Abstract

There has been growing progress on theoretical analyses for provably efficient learning in MDPs with linear function approximation, but much of the existing work has made strong assumptions to enable exploration by conventional exploration frameworks. Typically these assumptions are stronger than what is needed to find good solutions in the batch setting. In this work, we show how under a more standard notion of low inherent Bellman error, typically employed in least-square value iteration-style algorithms, we can provide strong PAC guarantees on learning a near optimal value function provided that the linear space is sufficiently “explorable”. We present a computationally tractable algorithm for the reward-free setting and show how it can be used to learn a near optimal policy for any (linear) reward function, which is revealed only once learning has completed. If this reward function is also estimated from the samples gathered during pure exploration, our results also provide same-order PAC guarantees on the performance of the resulting policy for this setting.

1 Introduction

Reinforcement learning (RL) aims to solve complex multi-step decision problems with stochastic outcomes framed as a Markov decision process (MDP). RL algorithms often need to explore large state and action spaces where function approximations become necessity. In this work, we focus on exploration with linear predictors for the action value function, which can be quite expressive [Sutton and Barto, 2018].

Existing guarantees for linear value functions Exploration has been widely studied in the tabular setting [Azar et al., 2017, Zanette and Brunskill, 2019, Efroni et al., 2019, Jin et al., 2018, Dann et al., 2019], but obtaining formal guarantees for exploration with function approximation appears to be a challenge even in the linear case. The minimal necessary and sufficient conditions to reliably learn a linear predictor are not fully understood even with access to a generative model [Du et al., 2019b]. We know that when the best policy is unique and the predictor is sufficiently accurate it can be identified [Du et al., 2019c, 2020], but in general we are interested in finding only near-optimal policies using potentially misspecified approximators.

To achieve this goal, several ideas from tabular exploration and linear bandits [Lattimore and Szepesvári, 2020] have been combined to obtain provably efficient algorithms in low-rank MDPs [Yang and Wang, 2020, Zanette et al., 2020a, Jin et al., 2020b] and their extension [Wang et al., 2019, 2020b]. We shall identify the core assumption of the above works as optimistic closure: all these
settings assume the Bellman operator maps any value function of the learner to a low-dimensional space $Q$ that the learner knows. When this property holds, we can add exploration bonuses because by assumption the Bellman operator maps the agent’s optimistically modified value function back to $Q$, which the algorithm can represent and use to propagate the optimism and drive the exploration. However, the optimistic closure is put as an assumption to enable exploration using traditional methods, but is stronger that what is typically required in the batch setting.

Towards batch assumptions This work is motivated by the desire to have exploration algorithms that we can deploy under more mainstream assumptions, ideally when we can apply well-known batch procedures like least square policy iteration (LSP1) [Lagoudakis and Parr 2003], and least square value iteration (LSVI) [Munos 2005].

LSP1 has convergence guarantees when the action value function of all policies can be approximated with a linear architecture [Lazaric et al., 2012], i.e., $Q^n$ is linear for all $\pi$; in this setting, Lattimore and Szepesvári [2020] recently use a design-of-experiments procedure from the bandit literature to obtain a provably efficient algorithm for finding a near optimal policy, but they need access to a generative model. LSVI, another popular batch algorithm, requires low inherent Bellman error [Munos and Szepesvári, 2008] [Chen and Jiang, 2019]. In this setting, Zanette et al. [2020b] present a near-optimal (with respect to noise and misspecification) regret-minimizing algorithm that operates online, but a computationally tractable implementation is not known. It is worth noting that both settings are more general than linear MDPs [Zanette et al., 2020b].

A separate line of research is investigating settings with low Bellman rank [Jiang et al., 2017] which was found to be a suitable measure of the learnability of many complex reinforcement learning problems. The notion of Bellman rank extends well beyond the linear setting.

The lack of computational tractability in the setting of Zanette et al. [2020b] and in the setting with low Bellman rank [Jiang et al., 2017] and of a proper online algorithm in [Lattimore and Szepesvári, 2020] highlight the hardness of these very general settings which do not posit additional assumptions on the linear value function class $Q$ beyond what is required in the batch setting.

Reward-free exploration We tackle the problem of designing an exploration algorithm using batch assumptions by adopting a pure exploration perspective: our algorithm can return a near optimal policy for any linear reward function that is revealed after an initial learning phase. It is therefore a probably approximately correct (PAC) algorithm. Reward-free exploration has been investigated in the tabular setting with an end-to-end algorithm [Jin et al., 2020a]. Hazan et al. [2018] design an algorithm for a more general setting through oracles that also recovers guarantees in the tabular domains. Others [Du et al., 2019a, Misra et al., 2020] also adopt the pure exploration perspective assuming a small but unobservable state space. More recently, reward free exploration has gained attention in the tabular setting [Kaufmann et al., 2020a, Tarbouriech et al., Ménard et al., 2020] as well as the context of function approximation Wainwright [2019], Agarwal et al. [2020].

Contribution This work makes two contributions. It presents a statistically and computationally efficient online PAC algorithm to learn a near-optimal policy 1) for the setting with low inherent Bellman error [Munos and Szepesvári, 2008] and 2) for reward-free exploration in the same setting.

From a technical standpoint, 1) implies we cannot use traditional exploration methodologies and 2) implies we cannot learn the full dynamics, which would require estimating all state-action-state transition models. Both goals are accomplished by driving exploration by approximating G-optimal experimental design [Lattimore and Szepesvári, 2020] in online reinforcement learning through randomization. Our algorithm returns a dataset of well chosen state-action-transition triplets, such that invoking the LSVI algorithm on that dataset (with a chosen reward function) returns a near optimal policy on the MDP with that reward function.

2 Preliminaries and Intuition

We consider an undiscounted $H$-horizon MDP [Puterman, 1994] $M = (S, A, p, r, H)$ defined by a possibly infinite state space $S$ and action space $A$. For every $t \in [H] = \{1, \ldots, H\}$ and state-action pair $(s, a)$, we have a reward function $r_t(s, a)$ and a transition kernel $p_t(\cdot | s, a)$ over the next state. A policy $\pi$ maps a $(s, a, t)$ triplet to an action and defines a reward-dependent action value function
We therefore define the inherent Bellman error directly in the transition operator $T$. This condition is more general than the low rank MDP setting or optimistic closure [Yang and Wang, 2020]. G-optimal design is a procedure [Kiefer and Wolfowitz, 1960] through randomization; the connection is briefly outlined in section 5.5.

Infeasible. Thus we relax this program by finding a policy that in most of the episodes makes at least $s, a$ and a parameter $θ ∈ Q$, which we assume to be the Euclidean ball with unit radius $B_l = \{θ ∈ \mathbb{R}^{d_l} \mid \|θ\|_2 ≤ 1\}$. This defines the value functional spaces as

$$Q_t = \{Q_t(s, a) = φ_t(s, a)\top θ, θ ∈ B_l\}, \quad V_t = \{V_t : \max_{a} φ_t(s, a)\top θ, θ ∈ B_l\}.$$

### Inherent Bellman error

The inherent Bellman error condition is typically employed in the analysis of LSVI [Munos and Szepesvari, 2008; Chen and Jiang, 2019]. It measures the closure of the prescribed functional space $Q$ with respect to the Bellman operator $T$, i.e., the distance of $TQ$ from $Q$ provided that $Q ⊆ \mathbb{R}$. In other words, low inherent Bellman error ensures that if we start with an action value function in $Q$ then we approximately remain in the space after performance of the Bellman update. For finite horizon MDP we can define the inherent Bellman error as:

$$\max_{Q_{t+1} ∈ Q_{t, 1}} \min_{Q_t ∈ \mathbb{Q}_t} \max_{(s, a)} \|Q_t - T(Q_t + 1)\|\|s, a\|. \quad (1)$$

When linear function approximations are used and the inherent Bellman error is zero, we are in a setting of low Bellman rank [Jiang et al., 2017], where the Bellman rank is the feature dimensionality. This condition is more general than the low rank MDP setting or optimistic closure [Yang and Wang, 2020; Jin et al., 2020b; Zanette et al., 2020a; Wang et al., 2019]; for a discussion of this see [Zanette et al., 2020b].

### Model-free reward-free learning

In the absence of reward signal, how should $Q_t$ look like? Define the reward-free Bellman operator $\mathcal{T}^{P}(Q_{t+1})(s, a) = \mathbb{E}_{s' \sim π_t(s, a)} \max_{a'} Q_{t+1}(s', a')$. It is essentially equivalent to measure the Bellman error either on the full Bellman operator $T$ or directly on the dynamics $T^{P}$ when the reward function is linear (see proposition 2 of Zanette et al., 2020b). We therefore define the inherent Bellman error directly in the transition operator $T^{P}$:

**Definition 1** (Inherent Bellman Error),

$$\mathcal{T}(Q_t, Q_t^{P}) \overset{\text{def}}{=} \max_{Q_{t+1} ∈ \mathbb{Q}_{t+1}} \min_{Q_t ∈ \mathbb{Q}_t} \max_{(s, a)} |Q_t - T^{P}(Q_{t+1})|\|s, a\|. \quad (2)$$

### Approximating G-optimal design

G-optimal design is a procedure [Kiefer and Wolfowitz, 1960] that identifies an appropriate sequence of features $φ_1, …, φ_n$ to probe to form the design matrix $Σ = \sum_{i=1}^{n} φ_i φ_i\top$ in order to uniformly reduce the maximum “uncertainty” over all the features as measured by $\max_{φ} \|φ\|_{Σ^{-1}}$, see appendix C. This principle has recently been applied to RL with a generative model [Lattimore and Szepesvari, 2020] to find a near optimal policy.

However, the basic idea has the following drawbacks in RL: 1) it requires access to a generative model; 2) it is prohibitively expensive as it needs to examine all the features across the full state-action space before identifying what features to probe. This work addresses these 2 drawbacks in reinforcement learning by doing two successive approximations to G-optimal design. The first approximation would be compute and follow the policy $π$ (different in every rollout) that leads to an expected feature $\overline{φ}_n$ in the most uncertain direction $\overline{σ}_n$ (i.e., the direction where we have the least amount of data). This solves problem 1 and 3 above, but unfortunately it turns out that computing such $π$ is computationally infeasible. Thus we relax this program by finding a policy that in most of the episodes makes at least some progress in the most uncertain direction, thereby addressing point 2 above. This is achieved through randomization; the connection is briefly outlined in section 5.5.

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1 This is an approximation to $G$-optimal design, because $π$ here is the policy that leads to the most uncertain direction $\overline{σ}_n$ rather than to the direction that reduces the uncertainty the most.
3 Algorithm

Moving from the high-level intuition to the actual algorithm requires some justification, which is left to section 5. Here instead we give few remarks about algorithm [1] first, the algorithm proceeds in phases \( p = 1, 2, \ldots \) and in each phase it focuses on learning in phase 2 it learns the dynamics at timestep 2).

**Algorithm 1** Forward Reward Agnostic Navigation with Confidence by Injecting Stochasticity (FRANCIS)

1. **Inputs**: failure probability \( \delta \in [0, 1] \), target precision \( \epsilon > 0 \), feature map \( \phi \).
2. Initialize \( \Sigma_{t=1} = \lambda I, \hat{\theta}_t = 0, \forall t \in [H], D = \emptyset; \) set \( c_0, c_r, c_\theta \in \mathbb{R} \) (see appendix), \( \lambda = 1 \).
3. for phase \( p = 1, 2, \ldots, H \) do
4. \( k = 1 \), set \( \sigma = \sigma_{\text{start}} \triangleq c_\sigma / (d_p \ln(d_p \alpha)) \).
5. while \( \sigma < c_\alpha H^2 (d_p + d_{p+1}) \ln(d_p \alpha) \) do
6. for \( i = 1, 2, \ldots , c_\alpha d_p^2 \sigma \) do
7. \( k = k + 1 \), receive starting state \( s_1 \sim \rho \).
8. \( \xi_p \sim \mathcal{N}(0, \Sigma_{p,k}^{-1}) \); \( R_{p,s,a} \triangleq \phi_p(s,a)^\top \xi_p \).
9. \( \pi \leftarrow \text{LSVI}(p, \rho, \hat{D}) \).
10. Run \( \pi \); \( \hat{D} \leftarrow \hat{D} \cup \{s_{p,k}, a_{p,k}, s_{p,k+1}\} \).
11. \( \phi_{p,k} \triangleq \phi_p(s_{p,k}, a_{p,k}); \Sigma_{p,k+1} \leftarrow \Sigma_{p,k} + \phi_{p,k}^\top \phi_{p,k} \).
12. end for
13. \( \sigma \leftarrow 2\sigma \).
14. end while
15. end for
16. return \( D \).

4 Main Result

Before presenting the main result is useful to define the average feature \( \bar{\varphi}_{\pi,t} = \mathbb{E}_{x_t \sim \pi} \varphi_t(x_t, \pi_t(x_t)) \) encountered at timestep \( t \) upon following a certain policy \( \pi \). In addition, we need a way to measure how “explorable” the space is, i.e., how easy it is to collect information in a given direction of the feature space using an appropriate policy. The explorability coefficient \( \nu \) measures how much we can align the expected feature \( \bar{\varphi}_{\pi,t} \) with the most challenging direction \( \theta \) to explore even if we use the best policy \( \pi \) for the task (i.e., the policy that maximizes this alignment). It measures how difficult it is to explore the most challenging direction, even if we use the best (and usually unknown) policy to do so. This is similar to a diameter condition in the work of Jaksch et al. [2010] in the features space, but different from ergodicity, which ensures that sufficient information can be collected by any policy. It is similar to the reachability parameter of Du et al. [2019a] and Misra et al. [2020], but our condition concerns the features rather than the state space and is unavoidable in certain settings (see discussion after the main theorem).

**Definition 2** (Explorability). \( \nu_t \triangleq \min_{\|\theta\|_2 = 1} \max_{\pi} \|\varphi_{\pi,t}^\top \theta\|; \quad \nu_{\text{min}} = \min_{t \in [H]} \nu_t. \)

**Theorem 4.1.** Assume \( \|\varphi_t(s,a)\|_2 \leq 1 \) and set \( \epsilon \) to satisfy \( \epsilon \geq \hat{O}(d_l H^2 Q_0 + Q_{t+1}) \) and \( \epsilon \leq \hat{O}(\nu_{\text{min}} / \sqrt{T}) \) for all \( t \in [H] \). FRANCIS terminates after \( \hat{O} \left( H^2 \sum_{t=1}^{H} \frac{d_t^2 + d_{t+1}^2}{\epsilon^2} + 1 \right) \) episodes.

Fix a reward function \( r_t(\cdot, \cdot) \) such that each state-action-successor state \( (s_{tk}, a_{tk}, s_{tk+1}^t) \) triplet in \( D \) (where \( t \in [H] \) and \( k \) is the episode index in phase \( t \)) is augmented with a reward \( r_{tk} = r_t(s_{tk}, a_{tk}) \).
batch with linear function approximation. This work provides one of the first end-to-end algorithms for provably efficient reward-free exploration with low Bellman rank (the Bellman rank being the dimensionality of the features). In addition, this setting allows us to model MDPs where the reward function satisfies (on the MDP with some parameters \( \theta_1^\ast \in \mathbb{R}^{d_1}, \ldots, \theta_H^\ast \in \mathbb{R}^{d_H} \))

\[
\forall (s, a, t) \quad \| \theta_t^\ast \|_2 \leq \frac{1}{H}, \quad r_t(s, a) = \phi_t(s, a)^\top \theta_t^\ast
\]

then with probability at least \( 1 - \delta \) the policy \( \pi \) returned by LSVI using the augmented dataset \( \mathcal{D} \) satisfies (on the MDP with \( r_t(\cdot, \cdot) \) as reward function)

\[
\mathbb{E}_{x_1 \sim \rho}(V_1^\ast - V_1)(x_1) \leq \epsilon.
\]

The full statement is reported in appendix D.6. The reward function \( r_t(\cdot, \cdot) \) could even be adaptively chosen after the algorithm has terminated. If the reward function is estimated from data then the theorem immediately gives same-order guarantees as a corollary. The dynamics error \( O(d_H H^2 Q_t, Q_{t+1}) \) is contained in \( \epsilon \).

The setting allows us to model MDPs where the reward function \( r_t(\cdot, \cdot) \) and \( V_1^\ast \) are in \([0,1]\). When applied to MDPs with rewards in \([0,1]\) (and value functions in \([0,H]\)), the input and output should be rescaled and the number of episodes to \( \epsilon \) accuracy should be multiplied by \( H^2 \).

The significance of the result lies in the fact that this is the first statistically and computationally efficient PAC algorithm for the setting of low inherent Bellman error; this is special case of the setting with low Bellman rank (the Bellman rank being the dimensionality of the features). In addition, this work provides one of the first end-to-end algorithms for provably efficient reward-free exploration with linear function approximation.

In table 1 we describe our relation with few relevant papers in the field. The purpose of the comparison is not to list the pros and cons of each work with respect to one another, as these works all operate under different assumptions, but rather to highlight what is achievable in different settings.

<table>
<thead>
<tr>
<th>Online?</th>
<th>Reward-agnostic?</th>
<th>Need optimistic closure?</th>
<th># episodes</th>
<th># computations</th>
</tr>
</thead>
<tbody>
<tr>
<td>This work</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
<td>( \frac{d^3 H^5}{\epsilon^2} )</td>
</tr>
<tr>
<td>G-optimal design + LSVI</td>
<td>No</td>
<td>Yes</td>
<td>No</td>
<td>( \frac{d^3 H^5}{\epsilon^2} )</td>
</tr>
<tr>
<td>[Jin et al. 2020b]</td>
<td>Yes</td>
<td>No</td>
<td>No</td>
<td>( \frac{d^3 H^5}{\epsilon^2} )</td>
</tr>
<tr>
<td>[Jiang et al. 2017]</td>
<td>Yes</td>
<td>No</td>
<td>Yes</td>
<td>( \frac{d^3 H^5}{\epsilon^2} )</td>
</tr>
<tr>
<td>[Jin et al. 2020a]</td>
<td>Yes</td>
<td>Yes</td>
<td>(tabular)</td>
<td>( \frac{H^5 S^2 A}{\epsilon^3} )</td>
</tr>
<tr>
<td>[Wang et al. 2020a]</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>( \frac{d^3 H^5}{\epsilon^2} )</td>
</tr>
</tbody>
</table>

Table 1: We consider the number of episodes to learn an \( \epsilon \)-optimal policy. We assume \( r \in [0,1] \) and \( Q^\ast \in [0,H] \), and rescale the results to hold in this setting. We neglect misspecification for all works. The column “optimistic closure” refers to the assumption that the Bellman operator projects any value function into a prescribed space (notably, low-rank MDPs of [Jin et al. 2020b]). For our work we assume \( \epsilon = \Omega(\nu_{\min}/\sqrt{d}) \). We recall that if an algorithm has regret \( \sqrt{K} \), with \( K \) the number of episodes then we can extract a PAC algorithm to return an \( \epsilon \)-optimal policy in \( \frac{d}{\epsilon^2} \) episodes. We evaluate [Jian et al. 2017] in our setting where the Bellman rank is \( d \) (the result has an explicit dependence on the number of actions, though this could be improved in the linear setting). G-optimal design is from the paper [Lattimore and Szepesvari 2020] which operates in infinite-horizon and assuming linearity of \( Q^\ast \) for all \( \pi \), so the same idea of G-optimal design was applied to our setting to derive the result and we report the number of required samples (as opposed to the number of episodes), see appendix C. For [Jin et al. 2020a] we ignore the \( \frac{d^3 H^5}{\epsilon^2} \) lower order term.

If the reward function \( r_t(\cdot, \cdot) \) satisfies for some parameters \( \theta_1^\ast \in \mathbb{R}^{d_1}, \ldots, \theta_H^\ast \in \mathbb{R}^{d_H} \)

\[
\forall (s, a, t) \quad \| \theta_t^\ast \|_2 \leq \frac{1}{H}, \quad r_t(s, a) = \phi_t(s, a)^\top \theta_t^\ast
\]

then with probability at least \( 1 - \delta \) the policy \( \pi \) returned by LSVI using the augmented dataset \( \mathcal{D} \) satisfies (on the MDP with \( r_t(\cdot, \cdot) \) as reward function)

\[
\mathbb{E}_{x_1 \sim \rho}(V_1^\ast - V_1)(x_1) \leq \epsilon.
\]

The full statement is reported in appendix D.6. The reward function \( r_t(\cdot, \cdot) \) could even be adversarially chosen after the algorithm has terminated. If the reward function is estimated from data then the theorem immediately gives same-order guarantees as a corollary. The dynamics error \( O(d_H H^2 Q_t, Q_{t+1}) \) is contained in \( \epsilon \).

The setting allows us to model MDPs where the reward function \( r_t(\cdot, \cdot) \) and \( V_1^\ast \) are in \([0,1]\). When applied to MDPs with rewards in \([0,1]\) (and value functions in \([0,H]\)), the input and output should be rescaled and the number of episodes to \( \epsilon \) accuracy should be multiplied by \( H^2 \).

The significance of the result lies in the fact that this is the first statistically and computationally efficient PAC algorithm for the setting of low inherent Bellman error; this is special case of the setting with low Bellman rank (the Bellman rank being the dimensionality of the features). In addition, this work provides one of the first end-to-end algorithms for provably efficient reward-free exploration with linear function approximation.

In table 1 we describe our relation with few relevant papers in the field. The purpose of the comparison is not to list the pros and cons of each work with respect to one another, as these works all operate under different assumptions, but rather to highlight what is achievable in different settings.

Is small Bellman error needed? As of writing, the minimal conditions that enable provably efficient learning with function approximation are still unknown [Du et al. 2019b]. In this work we focus on small Bellman error which is a condition typically used for batch analysis of LSVI [Munos 2005, Munos and Szepesvari 2008, Chen and Jiang 2019]. What is really needed for the functioning of FRANCIS is that vanilla LSVI outputs a good solution in the limit of infinite data on different (linear) reward functions: as long as LSVI can return a near-optimal policy for the given reward function given enough data, FRANCIS can proceed with the exploration. This requirement

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2FRANCIS requires only polynomial calls to LSVI and samples from a multivariate normal, see appendix D.7.
is really minimal, because even if the best dataset \( \mathcal{D} \) is collected through G-optimal design on a generative model (instead of using FRANCIS), LSVI must anyway be able to output a good policy on the prescribed reward function.

**Is explorability needed?** Theorem 4.1 requires \( \epsilon \leq \tilde{O}(\nu_{\text{min}}/\sqrt{d_t}) \). Unfortunately, a dependence on \( \nu_{\text{min}} \) turns out to be unavoidable in the more general setting we consider in the appendix; we discuss this in more detail in appendix E, but here we give some intuition regarding the explorability requirement.

FRANCIS can operate under two separate set of assumptions, which we call implicit and explicit regularity, see definition 6 (Reward Classes) in appendix and the main result in theorem 1.

Under implicit regularity we do not put assumptions on the norm of reward parameter \( \|\theta^r\|_2 \), but only a bound on the expected value of the rewards under any policy: \( |\mathbb{E}_{x_t \sim \pi_t} \pi_t(x_t)| \leq 1 \). This representation allows us to represent very high rewards \( \geq 1 \) in hard-to-reach states. It basically controls how big the value function can get. This setting is more challenging for an agent to explore even in the tabular setting and even in the case of a single reward function. If a state is hard to reach, the reward there can be very high, and a policy that tries to go there can still have high value. Under this implicit regularity assumption, the explorability parameter would show up for tabular algorithms as well (as minimum visit probability to any state under an appropriate policy).

By contrast, under explicit regularity (which concerns the result reported in theorem 4.1), we do make the classical assumption that bounds the parameter norm \( \|\theta^r\|_2 \leq 1/H \). In this case, the lower bound no longer applies, but the proposed algorithm still requires good “explorability” to proceed. Removing this assumption is left as future work.

## 5 Technical Analysis

For the proof sketch we neglect misspecification, i.e., \( \mathcal{I}(Q_t, Q_{t+1}) = 0 \). We say that a statement holds with very high probability if the probability that it does not hold is \( \ll \delta \).

### 5.1 Analysis of LSVI, uncertainty and inductive hypothesis

FRANCIS repeatedly calls LSVI on different randomized linearly-parameterized reward functions \( R_p \) and so we need to understand how the signal propagates. Let us begin by defining an uncertainty function in episode \( i \) of phase \( p \) using the covariance matrix \( \Sigma_p = \sum_{j=1}^{t-1} \phi_{pj}^\top \phi_{pj} + I \) on the observed features \( \phi_{pj} = \phi_p(s_{pj}, a_{pj}) \) at episode \( j \) of phase \( p \):

**Definition 3** (Max Uncertainty). \( \mathcal{U}^\phi_p(\sigma) \) \( \overset{\text{def}}{=} \max_{\pi, \|\theta^r\|_2 \leq \sigma} \sqrt{\|\phi_{\pi, p}\|_{\Sigma_p}^{-1}} \).

Let \( \Sigma_t \) denote the covariance matrix in timestep \( t \) once learning in that phase has completed, and likewise denote with \( \mathcal{U}^\phi_t(\sigma) \) the final value of the program of definition 3 once learning in phase \( t \) has completed (so using \( \Sigma_t \) in the definition); let \( \sqrt{\alpha_t} = \tilde{O}(\sqrt{d_t + d_{t+1}}) \) and \( \rho_p(s, a) = \phi_p(s, a)^\top \xi_p \).

**Lemma 1** (see appendix B.4). Assume \( \|\xi_p\|_2 \leq 1 \) and \( \lambda_{\text{min}}(\Sigma_t) = \Omega(H^2 \alpha_t) \) for all \( t \in [p-1] \). Then with very high probability \( LSVI(p, R_p, \mathcal{D}) \) computes a value function \( \hat{V} \) and a policy \( \pi \) s.t.

\[
|\mathbb{E}_{x_t \sim \rho} \hat{V}_t(x_t) - \hat{\phi}_{\pi, p}^\top \xi_p| \leq \sum_{t=1}^{p-1} \sqrt{\alpha_t \|\hat{\phi}_{\pi, t}\|_{\Sigma_t}^{-1}} = \sum_{t=1}^{p-1} \mathcal{U}^\phi_t(\alpha_t) = \text{Least-Square Error}.
\]

The least-square error in the above display can be interpreted as a planning error to propagate the signal \( \xi_p \); it also appears when LSVI uses the batch dataset \( \mathcal{D} \) to find the optimal policy on a given reward function after FRANCIS has terminated, and it is the quantity we target to reduce. Since \( \alpha_t \) is constant, we need to shrink \( \|\hat{\phi}_{\pi, p}\|_{\Sigma_p}^{-1} \) over any choice of \( \pi \) as much as possible by obtaining an appropriate feature matrix \( \Sigma_t \).

\(^3\text{G-optimal design does this optimally, but requires choosing the features, which is only possible if one has access to a generative model or in a bandit problem.} \)
A final error across all timesteps of order $\epsilon$ can be achieved when the algorithm adds at most $\epsilon/H$ error at every timestep. Towards this, we define an inductive hypothesis that the algorithm has been successful up to the beginning of phase $p$ in reducing the uncertainty encoded in $U_p^*$:

**Inductive Hypothesis 1.** At the start of phase $p$ we have $\sum_{i=1}^{p-1} U_i^*(\alpha_i) \leq \frac{p-1}{p} \epsilon$.

The inductive hypothesis critically ensures that the reward signal $\xi$ can be accurately propagated backward by LSVI, enabling navigation capabilities of FRANCIS to regions of uncertainty in phase $p$ (this justifies the phased design of FRANCIS).

5.2 Overestimating the maximum uncertainty through randomization

Assuming the inductive hypothesis, we want to show how to reduce the uncertainty in timestep $p$. Similar to how optimistic algorithms overestimate the optimal value function, here $E_{x_1,\sim p} V_1(x_1) \approx \phi_{\pi,p}^T \xi_{pi}$ should overestimate the current uncertainty in episode $i$ of phase $p$ encoded in $U_{pi}^*(\alpha_i)$. This is achieved by introducing a randomized reward signal $\xi_{pi} \sim N(0, \sigma_{pi}^{-1})$ at timestep $p$.

**Lemma 2** (Uncertainty Overestimation, appendix D.2). If $\xi_{pi} \sim N(0, \sigma_{pi}^{-1})$, $U_{pi}^*(\alpha_i) = \Omega(\epsilon)$, $\|\xi_{pi}\|_2 \leq 1$ and the inductive hypothesis holds then LSVI returns with some constant probability $q \in \mathbb{R}$ a policy $\pi$ such that $\phi_{\pi,p}^T \xi_{pi} \geq U_{pi}^*(\alpha_i)$.

The proof of the above lemma uses lemma 1. The condition $U_{pi}^*(\sigma) = \Omega(\epsilon)$ is needed: if the signal $\xi_{pi}$ or uncertainty $U_{pi}^*(\sigma)$ are too small relative to $\epsilon$ then the least-square error of order $\epsilon$ that occurs in LSVI is too large relative to the signal $\xi_{pi}$, and the signal cannot be propagated backwardly.

The lemma suggests we set $\sigma = \alpha_i$ to ensure $\phi_{\pi,p}^T \xi_{pi} \geq U_{pi}^*(\alpha_i)$ with fixed probability $q \in \mathbb{R}$. Unfortunately this choice would generate a very large $\|\xi_{pi}\|_2$ which violates the condition $\|\xi_{pi}\|_2 \leq 1$. In particular, the condition $\|\xi_{pi}\|_2 \leq 1$ determines how big $\sigma$ can be.

**Lemma 3** (see appendix D.1). If $\sigma = \tilde{O}(\lambda_{min}(\Sigma_{pi})/d_p)$ and $\xi_{pi} \sim N(0, \sigma_{pi}^{-1})$ then $\|\xi_{pi}\|_2 \leq 1$ with very high probability.

Since initially $\Sigma_{p1} = I$, the above lemma determines the initial value $\sigma \approx 1/d_p \ll \alpha_p$. This implies FRANCIS won’t be able to overestimate the uncertainty $U_{pi}^*(\alpha_i)$ initially.

The solution is to have the algorithm proceed in epochs. At the end of every epoch FRANCIS ensures $U_{pi}^*(\sigma) \leq \epsilon$, and that $\lambda_{min}(\Sigma_{pi})$ is large enough that $\sigma$ can be doubled at the beginning of the next epoch.

5.3 Learning an Epoch

Using lemma 2 we can analyze what happens within an epoch when $\sigma$ is fixed (assuming $\sigma$ is appropriately chosen to ensure $\|\xi_{pi}\|_2 \leq 1$ with very high probability). We first consider the average uncertainty as a measure of progress and derive the bound below by neglecting the small error from encountering the feature $\phi_{pi}$ (step (a) below) instead of the expected feature $\phi_{\pi,p}$ (identified by the policy $\pi$, played by FRANCIS in episode $i$), by using a high probability bound $\|\xi_{pi}\|_2 \approx \sqrt{d_p \sigma}$ and by using the elliptic potential lemma in [Abbasi-Yadkori et al., 2011] for the last step.

$$1/k \sum_{i=1}^k U_{pi}^*(\sigma) \leq 1/k \sum_{i=1}^k \phi_{\pi,p}^T \xi_{pi} \approx 1/k \sum_{i=1}^k \phi_{pi}^T \xi_{pi} \leq 1/k \sum_{i=1}^k \|\phi_{pi}\|_2 \Sigma_{pi}^{-1} \|\xi_{pi}\|_2 \leq \sqrt{d_p \sigma}/k. \quad (4)$$

The inequality $\phi_{\pi,p}^T \xi_{pi} \geq U_{pi}^*(\sigma)$ in the first step only holds for some of the episodes (since lemma 2 ensures the inequality with probability $q \in \mathbb{R}$), but this only affects the bound up to a constant with high probability. Since the uncertainty is monotonically decreasing, the last term $U_{pi,k}^*(\sigma)$ must be
This implies we need to use the explorability condition to allow $F_{p_k}(\sigma) \leq \epsilon$ with high probability using $O(d_p^2 \sigma^2 / \epsilon^2)$ trajectories. At the end of an epoch FRANCIS ensures $U_{p_k}(\sigma) \leq \epsilon$, but we really need $U_{p_k}(\alpha_p) \leq \epsilon$ to hold.

5.4 Learning a Phase

We need to use the explorability condition to allow FRANCIS to proceed to the next epoch:

**Lemma 5** (see appendix D.4). Let $k$ and $\bar{k}$ be the starting and ending episodes in an epoch. If $\epsilon = O(\nu_{\min} / \sqrt{d_p})$ and $U_{p_k}(\sigma) = O(\epsilon)$ then $\lambda_{\min}(\Sigma_{p_k}) \geq 2\lambda_{\min}(\Sigma_{p_{\bar{k}}})$.

Since the minimum eigenvalue for the covariance matrix has doubled, we can double $\sigma$ (i.e., inject a stronger signal) and still satisfy lemma [5] at this point FRANCIS enters into a new epoch. At the beginning of every epoch we double $\sigma$, and this is repeated until $\sigma$ reaches the final value $\sigma \approx H^2 \alpha_p$. There are therefore only logarithmically many epochs (in the input parameters).

**Lemma 6** (FRANCIS meets target accuracy at the end of a phase, see appendix D.4). When FRANCIS reaches the end of the last epoch in phase $p$ it holds that $\sigma \approx H^2 \alpha_p$ and $\epsilon \geq U_{p_k}(\sigma) = H U_{p_k}(\alpha_p)$.

This implies $U_{p_k}(\alpha_p) \leq \epsilon / H$, as desired. Furthermore, this is achieved in $O(d_p^2 H^2 \alpha_p / \epsilon)$ trajectories.

Since $U_{p_k}(\alpha_p) \leq \epsilon / H$ the inductive step is now proved; summing the number of trajectories over all the phases gives the final bound in theorem 4.1. At this point, an $\epsilon$-optimal policy can be extracted by LSVI on the returned dataset $D$ for any prescribed linear reward function.

5.5 Connection with G-optimal design

We briefly highlight the connection with G-optimal design. G-optimal design would choose a design matrix $\Sigma$ such that $\|\hat{\phi}_{\pi,p}\|_{\Sigma^{-1}}$ is as small as possible for all possible $\pi$. Since we cannot choose the features in the online setting, a first relaxation is to instead compute (and run) the policy $\pi$ that maximizes the program $U_{p_i}(\sigma)$ in every episode $i$. Intuitively, as the area of maximum uncertainty is reached, information is acquired there and the uncertainty is progressively reduced, even though this might not be the most efficient way to proceed from an information-theoretic standpoint. Such procedure would operate in an online fashion, but unfortunately it requires an intractable optimization in policy space. Nonetheless this is the first relaxation to G-optimal design. To obtain the second relaxation, it is useful to consider the alternative definition $U_{p_i}(\sigma) = \max_{\pi,\|\hat{\phi}\|_{\Sigma_{\pi,p}} \leq \sqrt{\sigma} \hat{\phi}_{\pi,p}} \hat{\phi}_{\pi,p}^T \Sigma_{\pi,p}^{-1} \hat{\phi}_{\pi,p}$. If we relax the constraint $\|\hat{\phi}\|_{\Sigma_{\pi}} \leq \sqrt{\sigma}$ to obtain $\|\hat{\phi}\|_{\Sigma_{\pi,p}} \leq \sqrt{\sigma \Sigma_{\pi,p}}$ then the feasible space is large enough that random sampling from the feasible set (and computing the maximizing policy by using LSVI) achieves the goal of overestimating the maximum of the unrelaxed program; in particular, sampling $\xi_{\pi,p} \sim N(0,\sigma\Sigma_{\pi,p}^{-1})$ satisfies the relaxed constraints with high probability and is roughly uniformly distributed in the constraint set.

6 Discussion

This works makes progress in relaxing the optimistic closure assumptions on the function class for exploration through a statistically and computationally efficient PAC algorithm. From an algorithmic standpoint, our algorithm is inspired by [Osband et al. 2016b], but from an analytical standpoint, it is justified by a design-of-experiments approach [Lattimore and Szepesvari 2020]. Remarkably, our approximations to make G-experimental design implementable online and with polynomial computational complexity only add a $d$ factor compared to G-optimal design. The proof technique is new to our knowledge both in principles and in execution, and can be appreciated in the appendix. We hope that the basic principle is general enough to serve as a foundation to develop new algorithms with even more general function approximators. The contribution to reward-free exploration [Jin et al. 2020b] to linear value functions is also a contribution to the field.
7 Broader Impact

This work is of theoretical nature and aims at improving our core understanding of reinforcement learning; no immediate societal consequences are anticipated as a result of this study.

Acknowledgment

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A Preliminaries

A.1 Symbols

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<td>$r_t(s,a)$</td>
<td>expected reward in $(s,a,t)$</td>
</tr>
<tr>
<td>$p_t(s,a)$</td>
<td>transition function in $(s,a,t)$</td>
</tr>
<tr>
<td>$s_{t+k}$</td>
<td>experienced state at timestep $t$ in episode $k$ in phase $t$</td>
</tr>
<tr>
<td>$a_{t+k}$</td>
<td>experienced action at timestep $t$ in episode $k$ in phase $t$</td>
</tr>
<tr>
<td>$r_k$</td>
<td>experienced reward at timestep $t$ in episode $k$ in phase $t$</td>
</tr>
<tr>
<td>$s_{t+1,k}$</td>
<td>experienced state at timestep $t+1$ in episode $k$ in phase $t$</td>
</tr>
<tr>
<td>$L_\phi$</td>
<td>upper bound on $\sup_{s,a,t}</td>
</tr>
<tr>
<td>$\phi_{tk}$</td>
<td>$\phi_t(s_{tk},a_{tk})$</td>
</tr>
<tr>
<td>$\Sigma_{tk}$</td>
<td>$\Sigma_{tk}$ matrix after FRANCIS has completed learning in phase $t$ (k is the last episode in that phase)</td>
</tr>
<tr>
<td>$T_t(Q_{t+1})(s,a)$</td>
<td>$r_t(s,a) + \mathbb{E}<em>{s' \sim p_t(s,a)} Q</em>{t+1}(s,a)$</td>
</tr>
<tr>
<td>$T_t^r(Q_{t+1})(s,a)$</td>
<td>any $\hat{\theta}<em>t(Q</em>{t+1}) \in B_t$ s.t. $\max_{\theta}</td>
</tr>
<tr>
<td>$\theta_t$</td>
<td>$\Sigma_{t-1}^{-1} \sum_{i=t}^{\infty} \phi_{ti}[V_{t+1}(s_{t+1,i})]$</td>
</tr>
<tr>
<td>$\pi_t$</td>
<td>policy played in episode $t$ of phase $t$</td>
</tr>
<tr>
<td>$Q_t(\theta)$</td>
<td>action value function $(s,a) \mapsto \phi_t(s,a)\top \theta$</td>
</tr>
<tr>
<td>$V_t(\theta)$</td>
<td>value function $s \mapsto \max_{\phi} \phi_t(s,a)\top \theta$</td>
</tr>
<tr>
<td>$\eta_t(\hat{V}_{t+1})$</td>
<td>$\hat{V}<em>{t+1}(s</em>{t+1,i}) - \mathbb{E}<em>{s' \sim p</em>{t}(s_{t+1,i})} \hat{V}_{t+1}(s')$</td>
</tr>
<tr>
<td>$\Delta_t^r(s,a)$</td>
<td>$r_t(s,a) - \phi_t(s,a)\top \theta^*_t$</td>
</tr>
<tr>
<td>$\Delta_t^\delta(s,a)$</td>
<td>$r_t(s,a) - \phi_t(s,a)\top \delta^*_t$</td>
</tr>
<tr>
<td>$\eta^\delta_t$</td>
<td>$r_t(s,a) - \phi_t(s,a)\top \delta^*_t$ (reward noise)</td>
</tr>
<tr>
<td>$T(Q_t,Q_{t+1})$</td>
<td>$\max_{Q_t \in Q_t} \min_{Q_{t+1} \in Q_{t+1}}</td>
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<tr>
<td>$E_t$</td>
<td>approximation error for the reward, see eq. 157</td>
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<tr>
<td>$k$</td>
<td>is an overestimate of the number of episodes and is used in the definition of the $\beta$’s below</td>
</tr>
<tr>
<td>$\delta'$</td>
<td>is used in the definition of the $\beta$’s below</td>
</tr>
<tr>
<td>$\beta^\delta_t$</td>
<td>$\sqrt{T} \times 2 \sqrt{d_t \ln \left(1 + L^2/d_t\right)} + d_{t+1} \ln(1 + 4R_{t+1}/(2L_\phi \sqrt{T}^2)) + \ln (\frac{1}{\delta}) + \frac{2}{2}$</td>
</tr>
<tr>
<td>$\beta_t^\delta$</td>
<td>$\sqrt{d_t \ln \left(1 + kL^2/d_t\right)} +</td>
</tr>
<tr>
<td>$D_p$</td>
<td>$d_p \ln(1 + kL^2/d_p)$</td>
</tr>
<tr>
<td>$\beta_t^\delta'$</td>
<td>$\beta_t^\delta$</td>
</tr>
<tr>
<td>$\sqrt{a_t}$</td>
<td>$3 \left(\sqrt{\delta^2} + \sqrt{\beta^2} + 2\right) = \tilde{O}(\sqrt{d_t + d_{t+1}})$</td>
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\footnote{This only applies if the reward function is learned from data; since we’re doing reward free exploration, it instead represents the reward used to populate the dataset $\mathcal{D}$ after FRANCIS has terminated.}

\footnote{In particular it can be set to be equal to $n(t)$ and is $\text{poly}(d_1, \cdots, d_H, H, \frac{1}{\gamma}, \frac{1}{\delta})$.}

\footnote{In particular it is $\frac{\text{poly}(d_1, \cdots, d_H, H, \frac{1}{\gamma}, \frac{1}{\delta})}{d_1}$.}
\( n(t) \) \( \overset{\text{def}}{=} \) number of samples collected in phase \( t \)
\[ \hat{\theta}_t \] \( \overset{\text{def}}{=} \Sigma_t^{-1} \sum_{i=1}^{n(t)} \phi_t \left( \hat{V}_{t+1}(s^+_{t+1,i}) \right) \)
\[ \tilde{\theta}_t \] \( \overset{\text{def}}{=} \Sigma_t^{-1} \sum_{i=1}^{n(t)} \phi_t \left( \tau_h \right) \)
\[ \tilde{\theta}_t^{PV} \] \( \overset{\text{def}}{=} \tilde{\theta}_t^2 + \hat{\theta}_t \)
\( R_t \) \( \overset{\text{def}}{=} \) radius at timestep \( t \) (but these will be all equal to 1 in the end)
\( R \) \( \overset{\text{def}}{=} \) \( R_1 = \cdots = R_H = 1 \)
\( q \) \( \overset{\text{def}}{=} \) \( \Phi(-3) \) (normal cdf evaluated at \(-3\))
\( C_k \) \( \overset{\text{def}}{=} \) \( \left\{ \max_{\pi, \eta \in \mathbb{R}^{dp}} \| \eta \| \Sigma_p \leq \sqrt{\sigma} \phi_T^\top \pi, \eta > \epsilon'' > \tau \right\} \)
\( \mathcal{E}_k \) \( \overset{\text{def}}{=} \) \( \left\{ \mathbb{E}_{x_1 \sim \rho} \hat{V}_{1k}(x_1) - \tau \geq \max_{\pi, \eta \in \mathbb{R}^{dp}} \| \eta \| \Sigma_p \leq \sqrt{\sigma} \phi_T^\top \pi, \eta \right\} \)
\( k(e, i) \) \( \overset{\text{def}}{=} \) episode in epoch \( e \) (of a certain phase) such that \( E_k \) happens for the \( i \)-th time.
\( \zeta_{pk(e, i)} \) \( \overset{\text{def}}{=} \) \( \phi_T^\top \pi_{k(e, i)} - \phi_T^\top p_{k(e, i)} \xi_{pk(e, i)} - \phi_T^\top p_{k(e, i)} \xi_{pk(e, i)} \)
\( A \) \( \overset{\text{def}}{=} \sqrt{\frac{8 \ln(1/\delta'')}{\gamma_t(\sigma)}} \)
\( \gamma_t(\sigma) \) \( \overset{\text{def}}{=} \sqrt{\frac{2\sigma_t d_t \ln \frac{2d_t}{\delta'\prime}}{\sigma}} \)
\( \pi_t(s) \) \( \overset{\text{def}}{=} \) indicates the action taken at timestep \( t \) by policy \( \pi \) in state \( s \)
\( \sigma_{\text{Start}} \) \( \overset{\text{def}}{=} \) \( 1/ \left( 8d_p \ln \frac{2d_p}{\delta'\prime} \right) \)
\( aB_t \) \( \overset{\text{def}}{=} \) \( \{ ax \mid x \in B_t \} \) for a positive real \( a \)
\( V_t^n \) \( \overset{\text{def}}{=} \) value function of policy \( \pi \) at timestep \( t \) on \( M \) once the reward function is fixed
\( V^* \) \( \overset{\text{def}}{=} \) optimal value function on \( M \) once the reward function is fixed
\( \pi^* \) \( \overset{\text{def}}{=} \) optimal policy on \( M \) once the reward function is fixed
\( c_{e, 1}, c_{e, 2}, c_{e, 3} \) \( \overset{\text{def}}{=} \) constants implicitly determined, see proof of theorem \( 1 \) and footnote in that page
A.2 Inherent Bellman Error

Definition 4 (Inherent Bellman Error and Best Approximator). Given two compact linear functional spaces $Q_t$ and $Q_{t+1}$, the inherent Bellman error at step $t$ is the maximum (in absolute value) residual

$$\mathcal{I}(Q_t, Q_{t+1}) \overset{\text{def}}{=} \max_{Q_{t+1} \in Q_{t+1}} \min_{Q_t \in Q_t} \max_{(s,a)} |Q_t - T^P_t(Q_{t+1})|(s,a)|.$$

The approximator $Q_t(Q_{t+1}) \in Q_t$ of $Q_{t+1} \in Q_{t+1}$ through $T^P_t$ is defined by its parameter $\hat{\theta}_t(Q_{t+1})$ as any solution $\hat{\theta}_t \in \mathcal{B}$, that verifies (always exists from the above display) for any $Q_{t+1} \in Q_{t+1}$

$$\max_{(s,a)} |\phi_t(s,a) \hat{\theta}_t(Q_{t+1}) - T^P_t(Q_{t+1})(s,a)| \leq \mathcal{I}(Q_t, Q_{t+1}) \tag{6}$$

The Bellman residual function $\Delta_{t,1}$ under policy $\pi$ is implicitly defined in the error decomposition below:

$$T^P_t(Q_{t+1})(s,a) \overset{\text{def}}{=} Q_t(Q_{t+1})(s,a) + \Delta_t(Q_{t+1})(s,a). \tag{7}$$

and it satisfies

$$\mathcal{I}(Q_t, Q_{t+1}) = \max_{(s,a)} |\Delta_t(Q_{t+1})(s,a)| \tag{8}$$

We briefly argue why we have the last equality in the above definition

$$\mathcal{I}(Q_t, Q_{t+1}) \geq \max_{Q_{t+1} \in Q_{t+1}} \max_{Q_t \in Q_t} |Q_t - T^P_t(Q_{t+1})|(s,a)| \geq \mathcal{I}(Q_t, Q_{t+1}) \tag{9}$$

$$= \max_{Q_{t+1} \in Q_{t+1}} \max_{(s,a)} |\Delta_t(Q_{t+1})(s,a)| \tag{10}$$

where the second step uses eq. (7).

We are going to use the following property throughout the appendix:

Proposition 1 (Positive Homogeneity of Inherent Bellman Error of System Dynamics). Let $\gamma$ be a positive scalar number. If

$$\max_{Q_{t+1} \in Q_{t+1}} \min_{Q_t \in Q_t} \max_{(s,a)} |Q_t - T^P_t(Q_{t+1})|(s,a)| \leq \mathcal{I}(Q_t, Q_{t+1}) \tag{11}$$

then

$$\max_{Q_t \in \gamma Q_t} \min_{Q_{t+1} \in \gamma Q_{t+1}} \max_{(s,a)} |Q_t - T^P_t(Q_{t+1})|(s,a)| \leq \gamma \mathcal{I}(Q_t, Q_{t+1}) \tag{12}$$

where

$$\gamma Q_t = \{ Q_t \mid Q_t(s,a) = \phi_t(s,a)^{\top} \theta, \|\theta\|_2 \leq \gamma R_t \}, \quad t \in \{ t, t+1 \} \tag{13}$$

This implies that if $\|\theta\|_2 \leq \gamma R_{t+1}$ then we can find a $\theta_t$ satisfying $|\theta_t(V_{t+1}(\theta_{t+1}))|_2 \leq \gamma R_t$.

Proof. Notice that when we write $\max_x f(x) \leq I$ (for a generic scalar function $f$, an element $x$ in a set, and a scalar $I$) we can replace the statement with $\forall x, f(x) \leq I$ and viceversa:

$$\max_x f(x) \leq I \iff \forall x, f(x) \leq I \tag{14}$$

Likewise:

$$\max_x \min_y f(x,y) \leq I \iff \forall x, \exists y : f(x,y) \leq I \tag{15}$$

We can recast the Bellman error condition as

$$\forall Q_{t+1} \in Q_{t+1}, \exists Q_t \in Q_t : \max_{(s,a)} |Q_t - T^P_t(Q_{t+1})|(s,a)| \leq \mathcal{I}(Q_t, Q_{t+1}) \tag{16}$$

Now consider the bijection

$$Q_t \in Q_t \leftrightarrow Q'_t = \gamma Q_t \in Q_t,$$

$$Q_{t+1} \in Q_{t+1} \leftrightarrow Q'_{t+1} = \gamma Q_{t+1} \in Q_{t+1}, \tag{17}$$

For infinite horizon MDPs, these normally coincide.
We have that the statement below holds if and only if
\[
\forall Q_{t+1} \in Q_{t+1}, \exists Q_{t} \in Q_{t} : \max_{(s,a)} [(Q_{t} - \mathcal{P}_{t}^{P}(Q_{t+1}))(s,a)] \leq \gamma \mathcal{I}(Q_{t}, Q_{t+1}) \tag{18}
\]
holds and the prior display implies that eq. (19) holds, and so does eq. (18) which is equivalent to eq. (12).
Finally to conclude the proof of the theorem notice that if \(\theta_{t+1} \in \gamma \mathcal{R}_{t+1}\) then we can find a \(\theta_{t} \in \gamma \mathcal{R}_{t}\) such that the Bellman error is at most \(\gamma \mathcal{I}(Q_{t}, Q_{t+1})\).
B Analysis of vanilla LSVI

We recall the popular LSVI protocol [Munos 2005] [Munos and Szepesvári 2008] operating on a batch dataset \( D \) of experienced state-action-reward-successor states. We use \( n(t) \) to denote the number of samples collected at a certain timestep \( t \). The regularization parameter is optional and defaults to \( \lambda = 1 \). The LSVI algorithm is used without reward from the dataset \( D \) when called by FRANCIS; instead a pseudoreward function \( R_p \) is prescribed in the last timestep.

Algorithm 2 LSVI(\( H, R_H, D; \lambda = 1 \)) - This is for use in FRANCIS with reward signal \( R_H \)

1: \textbf{Input}: horizon \( H \), dataset \( D \), regularization \( \lambda \).
2: Extract pseudo-reward parameter \( \xi_H \) from \( R_H \) function.
3: Set \( \hat{\theta}_H = \xi_H \).
4: for timestep \( t = H - 1, \ldots, 1 \) do
5: \quad Solve \( \hat{\theta}_t = \arg \min_{\theta} \sum_{k=1}^{n(t)} [\phi_t(s_{tk}, a_{tk})^\top \theta - \max_{a'} \phi_{t+1}(s_{t+1,k}^+, a')^\top \hat{\theta}_{t+1}]^2 + \lambda \|\theta\|_2^2 \)
6: end for
7: \textbf{Return} \( \pi : (s, t) \mapsto \arg \max_a \phi_t(s, a)^\top \hat{\theta}_t \).

Algorithm 3 LSVI(\( H, D; \lambda = 1 \)) - This is the regular batch algorithm

1: \textbf{Input}: horizon \( H \), dataset \( D \), regularization \( \lambda \).
2: Set \( \hat{\theta}_{H+1}^{R+PV} = 0 \).
3: for timestep \( t = H, H - 1, \ldots, 1 \) do
4: \quad Solve \( \hat{\theta}_t^{R+PV} = \arg \min_{\theta} \sum_{i=1}^{n(t)} [\phi_{t,i}(s_{t+1,i})^\top \theta - \max_{a'} \phi_{t+1,i}(s_{t+1,i}^+, a')^\top \hat{\theta}_{t+1}^{R+PV}]^2 + \lambda \|\theta\|_2^2 \)
5: end for
6: \textbf{Return} \( \pi : (s, t) \mapsto \arg \max_a \phi_t(s, a)^\top \hat{\theta}_t^{R+PV} \).

B.1 Single Step Error Decomposition

Lemma 7 (Analysis of Transition Error in Parameter Space). Let \( n(t) \) be the number of episodes where samples have been collected at timestep \( t \). If \( \hat{\theta}_t \) satisfies

\[
\hat{\theta}_t = \sum_{i=1}^{n(t)} \phi_{t,i} \left[ \hat{V}_{t+1}(s_{t+1,i}) \right]
\]

then it must also satisfy:

\[
\hat{\theta}_t = \hat{\theta}_t(\hat{V}_{t+1}) + \sum_{i=1}^{n(t)} \phi_{t,i} \left[ \Delta_t \left( \hat{V}_{t+1} \right) + \eta^i_t(\hat{V}_{t+1}) \right] - \lambda \hat{\theta}_t(\hat{V}_{t+1}) \]

(26)

Proof. Let \( \pi_{t,i} \) be the policy used to generate the rollouts of episode \( i \) of phase \( t \). Define the trajectory noise of episode \( i \) of phase \( t \) using the next-state value function \( \hat{V}_{t+1} \) as:

\[
\eta^i_t(\hat{V}_{t+1}) \overset{d}{=} \hat{V}_{t+1}(s_{t+1,i}) - \mathbb{E}_{s' \sim p(s_{t+1,i}, \pi_{t,i}(s_{t,i}))} \hat{V}_{t+1}(s').
\]

From eq. 25 we can rewrite the unique solution for \( \hat{\theta}_t \) as

\[
\hat{\theta}_t = \sum_{i=1}^{n(t)} \phi_{t,i} \left[ \mathbb{E}_{s' \sim p(s_{t+1,i}, \pi_{t,i}(s_{t,i}))} \hat{V}_{t+1}(s') + \eta^i_t(\hat{V}_{t+1}) \right]
\]

(28)

Recall the error decomposition of eq. 7 with \((s, a) = (s_{t,i}, \pi_{t,i}(s_{t,i})), \phi_{t,i} = \phi(s, a), \Delta_t = \Delta_t(s, a), \mathbb{E}_{s' \sim p(s_{t+1,i}, \pi_{t,i}(s_{t,i}))} \hat{V}_{t+1}(s') = \phi_t(\hat{\theta}_t(Q_{t+1})) + \Delta_t(Q_{t+1}) \)

(29)

where \( \hat{\theta}_t(Q_{t+1}) \in \mathcal{B}_t \).
Plugging back eq. (29) into eq. (28) gives:

\[
\hat{\theta}_t = \Sigma_t^{-1} \left( \sum_{i=1}^{n(t)} \phi_{ti} [\phi_{ti}^T \hat{\theta}_t(\hat{V}_{t+1}) + \Delta_{ti}(\hat{V}_{t+1}) + \eta_{ti}^T(\hat{V}_{t+1})] + \lambda \hat{\theta}_t(\hat{V}_{t+1}) - \lambda \hat{\theta}_t(\hat{V}_{t+1})^{\infty} \right)
\]  

(30)

\[
= \Sigma_t^{-1} \Sigma_t \hat{\theta}_t(\hat{V}_{t+1}) + \Sigma_t^{-1} \sum_{i=1}^{n(t)} \phi_{ti} [\Delta_{ti}(\hat{V}_{t+1}) + \eta_{ti}^T(\hat{V}_{t+1})] - \lambda \hat{\theta}_t(\hat{V}_{t+1})
\]  

(31)

\[
= \hat{\theta}_t(\hat{V}_{t+1}) + \Sigma_t^{-1} \sum_{i=1}^{n(t)} \phi_{ti} [\Delta_{ti}(\hat{V}_{t+1}) + \eta_{ti}^T(\hat{V}_{t+1})] - \lambda \hat{\theta}_t(\hat{V}_{t+1})
\]  

(32)

This proves the lemma. □
Lemma 8 (Analysis of Reward Error in Parameter Space). Let $n(t)$ be the number of episodes where samples have been collected at timestep $t$. If $\hat{\theta}^*_t$ satisfies

$$\hat{\theta}^*_t = \Sigma_t^{-1} \sum_{i=1}^{n(t)} \phi_{ti} r_{tk}$$ (33)

then it must also satisfy:

$$\hat{\theta}^*_t = \theta^*_t + \Sigma_t^{-1} \left( \sum_{i=1}^{n(t)} \phi_{ti} \left[ \eta^*_{ti} + \Delta^*_{ti} \right] - \lambda \theta^*_t \right)$$ (34)

**Proof.** Let $\pi_{ti}$ be the policy used to generate the rollouts of episode $i$ of phase $t$.

From eq. (33) we can rewrite the unique solution for $\hat{\theta}^*_t$ as (for the definitions of the symbols see table 2):

$$\hat{\theta}^*_t = \Sigma_t^{-1} \sum_{i=1}^{n(t)} \phi_{ti} \left[ r_{ti} (s_{ti}, a_{ti}) + \eta^*_{ti} \right]$$

$$= \Sigma_t^{-1} \left( \sum_{i=1}^{n(t)} \phi_{ti} \left[ \phi^\top_{ti} \theta^*_t + \Delta^*_{ti} + \eta^*_{ti} \right] + \lambda \theta^*_t - \lambda \theta^*_t \right)$$

$$= \theta^*_t + \Sigma_t^{-1} \left( \sum_{i=1}^{n(t)} \phi_{ti} \left[ \eta^*_{ti} + \Delta^*_{ti} \right] - \lambda \theta^*_t \right)$$ (35)

\(\square\)
B.2 Single Step Error Bounds

**Definition 5** (Good Event for LSVI). Assume $\sqrt{n(t)I(Q_t, Q_{t+1})} \leq \sqrt{\alpha_t}/3$ and $\sqrt{n(t)E_{t}} \leq \sqrt{\alpha_t}/3$. We say that LSVI (algorithms 2 and 3) is in the good event when the following bound holds for all $t \in [H]$ with $V_{t+1} \in \mathcal{V}_{t+1}$. The definition of the symbols are reported in table 2:

\[ \| \sum_{i=1}^{n(t)} \phi_{t_i} \Delta_{t_i}(\hat{V}_{t+1}) \|_{\Sigma_t^{-1}} \leq \sqrt{n(t)I(Q_t, Q_{t+1})} \]  
\[ \| \sum_{i=1}^{n(t)} \phi_{t_i} \eta_i^t(\hat{V}_{t+1}) \|_{\Sigma_t^{-1}} \leq \sqrt{\beta_t} \]  
\[ \lambda \| \hat{\theta}_t(\hat{V}_{t+1}) \|_{\Sigma_t^{-1}} \leq \sqrt{\lambda R_t} \]  
\[ \| \sum_{i=1}^{n(t)} \phi_{t_i} \Delta_{t_i}^r \|_{\Sigma_t^{-1}} \leq \sqrt{n(t)E_t} \]  
\[ \| \sum_{i=1}^{n(t)} \phi_{t_i} \eta_i^t \|_{\Sigma_t^{-1}} \leq \sqrt{\beta_t} \]  
\[ \lambda \| \theta_r^t \|_{\Sigma_t^{-1}} \leq \sqrt{\lambda \| \theta_r^t \|_2} \]

In addition, the above expressions with the relations in lemma 7 (Analysis of Transition Error in Parameter Space) and lemma 8 (Analysis of Reward Error in Parameter Space) imply:

\[ \| \hat{\theta}_t - \theta_r^t \|_{\Sigma_t} + \| \hat{\theta}_t - \hat{\theta}_t(\hat{V}_{t+1}) \|_{\Sigma_t} \]  
\[ \leq \sqrt{n(t)I(Q_t, Q_{t+1})} + \sqrt{n(t)E_t} + \sqrt{\beta_t} + \sqrt{\beta_t} + \sqrt{\lambda R_t} + \sqrt{\lambda \| \theta_r^t \|_2} \]  
\[ \leq \sqrt{\alpha_t} \]  

**Lemma 9** (Probability of Good Event for LSVI). There exists a parameter $\delta' = \frac{\delta}{\text{poly}(d_1, \ldots, d_H, H, 3)}$, such that the good event of definition 5 holds with probability at least $1 - \delta/2$.

**Proof.** Since $| \Delta_{t_i}(\hat{V}_{t+1}) | \leq I(Q_t, Q_{t+1})$, the projection bound (lemma 8 in [Zanette et al., 2020b]) gives the first inequality in the statement of the theorem. The second inequality is proved in lemma 21 (Transition Noise High Probability Bound) respectively. The third inequality follows from lemma 25 (Change of $L^2$-Norm). Since $| \Delta_{t_i} | \leq E_t$, the projection bound (lemma 8 in [Zanette et al., 2020b]) again gives the fourth inequality. The fifth inequality follows from theorem 2 in [Abbas-Yadkori et al., 2011] with 1-sub-Gaussian noise and the last inequality again follows from lemma 25 (Change of $L^2$-Norm). In particular it is possible to choose $\delta'$ (in the definition of the $\beta$'s) such that these statements jointly hold with probability at least $1 - \delta/2$ after a union bound over each statement and the timestep $H$. At this point the statement in eq. 42 follows deterministically by chaining with lemmas 7 and 8.

---

Note that if $\hat{V}_{t+1} \in R \times \mathcal{V}_{t+1}$ (the set $\mathcal{V}_{t+1}$ where all elements are scaled by the scalar $R$) then the bounds still hold provided that they are rescaled by $R$. 
B.3 Iterate Boundness

In this section we discuss the boundness in the value function parameter.

**Lemma 10 (Boundness at Intermediate Timesteps for algorithm 2).** On the good event for LSVI of definition 5 if

\[ \lambda_{\min}(\Sigma_t) \geq 4H^2\alpha_t, \quad \forall t \in [p-1] \]  

\[ \|\xi_t\|_2 \leq \frac{1}{2} \]  

then

\[ \|\hat{\theta}_t\|_2 \leq 1, \quad \forall t \in [p]. \]  

**Proof.** We proceed by induction, showing that \( \hat{\theta}_t \) due to errors can live in bigger and bigger balls, with radius starting from \( \frac{1}{2} \) at timestep \( p \) to radius 1 at timestep 1.

**Inductive Hypothesis 2.** \( \|\hat{\theta}_t\|_2 \leq (1 - \frac{t-1}{2H}) \).

The inductive statement clearly holds at \( t = p \) by hypothesis of the lemma; therefore we focus on the inductive step (notice that the induction goes from \( t = p \) down to \( t = 1 \), so the inductive step assumes the inductive hypothesis holds when written for \( t + 1 \).)

The inherent Bellman error definition (definition 4 \( \text{[Inherent Bellman Error and Best Approximator]} \)) and proposition 1 (positive homogeneity of inherent Bellman error of system dynamics) ensures

\[ \|\hat{\theta}_{t+1}\|_2 \leq (1 - \frac{t}{2H}) \rightarrow \|\hat{\theta}_t(V_{t+1}(\hat{\theta}_{t+1}))\|_2 \leq \left(1 - \frac{t}{2H}\right) \]  

(46)

In particular, the left statement is ensured by the inductive hypothesis for \( t + 1 \). Next, under the good event of definition 5 (good event for LSVI) we have that Lemma 25 (change of \( \Sigma \)-norm) ensures (writing \( \hat{\theta}_t = \theta_t(V_{t+1}(\hat{\theta}_{t+1})) \) for short)

\[ \sqrt{\alpha_t} \geq \|\hat{\theta}_t - \hat{\theta}_t\|_{\Sigma_t} \geq \lambda_{\min}(\Sigma_t)\|\hat{\theta}_t - \hat{\theta}_t\|_2 \]  

(47)

Solving for \( \|\hat{\theta}_t - \hat{\theta}_t\|_2 \) and using the lemma’s hypothesis gives

\[ \|\hat{\theta}_t - \hat{\theta}_t\|_2 \leq \frac{\sqrt{\alpha_t}}{2H\sqrt{\alpha_t}} = \frac{1}{2H}. \]  

(48)

Combined with the prior display, we deduce

\[ \|\hat{\theta}_t\|_2 \leq \|\hat{\theta}_t - \hat{\theta}_t\|_2 + \|\hat{\theta}_t\|_2 \leq 1 - \frac{t}{2H} + \frac{1}{2H} = 1 - \frac{t-1}{2H}. \]  

(49)

This shows the inductive step.

\[ \square \]
Lemma 11 (Boundness at Intermediate Timesteps for algorithm 3). Under the good event definition 5, fix a positive scalar $R$; if

$$
\lambda_{\min}(\Sigma_t) \geq AH^2 \alpha_t, \quad \forall t \in [H]
$$

(50)

$$
\|\theta^*_1\|_2 \leq \frac{R}{H}
$$

(51)

then

$$
\|\hat{\theta}^{R+PV}_{t+1}\|_2 = \|\hat{\theta}^*_R + \hat{\theta}_t\|_2 \leq 2R, \quad \forall t \in [H].
$$

(52)

Proof. We proceed by induction, showing that $\hat{\theta}^{R+PV}_t$ due to errors can live in bigger and bigger balls

Inductive Hypothesis 3. $\|\hat{\theta}^{R+PV}_t\|_2 \leq 2(1 - \frac{t-1}{H})R$.

The inductive statement clearly holds at $t = H + 1$; therefore we focus on the inductive step (notice that the induction goes from $t = H + 1$ down to $t = 1$, so the inductive step assumes the inductive hypothesis holds when written for $t + 1$).

The inherent Bellman error definition [definition 3][Inherent Bellman Error and Best Approximator] and proposition [Positive Homogeneity of Inherent Bellman Error of System Dynamics] ensures

$$
\|\hat{\theta}^{R+PV}_{t+1}\|_2 \leq 2 \left(1 - \frac{t}{H}\right) R \longrightarrow \|\hat{\theta}_t(V_{t+1}(\hat{\theta}^{R+PV}_{t+1}))\|_2 \leq 2 \left(1 - \frac{t}{H}\right) R
$$

(53)

In particular, the left statement is ensured by the inductive hypothesis for $t + 1$. Next, under the good event of definition 5[Good Event for LSVI] (with a scaling argument by $R$ on the $\|\cdot\|_2$ norm of the regressed parameter) we have that lemma 25[Change of $\Sigma$-Norm] ensures (writing $\hat{\theta}_t = \hat{\theta}_t(V_{t+1}(\hat{\theta}^{R+PV}_{t+1}))$ for short)

$$
R\sqrt{\alpha_t} \geq \left(\|\hat{\theta}^*_t - \theta^*_t\|_{\Sigma_t} + \|\hat{\theta}_t - \hat{\theta}_t\|_{\Sigma_t}\right) \geq \sqrt{\lambda_{\min}(\Sigma_t)} \left(\|\hat{\theta}^*_t - \theta^*_t\|_2 + \|\hat{\theta}_t - \hat{\theta}_t\|_2\right)
$$

(54)

Solving for $\left(\|\hat{\theta}^*_t - \theta^*_t\|_2 + \|\hat{\theta}_t - \hat{\theta}_t\|_2\right)$ and using the lemma’s hypothesis gives

$$
\left(\|\hat{\theta}^*_t - \theta^*_t\|_2 + \|\hat{\theta}_t - \hat{\theta}_t\|_2\right) \leq \frac{\sqrt{\alpha_t}}{2H} R \leq \frac{R}{2H}.
$$

(55)

Combined with the prior display, we deduce

$$
\|\hat{\theta}^{R+PV}_t\|_2 \leq \|\hat{\theta}^*_t - \theta^*_t\|_2 + \|\hat{\theta}_t - \hat{\theta}_t\|_2 + \|\theta^*_t\|_2 + \|\hat{\theta}_t\|_2 \leq \frac{R}{2H} + \frac{R}{H} + 2(1 - \frac{t}{H}) R \leq 2(1 - \frac{t-1}{H}) R.
$$

(56)

This shows the inductive step. \qed
B.4 Multi-Step Analysis: Error Bounds for LSVI

Lemma 12 (Telescopic Expansion). Under the good event of definition for algorithm the boundness of the iterate \( \hat{θ}_t \) is given by

\[
\| \hat{θ}_t \|_2 \leq 1, \quad t \in [p].
\]

Furthermore, for any policy \( π \)

\[
E_{x_t \sim ρ} \hat{Q}_t(x_t, π_t(x_t)) \geq - \sum_{t=1}^{p-1} \left[ I(Q_t, Q_{t+1}) + \sum_{a \in A} \hat{Q}_t(x_t, π_t(x_t)) \right] \]

and for the greedy policy \( \hat{π} \) with respect to \( \hat{Q} \), \( \pi_t(x_t) = \arg \max_a \hat{Q}_t(s, a) \) it additionally holds that

\[
E_{x_t \sim ρ} \hat{V}_t(x_t) \leq - \sum_{t=1}^{p-1} \left[ I(Q_t, Q_{t+1}) + \sum_{a \in A} \hat{Q}_t(x_t, π_t(x_t)) \right] + E_{x_t \sim ρ} \hat{V}_t(x_t).
\]

Proof. On the good event for LSVI of definition the boundness of the iterate \( \hat{θ}_t \) is given by lemma \( \| \hat{θ}_t \|_2 \leq 1 \) for algorithm we can use Cauchy-Schwartz to write:

\[
| \hat{θ}_t - \hat{θ}_t(\hat{V}_t) \| \leq \| \hat{θ}_t - \hat{θ}_t(\hat{V}_t) \| \leq \sqrt{\alpha_t} \| \hat{θ}_t \|_2 \| \hat{θ}_t \|_2
\]

Using the Inherent Bellman Error and Best Approximator we can write:

\[
| \hat{θ}_t - \hat{θ}_t(\hat{V}_t) \| \leq \| \hat{θ}_t - \hat{θ}_t(\hat{V}_t) \| \leq \sqrt{\alpha_t} \| \hat{θ}_t \|_2 \| \hat{θ}_t \|_2
\]

Combining the two expression gives:

\[
| E_{x_t \sim ρ} \hat{Q}_t(x_t, π_t(x_t)) - E_{x_t \sim ρ} \hat{V}_t(x_t+1) \| \leq \| \hat{θ}_t - \hat{θ}_t(\hat{V}_t) \| \leq \sqrt{\alpha_t} \| \hat{θ}_t \|_2 \| \hat{θ}_t \|_2
\]

To show the upper bound if \( π \) is the greedy policy with respect to \( \hat{Q} \) then we can equivalently write \( \hat{V}_t(x_t) = \hat{Q}_t(x_t, π_t(x_t)) \)

\[
| E_{x_t \sim ρ} \hat{V}_t(x_t) - E_{x_t \sim ρ} \hat{V}_t(x_t+1) \| \leq \sqrt{\alpha_t} \| \hat{θ}_t \|_2 \| \hat{θ}_t \|_2 + I(Q_t, Q_{t+1}).
\]

Induction now shows the upper bound.

To show the lower bound, for a generic policy \( \hat{V}_t(x_t) ≥ \hat{Q}_t(x_t, π(x_t)) \) and so

\[
E_{x_t \sim ρ} \hat{Q}_t(x_t, π(x_t)) \geq - \sqrt{\alpha_t} \| \hat{θ}_t \|_2 \| \hat{θ}_t \|_2 + I(Q_t, Q_{t+1}) + E_{x_t \sim ρ} \hat{V}_t(x_t+1)
\]

\[
≥ - \sqrt{\alpha_t} \| \hat{θ}_t \|_2 \| \hat{θ}_t \|_2 + I(Q_t, Q_{t+1}) + \hat{Q}_t(x_t+1, π(x_t+1)).
\]

Induction concludes. □
Proposition 2 (Batch LSVI Guarantees (algorithm [4])). Under the good event of Definition 5 (Good Event for LSVI) assume that

\[ \forall t \in [H], \quad \|\theta_t\|_2 \leq \frac{R}{H} \tag{72} \]

If \( \tilde{V} \) and \( \tilde{\pi}^* \) are the value function and policy returned by algorithm 5 then

\[
\mathbb{E}_{x_1 \sim \rho} \left( V_1^* - \tilde{V}_1 \right) (x_1) \leq \sum_{t=1}^{H} \left[ 2E_t + R \left( I(Q_t, Q_{t+1}) + \sqrt{\alpha_t}\|\tilde{\phi}_{x^*, T}\|_{\Sigma_t^{-1}} \right) \right] \\
\mathbb{E}_{x_1 \sim \rho} \left( \tilde{V}_1 - V_1^{\tilde{\pi}^*} \right) (x_1) \leq \sum_{t=1}^{H} \left[ 2E_t + R \left( I(Q_t, Q_{t+1}) + \sqrt{\alpha_t}\|\tilde{\phi}_{\tilde{\pi}^*, T}\|_{\Sigma_t^{-1}} \right) \right]. 
\tag{73}
\]

Proof. Boundness of the iterates \( \|\tilde{\theta}^t + \tilde{\theta}^{R+PV}\|_2 \) is ensured by Lemma 11 (Boundness at Intermediate Timesteps for algorithm [3]). Consider a generic timestep \( t \); using the Bellman equation and the fact that \( \tilde{V}_1(x_t) \geq Q_t(x_t, \tilde{\pi}_t(x_t)) \) gives

\[
\mathbb{E}_{x_t \sim \pi^*} \left( V_t^* - \tilde{V}_t \right) (x_t) \leq \mathbb{E}_{x_t \sim \pi^*} \left[ r_t(x_t, \pi_t(x_t)) + \mathbb{E}_{x_{t+1} \sim \pi^*} V_{t+1}^*(x_{t+1}) - \mathbb{E}_{x_t \sim \pi^*} \pi_t(x_t) \right] \left( \tilde{\theta}^t + \tilde{\theta}_t \right) \\
\leq E_t + \tilde{\phi}_{\pi^*, t}^T \tilde{\theta}_t^r + \mathbb{E}_{x_{t+1} \sim \pi^*} V_{t+1}^*(x_{t+1}) - \mathbb{E}_{x_t \sim \pi^*} \pi_t^T \left( \tilde{\theta}^t + \tilde{\theta}_t \right) 
\tag{74}
\]

Next, under the good event of Definition 5 we can write:

\[
\leq 2E_t + \tilde{\phi}_{\pi^*, t}^T \tilde{\theta}_t^r + \mathbb{E}_{x_{t+1} \sim \pi^*} V_{t+1}^*(x_{t+1}) - \tilde{\phi}_{\pi^*, t}^T \theta_t^r \\
- \mathbb{E}_{x_{t+1} \sim \pi^*} \tilde{V}_{t+1}(x_{t+1}) + R[I(Q_t, Q_{t+1}) + \sqrt{\alpha_t}\|\tilde{\phi}_{\pi^*, t}\|_{\Sigma_t^{-1}}] 
\tag{76}
\]

Induction gives the first statement.

Now again we start with the definition of expected feature and the Bellman equation:

\[
\mathbb{E}_{x_t \sim \pi^*} \left( \tilde{V}_t - V_t^{\pi^*} \right) (x_t) = \tilde{\phi}_{\pi^*, t}^T \tilde{\theta}_t^r - \mathbb{E}_{x_t \sim \pi^*} r_t(x_t, \tilde{\pi}_t(x_t)) - \mathbb{E}_{x_{t+1} \sim \pi^*} V_{t+1}^{\pi^*}(x_{t+1}) \\
\leq \tilde{\phi}_{\pi^*, t}^T \theta^r + E_t + R[I(Q_t, Q_{t+1}) + \sqrt{\alpha_t}\|\tilde{\phi}_{\pi^*, t}\|_{\Sigma_t^{-1}}] + \\
+ \mathbb{E}_{x_{t+1} \sim \pi^*} \tilde{V}_{t+1}(x_{t+1}) - \tilde{\phi}_{\pi^*, t}^T \theta_t^r + E_t + \mathbb{E}_{x_{t-1} \sim \pi^*} V_{t+1}^{\pi^*}(x_{t+1}). 
\tag{78}
\]

Induction again concludes. 

\[ \square \]
C  Design of Experiments

We show that obtaining $\|\overline{\phi}_{\pi,t}\|_{\Sigma_t^{-1}} \leq \sqrt{\frac{\epsilon'}{H\alpha_t}} = \epsilon'$ suffices; we assume $\mathcal{I}(Q_t, Q_{t+1}) = E_t = 0$ for simplicity as well as $d_1 = \cdots = d_H$. We immediately have that

$$\sum_{t=1}^{H} \sqrt{\alpha_t \|\overline{\phi}_{\pi,t}\|_{\Sigma_t^{-1}}} \leq H \times \sqrt{\frac{\epsilon'}{H\alpha_t}} \times \frac{\epsilon}{H\sqrt{\alpha_t}} = \epsilon. \quad (81)$$

Thus, summing the two equations in eq. (73) for any linear reward function with $\|\theta_t\|_2 \leq \frac{1}{H}$ ensures an $\epsilon$-optimal policy on that reward function is returned.

The Kiefer-Wolfowitz theorem in [Lattimore and Szepesvári, 2020] guarantees such reduction in $\|\overline{\phi}_{\pi,t}\|_{\Sigma_t^{-1}}$ using $\tilde{O}(d^2 + \frac{d}{\sqrt{\alpha_t}}) = \tilde{O}(d^2 + \frac{H^2\alpha_t}{\epsilon})$ samples at every level / timestep if $G$-optimal design is used. After sampling all levels and substituting the value for $\alpha_t$ in table 2 the sample complexity of doing $G$-optimal design becomes $\tilde{O}(d^2 + \frac{d^2H^3}{\epsilon})$.

Notice that this setting can model MDPs with rewards in $[0, 1/H]$ and value functions in $[0, 1]$; moving to the standard setting with rewards in $[0, 1]$ and value function in $[0, H]$ adds $H^2$ to the sample complexity to obtain an $\epsilon$-optimal policy.
D Analysis of FRANCIS

D.1 Generating Bounded Iterates

The following lemma ensures FRANCIS generates bounded iterates for an appropriate choice of \( \sigma \).

**Lemma 13** (Boundness at Exploratory Timestep). In episode \( k \) of phase \( p \), if

\[
\lambda_{\text{min}}(\Sigma_{pk}) \geq 8d_p \ln \frac{2d_p \ln 2}{\delta^p} \sigma \tag{82}
\]

\[
\xi_p \sim \mathcal{N}(0, \sigma^{-1} \Sigma_{pk}) \tag{83}
\]

then

\[
\|\xi_p\|_2 \leq \frac{1}{2} \tag{84}
\]

on the good event of definition 7 ([Good Event for FRANCIS]).

**Proof.** Directly by the choice of \( \sigma \) and the definition of good event for FRANCIS (see definition 7 ([Good Event for FRANCIS])).

D.2 Derandomization

The following lemma relates the sampling of the algorithm to a procedure that selects the policy / parameter leading to the area of highest (scaled) uncertainty.

**Lemma 14** (Derandomization). Outside of the failure event, assume that for any policy \( \pi \),

\[
\sum_{t=1}^{p-1} [ -I(Q_t, Q_{t+1}) + \sqrt{\alpha_t \| \phi_{\triangle \pi, t} \| \Sigma_{pk}^{-1}} ] \leq \tau \tag{85}
\]

for some scalar \( \tau > 0 \). Consider sampling

\[
\xi_p \sim \mathcal{N}(0, \sigma^{-1} \Sigma_{pk}), \tag{86}
\]

define \( R_p(s, a) = \phi_p(s, a) \top \xi_p \) and let \( \hat{V} \) be the value function computed by LSVI\((p, R_p \mathcal{D})\) (see algorithm 2). Then for a fixed constant \( q \in \mathbb{R} \)

\[
\max_{\pi, \eta \in \mathbb{R}^{d_p}} \| \eta \| \Sigma_{pk}^{-1} \leq \sqrt{\sigma \phi_{\triangle \pi,p} \top \eta} \tag{87}
\]

if

\[
\max_{\pi, \eta \in \mathbb{R}^{d_p}} \| \eta \| \Sigma_{pk}^{-1} \leq \sqrt{\sigma \phi_{\triangle \pi,p} \top \eta} \geq \tau \tag{88}
\]

\[
\|\xi_p\|_2 \leq \frac{1}{2}. \tag{89}
\]

**Proof.** Define the maximizer of the “scaled uncertainty” in a generic episode \( k \) of phase \( p \):

\[
\left( \triangle \pi, \triangle \eta \right) \overset{\text{def}}{=} \arg \max_{\pi, \eta \in \mathbb{R}^{d_p}} \| \phi_{\triangle \pi,p} \| \Sigma_{pk}^{-1} \leq \sqrt{\sigma} \|
\]

as the policy / parameter that maximizes the uncertainty.

Next, let \( \pi \) be the policy selected by the agent, through LSVI, corresponding to the sampled parameter \( \xi_p \) and let \( \hat{Q}, \hat{V} \) be the (action) value functions. Since \( \pi \) is the maximizing policy for \( \hat{Q} \), we must have:

\[
\mathbb{E}_{x_1 \sim \rho} \hat{V}(x_1) = \mathbb{E}_{x_1 \sim \rho} \hat{Q}(x_1, \pi_1(x_1)) \geq \mathbb{E}_{x_1 \sim \rho} \hat{Q}(x_1, \hat{\triangle} \pi_1(x_1)). \tag{91}
\]

In addition on the good event for LSVI [Telescopic Expansion] gives:

\[
\mathbb{E}_{x_1 \sim \rho} \hat{V}(x_1) \geq \mathbb{E}_{x_1 \sim \rho} \hat{Q}(x_1, \hat{\triangle} \pi_1(x_1)) \geq \sum_{t=1}^{p-1} \left[ I(Q_t, Q_{t+1}) - \sqrt{\alpha_t \| \phi_{\triangle \pi, t} \| \Sigma_{pk}^{-1}} \right] + \mathbb{E}_{x_p \sim \pi} \hat{Q}(x_p, \hat{\triangle} \pi_p(x_p)), \tag{92}
\]

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Subtracting $\tau$ to both sides and using the hypothesis gives
\[
E_{x_1 \sim \rho} \hat{V}_1(x_1) - \tau \geq -2\tau + (\hat{\sigma}_{\hat{\pi}, \rho})^T \xi_p. \tag{93}
\]

We can now proceed to bound the quantity of interest:
\[
P \left( E_{x_1 \sim \rho} \hat{V}_1(x_1) - \tau \geq (\hat{\sigma}_{\hat{\pi}, \rho})^T \hat{\eta} \right)
\geq P \left( -2\tau + \hat{\sigma}_{\hat{\pi}, \rho}^T \xi_p \geq (\hat{\sigma}_{\hat{\pi}, \rho})^T \hat{\eta} \right) \tag{94}
\geq P \left( \frac{-2\tau + \hat{\sigma}_{\hat{\pi}, \rho}^T \xi_p \geq (\hat{\sigma}_{\hat{\pi}, \rho})^T \hat{\eta}}{\frac{2\tau}{\max_{\phi, \|\eta\| \Sigma_{\pi k} \leq \sqrt{\sigma_{\pi, \rho}}}} \geq q} \right) \tag{95}
= P \left( \hat{\sigma}_{\hat{\pi}, \rho}^T \xi_p \geq 2\tau + (\hat{\sigma}_{\hat{\pi}, \rho})^T \hat{\eta} \right) \tag{96}
\]

Notice that $\xi_p$ is independent of $\hat{\sigma}_{\hat{\pi}}$ when conditioned on the $\Sigma_{\pi k}$. The last step is an application of [lemma 15] as long as the condition
\[
\tau \leq \max_{\phi, \|\eta\| \Sigma_{\pi k} \leq \sqrt{\sigma_{\pi, \rho}} \eta} \tag{97}
\]
is met. □
Lemma 15 (Uncertainty Overestimation). Let $\tau, \sigma$ be positive scalars, and let $\Sigma$ be an spd matrix and let

$$\xi \sim \mathcal{N}(0, \sigma \Sigma^{-1})$$

be the associated random vectors. For a fixed vector $\phi$ we have that

$$P\left(\phi^\top \xi \geq \max_{\phi, \eta \|\eta\|_\Sigma \leq \sqrt{\sigma}} \phi^\top \eta + 2\tau\right) \geq \Phi(-3) \overset{df}{=} q$$

where $\Phi(\cdot)$ is the normal CDF function as long as the condition

$$\tau \leq \max_{\phi, \|\phi\|_\Sigma \leq \sqrt{\sigma}} \sqrt{\sigma} \|\phi\|_\Sigma^{-1}$$

holds true.

Proof. Before we prove the statement, we notice that the equivalent expression $\max_{\phi, \|\eta\|_\Sigma \leq \sqrt{\sigma}} \phi^\top \eta = \sqrt{\sigma} \|\phi\|_\Sigma^{-1}$ can be found in chapter 19 of [Lattimore and Szepesvári, 2020] about the LinUCB algorithm, see also Lemma 26 [Linear Bandit Exploration Bonus]. For any fixed $\Sigma$, we have that $\xi \sim \mathcal{N}(0, \sigma \Sigma^{-1})$ is independent of $\phi$ by hypothesis, and so the inner product below is normally distributed

$$\phi^\top \xi \sim \mathcal{N}\left(0, \sigma \|\phi\|_\Sigma^{-1} \phi\right),$$

or equivalently

$$\phi^\top \xi \sim \mathcal{N}\left(0, \sigma \|\phi\|_\Sigma^{-1}^2\right).$$

Rescaling by its standard deviation leads to the following definition:

$$X \overset{df}{=} \frac{\phi^\top \xi}{\sqrt{\sigma} \|\phi\|_\Sigma^{-1}} \sim \mathcal{N}(0, 1).$$

The step below follows

$$P\left(\phi^\top \xi \geq \sqrt{\sigma} \|\phi\|_\Sigma^{-1} + 2\tau\right) = P\left(X \geq 1 + \frac{2\tau}{\sqrt{\sigma} \|\phi\|_\Sigma^{-1}}\right).$$

The rhs above is $\geq \Phi(-3)$ as long as

$$\tau \leq \sqrt{\sigma} \|\phi\|_\Sigma^{-1}.$$

The thesis follows from the definition of the normal CDF. $\square$
D.3 Learning an Epoch

The following lemma is key to our analysis and shows the number of episodes required to reduce the scaled uncertainty to the minimum allowable ($\approx \tau > 0$). In an epoch the value for $\sigma$ is fixed.

**Lemma 16** (Learning an Epoch). Let $k$ and $k^*$ be the starting and ending episodes in epoch $e$ of phase $p$. If the following statements hold:

1. For any policy $\pi$ it holds that $\sum_{i=1}^{p-1} \left[ T(Q_i, Q_{i+1}) + \sqrt{\gamma \alpha_t} \left\| \phi_{x_i, t}^\top \right\|_{\Sigma_t^{-1}} \right] \leq \tau$
2. $\lambda_{\min}(\Sigma_{pk}) \geq 8d_p \ln \frac{2d_p}{\epsilon''} \sigma$ (this ensures boundedness of $\|\phi_p\|_2$ in **Lemma 14** [Derandomization])
3. $\frac{\lambda_2}{\lambda_1} \leq 1$ (always satisfied by our choice $L_\phi = 1$ and $\lambda = 1$)
4. $\lambda > 1$ (always satisfied by our choice $\lambda = 1$)

then after at most

$$k_{\max} = k^* - k = \left\lfloor \frac{2}{1 - q} \times \left( \frac{\sqrt{\gamma}(p)D_p + A}{(\epsilon'')^2} \right)^2 \right\rfloor$$

episodes we must have

$$\max_{\pi, \eta \in \mathbb{R}^{d_p} : \|\eta\|_{\Sigma_p} \leq \sqrt{\sigma}} \psi_{\pi, p}^\top \eta \leq \epsilon''$$

on the good event **definition?** (Good Event for FRANCIS) provided that

$$\epsilon'' \geq \tau.$$  

**Proof.** First notice that if the eigenvalue condition is satisfied for at a given episode $k$ then it must be satisfied for all successive episodes $k \geq k$ since $\Sigma_{tk} \succeq \Sigma_{tk}$. In particular, define the events

$$\mathcal{C}_k \overset{\text{def}}{=} \left\{ \pi, \eta \in \mathbb{R}^{d_p} : \|\eta\|_{\Sigma_p} \leq \sqrt{\sigma} \right\}$$

$$\mathcal{E}_k \overset{\text{def}}{=} \left\{ \mathbb{E}_{x_1 \sim \rho} \tilde{V}_{nk}(x_1) - \tau \geq \max_{\pi, \eta \in \mathbb{R}^{d_p} : \|\eta\|_{\Sigma_p} \leq \sqrt{\sigma}} \psi_{\pi, p}^\top \eta \right\}.$$  

We examine what happens in those episodes where $\mathcal{E}_k$ occurs (notice that $\mathbb{P}(\mathcal{E}_k \mid \mathcal{C}_k) \geq q$ thanks to **Lemma 14** [Derandomization]).

Let $k(e, i)$ be the $i$-th consecutive episode index in epoch $e$ of phase $p$ such that $\mathcal{E}_k(e, i)$ occurs (so in $k(e, 1), k(e, 2), \ldots$ we have that $\mathcal{E}_k(e, 1), \mathcal{E}_k(e, 2)$ occurs). Since $\|\phi_{pk(e, i)}\|_2 \leq 1/2$ in the good event of definition? (Good Event for FRANCIS), we can use **Lemma 13** (Boundness at Exploratory Timestep) and **Lemma 12** [Telescopic Expansion] to write

$$\mathbb{E}_{x_1 \sim \rho} \tilde{V}_{nk}(x_1) - \tau \leq \phi_{pk(e, i)}^\top \xi_{pk(e, i)} + \zeta_{pk(e, i)}.$$  

where

$$\zeta_{pk(e, i)} = \phi_{pk(e, i)}^\top \xi_{pk(e, i)} - \phi_{pk(e, i)}^\top \xi_{pk(e, i)}.$$  

Let $i_{\max}$ be a fixed positive constant to be determined later. Taking average of the previous display up to $i_{\max}$ gives:

$$\frac{1}{i_{\max}} \sum_{i=1}^{i_{\max}} \mathbb{E}_{x_1 \sim \rho} \tilde{V}_{pk(e, i)}(x_1) - \tau \leq \frac{1}{i_{\max}} \sum_{i=1}^{i_{\max}} (\phi_{pk(e, i)}^\top \xi_{pk(e, i)} + \zeta_{pk(e, i)}).$$  

Under the good event of definition? (Good Event for FRANCIS), we have

$$\frac{1}{i_{\max}} \sum_{i=1}^{i_{\max}} \zeta_{pk(e, i)} \leq \frac{A}{\sqrt{i_{\max}}}$$  

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with $A = \mathcal{O}(1)$. For the remaining term, using Cauchy-Schwartz, and the fact that we are on the good event (see definition[7][Good Event for FRANCIS]) gives

$$
\frac{1}{i_{\text{max}}} \sum_{i = 1}^{i_{\text{max}}} \phi_{pk(e,i)} \xi_{pk(e,i)} \leq \frac{1}{i_{\text{max}}} \sum_{i = 1}^{i_{\text{max}}} \|\phi_{pk(e,i)}\|_{\Sigma_{pk(e,i)}^{-1}} \|\xi_{pk(e,i)}\|_{\Sigma_{pk(e,i)}} \sqrt{\gamma_i(\sigma)}
$$

After one more Cauchy-Schwartz we obtain the upper bound below:

$$
\leq \sqrt{\gamma_i(\sigma)} \frac{1}{i_{\text{max}}} \sum_{i = 1}^{i_{\text{max}}} \|\phi_{pk(e,i)}\|_{\Sigma_{pk(e,i)}^{-1}} \leq \sqrt{\gamma_i(\sigma)} \frac{1}{i_{\text{max}}} \sum_{i = 1}^{i_{\text{max}}} \|\phi_{pk(e,i)}\|_{\Sigma_{pk(e,i)}^{-1}}.
$$

We focus on the sum of squared features; by lemma[25][Change of Σ-Norm] and the lemma’s hypothesis

$$
\|\phi_{pk(e,i)}\|_{\Sigma_{pk(e,i)}^{-1}} \leq \frac{1}{\lambda} \|\phi_{pk(e,i)}\|_{2} \leq \frac{L^2}{\lambda} \leq 1
$$

and so the sum of squared features becomes9(using the elliptic potential lemma, see lemma 11 in Abbasi-Yadkori et al. [2011]):

$$
\sum_{i=1}^{i_{\text{max}}} \|\phi_{pk(e,i)}\|_{\Sigma_{pk(e,i)}^{-1}} = \sum_{i=1}^{i_{\text{max}}} \min\{1, \|\phi_{pk(e,i)}\|_{\Sigma_{pk(e,i)}^{-1}}\} \leq \ln \left( \frac{\det \Sigma_{pk(e,i_{\text{max}})}}{\det \Sigma_{pk}} \right) \leq \ln \det \Sigma_{pk(e,i_{\text{max}})}.
$$

The last step follows because $\Sigma_{pk(e,i_{\text{max}})} \geq \lambda I \succeq I$, so an upper bound to $\ln \det \Sigma_{pk(e,i_{\text{max}})}$ (see lemma 10 in Abbasi-Yadkori et al. [2011]). We can claim that an upper bound to eq. (113) is

$$
\leq A + \gamma(\sigma) \frac{D_p}{\sqrt{i_{\text{max}}}}.
$$

Since we’re summing over episode indexes where $E_{k(e,i)}$ holds, it follows that

$$
\frac{1}{i_{\text{max}}} \sum_{i = 1}^{i_{\text{max}}} \left[ \max_{\pi, \eta \in \mathcal{R}^d_{p}: \|\eta\|_{\Sigma_{pk(e,i)}} \leq \sqrt{\sigma}} \frac{-\phi_{\pi,p}^T \eta}{\sqrt{\sigma}} \right] \leq A + \gamma(\sigma) \frac{D_p}{\sqrt{i_{\text{max}}}}
$$

if each term in the summation in the lhs is $\geq \epsilon''$ (the condition is needed to apply lemma[14][Derandomization], if it does not hold the lemma’s thesis is satisfied). By lemma[17][Uncertainty Lemma]

$$
\max_{\pi, \eta \in \mathcal{R}^d_p: \|\eta\|_{\Sigma_{pk(e,i)}} \leq \sqrt{\sigma}} \frac{-\phi_{\pi,p}^T \eta}{\sqrt{\sigma}} \leq \max_{\pi, \eta \in \mathcal{R}^d_{p}: \|\eta\|_{\Sigma_{pk(e,i)}} \leq \sqrt{\sigma}} -\phi_{\pi,p}^T \eta
$$

Since the terms in the lhs of eq. (120) are strictly decreasing, the last one must be smaller than the average, which implies we must obtain

$$
\max_{\pi, \eta \in \mathcal{R}^d_{p}: \|\eta\|_{\Sigma_{pk(e,i)}} \leq \sqrt{\sigma}} -\phi_{\pi,p}^T \eta \leq \epsilon''
$$

after

$$
i_{\text{max}} \geq \left( \frac{\sqrt{\gamma_i(p)D_p + A}}{(\epsilon'')^2} \right)^2
$$

episodes provided that10

$$
\epsilon'' \geq \epsilon.
$$

We can finally compute how big $k_{\text{max}}$ (the total number of episodes in the epoch) needs to be: from definition[8][Good Event for FRANCIS] if

$$
k_{\text{max}} \geq \frac{1}{4} \times \frac{2 \ln \left( \frac{\delta}{\epsilon'} \right)}{1 - q}
$$

\[9\]notice that we are not accounting for the progress made in episodes where $E_k$ does not occur

\[10\]This condition is recurrent in this proof, and is used to invoke lemma[14][Derandomization] but if it doesn’t hold the thesis is automatically satisfied.
then we can write

$$\frac{i_{\text{max}}}{k_{\text{max}}} \geq \frac{1 - q}{2}.$$  \hspace{1cm} (126)

(recall $i_{\text{max}}$ is the the number of episodes where $\mathcal{E}_k$ occurs: $i_{\text{max}} = \sum_{k=1}^{K_{\text{max}}} \mathbb{I}\{\mathcal{E}_k \mid C_k\}$). Therefore, a total number of episodes

$$k_{\text{max}} = \left[ \frac{2}{1 - q} \times \left( \frac{\sqrt{\gamma_t(\rho)D_p} + A}{\epsilon''} \right)^2 \right]$$  \hspace{1cm} (127)

suffices (as this automatically satisfies eq. (125)).

Lemma 17 (Uncertainty Lemma). Let $\mathcal{K}$ and $\mathcal{K}$ be two generic episodes in an epoch $\mathcal{E}$ in phase $p$ such that $\mathcal{K} \geq k$. We have that

$$\max_{\pi, \eta \in \mathbb{R}^{d_p}}: \|\eta\|_{\Sigma_{\mathcal{K}}} \leq \sqrt{\phi^\top \pi, \Sigma_{\mathcal{K}} \phi} \leq \max_{\pi, \eta \in \mathbb{R}^{d_p}}: \|\eta\|_{\Sigma_{\mathcal{K}}} \leq \sqrt{\phi^\top \pi, \Sigma_{\mathcal{K}} \phi}.$$  \hspace{1cm} (128)

In addition, for positive real numbers $\rho_1 \leq \rho_2$ and a generic spd matrix $\Sigma$ we also have

$$\max_{\pi, \eta \in \mathbb{R}^{d_p}}: \|\eta\|_{\Sigma_{\mathcal{K}}} \leq \sqrt{\phi^\top \pi, \Sigma_{\mathcal{K}} \phi} = \sqrt{\frac{\rho_1}{\rho_2}} \max_{\pi, \eta \in \mathbb{R}^{d_p}}: \|\eta\|_{\Sigma_{\mathcal{K}}} \leq \sqrt{\phi^\top \pi, \Sigma_{\mathcal{K}} \phi}.$$  \hspace{1cm} (129)

Proof. Since $\Sigma_{\mathcal{K}} \geq \Sigma_{pk}$ (this notation means $\Sigma_{\mathcal{K}}$ is more positive definite than $\Sigma_{pk}$, more precisely $\phi^\top \Sigma_{\mathcal{K}} \phi \geq \phi^\top \Sigma_{pk} \phi$ for all $\phi$) we have the set inclusion

$$\{\eta \mid \|\eta\|_{\Sigma_{\mathcal{K}}} \leq \sqrt{\phi^\top \pi, \Sigma_{\mathcal{K}} \phi} \} \subseteq \{\eta \mid \|\eta\|_{\Sigma_{pk}} \leq \sqrt{\phi^\top \pi, \Sigma_{pk} \phi} \}$$  \hspace{1cm} (130)

Since we’re maximizing over a smaller set, the first result follows.

For the second statement, recall we can rewrite the programs in eq. (129) (see chapter 19 of [Lattimore and Szepesvári, 2020] about LinUCB or equivalently [Lemma 25](Linear Bandit Exploration Bonus)); here we identify the feature of an action in LinUCB with $\phi_{\pi, \mathcal{K}}$ as

$$\max_{\pi} \sqrt{\frac{\rho_1}{\rho_2}} \|\phi_{\pi, \mathcal{K}}\|_{\Sigma_{\mathcal{K}}}$$  \hspace{1cm} (131)

for the lhs and

$$\max_{\pi} \sqrt{\frac{\rho_1}{\rho_2}} \|\phi_{\pi, \mathcal{K}}\|_{\Sigma_{\mathcal{K}}}$$  \hspace{1cm} (132)

for the rhs, showing equality.
D.4 Learning a Phase

In this section we show how Francis learns a phase (i.e., the dynamics at a certain timestep) and compute the total number of episodes required to do so. This is where the explorability condition is used.

**Lemma 18 (Learning a Level).** Consider phase $p$ and let the following hypotheses hold

1. $\sum_{t=1}^{p-1} I(Q_t, Q_{t+1}) + \sqrt{\sigma_1} \|\bar{\phi}_{s,t}\|_{\Sigma^{-1}} \leq \tau$
2. $(\frac{\tau}{\epsilon})^2 \geq 2 \times 8d_p \ln \frac{2d_p}{\delta_H}$

Then after at most $e_{\text{max}} = O(1)$ and $e_{\text{max}}$ are defined in the proof

$$n(t) = \left[ \frac{2}{1 - q} \times \left( \sqrt{\gamma} (e_{\text{max}})(D_p + A)^2 \right) \times e_{\text{max}} = O \left( \frac{d_p^2 H^2\alpha_p}{\epsilon^2} \right) = O \left( \frac{d_p^2 H^2(d_p + d_{p+1})}{\epsilon^2} \right) \right]$$

episodes it must hold that

$$\max_{\sigma, \eta \in \mathbb{R}^d: ||\sigma||_{\Sigma_p} \leq \sqrt{\sigma_2}} \bar{\Phi}_{s,t}^\top \eta \leq \frac{\epsilon}{2H}$$

**Proof.** Let $\sigma_1, \sigma_2, \ldots$ be the sequences of the $\sigma$ parameter chosen in the different epochs, and additionally

$$\sigma_{\text{Start}} = \frac{1}{8d_p \ln \frac{2d_p}{\delta_H}}.$$  

We proceed by induction, with the following inductive hypothesis:

**Inductive Hypothesis 4.** In phase $p$ the following conditions hold

(a) $\lambda_{\text{min}}(\Sigma_{pk}(e,1)) \geq 8d_p \ln \frac{2d_p}{\delta_H} \sigma_e$ (at the beginning of epoch $e$)

(b) $\sigma_e = 2^{e-1} \sigma_{\text{Start}}$ (at the beginning of epoch $e$)

To show that the inductive hypothesis is satisfied in the base case ($e = 1$), notice that (b) holds by definition and (a) holds by setting $\lambda = 1$. Now we show the inductive step.

Since the inductive hypothesis satisfies the hypothesis of Lemma 16 (Learning an Epoch) on the good event [Definition 7 (Good Event for Francis)] it immediately follows that

$$\max_{\sigma, \eta \in \mathbb{R}^d: ||\sigma||_{\Sigma_p} \leq \sqrt{\sigma_2}} \bar{\Phi}_{s,t}^\top \eta \leq \epsilon'$$

after $k_{\text{max}}$ episodes (see Lemma 16 (Learning an Epoch)). Here in particular $k$ is the last episode of epoch $e$.

The explorability condition in Definition 2 (Explorability) implies that

$$\forall \eta \neq 0, \exists \pi \text{ such that } \bar{\Phi}_{s,t}^\top \eta \geq \nu_{\text{min}}.$$  

Consider the normalized vector $v$ corresponding to the minimum eigenvalue $q > 0$ for $\Sigma_{pk}$ and define:

$$\eta = qv.$$  

We’re interested in determining the maximum $q$ so that the constraint in the program eq. (136) is still satisfied, i.e., the condition below

$$\sigma_e \geq \|qv\|_{\Sigma_{pk}}^2 = (qv)^\top \Sigma_{pk} (qv) = q^2 \lambda_{\text{min}}(\Sigma_{pk})$$

gives the maximum value for $q$

$$q = \sqrt{\frac{\sigma_e}{\lambda_{\text{min}}(\Sigma_{pk})}}$$

in order for $qv$ to satisfy $\|qv\|_{\Sigma_{pk}} \leq \sqrt{\sigma_e}$. In other words, the $qv$ vector so defined is a feasible solution to the first program below, justifying one inequality:

$$\epsilon' \geq \max_{\sigma, ||\sigma||_{\Sigma_p} \leq \sqrt{\sigma_2}} \bar{\Phi}_{s,t}^\top \eta \geq \max_{\sigma} \bar{\Phi}_{s,t}^\top (qv) = \|qv\|_2 \max_{\pi} \bar{\Phi}_{s,t}^\top (qv)_{||qv||_2}$$

$$\geq \|qv\|_2 \nu_{\text{min}} = qv_{\text{min}} \geq \sqrt{\frac{\sigma_e}{\lambda_{\text{min}}(\Sigma_{pk})}} \nu_{\text{min}}.$$
Solving for $\lambda_{\min}$ gives:

$$\lambda_{\min}(\Sigma_{pk}) \geq \sigma_e \left( \frac{\nu_{\min}}{\epsilon} \right)^2 \geq \sigma_e \times 2 \times 8d_p \ln \frac{2d_p}{\delta^e} = \sigma_{e+1} \times 8d_p \ln \frac{2d_p}{\delta^e}$$ \hspace{1cm} (143)

Therefore the inductive hypothesis must hold for $e + 1$ as well, in other words, the statement in inductive hypothesis $4$ must hold for all $e$.

Now we determine the required value for $\rho$ at the end of the phase. We want to ensure

$$\max_{\pi, \eta \in \mathbb{R}^{dp}} \| \eta \|_{\Sigma_{pk}} \leq \sqrt{\alpha_p \phi^\top \pi, \eta} \leq \epsilon$$ \hspace{1cm} (144)

where now $k$ is the episode at the end of phase $p$. Since the inductive hypothesis holds in epoch $e$, lemma 16 ensures

$$\max_{\pi, \eta \in \mathbb{R}^{dp}} \| \eta \|_{\Sigma_{pk}} \leq \sqrt{\sigma_e \phi^\top \pi, \eta} \leq \epsilon''.$$ \hspace{1cm} (145)

We combine the above finding with a scaling argument given by lemma 17 [Uncertainty Lemma] that gives:

$$\max_{\pi, \eta \in \mathbb{R}^{dp}} \| \eta \|_{\Sigma_{pk}} = \sqrt{\sigma_e \phi^\top \pi, \eta} \leq \epsilon''.$$ \hspace{1cm} (146)

Requiring the above rhs to be $\leq \frac{\epsilon}{2H}$ gives a condition on the number of epochs $e_{\max}$ required ($e_{\max}$ is the number of epochs) and on $\sigma_{e_{\max}}$, setting $\epsilon'' = \epsilon$ gives

$$\sqrt{\frac{\sigma_e}{\sigma_{e_{\max}}} \frac{\alpha_p}{\epsilon}} \leq \frac{\epsilon}{2H} \rightarrow \sqrt{\frac{\sigma_{e_{\max}}}{\alpha_p}} \geq 2H$$ \hspace{1cm} (147)

$$\rightarrow \sigma_{e_{\max}} = 2^{e_{\max} - 1} \sigma_{Start} \geq 4H^2 \alpha_p$$ \hspace{1cm} (by induction) \hspace{1cm} (148)

$$\rightarrow 2^{e_{\max} - 1} \geq \frac{4H^2 \alpha_p}{\sigma_{Start}} \Rightarrow e_{\max} = \left[ 1 + \ln_2 \left( \frac{4H^2 \alpha_p}{\sigma_{Start}} \right) \right].$$ \hspace{1cm} (149)

In every epoch, $\epsilon'' = \epsilon$ and so the number of episodes necessary to achieve the required precision is (see lemma 16 [Learning an Epoch]):

$$\sum_{e=1}^{e_{\max}} \left[ \frac{2}{1 - q} \times \left( \sqrt{\gamma(e)D_p + A} \right)^2 \right]$$ \hspace{1cm} (150)

and since $\gamma(e)$ strictly increases with $e$ we can say that

$$\left[ \frac{2}{1 - q} \times \left( \sqrt{\gamma(e_{\max})D_p + A} \right)^2 \right] \times e_{\max}$$ \hspace{1cm} (151)

episodes suffices.
D.5 Learning to Navigate

In this section we show that FRANCIS “learns to navigate”, minimizing the least-square error in LSVI across timesteps.

Proposition 3 (Learning to Navigate). Assume that\[I(Q_t, Q_{t+1}) \leq \frac{\epsilon}{2H}\] (this is always satisfied by our assumptions on \(\epsilon\))

2. \((\frac{\epsilon}{2H})^2 \geq 2 \times 8d_p \ln \frac{2d_p}{\epsilon^2}\) (this is also always satisfied by our assumptions on \(\epsilon\))

Then after \(\tilde{O}\left(H^2 \sum_{t=1}^{H} \frac{d_t^2(d_t + d_{t+1})}{\epsilon^2}\right)\) episodes, outside of the failure event it holds that

\[\sum_{t=1}^{H} \left[I(Q_t, Q_{t+1}) + \sqrt{\alpha_t} ||\phi_{x,t}||_{\Sigma_t^{-1}}\right] \leq \epsilon, \quad \forall \pi\] (153)

and in particular

\[I(Q_t, Q_{t+1}) + \sqrt{\alpha_t} ||\phi_{x,t}||_{\Sigma_t^{-1}} \leq \frac{\epsilon}{H}, \quad \forall \pi, t \in [H].\] (154)

Proof. We proceed by induction over timesteps / phases \(p\):

Inductive Hypothesis 5 (Main Inductive Hypothesis). In phase \(p \in [H]\) it holds that

1. \(\sum_{t=1}^{p-1} I(Q_t, Q_{t+1}) + \sqrt{\alpha_t} ||\phi_{x,t}||_{\Sigma_t^{-1}} \leq \frac{\epsilon_{p-1}}{H} \epsilon\) (this ensures accuracy in LSVI)

2. \(\lambda_{\min}(\Sigma_t) \geq 4H^2 \alpha_t \quad t \in [p-1]\) (this ensures boundness of the iterates in LSVI)

The inductive hypothesis vacuously holds for \(p = 1\) (there is nothing to check). Now we show the inductive step. Assume the inductive hypothesis holds for a generic \(p - 1\), we want to show it still holds for \(p\). A direct application of [lemma 18](Learning a Level) gives (\(\Sigma_p\) is the covariance matrix after learning has completed):

\[\sqrt{\alpha_p} ||\phi_{x,p}||_{\Sigma_p^{-1}} \leq \frac{\epsilon}{2H}\] (155)

Adding \(I(Q_t, Q_{t+1}) \leq \frac{\epsilon}{2H}\) to both sides and adding the result to the equation in the inductive hypothesis proves the inductive step. The final number of episodes follows from summing the episodes needed in every phases according to [lemma 18](Learning a Level).
D.6 Solution Reconstruction (Main Result)

In this section we present our main result in a more formal way than in the main text; throughout the appendix the symbols are generally reported in table 2.

First, let us define the reward classes.

**Definition 6 (Reward Classes).** Consider an MDP \( \mathcal{M}(S, A, p, r, H) \) without any reward function. Fix a misspecification function \( \Delta^r(t, \cdot, \cdot) : S \times A \to \mathbb{R} \) for every \( t \in [H] \) which can depend on the state and action pair, and is subject to the constraint

\[
\forall (\pi, t) \quad |\mathbb{E}_{x_t \sim \pi} \Delta^r_t(x_t, \pi_t(x_t))| \overset{def}{=} |\Sigma^r_{\pi, t}| \leq E_t. \tag{157}
\]

Define the following class \( \mathcal{I} \) (Implicit Regularity) of (expected) reward functions \( (r_1, \ldots, r_H) \) on \( \mathcal{M} \), parameterized by \((\theta_1^r, \ldots, \theta_H^r)\) and satisfying \( \forall (s, a, t, \pi) \in S \times A \times [H] \times \Pi \) (here \( \Pi \) is the policy space):

1. \( r_t(s, a) = \phi_t(s, a) \top \theta_1^r + \Delta_t^r(s, a) \)
2. \( |\Delta_t^r(s, a)| \leq 1 \)
3. \( |\theta_1^r|_2 \leq \frac{1}{H} \).

In addition, define the following class \( \mathcal{E} \) (Explicit Regularity) of (expected) reward functions \( (r_1, \ldots, r_H) \) on \( \mathcal{M} \) parameterized by \((\theta_1^r, \ldots, \theta_H^r)\) satisfying \( \forall (s, a, t, \pi) \in S \times A \times [H] \times \Pi \):

1. \( r_t(s, a) = \phi_t(s, a) \top \theta_1^r + \Delta_t^r(s, a) \)
2. \( |\Delta_t^r(s, a)| \leq 1 \)
3. \( |\theta_1^r|_2 \leq \frac{1}{H} \).

Under explicit regularity the bound on \(|\theta_1^r|_2\) constrains the maximum value the reward can take; instead, under implicit regularity we do not have such requirement, as only the expectation is controlled. This implies the local reward can be much larger than the expectation, making this a much harder setting.

We are now ready to present the main result formally.

**Theorem 1 (Restating theorem 4.1 formally).** Consider an MDP \( \mathcal{M} \) and a feature extractor \( \phi \) satisfying \( \|\phi_t(s, a)\|_2 \leq 1 \) for every \((s, a) \in S \times A\) and fix two classes of reward functions \( \mathcal{I} \) and \( \mathcal{E} \) according to Definition 6 [Reward Classes].

Set \( \epsilon \) to satisfy \( \epsilon \geq \Omega(d_H (\mathcal{I}(Q_1, Q_{t+1}) + E_1)) \) and \( \epsilon \leq \tilde{O}(\nu_{\min}/\sqrt{d_t}) \) for all \( t \in [H] \).

**Francis** always terminates after \( \tilde{O} \left( H^2 \sum_{t=1}^{H} \frac{d_t^2 (d_t + d_{t+1})}{\epsilon^2} \right) \) episodes (with probability one), returning a dataset \( \mathcal{D} = \{(s_{t_k}, a_{t_k}, s^+_{t+1,k})\}_{t_k=1, \ldots, \nu_{\min}^\nu, k=1, \ldots, \tilde{O}^\nu} \) of the collected state-action-successor states \((s_{t_k}, a_{t_k}, s^+_{t+1,k})\) in episode \( k \in [\nu_{\min}] \) for each timestep \( t \in [H] \).

Now consider any reward function \( r \in \mathcal{E} \) or \( r \in \mathcal{I} \) and the MDP induced by that reward function \( \mathcal{M}(S, A, p, r, H) \), and replace each tuple \((s_{t_k}, a_{t_k}, s^+_{t+1,k})\) in \( \mathcal{D} \) with \((s_{t_k}, a_{t_k}, r_{t_k}, s^+_{t+1,k})\) where \( r_{t_k} \) satisfies

\[
r_{t_k} = r_t(s_{t_k}, a_{t_k}) + \eta^r \tag{158}
\]

where \( \eta^r \) is 1-sub-Gaussian noise.

Then with probability at least \( 1 - \delta \), the batch LSVI algorithm run on \( \mathcal{D} \) (see algorithm 3) returns a policy \( \pi \) such that on \( \mathcal{M} \)

\[
\mathbb{E}_{x_1 \sim \rho}(V^\pi_1(x_1) - V^\pi_1(x_1)) \leq \frac{\epsilon}{\nu_{\min}}. \tag{159}
\]

if \( r \in \mathcal{I} \) and

\[
\mathbb{E}_{x_1 \sim \rho}(V^\pi_1(x_1) - V^\pi_1(x_1)) \leq \epsilon. \tag{160}
\]

if \( r \in \mathcal{E} \).

We have expressed the theorem in its full generality, but if the reward function is prescribed a posteriori through an oracle then we expect the noise \( \eta^r \) in eq. (158) to be absent. In general, if the reward function is prescribed a posteriori then it should be prescribed as a linear function (in the chosen features) to avoid any additional error in the LSVI procedure. Finally the reward misspecification \( \Delta_t^r(\cdot, \cdot) \) can depend on the parameter \( \theta \) if it is a Lipshitz function of \( \theta \). Alternatively, if it is a discontinuous function of \( \theta \) then same-order guarantees are still recovered if eq. (157) is replaced with \( \mathbb{E}(\pi, s, a \mid r) \tau(s, a) \leq E_t \).
Proof. (of the main result) Let \( n(t) \) the number of samples collected at each level (notice that we only store one sample every trajectory, so the number of samples equals the number of trajectories / number of episodes), according to [lemma 18][Learning a Level] Using the assumptions on \( \epsilon \) (these conditions are used in the good event for LSVI in [definition 5][Good Event for LSVI]) we can ensure:

\[
\sqrt{\frac{n(t)}{\alpha t}} E_t = \sqrt{\frac{n(t)}{\alpha t}} E_t \sqrt{\alpha t} = \tilde{O} \left( \frac{d t H \sqrt{\alpha t}}{\sqrt{\alpha t \epsilon}} \right) E_t \sqrt{\alpha t} \leq \sqrt{\alpha t / 3} \tag{161}
\]

\[
\sqrt{n(t)} I(Q_t, Q_{t+1}) = \sqrt{n(t)} I(Q_t, Q_{t+1}) \sqrt{\alpha t} = \tilde{O} \left( \frac{d t H \sqrt{\alpha t}}{\sqrt{\alpha t \epsilon}} \right) I(Q_t, Q_{t+1}) \sqrt{\alpha t} \leq \sqrt{\alpha t / 3}. \tag{162}
\]

We assume we are in the good event for FRANCIS, see [definition 7][Good Event for FRANCIS] which occurs with probability \( 1 - \delta \) according to [lemma 19][Probability of Good Event for FRANCIS]. We apply [proposition 3][Learning to Navigate] which gives the stated number of episodes to termination and the condition satisfied by the samples in the dataset \( D \) (through the covariance matrices \( \Sigma^{-1}_t \)):

\[
\sum_{t=1}^{H} \left[ I(Q_t, Q_{t+1}) + \sqrt{\alpha t} \| \hat{\theta}_{x,t} \|_{\Sigma^{-1}_t} \right] \leq \epsilon, \quad \forall \pi \tag{163}
\]

Now, under implicit regularity [lemma 20][Reward Boundness] ensures (the lemma requires \( E_t \leq \frac{1}{H} \), which is always satisfied since we must have \( \epsilon < 1 \) to produce any useful result, and from the theorem hypothesis \( E_t \leq \epsilon / (d t H) \leq 1 / H \))

\[
\| \theta_t^r \|_2 \leq \frac{2}{H \nu_{\min}} \text{ def } = \frac{R}{H}. \tag{164}
\]

Finally, [proposition 2][Batch LSVI Guarantees (algorithm 3)] ensures that LSVI in algorithm 3 returns a value function \( V \) and policy \( \hat{\pi}^r \) such that

\[
E_{x_1 \sim \rho} \left( V_t^* - \hat{V}_t \right)(x_1) \leq \sum_{t=1}^{H} \left[ 2 E_t + R \left( I(Q_t, Q_{t+1}) + \sqrt{\alpha t} \| \hat{\theta}_{\hat{x},t} \|_{\Sigma^{-1}_t} \right) \right]
\]

\[
E_{x_1 \sim \rho} \left( \hat{V}_t - \hat{V}_t^{\hat{\pi}^r} \right)(x_1) \leq \sum_{t=1}^{H} \left[ 2 E_t + R \left( I(Q_t, Q_{t+1}) + \sqrt{\alpha t} \| \hat{\theta}_{\hat{x},t} \|_{\Sigma^{-1}_t} \right) \right]. \tag{165}
\]

Using eq. (163) (and recalling \( E_t \leq \epsilon \) by hypothesis of the theorem) to further simplify it we obtain:

\[
E_{x_1 \sim \rho} \left( V_t^* - \hat{V}_t \right)(x_1) \leq 2 R \epsilon
\]

\[
E_{x_1 \sim \rho} \left( \hat{V}_t - \hat{V}_t^{\hat{\pi}^r} \right)(x_1) \leq 2 R \epsilon.
\]

Summing the two expression gives:

\[
E_{x_1 \sim \rho} \left( V_t^* - \hat{V}_t^{\hat{\pi}^r} \right)(x_1) \leq 4 R \epsilon.
\]

Rescaling \( \epsilon \) by 4 and substituting the value for \( R \) gives the thesis under implicit regularity.

Under explicit regularity the steps are the same, but now

\[
\| \theta_t^r \|_2 \leq \frac{1}{H} \text{ def } = \frac{R}{H} \tag{166}
\]

is explicitly prescribed, and the thesis immediately follows.

The generality of the main result allows us to immediately obtain the following corollary:

**Corollary 1** (Learning a Prescribed Reward Function during the Execution). Under the same assumptions as theorem 2 assume the reward function \( r \in \mathcal{R} \) or \( r \in \mathcal{R} \) is prescribed before the execution of FRANCIS and

\[
r_{t,k} = r_t(s_{t,k}, a_{t,k}) + \eta^r \tag{167}
\]

where \( \eta^r \) is 1-sub-Gaussian noise. Assume \( (s_{t,k}, a_{t,k}, r_{t,k}, \eta^r_{t+1,k}) \) is stored in the dataset \( D \).
Then with probability at least $1 - \delta$, the batch LSVI algorithm run on $D$ (see algorithm 3) returns a policy $\pi$ such that on $\mathcal{M}$

$$
\mathbb{E}_{x_1 \sim \rho}(V_1^\pi - V_1^\pi)(x_1) \leq \frac{\epsilon}{\nu_{\min}}.
$$

(168)

if $r \in \mathcal{I}$ and

$$
\mathbb{E}_{x_1 \sim \rho}(V_1^\pi - V_1^\pi)(x_1) \leq \epsilon.
$$

(169)

if $r \in \mathcal{J}$ and

$$
\mathbb{E}_{x_1 \sim \rho}(V_1^\pi - V_1^\pi)(x_1) \leq \epsilon.
$$

(169)

if $r \in \mathcal{E}$. 

D.7 Computational Complexity

Theorem 4.1 gives a bound on the number of episodes to termination. In every episode, a multivariate normal vector is sampled (which can be done efficiently) and LSVi is invoked.

Assume $d_1 = \cdots = d_H = d$ for simplicity; a naive implementation would factorize and store the new covariance matrix at the end of a phase (total of $\tilde{O}(Hd^3)$ work across all phases); after this, computing the $\hat{\theta}_t$’s requires $\tilde{O}(H(d^2 + Ad) \times n_{\text{episodes}})$ computations at every episode where $n_{\text{episodes}}$ is the total number of episodes at termination given in theorem 4.1.
Definition 7 (Good Event for FRANCIS). We say the good event for FRANCIS occurs if for all timesteps $t \in [H]$ or phases $p \in [H]$ and episodes $k$ in that phase the following bounds jointly hold and we are in the good event for LSVI (see definition 5 (Good Event for LSVI)).

$$\left| \frac{1}{i_{\max}} \sum_{i=1}^{i_{\max}} \zeta_{p,k(e,i)} \right| \leq \sqrt{\frac{2(2L_{\phi} R(t))^{2} \ln \left( \frac{1}{\delta} \right)}{i_{\max}}} = \sqrt{\frac{8 \ln \left( \frac{1}{\delta} \right)}{i_{\max}}} \overset{df}{=} A \sqrt{i_{\max}}$$

(170)

$$\|\xi_{t,k(e,i)}\|_{2} \overset{df}{=} \sqrt{\frac{2 \sigma_{t} d_{t} \ln \frac{2 d_{t}}{\delta'}}{\lambda_{\min}(\Sigma_{p,k(e,i)})}} \ln \frac{2 d_{t}}{\delta'}$$

(171)

$$\frac{1}{k_{\max}} \sum_{k=1}^{k_{\max}} \mathbb{1}\{E_{k} | C_{k}\} \geq (1-q) - \sqrt{\frac{2 \ln \left( \frac{1}{\delta'} \right)}{k_{\max}}}$$

(172)

Lemma 19 (Probability of Good Event for FRANCIS). There exists a parameter $\delta'' = \text{poly}(d_{1}, \ldots, d_{H}, H, \frac{1}{\epsilon})$, such that the good event of definition 7 holds with probability at least $1 - \delta$.

Proof. The first and fourth inequality follow from lemma 24 (Azuma-Hoeffding Inequality). The second and third inequality follow from lemma 22 (Large Deviation Multivariate Normal). In particular, a union bound over the statements, over $H$ and over the number of episodes ensures all statements jointly hold at any point during the execution of the program; from this, the value for $\delta''$ can be determined. 

Some symbols, like $i_{\max}, k_{\max}$ are defined directly in the lemma where the bound is used.
E Lower Bound

We sketch the lower bound to highlight that explorability is required.

**Proposition 4** (Lower Bound on Explorability Dependence under Implicit Regularity). There exists an MDP and a feature map \( \phi : (s, a) \mapsto \phi(s, a) \in \mathbb{R}^2 \) with explorability parameter \( \nu_{\text{min}} \) and a reward function such that:

\[
\forall (\pi, t) \quad r_t(s, a) = \phi_t(s, a)^\top \theta^*_\pi, \quad |E_{x_t \sim \pi} r_t(x_t, \pi_t(x_t))| \leq 1
\]

and yet no reinforcement learning agent without knowledge of \( \theta^* \) can return an \( \epsilon \)-optimal policy for \( \epsilon \leq \nu_{\text{min}} \leq \frac{1}{4} \) in less than \( \Omega(1/(\epsilon \nu_{\text{min}})^2) \) trajectories with probability higher than \( 2/3 \).

Notice that the proposition above is for a fixed (but unknown) deterministic reward function; this is thus a special case of the reward-free learning setting we consider, implying that the hardness is due to the implicit regularity conditions rather than to reward-free learning.

The proof essentially uses a multi-armed bandit lower bound where the noise is \( 1/\nu_{\text{min}} \) scaled by the inverse explorability coefficient \( \pi/\epsilon \).

**Proof.** We construct the MDP as follows: there is a single starting state \( s_{\text{start}} \) with two actions \( a_L \) and \( a_R \), and the identity feature \( \phi_0(s_{\text{start}}, a_L) = e_1, \phi_0(s_{\text{start}}, a_R) = e_2 \), where \( e_1, e_2 \) are canonical vectors in \( \mathbb{R}^2 \). Now fix a scalar \( \epsilon \in [-\nu_{\text{min}}, \nu_{\text{min}}/2] \):

1. action \( a_L \) gives an immediate reward \(-1/2\) and leads to state \( s_{L1} \) with probability \( 1/2 + \nu_{\text{min}} \) and to \( s_{L2} \) with probability \( 1/2 - \nu_{\text{min}} \). The feature map reads \( \phi_2(s_{L1}) = e_1 + \phi_2(s_{L2}) = -e_1 \) in the only action available in each state.

2. action \( a_R \) gives an immediate reward \(-1/2\) and leads to state \( s_{R1} \) with probability \( 1/2 + \nu_{\text{min}} + \epsilon \) and to \( s_{R2} \) with probability \( 1/2 - \nu_{\text{min}} - \epsilon \). The feature map reads \( \phi_2(s_{R1}) = e_2 \) and \( \phi_2(s_{R2}) = -e_2 \)

In this MDP there are only two distinct policies: \( \pi_L \) that selects \( a_L \) first and then the only available action in either \( s_{L1} \) or \( s_{L2} \), and \( \pi_R \) that selects \( a_R \) first and then the only available action in either \( s_{R1} \) or \( s_{R2} \). Therefore, this is equivalent to a multiarmed bandit problem with reward \(-1/2 + \phi_{\pi_L,2}^\top \theta_L^* \) for \( \pi_L \) and \(-1/2 + \phi_{\pi_R,2}^\top \theta_L^* \) for \( \pi_R \). The minimum explorability coefficient is \( (\nu_{\text{min}} = 1 \text{ at timestep 1}) \)

\[
\min_{\theta \neq 0} \max_{\pi} \| \phi_{\pi,2}^\top \theta \|_2 \geq \frac{1}{2} + \nu_{\text{min}} - \nu_{\text{min}} 2
\]

\[
\text{corresponding to policy } \pi_R \text{ (this can be computed by inspection; notice that } \pi_L \text{ yields the same } \nu_{\text{min}} \text{). Now consider the reward parameter } \theta_L^* \text{ is } 1/(\nu_{\text{min}} \times [1/2, 1/2]); \text{ the expected reward at timestep 2 under policy } \pi_R \text{ is } E_{x_2 \sim \nu_{\text{min}}} r_2(x_2) = \nu_{\text{min}} \times 1/(\nu_{\text{min}}) \leq 1 \text{ which satisfies the assumptions of the lemma. At the same time } E_{x_2 \sim \nu_{\text{min}}} r_2(x_2) = \nu_{\text{min}} \times 1/(\nu_{\text{min}}) \leq 1. \text{ This implies the random return } -1/2 + \phi_2(s)^\top \theta_L^* \text{ with } s \sim p_1(s_{L1}, a_L) \text{ is a scaled and shifted Bernoulli random variable with mean zero, taking the values } -1/2 + 1/(\nu_{\text{min}}) \text{ and } -1/2 - 1/(\nu_{\text{min}}). \text{ Since the standard deviation of this random variables (with } \nu_{\text{min}} \leq 1/4) \text{ is } \Omega(1/\nu_{\text{min}}), \text{ this random variable must be } \Omega(1/\nu_{\text{min}}) \text{-sub-Gaussian}. \text{ The same reasoning applies to } -1/2 + \phi_2(s)^\top \theta_L^* \text{ with } s \sim p_1(s_{R1}, a_R). \text{ Notice that both expectations are at most 1.}

Solving this class of problems (parameterized by \( \epsilon \)), i.e., identifying an \( |\epsilon|/2 \)-optimal policy is equivalent to solving a multiarmed bandit problem with 2 actions (corresponding to the policies \( \pi_L \) and \( \pi_R \)). This construction is exactly the same as theorem 2 from [Krishnamurthy et al., 2016] with shifted Bernoulli random variables that are scaled by the inverse explorability coefficient \( 1/\nu_{\text{min}} \). This implies that a sample complexity \( \Omega(1/(\nu_{\text{min}}|\epsilon|)^2) \) is required to output an \( |\epsilon|/2 \)-optimal policy with probability > 2/3.

---

14 See for example exercise 2.5 in [Wainwright, 2019].
Lemma 20 (Reward Boundness). If we assume that

\[ \forall \pi \quad |E_{x_t \sim \pi} r_t(x_t, \pi_t(x_t))| \leq \frac{1}{H} \]  

(176)

and \( \exists \theta_t^e \in \mathbb{R}^{d_t} \) such that \( |E_{x_t \sim \pi} r_t(x_t, \pi_t(x_t)) - \Phi_{x_t}^T \theta_t^e| \leq E_t \leq \frac{1}{H} \)  

(177)

then it follows that

\[ \|\theta_t^e\|_2 \leq \frac{2}{H \nu_t}. \]  

(178)

Proof. From the hypothesis it follows

\[ \frac{2}{H} \geq |\Phi_{x_t}^T \theta_t^e| = \|\theta_t^e\|_2 \times |\Phi_{x_t}^T \theta_t^e| \frac{\theta_t^e}{\|\theta_t^e\|_2} \]  

(179)

in particular this must hold for the policy \( \pi \) that maximizes the above display. Therefore, after taking \( \max_{\pi} \), take \( \min_{\|\theta\|_2 = 1} \) to obtain (using \textbf{Definition 2} [Explorability]):

\[ \geq \|\theta_t^e\|_2 \times \min_{\|\theta\|_2 = 1} \max_{\|\Phi_{x_t}^T \theta\|} = \|\theta_t^e\|_2 \nu_t. \]  

(180)

Rearranging

\[ \|\theta_t^e\|_2 \leq \frac{2}{H \nu_t}. \]  

(181)
F.1 High Probability Bounds

Lemma 21 (Transition Noise High Probability Bound). If $\lambda = 1$ and $R = 2L_\phi \mathcal{R}_{t+1}$ with probability at least $1 - \delta'$ it holds that $\forall V_{t+1} \in \mathcal{V}_{t+1}$:

$$\left\| \sum_{i=1}^{k-1} \phi_t (V_{t+1}(s_{t+1,k}^+)) - \mathbb{E}_{s' \sim p(s_{i,k},a_k)} V_{t+1}(s') \right\|_{\Sigma_{t+1}} \leq \sqrt{\beta_t^2}$$

where:

$$\sqrt{\beta_t^2} \overset{def}{=} \sqrt{2} \times 2 \frac{d_i}{2} \ln \left( 1 + \frac{L_k^2 k/d_i} {d_i} \right) + d_{t+1} \ln (1 + 4 \mathcal{R}_{t+1} / (2L_\phi \sqrt{k})) + \ln \left( \frac{1}{\delta'} \right) + 2. \quad (183)$$

Proof. Since the statement needs to hold for every $V_{t+1} \in \mathcal{V}_{t+1}$, we start by constructing an $\epsilon$-cover for set $\mathcal{V}_{t+1}$ using the supremum distance. To achieve this, we construct an $\epsilon$-cover for the parameter $\theta \in \mathcal{B}_{t+1}$ using the “Covering Number of Euclidean Ball” lemma in [Zanette et al., 2020b]. This ensures that there exists a set $D_{t+1} \subseteq \mathcal{B}_{t+1}$, containing $(1 + 2 \mathcal{R}_{t+1} / \epsilon')^d_{t+1}$ vectors $\theta_{t+1}$ that well approximates any $\theta_{t+1} \in \mathcal{B}_{t+1}$:

$$\exists \mathcal{D}_{t+1} \subseteq \mathcal{B}_{t+1} \text{ such that } \forall \theta_{t+1} \in \mathcal{B}_{t+1}, \exists \theta_{t+1} \in \mathcal{D}_{t+1} \text{ such that } \|\theta_{t+1} - \theta_{t+1}\|_2 \leq \epsilon'. \quad (184)$$

Let $\hat{V}_{t+1}(s) \overset{def}{=} \max_a \phi_{t+1}(s,a)^\top \theta$, where $\hat{\theta} = \arg\min_{\theta' \in \mathcal{D}_{t+1}} \|\theta' - \theta\|_2$ and consider $V_{t+1} \in \mathcal{V}_{t+1}$. For any fixed $s \in \mathcal{S}$ we have that:

$$|\hat{V}_{t+1} - \hat{V}_{t+1}(s)| = |\max_{a'} \phi_{t+1}(s,a')^\top \theta_{t+1} - \max_{a''} \phi_{t+1}(s,a'')^\top \theta_{t+1}|$$

$$\leq \max_a |\phi_{t+1}(s,a)^\top (\theta_{t+1} - \hat{\theta})|$$

$$\leq \max_a \|\phi_{t+1}(s,a)\|_2 \|\theta_{t+1} - \theta_{t+1}\|_2$$

$$\leq L_\phi \epsilon'. \quad (185)$$

By using the triangle inequality we can write:

$$\left\| \sum_{i=1}^{k-1} \phi_t (V_{t+1}(s_{t+1,k}^+)) - \mathbb{E}_{s' \sim p(s_{i,k},a_k)} V_{t+1}(s') \right\|_{\Sigma_{t+1}}$$

$$\leq \left\| \sum_{i=1}^{k-1} \phi_t (\hat{V}_{t+1}(s_{t+1,k}^+)) - \mathbb{E}_{s' \sim p(s_{i,k},a_k)} \hat{V}_{t+1}(s') \right\|_{\Sigma_{t+1}} +$$

$$\left\| \sum_{i=1}^{k-1} \phi_t \left( \hat{V}_{t+1}(s_{t+1,k}^+)) - \mathbb{E}_{s' \sim p(s_{i,k},a_k)} \hat{V}_{t+1}(s') \right) \right\|_{\Sigma_{t+1}}$$

$$\leq \left\| \sum_{i=1}^{k-1} \phi_t \left( \hat{V}_{t+1}(s_{t+1,k}^+)) - \mathbb{E}_{s' \sim p(s_{i,k},a_k)} \hat{V}_{t+1}(s') \right) \right\|_{\Sigma_{t+1}} +$$

$$\left\| \sum_{i=1}^{k-1} \phi_t \left( \hat{V}_{t+1}(s_{t+1,k}^+)) - \mathbb{E}_{s' \sim p(s_{i,k},a_k)} \hat{V}_{t+1}(s') \right) \right\|_{\Sigma_{t+1}} \quad (186)$$

Each of the last two terms above can be written for some $b_i$’s (different for each of the two terms) as $\left\| \sum_{i=1}^{k-1} \phi_t b_i \right\|_{\Sigma_{t+1}}$. The projection lemma, (lemma 8 from Zanette et al., 2020b) ensures:

$$\left\| \sum_{i=1}^{k-1} \phi_t b_i \right\|_{\Sigma_{t+1}} \leq L_\phi \epsilon' \sqrt{k} \quad (187)$$

We have used eq. 185 to bound the $b_i$’s. Now we examine the first term of the rhs in equation in eq. 185. In particular, we bound that term for a generic $\hat{V}_{t+1}$ and then do a union bound over all possible $V_{t+1}$, which are generated by finitely many $\hat{\theta}_{t+1} \in \mathcal{D}_{t+1}$ as explained before. We obtain that:

$$\mathbb{P} \left( \bigcup_{\hat{\theta}_{t+1} \in \mathcal{D}_{t+1}} C(\hat{\theta}_{t+1}) \right) \leq \sum_{\hat{\theta}_{t+1} \in \mathcal{D}_{t+1}} \mathbb{P} \left( C(\hat{\theta}_{t+1}) \right) \leq (1 + 2 \mathcal{R}_{t+1} / \epsilon')^{d_{t+1} \delta''} \overset{def}{=} \delta' \quad (188)$$
where $C$ is the event reported below (along with $\delta''$) and the last inequality above follows from Theorem 1 in [Abbasi-Yadkori et al., 2011] (the random variables $\hat{V}_{t+1}(\cdot)$ and $\hat{V}_{t+1}(\cdot)$ are $R = 2L_\phi R_{t+1}$-subgaussian by construction):

$$C(\theta_{t+1}) \overset{\Delta}{=} \left\{ \left\| \sum_{i=1}^{k-1} \phi_{ti} \left( \hat{V}_{t+1,i} - E_{s' \sim p(s_{t+1:a_{t+1}})} \hat{V}(s') \right) \right\|_2^2 > 2 \times (R)^2 \ln \left( \frac{\det(\Sigma_t)^{\frac{1}{2}} \det(\Lambda^{\frac{1}{2}})}{\delta''} \right) \right\}.$$  \hspace{1cm} (189)

In particular, we set

$$\delta'' = \frac{\delta'}{1 + 2R_{t+1}/\epsilon'} d_{t+1}$$ \hspace{1cm} (190)

from the prior display and so with probability $1 - \delta'$ (after a union bound over all possible $\theta_{t+1} \in D_{t+1}$) we have upper bounded eq. (186) by:

$$R \sqrt{2 \ln \left( \frac{\det(\Sigma_t)^{\frac{1}{2}} \det(\Lambda^{\frac{1}{2}})}{\delta'} \right) + \frac{d_{t+1}}{\delta'}} + 2L_\phi \epsilon' \sqrt{k}.$$ \hspace{1cm} (191)

If we now pick

$$\epsilon' = \frac{R}{2L_\phi \sqrt{k}}$$ \hspace{1cm} (192)

we get:

$$R \sqrt{2 \ln \left( \frac{\det(\Sigma_t)^{\frac{1}{2}} \det(\Lambda^{\frac{1}{2}}) (1 + 2R_{t+1}/\epsilon') d_{t+1}}{\delta'} \right)} + R$$ \hspace{1cm} (193)

$$= \sqrt{2} R \sqrt{\frac{1}{2} \ln(\det(\Sigma_t)) - \frac{d_{t+1}}{2} \ln(\Lambda) + d_{t+1} \ln(1 + 2R_{t+1}/\epsilon') + \ln \left( \frac{1}{\delta'} \right) + R}.$$ \hspace{1cm} (194)

Finally, using the Determinant-Trace Inequality (see lemma 10 of [Abbasi-Yadkori et al., 2011]) we obtain

$$\det(\Sigma_t) \leq (\lambda + L_\phi^2 k / d_t)^{d_{t+1}}$$ and so (with $\lambda = 1$)

$$\leq \sqrt{2} \times 2 \sqrt{\frac{d_{t+1}}{2} \ln \left( 1 + L_\phi^2 k / d_t \right) + d_{t+1} \ln(1 + 4R_{t+1}/(2L_\phi \sqrt{k})) + \ln \left( \frac{1}{\delta'} \right)} + 2 \overset{\Delta}{=} \sqrt{2} \beta_t.$$ \hspace{1cm} (195)
F.2 Known Results

**Lemma 22** (Large Deviation Multivariate Normal). Let \( \Sigma \in \mathbb{R}^{d \times d} \) be an spd matrix with minimum eigenvalue \( \lambda > 0 \) and let

\[
\xi \sim \mathcal{N} (0, \sigma \Sigma^{-1})
\]

for a positive scalar \( \sigma \). For any fixed \( \phi \in \mathbb{R}^d \) with probability at least \( 1 - \delta' \):

\[
|\phi^\top \xi|^2 \leq \frac{\sigma^2 \|\phi\|_2^2}{\lambda} \left( 2d \ln \frac{2d}{\delta'} \right)
\]

and so by choosing \( \phi = \frac{\xi}{\|\xi\|_2} \) when \( \xi \neq 0 \) it holds that

\[
\|\xi\|_2 \leq \sqrt{\sigma \lambda \left( 2d \ln \frac{2d}{\delta'} \right)}.
\]

Under the same event it holds that

\[
\|\xi\|_{\Sigma} \leq \sqrt{\sigma \left( 2d \ln \frac{2d}{\delta'} \right)}.
\]

**Proof.** If

\[
\xi \sim \mathcal{N} (0, \sigma \Sigma^{-1})
\]

it follows that

\[
\frac{1}{\sqrt{\sigma}} \Sigma^{1/2} \xi \sim \mathcal{N} (0, I)
\]

where \( I \) is the identity matrix on \( \mathbb{R}^d \). Therefore

\[
\frac{1}{\sigma} \|\xi\|_2^2 = \left( \frac{1}{\sqrt{\sigma}} \xi^\top \Sigma^{1/2} \right)^\top \left( \frac{1}{\sqrt{\sigma}} \Sigma^{1/2} \xi \right) \sim \chi_d^2
\]

where \( \chi_d^2 \) is the chi-square distribution with \( d \) degrees of freedom. From **Lemma 23** (\( \chi \)-square lemma) we can compute a high probability bound for the above random variable (this also proves the last statement):

\[
|\phi^\top \xi|^2 \leq \|\phi\|_2^2 \|\xi\|_2^2 \leq \|\phi\|_2^2 \frac{\sigma}{\lambda} \|\xi\|_2^2 \leq \frac{\sigma^2 \|\phi\|_2^2}{\lambda} \left( 2d \ln \frac{2d}{\delta'} \right)
\]

with probability at least \( 1 - \delta' \).

\[\square\]

**Lemma 23** (\( \chi \)-square lemma). Let \( X^2 \sim \chi_d^2 \) be a random variable that follows the chi-square distribution with \( d \) degrees of freedom. With probability at least \( 1 - \delta' \):

\[
X^2 \leq 2d \ln \frac{2d}{\delta'}.
\]

**Proof.** Let \( X_i \sim \mathcal{N}(0, 1), i \in [d] \). If \( X_i \in [-a, a], \forall i \in [d] \) then it must follow that \( \sum_{i \in [d]} X_i^2 \leq da^2 \).

Thus:

\[
P(X^2 = \sum_{i \in [d]} X_i^2 \geq da^2) \leq P(\exists i \in [d], X_i \notin [-a, a]) = P(\cup_{i \in [d]} X_i \notin [-a, a]) \leq d P(X_i \notin [-a, a]) \leq 2de^{-a^2/2}.
\]

Requiring the rhs above to be \( \leq \delta' \) gives

\[
a^2 = 2 \ln \frac{2d}{\delta'}.
\]

\[\square\]

**Lemma 24** (Azuma-Hoeffding Inequality). Let \( X_i \) be a martingale difference sequence such that \( X_i \in [-A, A] \) for some \( A > 0 \). Then with probability at least \( 1 - \delta' \) it holds that:

\[
\left| \sum_{i=1}^n X_i \right| \leq \sqrt{2A^2 n \ln \left( \frac{1}{\delta'} \right)}.
\]
Proof. The Azuma inequality reads:

\[
P \left( \left| \sum_{i=1}^{n} X_i \right| \geq t \right) \leq e^{-\frac{2t^2}{4A^2 n}},
\]

see for example [Wainwright, 2019]. From here setting the rhs equal to \( \delta' \) gives:

\[
t \overset{\text{def}}{=} \sqrt{2A^2 n \ln \left( \frac{1}{\delta'} \right)}.
\]

Lemma 25 (Change of \( \Sigma \)-Norm). For a compatible vector \( x \in \mathbb{R}^d \) and an spd matrix \( \Sigma \in \mathbb{R}^{d \times d} \) with minimum eigenvalue \( \lambda_{\min}(\Sigma) \) we have

\[
\|x\| \geq \sqrt{\lambda_{\min}(\Sigma)} \|x\|_2 (208)
\]

\[
\|x\|_{\Sigma^{-1}} \leq \frac{1}{\sqrt{\lambda_{\min}(\Sigma)}} \|x\|_2. (209)
\]

Proof. We show one inequality (the other is identical). Consider the eigendecomposition of \( \Sigma \) with orthonormal eigenvectors \( v_i \)'s and eigenvalues \( \lambda_i \)'s:

\[
\Sigma^{-1} = \sum_{i=1}^{d} \lambda_i^{-1} v_i v_i^T (210)
\]

We can write:

\[
\|x\|_{\Sigma^{-1}}^2 = x^T \Sigma^{-1} x (211)
\]

\[
= x^T \left( \sum_{i=1}^{d} \lambda_i^{-1} v_i v_i^T \right) x (212)
\]

\[
= \sum_{i=1}^{d} \frac{1}{\lambda_i} (v_i^T x)^2 (213)
\]

\[
\leq \frac{1}{\lambda_{\min}(\Sigma)} \sum_{i=1}^{d} (v_i^T x)^2 (214)
\]

\[
= \frac{1}{\lambda_{\min}(\Sigma)} \|x\|_2^2. (215)
\]

Lemma 26 (Linear Bandit Exploration Bonus). For an spd matrix \( \Sigma \), the equality below holds whenever the operations make sense:

\[
\max_{\phi, \|\eta\|_{\Sigma^{-1}} \leq \sqrt{\sigma}} \phi^T \eta = \sqrt{\sigma} \|\phi\|_{\Sigma^{-1}} (216)
\]

Proof. Choose \( \eta = \Sigma^{-1} \phi \frac{\sqrt{\sigma}}{\|\phi\|_{\Sigma^{-1}}} \), which satisfies the constraint

\[
||\Sigma^{-1} \phi \frac{\sqrt{\sigma}}{\|\phi\|_{\Sigma^{-1}}} \|_{\Sigma} = ||\phi \frac{\sqrt{\sigma}}{\|\phi\|_{\Sigma^{-1}}} \|_{\Sigma^{-1}} \sqrt{\sigma} = \sqrt{\sigma} (217)
\]

and gives an objective value

\[
\max_{\phi, \|\eta\|_{\Sigma^{-1}} \leq \sqrt{\sigma}} \phi^T \eta \geq \phi^T \Sigma^{-1} \phi \frac{\sqrt{\sigma}}{\|\phi\|_{\Sigma^{-1}}} = \sqrt{\sigma} ||\phi||_{\Sigma^{-1}} (218)
\]

On the other hand, Cauchy-Schwartz ensures:

\[
\max_{\phi, \|\eta\|_{\Sigma^{-1}} \leq \sqrt{\sigma}} \phi^T \eta \leq ||\phi||_{\Sigma^{-1}} ||\eta||_{\Sigma} = \sqrt{\sigma} ||\phi||_{\Sigma^{-1}} (219)
\]