac{H}{n(s, a)}},
\]

where \(c_0\) is an absolute constant only depends on \(c_1\). Thus we know that, with \(n(s, a)\) samples from \(P(s, a)\), we must suffer from an estimation error of \(\Omega\left(\sqrt{\frac{H}{n(s, a)}}\right)\) with probability at least 0.25.

Notice that we can set \(n(s, a) = nd_{m}\), thus finishes the proof. \(\square\)

### B.2 Lower Bound of Offline Policy Improvement

We can further show the lower bound of offline improvement for finite horizon time-homogeneous MDP, based on the hard instance we mentioned above.

Proof. We introduce additional states \(s_0\) and \(s_3\) in the previous hard instance, with the transition from \(s_0\), \(P(s_1|s_0, a_1) = 1\), \(P(s_3|s_0, a_1) = 1\), \(\forall a \neq a_1\), and \(s_3\) is an absorbing state with total reward in \(H\) steps (i.e. \(V_1(s_3)\)) as \(V(p_1) + c_0 \sqrt{\frac{H}{nd_{m}}}\). We always start from \(s_0\), and we need to choose the action at \(s_0\). Notice that, if \(p = p_2\), then the optimal arm is \(a_1\), while if \(p = p_1\), then the optimal arm is not \(a_1\), both with a sub-optimal gap of at least \(c_0 \sqrt{\frac{H}{nd_{m}}}\), which finishes the proof. \(\square\)

### C Proof for Linear MDP with Anchor Points

#### C.1 Proof for Offline Policy Evaluation

Notice that, the value difference lemma holds for any kinds of MDP. Thus, if we have Bernstein-type concentration for \(\hat{r}(s, a) - r(s, a)\) and \(\sum_{s'}(P(s'|s, a) - \hat{P}(s'|s, a))V(s')\), we can adopt the techniques for tabular MDP and obtain the desired results.
Lemma 7. With probability at least $1 - \delta$, we have that $\forall (s, a)$,

$$|\hat{r}(s, a) - r(s, a)| \leq \sqrt{\frac{r(s, a)}{nd_m}} + \frac{\tau}{nd_m}$$

where $\tau$ absorbs the logarithm factors $\log(\text{poly}(d)/\delta)$.

Proof. Notice that

$$|\hat{r}(s, a) - r(s, a)| \leq \sum_{k \in \mathcal{K}} \lambda_{s,a}^k |\hat{r}(s_k, a_k) - r(s_k, a_k)|$$

$$\leq \sum_{k \in \mathcal{K}} \lambda_{s,a}^k \left( \sqrt{\frac{r(s_k, a_k)}{nd_m}} + \frac{1}{nd_m} \right)$$

$$\leq \sqrt{\sum_{k \in \mathcal{K}} \lambda_{s,a}^k r(s_k, a_k)} + \frac{\tau}{nd_m}$$

$$= \sqrt{\frac{r(s, a)}{nd_m}} + \frac{\tau}{nd_m},$$

where the second inequality is due to the Bernstein’s inequality on each $(s_k, a_k)$ with $\tau = \log(2d/\delta)$ and $r(s_k, a_k) \in [0, 1]$, and the third inequality is due to Cauchy-Schwartz inequality and $\sum_{k \in \mathcal{K}} \lambda_{s,a}^k = 1$. \hfill \Box

Lemma 8. Suppose $V(s')$ is independent from $\hat{P}(s'|s, a)$, then with probability at least $1 - \delta$, we have that $\forall (s, a)$

$$\left| \sum_{s'} (P(s'|s, a) - \hat{P}(s'|s, a))V(s') \right| \leq \sqrt{\frac{\text{Var}_{P(s,a)}V(s')\tau}{nd_m}} + \frac{\tau}{nd_m},$$

where $\tau$ absorbs the logarithm factors $\log(\text{poly}(d)/\delta)$.

Proof. First, we have that

$$\left| \sum_{s'} (P(s'|s, a) - \hat{P}(s'|s, a))V(s') \right|$$

$$\leq \sum_{k \in \mathcal{K}} \lambda_{s,a}^k \left| \sum_{s'} (P(s'|s_k, a_k) - \hat{P}(s'|s_k, a_k))V(s') \right|$$

$$\leq \sum_{k \in \mathcal{K}} \lambda_{s,a}^k \left( \sqrt{\frac{\text{Var}_{P(s_k,a_k)}V(s')\tau}{nd_m}} + \frac{\tau}{nd_m} \right)$$

$$\leq \sqrt{\sum_{k \in \mathcal{K}} \lambda_{s,a}^k \text{Var}_{P(s_k,a_k)}V(s')\tau} + \frac{\tau}{nd_m}. $$

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where the first inequality is due to the Bernstein’s inequality on each \((s_k, a_k)\) with \(\epsilon = \log(2d/\delta)\) and the last inequality is due to Cauchy-Schwartz inequality and \(\sum_{k \in \mathcal{K}} \lambda_k^{s,a} = 1\). Notice that

\[
\sum_{k \in \mathcal{K}} \lambda_k^{s,a} \text{Var}_{P(s_k, a_k)} V(s') = \sum_{k \in \mathcal{K}} \lambda_k^{s,a} \left( \sum_{s'} P(s'|s_k, a_k) V(s')^2 - \left( \sum_{s'} P(s'|s_k, a_k) V(s') \right)^2 \right)
\leq \sum_{s'} P(s'|s, a) V(s')^2 - \left( \sum_{k \in \mathcal{K}} \lambda_k^{s,a} P(s'|s_k, a_k) V(s') \right)^2
\]

where the inequality is due to Cauchy-Schwartz inequality and \(\sum_{k \in \mathcal{K}} \lambda_k^{s,a} = 1\). Substitute this term back and we conclude the proof. \(\square\)

Notice that, Lemma 8 simultaneously holds for all of the \((s, a)\) if all of the concentration on the anchor points hold. Hence we can apply the analysis for tabular MDP and obtain the desired results, which finishes the proof for offline policy evaluation on linear MDP with anchor points.

### C.2 Proof for Offline Policy Optimization

Here we need a Bernstein-type concentration for \(\sum_{s'} (P(s'|s, a) - \hat{P}(s'|s, a)) V(s')\), when \(V(s')\) and \(\hat{P}(s'|s, a)\) are correlated. A naive application of Lemma 2 will introduce a \(|S|\) factor in the higher order term, which is not satisfactory, as \(|S|\) can be exponentially large. Here we use another method to replace this dependency on \(|S|\) with the feature dimension \(d\).

**Lemma 9.** Suppose \(\hat{V}(s')\) is independent from \(\hat{P}(s'|s, a)\), then with probability \(1 - \delta\), we have that

\[
\left| \sum_{s'} (P(s'|s, a) - \hat{P}(s'|s, a)) V(s') \right| \leq \sqrt{\text{Var}_{P(s,a)} V(s') \frac{\epsilon}{n d_m}} + \frac{\epsilon}{n d_m} + \|\hat{V} - V\|_\infty \left( 1 + \sqrt{\frac{\epsilon}{n d_m}} \right),
\]

where \(\epsilon\) absorbs the logarithm factors \(\log(\text{poly}(d)/\delta)\).

**Proof.** Notice that

\[
\left| \sum_{s'} (P(s'|s, a) - \hat{P}(s'|s, a)) V(s') \right| \leq \sum_{s'} \left| (P(s'|s, a) - \hat{P}(s'|s, a)) V(s') \right| + \sum_{s'} (P(s'|s, a) - \hat{P}(s'|s, a)) (V(s') - \hat{V}(s'))
\]

\[
\leq \sqrt{\text{Var}_{P(s,a)} V(s') \frac{\epsilon}{n d_m}} + \frac{\epsilon}{n d_m} + \|\hat{V} - V\|_\infty
\]

\[
\leq \sqrt{\text{Var}_{P(s,a)} V(s') \frac{\epsilon}{n d_m}} + \sqrt{\text{Var}_{P(s,a)} (\hat{V}(s') - \hat{V}(s')) \frac{\epsilon}{n d_m}} + \frac{\epsilon}{n d_m} + \|\hat{V} - V\|_\infty
\]

\[
\leq \sqrt{\text{Var}_{P(s,a)} V(s') \frac{\epsilon}{n d_m}} + \frac{\epsilon}{n d_m} + \|\hat{V} - V\|_\infty \left( 1 + \sqrt{\frac{\epsilon}{n d_m}} \right),
\]

where the first inequality is due to the triangle inequality, the second inequality is due to Lemma 8 and algebra, the third inequality is due to the triangle inequality for the variance, i.e. \(\sqrt{\text{Var}(X + Y)} \leq \sqrt{\text{Var}(X)} + \sqrt{\text{Var}(Y)}\), and the last inequality is due to the fact that \(\text{Var}(V) \leq \|V\|_\infty\). \(\square\)

With Lemma 9, we can construct an \(\epsilon\)-net (under \(\ell_\infty\) norm) for \(V\) to obtain Bernstein-type concentration. For tabular MDP, this \(\epsilon\)-net is of size \(O(\epsilon^{-|S|})\), which leads to the same result of Lemma 2.
However, in linear MDP, $Q$ follows a linear form $Q(s, a) = \phi(s, a)^\top w_Q$, thus the $V$ we consider lies in a $d$-dimensional manifolds, and the size of the $\epsilon$-net we exactly need is $O(\epsilon^{-d})$. This observation leads to the following corollary:

**Corollary 1.** For any $V(s')$, with probability $1 - \delta$, we have that

$$\left| \sum_{s'} (P(s'|s, a) - \hat{P}(s'|s, a)) V(s') \right| \leq \sqrt{\frac{d \text{Var}_{P(s,a)} V(s')}{nd_m}} + \frac{d \epsilon}{nd_m} + \epsilon \left(1 + \sqrt{\frac{d \epsilon}{nd_m}}\right),$$

where $\epsilon$ absorbs the logarithm factors, $\log(\text{poly}(d, 1/\epsilon)/\delta)$.

Here the additional $d$ comes from the logarithm of the size of $\epsilon$-net. We can further choose $\epsilon = \frac{d}{nd_m}$ and absorb $\epsilon \left(1 + \sqrt{\frac{d \epsilon}{nd_m}}\right)$ into $\frac{d \epsilon}{nd_m}$, then apply the analysis for tabular MDP and replace $S$ with $d$ to obtain the desired result, hence conclude the proof.

**D Step-by-Step Solving for $\Delta_1$ and $v^{\pi^*} - \hat{v}^{\pi^*}$**

**D.1 Explicit Bound for $\Delta_1$**

Notice that, for some absolute constant $c$,

$$\Delta_1 \leq c \sqrt{\frac{l}{Kd_m} \left( \frac{l}{Kd_m} + \Delta_1 + v^{\pi} \right)} + \frac{l}{Kd_m},$$

which means

$$c\Delta_1 \leq \sqrt{\frac{cl}{Kd_m} \left( \frac{cl}{Kd_m} + c\Delta_1 + cv^{\pi} \right)} + \frac{cl}{Kd_m},$$

$$\leq \frac{\Delta_1}{2} + \frac{cv^{\pi}}{2} + \frac{2cl}{Kd_m},$$

thus we have that

$$\Delta_1 \leq v^{\pi} + \frac{4l}{Kd_m}.$$  

Substitute back, suppose $N = \tilde{\Omega} \left( \frac{1}{\pi_m} \right)$, then with Assumption 1, we have that

$$\Delta_1 \leq c \sqrt{\frac{l}{Kd_m} \left( \frac{5l}{Kd_m} + 2v^{\pi} \right)} + \frac{l}{Kd_m} = O \left( \sqrt{\frac{l}{Kd_m} + \frac{l}{Kd_m}} \right).$$

**D.2 Explicit bound for $v^{\pi^*} - \hat{v}^{\pi^*}$**

Notice that, for some absolute constant $c$, we have that

$$\Delta_2 \leq c \sqrt{\frac{S_l}{Kd_m} \left( \frac{S_l}{Kd_m} + v^{\pi^*} - v^{\pi^*} + \Delta_2 \right)} + \frac{S_l}{Kd_m},$$

which means

$$c\Delta_2 \leq \sqrt{\frac{cS_l}{Kd_m} \left( \frac{cS_l}{Kd_m} + c(v^{\pi^*} - v^{\pi^*}) + c\Delta_2 \right)} + \frac{cS_l}{Kd_m},$$

$$\leq \frac{c(v^{\pi^*} - v^{\pi^*})}{2} + \frac{c\Delta_2}{2} + \frac{2cS_l}{Kd_m},$$

thus we have that

$$\Delta_2 \leq (v^{\pi^*} - v^{\pi^*}) + \frac{4S_l}{Kd_m}.$$  

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We then substitute back, and know that
\[ \Delta^2 \leq c \sqrt{\frac{S_t}{Kd_m}} \left( 2 (v^{\pi} - \hat{v}^{\pi}) + \frac{5S_t}{Kd_m} \right) + \frac{S_t}{Kd_m}. \]

Furthermore, for another absolute constant \( c' \), we have that
\[
\begin{align*}
v^{\pi} - \hat{v}^{\pi} &\leq c' \left( \sqrt{\frac{t}{Kd_m}} + \frac{c}{Kd_m} \right) + \Delta^2 \\
&\leq c' \sqrt{\frac{t}{Kd_m}} + c \sqrt{\frac{S_t}{Kd_m}} \left( 2 (v^{\pi} - \hat{v}^{\pi}) + \frac{5S_t}{Kd_m} \right) + (c' + 1) \frac{S_t}{Kd_m},
\end{align*}
\]
which means
\[
\left( \sqrt{2 (v^{\pi} - \hat{v}^{\pi}) + \frac{5S_t}{Kd_m}} - c \sqrt{\frac{S_t}{Kd_m}} \right)^2 \leq 2 c' \sqrt{\frac{t}{Kd_m}} + \frac{(c^2 + 2c' + 2)S_t}{Kd_m},
\]
that can be translated to the bound
\[
\begin{align*}
v^{\pi} - \hat{v}^{\pi} &\leq \left( \sqrt{2 c' \sqrt{\frac{t}{Kd_m}} + \frac{(c^2 + 2c' + 2)S_t}{Kd_m}} \right)^2 - \frac{5S_t}{Kd_m} \\
&\leq 2 \left( \sqrt{2 c' \sqrt{\frac{t}{Kd_m}} + \frac{(c^2 + 2c' + 2)S_t}{Kd_m}} \right)^2 \\
&= O \left( \sqrt{\frac{t}{Kd_m}} + \frac{S_t}{Kd_m} \right)
\end{align*}
\]
where we use \( \sqrt{a} + \sqrt{b} \leq \sqrt{2(a + b)}. \)

E Proof of Technical Lemmas

E.1 Bernstein’s Inequality

Lemma 10 (Bernstein’s Inequality). Let \( \{X_i\}_{i=1}^n \) be i.i.d random variables from \( X \) with values bounded in \([0, 1]\), then with probability at least \( 1 - \delta \), we have that
\[
\left| \sum_{i=1}^n X_i - \mathbb{E}[X] \right| \leq \sqrt{\frac{2 \text{Var}(X) \log \frac{2}{\delta}}{n}} + \frac{\log \frac{2}{\delta}}{3n},
\]
where \( \text{Var}(X) \) is the variance of \( X \).

For the proof of Bernstein’s inequality, we refer the interested reader to Wainwright [50].
E.2 Proof of Lemma 1

Proof. We have that

\[
\sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h^2(s, a) \text{Var}_{P(s, a)} \left( V_{h+1}(s')^{2^i} \right)
\]

\[
= \sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h^2(s, a) \left[ \sum_{s'} P(s'|s, a)V_{h+1}(s')^{2^i} - \left( \sum_{s'} P(s'|s, a)V_{h+1}(s') \right)^2 \right]
\]

\[
= \sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h^2(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}(s')^{2^i} + V_h(s) - \left( \sum_{s'} P(s'|s, a)V_{h+1}(s') \right)^2 \right]
\]

\[
- \sum_{s} \mu(s) V_1(s)^{2^i+1}
\]

\[
\leq \sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h^2(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}(s')^{2^i} \right]
\]

\[
+ \sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h^2(s, a) \left[ V_h(s)^{2^i+1} - \left( \sum_{s'} P(s'|s, a)V_{h+1}(s') \right)^2 \right]
\]

\[
\leq \sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h^2(s, a) \left[ \sum_{s'} \left( P(s'|s, a) - \hat{P}(s'|s, a) \right) V_{h+1}(s')^{2^i+1} \right]
\]

\[
+ 2^{i+1} \sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h^2(s, a) \left[ V_h(s) - \left( \sum_{s'} P(s'|s, a)V_{h+1}(s') \right) \right],
\]

where in the second step we use the fact that

\[
\sum_{s, a} \hat{\xi}_h^2(s, a) \left[ \sum_{s'} \hat{P}(s'|s, a)V_{h+1}(s')^{2^i+1} \right]
\]

\[
= \sum_{s'} \left[ \sum_{s, a} \hat{\xi}_h^2(s, a) \hat{P}(s'|s, a) \right] V_{h+1}(s')^{2^i+1}
\]

\[
= \sum_{s'} \hat{\xi}_h^2(s') V_{h+1}(s')^{2^i+1}
\]

\[
= \sum_{s', a'} \hat{\xi}_h^2(s', a') V_{h+1}(s')^{2^i+1}.
\]

We drop the \( \sum_{s} \mu(s)V_1^2(s)^{2^i+1} \) and use the convexity of \( x^{2^i} \) in the third step, and the last step is indicated by the assumption that \( V_h(s) \leq 1 \), \( \forall h \in [H], s \in S \).

Moreover, we have that

\[
\sum_{h \in [H]} \sum_{s, a} \hat{\xi}_h^2(s, a) \left[ \sum_{s'} V_h(s) - \sum_{s'} \hat{P}(s'|s, a)V_{h+1}(s') \right]
\]

\[
= \sum_{h \in [H]} \sum_{s} \hat{\xi}_h^2(s)V_h(s) - \sum_{h \in [H] \setminus \{1\}} \sum_{s} \hat{\xi}_h^2(s)V_h(s)
\]

\[
= \sum_{s} \mu(s) V_1(s).
\]

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So we can conclude that

\[
\sum_{h \in H} \sum_{s,a} \hat{\xi}_h(s,a) \left[ V_h(s') - \left( \sum_{s'} P(s'|s,a) V_{h+1}(s') \right) \right] \\
= \sum_{h \in H} \sum_{s,a} \hat{\xi}_h(s,a) \sum_{s'} \left[ \hat{P}(s'|s,a) - P(s'|s,a) \right] V_{h+1}(s') + \sum_{s} \mu(s)V_1(s),
\]

thus finish the proof. \(\square\)

### E.3 Proof of Lemma 2

**Proof.**

\[
\left| \sum_{s'} \left( P(s'|s,a) - \hat{P}(s'|s,a) \right) V_h(s') \right| \\
= \sum_{s'} \left( P(s'|s,a) - \hat{P}(s'|s,a) \right) \left( V_h(s') - \sum_{s'} P(s'|s,a) V_h(s') \right) \\
\leq \sum_{s'} \sqrt{\frac{P(s'|s,a)\iota}{n(s,a)}} \left| V_h(s') - \sum_{s'} P(s'|s,a) V_h(s') \right| + \frac{S\iota}{n(s,a)} \\
\leq \sqrt{\text{Var}_{P(s,a)}(V_h(s'))} \iota + \frac{S\iota}{n(s,a)},
\]

where the first equality is due to the fact that \(\sum_{s'} P(s'|s,a) = \sum_{s'} \hat{P}(s'|s,a) = 1\), the second inequality is due to Bernstein’s inequality on each \(s'\) and Assumption 1, and the last inequality holds by Cauchy-Schwarz inequality. \(\square\)

### E.4 Proof of Lemma 3

**Proof.** Lemma 3 have been shown in Dann et al. [51, Lemma E.15]. Here we include the proof for completeness.

\[
v^\pi - \tilde{v}^\pi \\
= \sum_{s,a} \hat{\xi}_1^\pi(s,a)(Q^\pi_1(s,a) - \hat{Q}_1^\pi(s,a)) \\
= \sum_{s,a} \hat{\xi}_1^\pi(s,a) \left( r(s,a) - \tilde{r}(s,a) + \sum_{s'} [P(s'|s,a)V_2^\pi(s')] - \sum_{s'} [\hat{P}(s'|s,a)\hat{V}_2^\pi(s')] \right) \\
= \sum_{s,a} \hat{\xi}_1^\pi(s,a) \left( r(s,a) - \tilde{r}(s,a) + \sum_{s'} \left( [P(s'|s,a) - \hat{P}(s'|s,a)]V_2^\pi(s') \right) \right) + \sum_{s} \hat{\xi}_2^\pi(s)(V_2^\pi(s) - \hat{V}_2^\pi(s)) \\
= \cdots \\
= \sum_{h \in [H]} \sum_{s,a} \hat{\xi}_h^\pi(s,a) \left( r(s,a) - \tilde{r}(s,a) + \sum_{s'} \left( [P(s'|s,a) - \hat{P}(s'|s,a)]V_{h+1}^\pi(s') \right) \right) \\
\square
E.5 Proof of Lemma 4

Proof. By Bernstein’s inequality, we know that with high probability,
\[
\left| \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) [r(s,a) - \hat{r}(s,a)] \right| 
\leq \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) \left[ \sqrt{\frac{r(s,a)}{n(s,a)}} + \frac{\lambda}{n(s,a)} \right] 
\leq \sqrt{\frac{\lambda}{Kd_m}} \sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) r(s,a) + \frac{\lambda}{Kd_m},
\]
where the first inequality use the fact that \( r(s,a) \leq 1 \) and the second inequality is Cauchy-Schwarz inequality associated with Assumption 2. With Assumption 1, we know that
\[
\sum_{h \in [H]} \sum_{s,a} \xi_h^\pi(s,a) r(s,a) \leq 1,
\]
as each trajectory that can generated by \( \hat{M} \) can be generated by \( M \).

E.6 Proof of Lemma 5

Proof. Lemma 5 is similar to the Lemma 11 in Zhang et al. [27], and here we provide a simplified proof. Define
\[
i_0 := \max \{ i \mid 2^{2i} \lambda_1^2 \leq \lambda_1 \mathbb{V}(i) \} \cup \{0\},
\]
which is the largest integer satisfies the inequality \( 2^{2i} \lambda_1^2 \leq \lambda_1 \mathbb{V}(i) \). We will make the recursion at most \( i_0 \) times. As \( \mathbb{V}(i) \) is upper bounded by \( H \), we know \( i_0 = O \left( \log \frac{H \lambda_1}{\lambda_2^2} \right) \). If \( i_0 = 0 \), we have that
\[
\mathbb{V}(1) \leq \frac{4 \lambda_2^2}{\lambda_1},
\]
Otherwise,
\[
2^{2i_0} \lambda_2^2 \lesssim 2 \lambda_1 \mathbb{V}(i_0) 
\lesssim \lambda_1 \left( \sqrt{\lambda_1 \mathbb{V}(i_0 + 1)} + \lambda_1 + 2^{i_0+1} \lambda_2 \right) 
\lesssim \lambda_1 \left( \sqrt{2^{2i_0+2} \lambda_2^2 + \lambda_1 + 2^{i_0+1} \lambda_2} \right) 
= \lambda_1 (2^{i_0+2} \lambda_2 + \lambda_1),
\]
we have that
\[
(2^{i_0} \lambda_2 - 2 \lambda_1)^2 \leq 5 \lambda_1^2,
\]
which means
\[
2^{i_0} \lambda_2 \leq (\sqrt{5} + 2) \lambda_1,
\]
so \( \forall i \leq i_0 \), we have that
\[
\mathbb{V}(i) \leq \sqrt{\lambda_1 \mathbb{V}(i + 1)} + (\sqrt{5} + 3) \lambda_1.
\]
As
\[
\mathbb{V}(i_0) \leq 2^{i_0+2} \lambda_2 + \lambda_1 \leq (4\sqrt{5} + 9) \lambda_1,
\]
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and when $V(i + 1) \geq \left( \frac{1}{2} + \sqrt{\frac{\sqrt{5} + 13}{4}} \right) \lambda_1$, we have $V(i) \leq V(i + 1)$, when $V(i + 1) \leq \left( \frac{1}{2} + \sqrt{\frac{\sqrt{5} + 13}{4}} \right) \lambda_1$, $V(i) \leq \left( \frac{1}{2} + \sqrt{\frac{\sqrt{5} + 13}{4}} \right) \lambda_1$. So $V(1) \leq (4\sqrt{5} + 9)\lambda_1$. Combine the two cases, we have that

$$V(1) \leq \max \left\{ \frac{4\lambda_2^2}{\lambda_1}, (4\sqrt{5} + 9)\lambda_1 \right\},$$

which means

$$V(0) \leq \max \left\{ 4\lambda_2 + \lambda_1, \left( \sqrt{4\sqrt{5} + 9} + 1 \right) \lambda_1 + 2\lambda_2 \right\} \leq 6(\lambda_1 + \lambda_2),$$

which concludes the proof. $\square$

F Further Discussion

F.1 Time-Homogenous vs. Time-Inhomogenous

F.1.1 Offline Policy Evaluation

One can notice that the lower bound of offline policy evaluation under time-homogenous setting (shown above) and time-inhomogenous setting (that can be simplified from the Cramer-Rao lower bound of [13]) are identical. It can be surprising at the first glance, as one common belief for reinforcement learning is that the time-homogenous MDP should be easier than the time-inhomogenous MDP. However, we remark that, this argument does not hold for the policy evaluation. This is due to the fact that in time-homogenous setting, as the error from the estimation of transition will be accumulated along the horizon, we need $\Omega(H)$ samples for each transition to make sure that the accumulated error should be $O(1)$. However, in the time-inhomogenous setting, the error from the estimation of each level transition is probably not accumulated, which means we only need $\Omega(1)$ samples for each transition at each level to make sure the final error to be $O(1)$. This can also be seen in the analysis of offline policy evaluation under time-inhomogenous setting in [19, 24], where the authors decouple the error from each sample $(s_h, a_h, s_{h+1})$, which forms a martingale difference sequence with each step variance $O(1)$ that can then apply the Freedman’s inequality to obtain a tight bound. Under time-inhomogenous setting, we notice that the error from each sample $(s, a, s')$ can only form a martingale difference sequence with each step variance $O(H)$, which will not lead to a tighter bound. We want to emphasize that the results in [19, 24] does not directly indicate results in this paper.

Moreover, we notice that the analysis in [19, 24] can be translated to a value-dependent bound and thus can be horizon-free under Assumption 1, with the following lemma:

Lemma 11. Under Assumption 1, we have that

$$\sum_{h \in [H]} \sum_{s,a} \xi^\pi_h(s, a) \text{Var}(r(s, a) + V^\pi_{h+1}(s')) \leq 3v^\pi,$$

where the term at the left hand side is exactly the variance term of MIS estimator considered in [19, Lemma 3.4]
We remark that [28] provides a value-dependent bound for the policy evaluation under infinite-horizon where we use the following fact:

This shows that the offline policy evaluation under time-inhomogeneous setting is also horizon-free, which also matches the intuition that if the density ratio can be lower bounded, then we can construct a trajectory-wise importance sampling (IS) estimator that only depends on the number of episodes. Notice that, in [19, 24], the higher order term has an additional \( \sqrt{SA} \) factor in sample complexity, thus can never be horizon-free. We also want to remark that, this is due to the time-homogeneous nature of our setting, which makes \( V_{h}^{\pi^{*}} \) heavily depends on \( \hat{P} \). On the other hand, in time-inhomogeneous setting, \( V_{h}^{\pi^{*}} \) only depends on \( \hat{P}_{k}(s, a) \) for \( k > h \), thus can directly apply Bernstein’s inequality when bounding the term \( \left( \hat{P}_{h}(s, a) - P_{h}(s, a) \right) V_{h+1}^{\pi^{*}} \), which will not introduce additional \( S \) factor. The best known result for finite horizon time-homogenous MDP [27] also has this additional \( S \) factor, and how to eliminate this additional \( S \) factor remains an open problem.

### F.2 Finite-Horizon vs. Infinite-Horizon

#### F.2.1 Offline Policy Evaluation

We remark that [28] provides a value-dependent bound for the policy evaluation under infinite-horizon generative model setting that accommodates full range of \( \epsilon \). We obtain the similar results in finite-
horizon setting that can accommodate full range of $\epsilon$, however, with a different and probably simpler analysis.

F.2.2 Offline Policy Optimization

[28] also provides a minimax-optimal sample complexity bound up to logarithmic factors for policy optimization with generative model that accommodates full range of $\epsilon$. We notice that such kinds of analysis cannot be directly applied to the finite-horizon setting, as their “absorbing MDP” technique cannot be directly applied to the finite-horizon MDP, due to the difference of time-homogenous value function in infinite-horizon setting and time-inhomogenous value function in finite-horizon setting, which has been pointed out by [1]. And thus most of the existing work does not provide a sample complexity bound that can match the lower bound. To the best of our knowledge, our work first provide a sample complexity bound that match the lower bound up to logarithmic factors and an high-order term.

F.3 With General Function Approximation

There are also works considering the offline policy optimization with general function approximation under different kinds of function class assumption like realizability and completeness [e.g. 52–54], which generally do not imply tight bounds under certain scenarios like e.g. tabular MDP. We leave the extension to general function approximation as future work.

F.4 Without Sufficient Exploratory Data

Recently, [45, 46] also introduces another perspective on performing offline policy optimization within a local policy set when the offline data is not sufficient exploratory, which is different from the global policy optimization we consider here. We want to emphasize that, if we want to approach the global optimal policy, our assumption on good data coverage i.e. Assumption 2 is necessary. Otherwise, we will suffer from the error from under-explored state-action pair, as Theorem 2 and Theorem 4 suggests.