A Analysis of Existing Algorithms

Let $f^{*,*}$ denote a function that incorporates an attacker strategy. When k=0, $f^{\text{CH, IS}}(D, w_y, g_y, k)$ is the result of applying the CH inequality to the IS weighted returns, obtained from D, which additionally includes k copies of a trajectory with an IS weight of w_y and return of g_y . Notice that $f^{*,*}$ is written in terms of IS weights. The following defines $f^{\text{CH, WIS}}$, written in terms of IS weights, when k=0:

$$f^{\text{CH, WIS}}(D, w_y, g_y, 0) = \frac{1}{\sum_{i=1}^n w_i} \sum_{i=1}^n w_i g_i - b \sqrt{\frac{\ln(1/\delta)}{2n}}.$$

For the rest of the paper, we use the following notation. Let $\mathcal{I}=\{I:\exists a\in\mathcal{A},\exists s\in\mathcal{S},I=\prod_{t=0}^{\tau-1}\pi_e(A_t=a,S_t=s)/\pi_b(A_t=a,S_t=s)\}$, i.e., the set of all IS weights that could be obtained from policies π_e and π_b . The maximum and minimum IS weight is denoted by $i^*=\max(\mathcal{I})$ and $i^{\min}=\min(\mathcal{I})$, respectively. For shorthand, let the sum of IS weights in D be written as $\beta=\sum_{i=1}^n w_i$. Also, we assume that $\beta>0$ to ensure that WIS is well-defined.

Next, we define a new term to describe how an attacker can increase the $1-\delta$ confidence lower bound on the mean of a bounded and real-valued random variable. We say that $f^{*,*}$ is adversarially monotonic given its inputs, if an attacker can maximize $f^{*,*}$ by maximizing the value of the added samples. For brevity, we say that $f^{*,*}$ is adversarially monotonic.

Definition 1. $f^{*,*}$ is adversarially monotonic for n > 1, k > 0, π_b , π_e and D if both

- 1. There exists two constants $p \ge 0$ and $q \in [0,1]$, with $pq \in [0,i^*]$, such that $f^{*,*}(D,p,q,k) \ge f^{*,*}(D,p,q,0)$, i.e., adding k copies of pq does not decrease f;
- 2. $\frac{\partial}{\partial g_y} f^{*,*}(D, i^*, g_y, k) \geq 0$ and $\frac{\partial}{\partial w_y} f^{*,*}(D, w_y, 1, k) \geq 0$, with no local maximums, i.e., f is a non-decreasing function w.r.t. the IS weight and return added by the attacker, respectively.

Definition 1 means that $f^{*,*}$ is maximized when w_y and g_y is maximized. In other words, the optimal strategy is to add k copies of the trajectory with the maximum IS weight and return. Notice that $f^{*,*}$ does not incorporate all possible attack functions, \mathcal{M} : specifically, the set of attacks, where the attacker can choose to add k different trajectories, is omitted. As described in Theorem 1, to perform a worst-case analysis, only the optimal attack must be incorporated as part of $f^{*,*}$.

In the following two lemmas, we show that a couple well-known Seldonian algorithms are adversarially monotonic.

Lemma 1. *Under Assumptions 1, 2 and 3, f^{CH, IS} is adversarially monotonic.*

Proof. Let $w_y \geq \frac{1}{n} \sum_{i=1}^n w_i g_i + \frac{(n+k)}{k} \left(b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - b \sqrt{\frac{\ln(1/\delta)}{2n}} \right)$ and $g_y = 1$. To show that $w_y g_y \in [0,i^*]$ as stated in (1) in Definition 1, it must be that $w_y \in [0,i^*]$. For all $i \in \{1,\ldots,n\}$, $w_i g_i \in [0,i^*]$. Thus, for any given dataset, $0 \leq 1/n \sum_{i=1}^n w_i g_i \leq i^*/n$. Using this fact, for any given D, the range of w_y is

$$\frac{1}{n} \sum_{i=1}^{n} (0) + \frac{(n+k)}{k} \left(b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - b \sqrt{\frac{\ln(1/\delta)}{2n}} \right) \le w_y \le \frac{1}{n} \sum_{i=1}^{n} (i^*) + \frac{(n+k)}{k} \left(b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - b \sqrt{\frac{\ln(1/\delta)}{2n}} \right) \\
\frac{b(n+k)}{k} \left(\sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - \sqrt{\frac{\ln(1/\delta)}{2n}} \right) \le w_y \le \frac{i^*}{n} + \underbrace{\frac{b(n+k)}{k} \left(\sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - \sqrt{\frac{\ln(1/\delta)}{2n}} \right)}_{0} \le i^*.$$

Therefore, w_y can be selected such that $w_y g_y \in [0, i^*]$. It follows that

$$\begin{split} f^{\text{CH, IS}}(D, w_y, 1, k) = & \frac{1}{n+k} \sum_{i=1}^n w_i g_i + \frac{k}{n+k} (w_y)(1) - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} \\ \geq & \frac{1}{n+k} \sum_{i=1}^n w_i g_i + \frac{k}{n+k} \left(\frac{1}{n} \sum_{i=1}^n w_i g_i + \frac{(n+k)}{k} \left(b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - b \sqrt{\frac{\ln(1/\delta)}{2n}} \right) \right) - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} \\ = & \frac{1}{n} \sum_{i=1}^n w_i g_i - b \sqrt{\frac{\ln(1/\delta)}{2n}} \\ = & f^{\text{CH, IS}}(D, w_y, g_y, 0). \end{split}$$

Next, we show that (2) in Definition 1 holds.

$$\begin{split} \frac{\partial}{\partial w_y} f^{*,\star}(D,w_y,g_y,k) &= \frac{\partial}{\partial w_y} \bigg(\sum_{i=1}^n \frac{w_i g_i}{n+k} \bigg) + \frac{k w_y g_y}{n+k} - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} \\ &= \frac{k g_y}{n+k} \\ \frac{\partial}{\partial w_y} f^{*,\star}(D,w_y,1,k) &= \frac{k}{n+k}. \\ \frac{\partial}{\partial g_y} f^{*,\star}(D,w_y,g_y,k) &= \frac{\partial}{\partial g_y} \bigg(\sum_{i=1}^n \frac{w_i g_i}{n+k} \bigg) + \frac{k w_y g_y}{n+k} - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} \\ &= \frac{k w_y}{n+k} \\ \frac{\partial}{\partial g_y} f^{*,\star}(D,i^*,g_y,k) &= \frac{k i^*}{n+k}. \end{split}$$

Notice that both partial derivatives are non-negative when $g_y=1$ and $w_y=i^*$, respectively. To find any critical points, the following equations are solved simultaneously: $\partial/\partial g_y f^{\text{CH, WIS}}(D,w_y,g_y,k)=0$ and $\partial/\partial w_y f^{\text{CH, WIS}}(D,w_y,g_y,k)=0$. Notice that points along the line $(w_g,0)$ and $(0,g_y)$ are all critical points. The following partial derivatives are computed to classify these points:

$$\begin{split} &\frac{\partial}{\partial (w_y)^2}(D,w_y,g_y,k)=&0.\\ &\frac{\partial}{\partial (g_y)^2}(D,w_y,g_y,k)=&0.\\ &\frac{\partial}{\partial g_yw_y}(D,w_y,g_y,k)=&\frac{k}{n+k}. \end{split}$$

Using the second partial derivative test, the critical points are substituted into the following equation:

$$\frac{\partial}{\partial (w_y)^2} \cdot \frac{\partial}{\partial (g_y)^2} - \left(\frac{\partial}{\partial g_y w_y}\right)^2 = -\left(\frac{k}{n+k}\right)^2,$$

which is less than zero. Therefore, points along the line $(w_g,0)$ and $(0,g_y)$ are saddle points. \Box

Lemma 2. Under Assumptions 1 and 2, $f^{CH, WIS}$ is adversarially monotonic.

Proof. First, we show that (1) in Definition 1 holds with $g_y = 1$ and $w_y = 0$.

$$f^{\text{CH, WIS}}(D, w_y, g_y, k) = \frac{1}{kw_y + \beta} \left(kw_y g_y + \sum_{i=1}^n w_i g_i \right) - b\sqrt{\frac{\ln(1/\delta)}{2(n+k)}}$$

$$f^{\text{CH, WIS}}(D, 0, 1, k) = \frac{1}{\beta} \sum_{i=1}^n w_i g_i - b\sqrt{\frac{\ln(1/\delta)}{2(n+k)}}$$

$$> \frac{1}{\beta} \sum_{i=1}^n w_i g_i - b\sqrt{\frac{\ln(1/\delta)}{2n}}$$

$$= f^{\text{CH, WIS}}(D, w_y, g_y, 0),$$
(1)

where (1) follows from $b\sqrt{\frac{\ln(1/\delta)}{2n}} > b\sqrt{\frac{\ln(1/\delta)}{2(n+k)}}$. Second, we show that (2) in Definition 1 holds.

$$\begin{split} \frac{\partial}{\partial w_{y}} f^{\text{CH, WIS}}(V, w_{y}, g_{y}, k) &= -\frac{k \sum_{i=1}^{n} w_{i} g_{i}}{(k w_{y} + \beta)^{2}} - \frac{k^{2} w_{y} g_{y}}{(k w_{y} + \beta)^{2}} + \frac{k g_{y}}{(k w_{y} + \beta)} \\ &= -\frac{k \sum_{i=1}^{n} w_{i} g_{i}}{(k w_{y} + \beta)^{2}} - \frac{k^{2} w_{y} g_{y}}{(k w_{y} + \beta)^{2}} + \frac{k g_{y} (k w_{y} + \beta)}{(k w_{y} + \beta)^{2}} \\ &= -\frac{k \sum_{i=1}^{n} w_{i} g_{i}}{(k w_{y} + \beta)^{2}} + \frac{k g_{y} \beta}{(k w_{y} + \beta)^{2}} \\ &= -\frac{k \sum_{i=1}^{n} w_{i} g_{i}}{(k w_{y} + \beta)^{2}} + \frac{k \sum_{i=1}^{n} w_{i} g_{y}}{(k w_{y} + \beta)^{2}} \\ &= \frac{k}{(\beta + k w_{y})^{2}} \sum_{i=1}^{n} w_{i} (g_{y} - g_{i}) \\ &\frac{\partial}{\partial w_{y}} f^{\text{CH, WIS}}(V, w_{y}, 1, k) = \frac{k}{(\beta + k w_{y})^{2}} \sum_{i=1}^{n} w_{i} (1 - g_{i}). \end{split} \tag{2}$$

Notice that (2) is non-negative: 1) When $g_y=1$, (2) is positive as long as there exists at least one $g_i<1$ for $i\in\{1,\ldots,n\}$; 2) If all $g_i=1$ in D, then (2) is zero. The following is the derivative of $f^{\text{CH,WIS}}(D,w_y,g_y,k)$ w.r.t. g_y :

$$\begin{split} &\frac{\partial}{\partial g_{y}}f^{\text{CH, WIS}}(D, w_{y}, g_{y}, k) = \frac{kw_{y}}{(\beta + kw_{y})} \\ &\frac{\partial}{\partial g_{y}}f^{\text{CH, WIS}}(D, i^{*}, g_{y}, k) = \frac{ki^{*}}{(\beta + ki^{*})}, \end{split} \tag{3}$$

which is also non-negative. To find any critical points, the following equations are solved simultaneously: $\partial/\partial g_y f^{\text{CH, WIS}}(D, w_y, g_y, k) = 0$ and $\partial/\partial w_y f^{\text{CH, WIS}}(D, w_y, g_y, k) = 0$. Notice that (3) is zero when $w_y = 0$. Plugging $w_y = 0$ into $\partial/\partial w_y f^{\text{CH, WIS}}(D, w_y, g_y, k) = 0$, and then solving for g_y , yields the x coordinate of a critical point.

$$\frac{k}{(\beta + k(0))^2} \sum_{i=1}^{n} w_i (g_y - g_i) = 0$$

$$\frac{k}{\beta^2} \sum_{i=1}^{n} w_i (g_y - g_i) = 0$$

$$g_y \sum_{i=1}^{n} w_i - \sum_{i=1}^{n} w_i g_i = 0$$

$$g_y = \frac{\sum_{i=1}^{n} w_i g_i}{\beta}.$$

The following partial derivatives are computed to classify whether $(0, \sum_{i=1}^{n} w_i g_i / \beta)$ is a minimum, maximum or saddle point:

$$\frac{\partial}{\partial (w_y)^2}(D, w_y, g_y, k) = \frac{-2k^2}{(\beta + kw_y)^3} \sum_{i=1}^n w_i (g_y - g_i).$$

$$\frac{\partial}{\partial (g_y)^2}(D, w_y, g_y, k) = 0.$$

$$\frac{\partial}{\partial g_y w_y}(D, w_y, g_y, k) = \frac{\partial}{\partial w_y} \frac{kw_y}{(\beta + kw_y)}$$

$$= \frac{k\beta}{(\beta + kw_y)^2}.$$

Using the second partial derivative test, the critical point is substituted into the following equation:

$$\frac{\partial}{\partial (w_y)^2} \cdot \frac{\partial}{\partial (g_y)^2} - \left(\frac{\partial}{\partial g_y w_y}\right)^2 = 0 - \left(\frac{k\beta}{(\beta + k(0))^2}\right)^2$$
$$= -\left(\frac{k\beta}{\beta^2}\right)^2,$$

which is less that zero. Therefore, $(w_y = 0, g_y = \sum_{i=1}^n w_i g_i/\beta)$ is a saddle point.

Next, we describe the trajectory that must be added to D to execute the optimal attack.

Definition 2 (Optimal Attack). An optimal attack strategy for k > 0 is to select

$$\underset{H \in \mathcal{H}_{\pi_e}}{\arg \max} f^{*,*}(D, w_y = w(H, \pi_e, \pi_b), g_y = g(H), k).$$

Definition 3 (Optimal Trajectory). Given that a maximum exists, let $(a', s') \in \underset{a \in \mathcal{A}, s \in \mathcal{S}}{\arg \max} \frac{\pi_e(a, s)}{\pi_b(a, s)}$.

If
$$\frac{\pi_e(a,s)}{\pi_b(a,s)} > 1$$
, let $H^* = \{S_0 = s', A_0 = a', R_0 = 1, \dots, S_{\tau-1} = s', A_{\tau-1} = a', R_{\tau-1} = 1\}$. Otherwise, let $H^* = \{S_0 = s', A_0 = a', R_0 = 1\}$.

Theorem 1. For any adversarially monotonic off-policy estimator, the optimal attack strategy is to add k repetitions of H^* to D.

Proof. An optimal attack strategy is equivalent to

$$\arg \max_{H \in \mathcal{H}_{\pi_e}} f^{*,*} \big(D, w(H, \pi_e, \pi_b), g(H), k \big) = \arg \max_{i^* \in \mathcal{I}, g^* \in [0, 1]} f^{*,*} \big(D, i^*, g^*, k \big).$$

For any off-policy estimator that is adversarially monotonic, by (1) of Definition 1, there exists a pq such that

$$f^{*,*}(D, p, q, k) > f^{*,*}(D, p, q, 0).$$

A return that maximizes $f^{*,\star}(D,w_y,g_y,k)$ implies that

$$\max_{g^* \in [0,1]} f^{*,*}(D, p, g^*, k) \ge f^{*,*}(D, p, q, k).$$

 $f^{\rm CH,\, IS}$ and $f^{\rm CH,\, WIS}$ are non-decreasing w.r.t. the return. Therefore,

$$\arg \max_{g^* \in [0,1]} f^{*,*}(D, p, g^*, k) = \max_{g^* \in [0,1]} g^*.$$

Setting $g^* = 1$, an importance weight that maximizes $f^{*,*}(D, w_y, 1, k)$ implies that

$$\max_{i^* \in \mathcal{I}} f^{*,\star}(D, i^*, 1, k) \ge f^{*,\star}(D, p, 1, k).$$

 $f^{\mathrm{CH,\,IS}}$ and $f^{\mathrm{CH,\,WIS}}$ are also non-decreasing w.r.t. the importance weight. So,

$$\operatorname*{arg\,max}_{i^* \in \mathcal{I}} f^{*,\star}(D,i^*,1,k) = \operatorname*{max}_{i^* \in \mathcal{I}} i^*.$$

Since the IS weight is a product of ratios over the length of a trajectory, the ratio at a single time step is maximized.

$$\max_{i^* \in \mathcal{I}} i^* = \max_{a \in \mathcal{A}, s \in \mathcal{S}} \prod_{t=0}^{\tau-1} \frac{\pi_e(A_t = a, S_t = s)}{\pi_b(A_t = a, S_t = s)}$$

$$= \begin{cases} \left(\max_{a \in \mathcal{A}, s \in \mathcal{S}} \frac{\pi_e(a, s)}{\pi_b(a, s)} \right)^{\tau} & \text{if } \max_{a \in \mathcal{A}, s \in \mathcal{S}} \frac{\pi_e(a, s)}{\pi_b(a, s)} > 1, \\ \max_{a \in \mathcal{A}, s \in \mathcal{S}} \frac{\pi_e(a, s)}{\pi_b(a, s)} & \text{otherwise.} \end{cases}$$

To create H^* , if the ratio at a single time step is greater than 1, a' and s' is repeated for the maximum length of the trajectory, τ ; otherwise, a' and s' is repeated only for a single time step. Thus, H^* represents the trajectory with the largest return and importance weight.

Next, we show how Equations (2) and (1), that define quasi- α -security and α -security, respectively, apply to $L^{*,*}$. Specifically, we show that a safety test using $L^{*,*}$ as a metric is a valid safety test that first predicts the performance of π_e using D, and then bounds the predicted performance with high probability. If $L^{*,*}(\pi_e,D)>J(\pi_b)$, the safety test returns True; otherwise it returns False.

Lemma 3. A safety test using
$$L^{*,*}$$
 is quasi- α -secure if $\forall m \in \mathcal{M}$, $\Pr\left(L^{*,*}(\pi_e, m(D, k)) > J(\pi_b) + \alpha\right) \leq \Pr\left(L^{*,*}(\pi_e, D) > J(\pi_b)\right)$.

Proof. For $x \in \mathbb{N}^+$, let $\mathcal{P}: \Pi \times D_n^{\pi_b} \to \mathbb{R}^x$ denote any function to predict the performance of some $\pi_e \in \Pi$, using data D collected from π_b . Also, let $\mathcal{B}: \mathbb{R}^x \times [0,1] \to \mathbb{R}$ denote any function that bounds performance with high probability, $1-\delta$, where $\delta \in [0,1]$. Starting with the definition of quasi- α -security, we have that $\forall m \in \mathcal{M}$,

$$\begin{split} \Pr\Big(\varphi\big(\pi_e, m(D, k), J(\pi_b) + \alpha\big) &= \mathsf{True}\Big) \leq \Pr\Big(\varphi\big(\pi_e, D, J(\pi_b)\big) = \mathsf{True}\Big) \\ &\iff \Pr\Big(\mathcal{B}\big(\mathcal{P}(\pi_e, m(D, k)), \delta\big) > J(\pi_b) + \alpha\Big) \leq \Pr\Big(\mathcal{B}\big(\mathcal{P}(\pi_e, D), \delta\big) > J(\pi_b)\Big) \\ &\iff \Pr\Big(L^{\star, *}\big(\pi_e, m(D, k)\big) > J(\pi_b) + \alpha\Big) \leq \Pr\Big(L^{\star, *}\big(\pi_e, D\big) > J(\pi_b)\Big). \end{split}$$

Lemma 4. A safety test using $L^{*,*}$ is α -secure if $\forall m \in \mathcal{M}$, $\Pr\left(L^{*,*}(\pi_e, m(D, k)) > J(\pi_b) + \alpha\right) < \delta$.

 \Box

Proof. For $x \in \mathbb{N}^+$, let $\mathcal{P}: \Pi \times D_n^{\pi_b} \to \mathbb{R}^x$ denote any function to predict the performance of some $\pi_e \in \Pi$, using data D collected from π_b . Also, let $\mathcal{B}: \mathbb{R}^x \times [0,1] \to \mathbb{R}$ denote any function that bounds performance with high probability, $1-\delta$, where $\delta \in [0,1]$. Starting with the definition of α -security, we have that $\forall m \in \mathcal{M}$,

$$\Pr\left(\varphi\left(\pi_{e}, m(D, k), J(\pi_{b}) + \alpha\right) = \text{True}\right) < \delta$$

$$\iff \Pr\left(\mathcal{B}\left(\mathcal{P}(\pi_{e}, m(D, k)), \delta\right) > J(\pi_{b}) + \alpha\right) < \delta$$

$$\iff \Pr\left(L^{\star, *}\left(\pi_{e}, m(D, k)\right) > J(\pi_{b}) + \alpha\right) < \delta.$$

In Lemma 5, we describe a condition that must hold in order to compute a valid α . The condition states that a valid α must be equal to or greater than the largest increase in the $1-\delta$ confidence lower bound on $J(\pi_e)$ across all datasets $D \in \mathcal{D}_n^{\pi_b}$ and all attack strategies (i.e., the optimal attack).

Lemma 5. A safety test using $L^{*,*}$ is quasi- α -secure or α -secure if $\forall D \in \mathcal{D}_n^{\pi_b}$ and $\forall m \in \mathcal{M}$, $L^{*,*}(\pi_e, m(D, k)) \leq L^{*,*}(\pi_e, D) + \alpha$.

Proof. If $L^{*,*}(\pi_e, m(D, k)) \leq L^{*,*}(\pi_e, D) + \alpha$, then

$$L^{*,\star}(\pi_e, D) \ge L^{*,\star}(\pi_e, m(D, k)) - \alpha. \tag{4}$$

A safety test checks whether $L^{*,\star}(\pi_e,D) > J(\pi_b)$. When (4) holds $\forall D \in \mathcal{D}_n^{\pi_b}$ and $\forall m \in \mathcal{M}$,

$$\Pr(L^{*,\star}(\pi_e, D) > J(\pi_b)) \ge \Pr(L^{*,\star}(\pi_e, m(D, k)) - \alpha > J(\pi_b)), \tag{5}$$

and hence via algebra that

$$\Pr(L^{*,*}(\pi_e, m(D, k)) > J(\pi_b) + \alpha) \le \Pr(L^{*,*}(\pi_e, D) > J(\pi_b)),$$

which, by Lemma (3), implies that a safety test using $L^{*,*}$ is quasi- α -secure. In the case of α -security, by Assumption 3, we require a "safe" safety test. That is,

$$\Pr(L^{*,*}(\pi_e, D) > J(\pi_b)) < \delta. \tag{6}$$

From the transitive property of \geq , we can conclude from (5) and (6) that

$$\Pr(L^{*,\star}(\pi_e, m(D, k)) - \alpha > J(\pi_b)) < \delta,$$

and hence via algebra that

$$\Pr(L^{*,\star}(\pi_e, m(D, k))) > J(\pi_b) + \alpha) < \delta,$$

which, by Lemma (4), implies that a safety test using $L^{*,*}$ is α -secure.

B Proof of Theorem 1

The result of (5) for the estimator that uses CH and IS is the following:

$$\begin{split} \alpha' &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} f^{\text{CH, IS}}(D, i^*, 1, k) - L^{\text{CH, IS}}(\pi_e, D) \\ &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} \frac{1}{n+k} \sum_{i=1}^n w_i g_i + \frac{k}{n+k} (i^*)(1) - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - \left(\frac{1}{n} \sum_{i=1}^n w_i g_i - b \sqrt{\frac{\ln(1/\delta)}{2n}}\right) \\ &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} b \sqrt{\frac{\ln(1/\delta)}{2n}} - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)} \left(i^* - \frac{\sum_{i=1}^n w_i g_i}{n}\right). \end{split}$$

Recall that b represents the upper bound of all IS weighted returns. Let $b = i^*$, and $g_i = 0$ for all $i \in \{1, ..., n\}$.

$$\alpha' = i^* \sqrt{\frac{\ln(1/\delta)}{2n}} - i^* \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)} (i^* - 0)$$
$$= i^* \left(\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)} \right).$$

The result of (5) for the estimator that uses CH and WIS is the following:

$$\alpha' = \max_{D \in \mathcal{D}_n^{\mathcal{H}}} f^{\mathsf{CH, WIS}}(D, i^*, 1, k) - L^{\mathsf{CH, WIS}}(\pi_e, D)$$

$$\begin{split} &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} \frac{1}{ki^* + \sum_{i=1}^n w_i} \bigg(\sum_{i=1}^n w_i g_i + k(i^*)(1) \bigg) - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - \bigg(\frac{1}{\sum_{i=1}^n w_i} \sum_{i=1}^n w_i g_i - b \sqrt{\frac{\ln(1/\delta)}{2n}} \bigg) \\ &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} b \sqrt{\frac{\ln(1/\delta)}{2n}} - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{ki^*}{(ki^* + \beta)} \bigg(1 - \frac{\sum_{i=1}^n w_i g_i}{\beta} \bigg). \end{split}$$

Let $g_i = 0$ for all $i \in \{1, ..., n\}$. Also, notice that b = 1 because importance weighted returns are in range [0, 1] for WIS.

$$\alpha' = \max_{D \in \mathcal{D}_n^{\mathcal{H}}} \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{ki^*}{(ki^* + \beta)}.$$

Recall that $\beta \neq 0$. So, let $w_i = 0$ for all $i \in \{1, \dots, n-1\}$ and $w_n = i^{\min}$.

$$\alpha' = \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{ki^*}{(i^{\min} + ki^*)}.$$

C Panacea: An Algorithm for Safe and Secure Policy Improvement

Table 1: α -security of Panacea.

Estimator	α
CH, IS	$c\left(\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)}\right)$
CH, WIS	$\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{kc}{(i^{\min}+kc)}$

Algorithm 1 Panacea (D, π_e, α, k)

- 1: Compute c, using α and k, given estimator
- 2: for $H \in D$ do
- 3: **if** IS weight computed using H is greater than c **then**
- 4: Set IS weight to c
- 5: return clipped D

C.1 Proof of Corollary 1

Let α' and k' denote the user-specified inputs to Panacea. Based on Table 1, $c^{\text{CH, IS}} = \alpha'/\left(\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)}\right)$ if k' = k. Recall that b is the upper bound on all IS weighted returns. Due to clipping, $b = c^{\text{CH, IS}}$, and let $g_i = 0$ for all $i \in \{1, \dots, n\}$. The result of (6) for the estimator that uses CH and IS is the following:

$$\begin{split} \max_{D \in \mathcal{D}_n^{\mathcal{H}}} f^{\text{CH, IS}} &(\text{Panacea}(D, c^{\text{CH, IS}}), c^{\text{CH, IS}}, 1, k) - L^{\text{CH, IS}} \big(\pi_e, \text{Panacea}(D, c^{\text{CH, IS}}) \big) \\ &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} \frac{1}{n+k} \sum_{i=1}^n w_i g_i + \frac{k}{n+k} (c^{\text{CH, IS}}) (1) - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - \left(\frac{1}{n} \sum_{i=1}^n w_i g_i - b \sqrt{\frac{\ln(1/\delta)}{2n}} \right) \\ &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} b \sqrt{\frac{\ln(1/\delta)}{2n}} - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)} \Big(c^{\text{CH, IS}} - \frac{\sum_{i=1}^n w_i g_i}{n} \Big) \\ &= c^{\text{CH, IS}} \sqrt{\frac{\ln(1/\delta)}{2n}} - c^{\text{CH, IS}} \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)} \Big(c^{\text{CH, IS}} - 0 \Big) \\ &= c^{\text{CH, IS}} \left(\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)} \right) \\ &= \frac{\alpha'}{\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)}} \cdot \left(\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)} \right) \\ &= \alpha'. \end{split}$$

For WIS, recall that no matter how the clipping weight is set, $b \leq 1$ because importance weighted returns are in range [0,1], and $\beta \neq 0$. So, let $w_i = 0$ for all $i \in \{1,\ldots,n-1\}$ and $w_n = i^{\min}$. Also, let $g_i = 0$ for all $i \in \{1,\ldots,n\}$. Based on Table 1, $c^{\text{CH,WIS}} = i^{\min} \left(\alpha' - \sqrt{\frac{\ln(1/\delta)}{2n}} + \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}\right)/k \left(1 - \alpha' + \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}\right)$ if k' = k. The result of (6) for the estimator that uses

CH and WIS is the following:

$$\begin{split} &\max_{D \in \mathcal{D}_n^{\mathcal{H}}} f^{\text{CH, WIS}}(\text{Panacea}(D, c^{\text{CH, WIS}}), c^{\text{CH, WIS}}, 1, k) - L^{\text{CH, WIS}}(\pi_e, \text{Panacea}(D, c^{\text{CH, WIS}})) \\ &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} \frac{1}{kc^{\text{CH, WIS}} + \sum_{i=1}^n w_i} \left(\sum_{i=1}^n w_i g_i + k(c^{\text{CH, WIS}})(1) \right) - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} - \left(\frac{1}{\sum_{i=1}^n w_i} \sum_{i=1}^n w_i g_i - b \sqrt{\frac{\ln(1/\delta)}{2n}} \right) \\ &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} b \sqrt{\frac{\ln(1/\delta)}{2n}} - b \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{kc^{\text{CH, WIS}}}{(kc^{\text{CH, WIS}} + \beta)} \left(1 - \frac{\sum_{i=1}^n w_i g_i}{\beta} \right) \\ &\leq \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{kc^{\text{CH, WIS}}}{(kc^{\text{CH, WIS}} + i^{\min})} \\ &= \left(\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} \right) + \frac{\left(\alpha' - \sqrt{\frac{\ln(1/\delta)}{2n}} + \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}\right)}{\left(\alpha' - \sqrt{\frac{\ln(1/\delta)}{2n}} + \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}\right) + \left(1 - \alpha' + \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}\right)} \\ &= \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \alpha' - \sqrt{\frac{\ln(1/\delta)}{2n}} + \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} \\ &= \alpha'. \end{split}$$

C.2 Proof of Corollary 2

Let α and k' denote the user-specified inputs to Panacea. If k' = k, i.e., the user inputs the correct number of trajectories added by the attacker, the result of (6) for the estimator that uses CH and IS is the following:

$$\begin{split} \alpha &= \max_{D \in \mathcal{D}_n^{\mathcal{H}}} f^{\text{CH, IS}}(\text{Panacea}(D, c), c, 1, k) - L^{\text{CH, IS}} \left(\pi_e, \text{Panacea}(D, c)\right) \\ \alpha &= c \left(\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)}\right) \\ c &= \frac{\alpha}{\left(\sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{k}{(n+k)}\right)}. \end{split}$$

If k' = k, the result of (6) for the estimator that uses CH and WIS is the following:

$$\max_{D \in \mathcal{D}_n^{\mathcal{H}}} f^{\mathrm{CH,\,WIS}}(\mathrm{Panacea}(D,c),c,1,k) - L^{\mathrm{CH,\,WIS}}\big(\pi_e,\mathrm{Panacea}(D,c)\big) \leq \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{kc}{(kc+i^{\min})}. \tag{7}$$

Setting the right-hand side of (7) to α , and solving for c equals:

$$\alpha = \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}} + \frac{kc}{(i^{\min} + kc)}$$

$$\frac{kc}{(i^{\min} + kc)} = \alpha - \sqrt{\frac{\ln(1/\delta)}{2n}} + \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}$$

$$kc - kc\alpha + kc\sqrt{\frac{\ln(1/\delta)}{2n}} - kc\sqrt{\frac{\ln(1/\delta)}{2(n+k)}} = i^{\min}\alpha - i^{\min}\sqrt{\frac{\ln(1/\delta)}{2n}} + i^{\min}\sqrt{\frac{\ln(1/\delta)}{2(n+k)}}$$

$$kc\left(1 - \alpha + \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}\right) = i^{\min}\alpha - i^{\min}\sqrt{\frac{\ln(1/\delta)}{2n}} + i^{\min}\sqrt{\frac{\ln(1/\delta)}{2(n+k)}}$$

$$c = \frac{i^{\min}\left(\alpha - \sqrt{\frac{\ln(1/\delta)}{2n}} + \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}\right)}{k\left(1 - \alpha + \sqrt{\frac{\ln(1/\delta)}{2n}} - \sqrt{\frac{\ln(1/\delta)}{2(n+k)}}\right)}.$$