
Supplement to “Cone-Constrained Principal Component Analysis”

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The supplement is organized as follows. In Section 1 we give an overview of the setting considered in the main article, the notation used and cover some useful general facts. In Sections 2, 3 and 4 we provide the proofs of Theorems 1 2 and 3 of the main article.

1 Preliminaries

Recall the setting in Problem (3) in the main article, we let $\mathbf{X} \in \mathbb{R}^{n \times n}$ be a symmetric matrix and $\mathcal{C}_n \subseteq \mathbb{R}^n$ be a closed cone (not necessarily convex). We consider the following optimization problem:

$$\begin{aligned} & \text{maximize } \langle \mathbf{v}, \mathbf{X}\mathbf{v} \rangle, \\ & \text{subject to } \mathbf{v} \in \mathcal{C}_n, \quad \|\mathbf{v}\|_2 = 1. \end{aligned} \tag{1}$$

We denote by $\lambda_{\max}(\mathbf{X}; \mathcal{C}_n)$ the value of this problem. For instance $\lambda_{\max}(\mathbf{X}; \mathbb{R}^n)$ is the largest eigenvalue of \mathbf{X} . Note that this optimization problem is –in general– NP-hard. We denote by \mathcal{C}_n^* the dual cone of \mathcal{C}_n :

$$\mathcal{C}_n^* = \{ \mathbf{x} : \forall \mathbf{z} \in \mathcal{C}_n, \langle \mathbf{x}, \mathbf{z} \rangle \geq 0 \},$$

and by \mathcal{C}_n° its polar cone $\mathcal{C}_n^\circ = -\mathcal{C}_n^*$.

Further, we let $\mathcal{C}_{n,\mathbf{x}} = \text{cone}\{\mathbf{y} - \mathbf{x} : \mathbf{y} \in \mathcal{C}_n\}$ be the tangent cone of \mathcal{C}_n at $\mathbf{x} \in \mathcal{C}_n$ (in particular $\mathcal{C}_{n,0} = \mathcal{C}_n$), and $\mathcal{C}_{n,\mathbf{x}}^*$ its dual (i.e. $\mathcal{C}_{n,\mathbf{x}}^* = (\mathcal{C}_{n,\mathbf{x}})^*$). Finally, we let $\mathbb{S}^{n-1} = \{\mathbf{x} : \|\mathbf{x}\|_2 = 1\}$ be the unit sphere in n dimensions.

Given a closed convex set $K \subseteq \mathbb{R}^n$, we let $P_K : \mathbb{R}^n \rightarrow \mathbb{R}^n$ denote the orthogonal projection onto K . Given vectors $\mathbf{v}_1, \mathbf{v}_2, \dots$, $\text{span}(\mathbf{v}_1, \mathbf{v}_2, \dots)$ denotes their linear span. For a linear subspace V , V^\perp denotes its orthogonal complement.

The following definitions will be useful in the sequel.

Definition 1.1. A local maximum of the problem (1) is a point $\mathbf{v}^* \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$, $\|\mathbf{v}^*\|_2 = 1$ for which there exists an $\varepsilon > 0$ such that $\langle \mathbf{v}, \mathbf{X}\mathbf{v} \rangle \leq \langle \mathbf{v}^*, \mathbf{X}\mathbf{v}^* \rangle$ for all $\mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$ within a neighborhood of \mathbf{v}^* : $\|\mathbf{v} - \mathbf{v}^*\|_2 \leq \delta$.

Definition 1.2. Let \mathcal{C}_n be a closed convex cone. For any $\mathbf{x} \in \mathbb{R}^n$, the Moreau decomposition of \mathbf{x} is defined as $\mathbf{x} = P_{\mathcal{C}_n}(\mathbf{x}) + P_{\mathcal{C}_n^\circ}(\mathbf{x})$. Further we have that $\langle P_{\mathcal{C}_n}(\mathbf{x}), P_{\mathcal{C}_n^\circ}(\mathbf{x}) \rangle = 0$.

1.1 General facts

Let \mathbf{v}^* denote a local maximizer of (1). We begin with the following remark that characterizes the tangent cone at a local maximizer.

Remark 1.3. We have that $\langle \mathbf{v}^*, \mathbf{v} \rangle = 0$ for all $\mathbf{v} \in \mathcal{C}_{n,\mathbf{v}^*}^*$.

Proof. If $\mathbf{v}^* = 0$, the result is trivial. If $\mathbf{v}^* \neq 0$, we have that $\mathbf{v}^* - \mathbf{v}^* = 0 \in \mathcal{C}_n$ and $\mathbf{v}^* - (-\mathbf{v}^*)2\mathbf{v}^* \in \mathcal{C}_n$. Hence $\{\mathbf{v}^*, -\mathbf{v}^*\} \subseteq \mathcal{C}_{n,\mathbf{v}^*}$ and as a consequence for $\mathbf{v} \in \mathcal{C}_{n,\mathbf{v}^*}^*$, $\langle \mathbf{v}^*, \mathbf{v} \rangle \geq 0$ and $\langle -\mathbf{v}^*, \mathbf{v} \rangle \geq 0$ imply $\langle \mathbf{v}^*, \mathbf{v} \rangle = 0$. \square

The following proposition characterizes the value at the local maxima of Problem (1).

Proposition 1.4. *If the point $\mathbf{v}^* \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$ is a local maximum of the problem (1), then there exists $\boldsymbol{\mu}^* \in \mathcal{C}_{n,\mathbf{v}^*}^*$ and $\lambda^* \in \mathbb{R}$ such that*

$$\mathbf{X}\mathbf{v}^* = \lambda^* \mathbf{v}^* - \boldsymbol{\mu}^*, \quad (2)$$

$$\lambda^* \geq \sup \left\{ \langle \mathbf{v}, \mathbf{X}\mathbf{v} \rangle : \mathbf{v} \in \mathcal{C}_{n,\mathbf{v}^*} \cap \text{span}(\boldsymbol{\mu}^*)^\perp, \|\mathbf{v}\|_2 = 1 \right\}. \quad (3)$$

Vice-versa, if the above conditions (2), (3) hold for some $\mathbf{v}^ \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$ with the last inequality being strict, then \mathbf{v}^* is a local maximum. Further, if \mathcal{C}_n is convex Eq. (2) is the Moreau decomposition of $\mathbf{X}\mathbf{v}^*$ with respect to \mathcal{C}_n . In particular*

$$\lambda^* \mathbf{v}^* = \text{P}_{\mathcal{C}_{n,\mathbf{v}^*}}(\mathbf{X}\mathbf{v}^*) = \text{P}_{\mathcal{C}_n}(\mathbf{X}\mathbf{v}^*), \quad \boldsymbol{\mu}^* = -\text{P}_{\mathcal{C}_{n,\mathbf{v}^*}^\circ}(\mathbf{X}\mathbf{v}^*) = -\text{P}_{\mathcal{C}_n^\circ}(\mathbf{X}\mathbf{v}^*). \quad (4)$$

Proof. Note that, for any $\mathbf{v} \in \text{relint}(\mathcal{C}_{n,\mathbf{v}^*})$ and $\varepsilon \geq 0$ we know that $\varepsilon \mathbf{v} \in \mathcal{C}_{n,\mathbf{v}^*}$. Therefore, by definition of the tangent cone, $\mathbf{v}(\varepsilon) \equiv \mathbf{v}^* + \varepsilon \mathbf{v} \in \mathcal{C}_n$. Then by local optimality of \mathbf{v}^* , for all ε small enough we have

$$\langle \mathbf{v}(\varepsilon), \mathbf{X}\mathbf{v}(\varepsilon) \rangle \leq \langle \mathbf{v}^*, \mathbf{X}\mathbf{v}^* \rangle \|\mathbf{v}(\varepsilon)\|^2.$$

Expanding both sides and letting $\lambda^* = \langle \mathbf{v}^*, \mathbf{X}\mathbf{v}^* \rangle$, we get

$$\begin{aligned} 2\varepsilon \langle \mathbf{v}, \mathbf{X}\mathbf{v}^* \rangle + \varepsilon^2 \langle \mathbf{v}, \mathbf{X}\mathbf{v} \rangle &\leq 2\lambda^* \varepsilon \langle \mathbf{v}^*, \mathbf{v} \rangle + \varepsilon^2 \lambda^* \|\mathbf{v}\|^2 \\ \text{whence } \langle \mathbf{v}, \mathbf{X}(\mathbf{v}^* - \lambda^* \mathbf{v}^*) \rangle &\leq \frac{\varepsilon}{2} (\|\mathbf{v}\|^2 - \langle \mathbf{v}, \mathbf{X}\mathbf{v} \rangle). \end{aligned} \quad (5)$$

With $\boldsymbol{\mu}^* \equiv -\mathbf{X}\mathbf{v}^* + \lambda^* \mathbf{v}^*$, and taking $\varepsilon \rightarrow 0$ we obtain:

$$\langle \mathbf{v}, \boldsymbol{\mu}^* \rangle \geq 0 \quad \text{for all } \mathbf{v} \in \text{relint}(\mathcal{C}_{n,\mathbf{v}^*})$$

Since \mathcal{C}_n is closed, so is $\mathcal{C}_{n,\mathbf{v}^*}$ and hence the above holds for every $\mathbf{v} \in \mathcal{C}_{n,\mathbf{v}^*}$. This implies that $\boldsymbol{\mu}^* \in \mathcal{C}_{n,\mathbf{v}^*}^*$. To prove Eq. (3) use $\mathbf{v} \in \mathcal{C}_{n,\mathbf{v}^*} \cap \text{span}(\boldsymbol{\mu}^*)^\perp$ in Eq. (5).

Before proving the converse, we consider first the case when \mathcal{C}_n is convex. By convexity of \mathcal{C}_n , we have that $\mathcal{C}_n \subseteq \mathcal{C}_{n,\mathbf{v}^*}$. Further since $\mathbf{v}^* \in \mathcal{C}_n$, we have that $\text{P}_{\mathcal{C}_{n,\mathbf{v}^*}}(\mathbf{X}\mathbf{v}^*) = \text{P}_{\mathcal{C}_n}(\mathbf{X}\mathbf{v}^*)$. Also, Remark 1.3 implies $\langle \boldsymbol{\mu}^*, \mathbf{v}^* \rangle = 0$. Together, these imply the Moreau decomposition claim.

In order to prove the converse, let $\mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$ and note that –as a consequence $\mathbf{w} \equiv \mathbf{v} - \mathbf{v}^* \in \mathcal{C}_{n,\mathbf{v}^*}$ with $2\langle \mathbf{w}, \mathbf{v}^* \rangle = -\|\mathbf{w}\|_2^2$. We then have

$$\begin{aligned} \langle \mathbf{v}, \mathbf{X}\mathbf{v} \rangle - \langle \mathbf{v}^*, \mathbf{X}\mathbf{v}^* \rangle &= 2\langle \mathbf{w}, \mathbf{X}\mathbf{v}^* \rangle + \langle \mathbf{w}, \mathbf{X}\mathbf{w} \rangle \\ &= 2\langle \mathbf{w}, \lambda^* \mathbf{v}^* - \boldsymbol{\mu}^* \rangle + \langle \mathbf{w}, \mathbf{X}\mathbf{w} \rangle \\ &= -2\langle \boldsymbol{\mu}^*, \mathbf{w} \rangle + \langle \mathbf{w}, (\mathbf{A} - \lambda^* \mathbf{I}) \mathbf{w} \rangle. \end{aligned} \quad (6)$$

Letting $\mathbf{M} = \mathbf{X} - \lambda^* \mathbf{I}$, and $V \equiv \text{span}(\boldsymbol{\mu}^*)^\perp$, we know by assumption that $\langle \mathbf{w}, \mathbf{M}\mathbf{w} \rangle \leq -2\Delta \|\mathbf{w}\|_2^2$ for all $\mathbf{w} \in \mathcal{D} \equiv \mathcal{C}_{n,\mathbf{v}^*} \cap V$, for some $\Delta > 0$. For a general $\mathbf{w} \in \mathcal{C}_{n,\mathbf{v}^*}$, let $\mathbf{w}_0 \equiv \text{P}_{\mathcal{D}}(\mathbf{w})$ and $\tilde{\mathbf{v}} \equiv \mathbf{w} - \mathbf{w}_0$. We then have

$$\begin{aligned} \langle \mathbf{w}, \mathbf{M}\mathbf{w} \rangle &= \langle \mathbf{w}_0, \mathbf{M}\mathbf{w}_0 \rangle + 2\langle \tilde{\mathbf{v}}, \mathbf{M}\mathbf{w}_0 \rangle + \langle \tilde{\mathbf{v}}, \mathbf{M}\tilde{\mathbf{v}} \rangle \\ &\leq -2\Delta \|\mathbf{w}_0\|_2^2 + 2\|\mathbf{M}\|_2 \|\tilde{\mathbf{v}}\|_2 \|\mathbf{w}_0\|_2 + \|\mathbf{M}\|_2 \|\tilde{\mathbf{v}}\|_2^2 \\ &\leq -2\Delta \|\mathbf{w}_0\|_2^2 + \gamma \|\mathbf{M}\|_2 \|\mathbf{w}_0\|_2^2 + (1 + \gamma^{-1}) \|\mathbf{M}\|_2 \|\tilde{\mathbf{v}}\|_2^2, \end{aligned}$$

where the last inequality follows for any $\gamma > 0$ by $2ab \leq \gamma a^2 + b^2/\gamma$. Setting $\gamma = \Delta/\|\mathbf{M}\|_2$ and $c = \|\mathbf{M}\|_2(1 + \gamma^{-1})$, we get

$$\langle \mathbf{w}, \mathbf{M}\mathbf{w} \rangle \leq -\Delta \|\mathbf{w}_0\|_2^2 + c \|\tilde{\mathbf{v}}\|_2^2. \quad (7)$$

Now we claim that, for any $\delta > 0$ there exists $L = L(\mathcal{C}_{n,\mathbf{v}^*}, \boldsymbol{\mu}^*, \delta) \geq 0$ such that, for all $\mathbf{w} \in \mathcal{C}_{n,\mathbf{v}^*}$,

$$\|\tilde{\mathbf{v}}\|_2^2 \leq \delta \|\mathbf{w}_0\|_2^2 + L \langle \boldsymbol{\mu}^*, \mathbf{w} \rangle^2. \quad (8)$$

Before proving this claim, let us show that it indeed implies the desired thesis. Setting $\delta = \Delta/c$, and $\tilde{L} = cL(\mathcal{C}_{n,\mathbf{v}^*}, \boldsymbol{\mu}^*, \Delta/c)$, we conclude from Eq. (7) that

$$\langle \mathbf{w}, \mathbf{M}\mathbf{w} \rangle \leq \tilde{L} \langle \boldsymbol{\mu}^*, \mathbf{w} \rangle^2.$$

Substituting this estimate in Eq. (6), we get

$$\langle \mathbf{v}, \mathbf{X} \mathbf{v} \rangle - \langle \mathbf{v}^*, \mathbf{X} \mathbf{v}^* \rangle \leq -2\langle \boldsymbol{\mu}^*, \mathbf{w} \rangle + \tilde{L}\langle \boldsymbol{\mu}^*, \mathbf{w} \rangle^2. \quad (9)$$

Hence, for all \mathbf{v} such that $\|\mathbf{v} - \mathbf{v}^*\|_2 = \|\mathbf{w}\|_2 \leq 1/(\|\boldsymbol{\mu}^*\|_2 \tilde{L})$ we have $0 \leq \langle \boldsymbol{\mu}^*, \mathbf{w} \rangle \leq 1/\tilde{L}$ and therefore

$$\langle \mathbf{v}, \mathbf{X} \mathbf{v} \rangle - \langle \mathbf{v}^*, \mathbf{X} \mathbf{v}^* \rangle \leq -\langle \boldsymbol{\mu}^*, \mathbf{w} \rangle \leq 0,$$

which completes our proof that \mathbf{v}^* is a local maximum.

We are left with the task of proving the claim (8). Notice that, by scaling both sides, it is sufficient to prove it under the additional assumption $\|\mathbf{w}\|_2 = 1$. Fix $\delta > 0$ and assume by contradiction that the claim is false. Then, for each $k \in \mathbb{N}$ there exists $\mathbf{w}^{(k)} \in \mathcal{C}_{n, \mathbf{v}^*}$, $\|\mathbf{w}^{(k)}\|_2 = 1$ such that, letting $\mathbf{w}_0^{(k)} \equiv \mathcal{P}_{\mathcal{D}}(\mathbf{w}^{(k)})$ and $\tilde{\mathbf{v}}^{(k)} \equiv \mathbf{w}^{(k)} - \mathbf{w}_0^{(k)}$ we get

$$\|\tilde{\mathbf{v}}^{(k)}\|_2^2 > \delta \|\mathbf{w}_0^{(k)}\|_2^2 + k\langle \boldsymbol{\mu}^*, \mathbf{w}^{(k)} \rangle^2. \quad (10)$$

Since $\{\mathbf{w}^{(k)}\}$ is a subset of the compact set $\mathcal{C}_{n, \mathbf{v}^*} \cap \mathbb{S}^{n-1}$, we can assume (by passing to a convergent subsequence), that $\mathbf{w}^{(k)} \rightarrow \mathbf{w}^{(\infty)}$. Since $\mathcal{C}_{n, \mathbf{v}^*} \cap \mathbb{S}^{n-1}$ is closed, $\mathbf{w}^{(\infty)} \in \mathcal{C}_{n, \mathbf{v}^*} \cap \mathbb{S}^{n-1}$. Further $\langle \boldsymbol{\mu}^*, \mathbf{w}^{(\infty)} \rangle^2 \leq \lim_{k \rightarrow \infty} \|\tilde{\mathbf{v}}^{(k)}\|_2^2/k \leq \lim_{k \rightarrow \infty} 4/k = 0$. Hence $\mathbf{w}^{(\infty)} \in \mathcal{D}$ i.e. $\mathbf{w}_0^{(\infty)} = \mathbf{w}^{(\infty)}$, $\tilde{\mathbf{v}}^{(\infty)} = 0$. Further taking the limit $k \rightarrow \infty$ in (10) we get

$$\|\tilde{\mathbf{v}}^{(\infty)}\|_2^2 \geq \delta \|\mathbf{w}_0^{(\infty)}\|_2^2$$

i.e. $0 \geq \delta$. Since we chose $\delta > 0$ this gives the desired contradiction. \square

The above proof implies a quantitative bound on the radius of the neighborhood within which \mathbf{v}^* is an optimum.

Corollary 1.5. *Assume the conditions*

$$\begin{aligned} \mathbf{X} \mathbf{v}^* &= \lambda^* \mathbf{v}^* - \boldsymbol{\mu}^*, \\ \lambda^* - 2\Delta &\geq \sup \left\{ \langle \mathbf{v}, \mathbf{X} \mathbf{v} \rangle : \mathbf{v} \in \mathcal{C}_{n, \mathbf{v}^*} \cap \text{span}(\boldsymbol{\mu}^*)^\perp, \|\mathbf{v}\|_2 = 1 \right\}, \end{aligned}$$

hold, and further assume that $L(\mathcal{C}_{n, \mathbf{v}^*}, \boldsymbol{\mu}^*, \delta) \geq 0$ is such that, for all $\mathbf{w} \in \mathcal{C}_{n, \mathbf{v}^*}$,

$$\|\tilde{\mathbf{v}}\|_2^2 \leq \delta \|\mathbf{w}_0\|_2^2 + L\langle \boldsymbol{\mu}^*, \mathbf{w} \rangle^2.$$

Let $\mathbf{M} = \mathbf{X} - \lambda^* \mathbf{I}$, $c = \Delta \|\mathbf{M}\|_2 (\Delta + \|\mathbf{M}\|_2)$, and $\tilde{L} = cL(\mathcal{C}_{n, \mathbf{v}^*}, \boldsymbol{\mu}^*, \Delta/c)$. Then \mathbf{v}^* is a global maximum of the optimization problem (1) within a neighborhood $\text{Ball}(\mathbf{v}^*, 1/(\tilde{L}\|\boldsymbol{\mu}^*\|_2))$.

2 Proof of Theorem 1

Recall the model assumptions for the data \mathbf{X} :

$$\mathbf{X} = \beta \mathbf{v}_0 \mathbf{v}_0^\top + \mathbf{Z}.$$

Here we assume that $\mathbf{v}_0 \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$ and $\mathbf{Z}_{ij} \sim \mathcal{N}(0, 1/n)$ are independent, up to symmetry. We first define the following useful quantities:

Definition 2.1. *For a cone $\mathcal{C}_n \subseteq \mathbb{R}^n$ we define its normalized Gaussian width as:*

$$\omega(\mathcal{C}_n) \equiv \frac{1}{\sqrt{n}} \mathbb{E} \left\{ \sup_{\mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1}} \langle \mathbf{g}, \mathbf{v} \rangle \right\},$$

where $\mathbf{g} \sim \mathcal{N}(0, \mathbf{I}_n)$.

Definition 2.2. *For a cone $\mathcal{C}_n \subseteq \mathbb{R}^n$, we define its packing number $N(\mathcal{C}_n, \varepsilon)$ as the size of the maximal subset X of $\mathcal{C}_n \cap \mathbb{S}^{n-1}$ such that for every $\mathbf{v}_1, \mathbf{v}_2 \in X$, $\|\mathbf{v}_1 - \mathbf{v}_2\| > \varepsilon$.*

We have the following useful facts:

Lemma 2.3. *There exist universal constants c_1, c_2 such that:*

$$c_1 \sup_{\varepsilon} \varepsilon \sqrt{\log N(\mathcal{C}_n, \varepsilon)} \leq \sqrt{n} \omega(\mathcal{C}_n) \leq c_2 \inf_{\varepsilon \leq 1} \{2\sqrt{n}\varepsilon(1 + \sqrt{\log(2/\varepsilon)}) + (2 - \varepsilon)\sqrt{\log N(\mathcal{C}_n, \varepsilon)}\}.$$

Proof. The left hand inequality is the Sudakov minoration inequality. For the latter, we first employ Dudley inequality:

$$\sqrt{n} \omega(\mathcal{C}_n) \leq c_3 \int_0^\infty \sqrt{\log N(\mathcal{C}_n, \varepsilon)} d\varepsilon,$$

for a universal constant c_3 . We know that the diameter of \mathbb{S}^{n-1} is 2. Further as $\mathcal{C}_n \cap \mathbb{S}^{n-1} \subseteq \mathbb{S}^{n-1}$ using a standard volume packing argument for \mathbb{S}^{n-1} [] we have that:

$$N(\mathcal{C}_n, \varepsilon) \leq \left(1 + \frac{2}{\varepsilon}\right)^n.$$

Thus for any $0 \leq \varepsilon \leq 1$:

$$\sqrt{n} \omega(\mathcal{C}_n) \leq c_3 \left[\int_0^\varepsilon \sqrt{n \log \left(1 + \frac{2}{u}\right)} du + \int_\varepsilon^2 \sqrt{N(\mathcal{C}_n, \varepsilon)} du \right].$$

We simplify the first integral:

$$\begin{aligned} \int_0^\varepsilon \sqrt{\log \left(1 + \frac{2}{u}\right)} du &\leq 2 \int_0^\varepsilon \sqrt{\log \left(\frac{2}{u}\right)} du \\ &= 2 \int_{\sqrt{\log(2/\varepsilon)}}^\infty (4y^2) \exp(-y^2) dy \\ &\leq 2\varepsilon \left(1 + \sqrt{\log \left(\frac{2}{\varepsilon}\right)}\right), \end{aligned}$$

using standard integration by parts. Since $N(\mathcal{C}, \varepsilon)$ is monotone nonincreasing in ε , we have that the second integral is bound above by $(2 - \varepsilon)N(\mathcal{C}_n, \varepsilon)$. These estimates imply the thesis using the observation that ε is arbitrarily chosen. □

The following lemma is proved in [ALMT13].

Lemma 2.4. *For any closed convex cone:*

$$\omega(\mathcal{C}_n)^2 \leq \delta(\mathcal{C}_n) \leq \omega(\mathcal{C}_n)^2 + \frac{1}{n}.$$

We can now prove Theorem 1

Proof of Theorem 1. Let X denote a maximal ε -net of $\mathcal{C}_n \cap \mathbb{S}^{n-1}$ as in Lemma 2.3, for ε to be fixed later in the proof. Hence, for any $\mathbf{v}_i, \mathbf{v}_j \in X$ distinct, we have $\|\mathbf{v}_i - \mathbf{v}_j\| > \varepsilon$ or, equivalently, $\langle \mathbf{v}_i, \mathbf{v}_j \rangle < 1 - \varepsilon^2/2$. Note that the maximality of X implies also that it is an ε -cover of $\mathcal{C}_n \cap \mathbb{S}^{n-1}$, i.e. for any $\mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$ $\min_{\mathbf{v}' \in X} \|\mathbf{v} - \mathbf{v}'\| \leq \varepsilon$.

Let \mathbf{v}_0 be uniformly distributed in the set X . For an estimator $\hat{\mathbf{v}}(\mathbf{X}) \in \mathcal{V}$, we define $G(\hat{\mathbf{v}}(Y)) = \arg \min_{\mathbf{v} \in X} \|\hat{\mathbf{v}}(\mathbf{X}) - \mathbf{v}\|_2$. We now bound the probability of the error event $\{G(\hat{\mathbf{v}}(\mathbf{X})) \neq \mathbf{v}_0\}$. By definition of $G(\hat{\mathbf{v}}(\mathbf{X}))$, the event $G(\hat{\mathbf{v}}(\mathbf{X})) \neq \mathbf{v}_0$ implies that there exists $\mathbf{v}_i \in X, \mathbf{v}_i \neq \mathbf{v}_0$ such that $\|\hat{\mathbf{v}}(\mathbf{X}) - \mathbf{v}_i\|_2 \leq \|\hat{\mathbf{v}}(\mathbf{X}) - \mathbf{v}_0\|_2$. This along with triangle inequality implies that $\varepsilon < \|\mathbf{v}_i - \mathbf{v}_0\| \leq 2\|\hat{\mathbf{v}}(\mathbf{X}) - \mathbf{v}_0\|$, i.e. $\|\hat{\mathbf{v}}(\mathbf{X}) - \mathbf{v}_0\|_2 > \varepsilon/2$. By Markov inequality we have:

$$\begin{aligned} \mathbb{P}\{G(\hat{\mathbf{v}}(\mathbf{X})) \neq \mathbf{v}_0\} &\leq \mathbb{P}\{\|\hat{\mathbf{v}}(\mathbf{X}) - \mathbf{v}_0\|_2 > \varepsilon/2\} \\ &\leq 4 \frac{\mathbb{E}\{\|\hat{\mathbf{v}}(\mathbf{X}) - \mathbf{v}_0\|_2^2\}}{\varepsilon^2} \\ &= \frac{8(1 - \mathbb{E}\{\langle \hat{\mathbf{v}}(\mathbf{X}), \mathbf{v}_0 \rangle\})}{\varepsilon^2} \\ &\leq \frac{8R(\hat{\mathbf{v}}(\mathbf{X}); \mathbf{v}_0)}{\varepsilon^2}, \end{aligned} \tag{11}$$

By Fano's inequality we have that:

$$\mathbb{P}\{G(\widehat{\mathbf{v}}(\mathbf{X})) \neq \mathbf{v}_0\} \geq 1 - \frac{\gamma + \log 2}{\log |X|},$$

where $\gamma = \max_{x \neq x'} D(P_{\mathbf{v}} \| P_{\mathbf{v}'})$ where $D(\cdot \| \cdot)$ is the Kullback-Liebler divergence and $P_{\mathbf{v}}$ denotes the law of \mathbf{X} conditional on $\mathbf{v}_0 = \mathbf{v}$. Conditional on $\mathbf{v}_0 = \mathbf{v}$, \mathbf{X} has mean $\mathbf{v}\mathbf{v}^\top$ and has Gaussian entries with variance $1/n$. A standard calculation implies that:

$$\begin{aligned} D(P_{\mathbf{v}} \| P_{\mathbf{v}'}) &\leq n\beta^2 \|\mathbf{v}\mathbf{v}^\top - \mathbf{v}'\mathbf{v}'^\top\|_F^2 \\ &= 2n\beta^2(1 - \langle \mathbf{v}, \mathbf{v}' \rangle^2) \\ &\leq 2n\beta^2. \end{aligned}$$

We have using Lemma 2.3 that $\log |X| \geq cn(\omega(\mathcal{C}_n) - 2\varepsilon(1 + \sqrt{\log(2/\varepsilon)}))^2$. Combining this with Fano's inequality above and Eq.(11):

$$R(\widehat{\mathbf{v}}(\mathbf{X}); \mathbf{v}_0) \geq \frac{\varepsilon^2}{8} \left(1 - \frac{2\beta^2}{c(\omega(\mathcal{C}_n) - 2\varepsilon(1 + \sqrt{\log(2/\varepsilon)}))^2} \right).$$

Further, the minimax risk satisfies $R(\mathcal{C}_n) \geq R(\widehat{\mathbf{v}}(\mathbf{X}); \mathbf{v}_0)$ for some estimator $\widehat{\mathbf{v}}(\mathbf{X})$ and the above bound is uniformly true for all estimators. Using $\varepsilon_* = 1/4\omega(\mathcal{C}_n)/\sqrt{\log(1/\omega(\mathcal{C}_n))}$ implies that $2\varepsilon(1 + \sqrt{\log(2/\varepsilon)}) \leq \omega(\mathcal{C}_n)/2$. Hence:

$$R(\mathcal{C}_n) \geq c' \frac{\omega(\mathcal{C}_n)^2}{\log(1/\omega(\mathcal{C}_n))},$$

when $\beta \leq c''\omega(\mathcal{C}_n)$ for some universal constants c', c'' . Applying Lemma 2.4 and the constraint $\omega(\mathcal{C}_n) \geq \sqrt{2/n}$ implies the result for appropriately adjusted c', c'' . □

3 Proof of Theorem 2

We first prove the following useful lemma based on standard Gaussian comparisons.

Lemma 3.1. *Assume that the noise \mathbf{Z} has i.i.d. $N(0, 1/n)$ entries up to symmetry. Then, with probability at least $1 - n^{-5}$*

$$\lambda_1(\mathbf{Z}; \mathcal{C}_n) \leq 2\sqrt{\delta(\mathcal{C}_n)} + \sqrt{\frac{10 \log n}{n}}. \quad (12)$$

Proof. We will apply the Sudakov-Fernique inequality (see e.g. [Vit00, Theorem 1]) to the two processes $\{\mathbf{M}(\mathbf{v})\}, \{\mathbf{V}(\mathbf{v})\}$ indexed by $x \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$ defined as follows:

$$\mathbf{M}(\mathbf{v}) \equiv \langle \mathbf{v}, \mathbf{Z}\mathbf{v} \rangle \quad \text{and} \quad \mathbf{V}(\mathbf{v}) \equiv 2\langle \mathbf{v}, \mathbf{g} \rangle,$$

for a random vector $\mathbf{g} \sim N(0, \mathbf{I}_n/n)$ and \mathbf{Z} being a standard normal matrix. Basic algebra gives for all \mathbf{v} , $\mathbb{E}\mathbf{M}(\mathbf{v}) = \mathbb{E}\mathbf{V}(\mathbf{v}) = 0$ and

$$\mathbb{E}\{[\mathbf{M}(\mathbf{v}_1) - \mathbf{M}(\mathbf{v}_2)]^2\} = \frac{4}{n}(1 - \langle \mathbf{v}_1, \mathbf{v}_2 \rangle^2), \quad \mathbb{E}\{[\mathbf{V}(\mathbf{v}_1) - \mathbf{V}(\mathbf{v}_2)]^2\} = \frac{8}{n}(1 - \langle \mathbf{v}_1, \mathbf{v}_2 \rangle).$$

Hence, by using the fact that for $a \in [-1, 1]$, the inequality $1 - a^2 \leq 2(1 - a)$ holds, we get $\mathbb{E}\{[\mathbf{M}(\mathbf{v}_1) - \mathbf{M}(\mathbf{v}_2)]^2\} \leq \mathbb{E}\{[\mathbf{V}(\mathbf{v}_1) - \mathbf{V}(\mathbf{v}_2)]^2\}$. We conclude using Sudakov-Fernique comparison lemma that:

$$\mathbb{E} \max \{ \langle \mathbf{v}, \mathbf{Z}\mathbf{v} \rangle : \mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1} \} \leq 2 \mathbb{E} \max \{ \langle \mathbf{v}, \mathbf{g} \rangle : \mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1} \}$$

By using Proposition 10.1 from [ALMT13], $\mathbb{E} \max \{ \langle \mathbf{v}, \mathbf{g} \rangle : \mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1} \} \leq \sqrt{\delta(\mathcal{C}_n)}$, we get that

$$\mathbb{E}\lambda_1(\mathbf{Z}; \mathcal{C}_n) \leq 2\sqrt{\delta(\mathcal{C}_n)}. \quad (13)$$

By using the fact that $\mathbf{Z} \mapsto \max \{ \langle \mathbf{v}, \mathbf{Z}\mathbf{v} \rangle : \mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1} \}$ is 1-Lipschitz and concentration inequality on $\max \{ \langle \mathbf{v}, \mathbf{Z}\mathbf{v} \rangle : \mathbf{v} \in \mathcal{C}_n \cap \mathbb{S}^{n-1} \}$, with probability at least $1 - \exp\{-t^2 n/2\}$, we have

$$\lambda_1(\mathbf{Z}; \mathcal{C}_n) \leq 2\sqrt{\delta(\mathcal{C}_n)} + t .$$

Take $t = \sqrt{(10 \log n)/n}$ and the claim follows. \square

We can now prove Theorem 2

Proof. By optimality of $\hat{\mathbf{v}}^{\text{ML}}$, we have

$$\begin{aligned} \langle \hat{\mathbf{v}}^{\text{ML}}, \mathbf{X}\hat{\mathbf{v}}^{\text{ML}} \rangle &\geq \langle \mathbf{v}_0, \mathbf{X}\mathbf{v}_0 \rangle = \beta + \langle \mathbf{v}_0, \mathbf{Z}\mathbf{v}_0 \rangle \\ &\geq \beta - \lambda_1(\mathbf{Z}, \mathcal{C}_n). \end{aligned}$$

On the other hand

$$\langle \hat{\mathbf{v}}^{\text{ML}}, \mathbf{X}\hat{\mathbf{v}}^{\text{ML}} \rangle = \beta \langle \mathbf{v}_0, \hat{\mathbf{v}}^{\text{ML}} \rangle^2 + \langle \hat{\mathbf{v}}^{\text{ML}}, \mathbf{Z}\hat{\mathbf{v}}^{\text{ML}} \rangle \leq \beta \langle \mathbf{v}_0, \hat{\mathbf{v}}^{\text{ML}} \rangle^2 + \lambda_{\max}(\mathbf{Z}; \mathcal{C}_n).$$

The claim that $\mathcal{R}_{\mathcal{C}_n}(\hat{\mathbf{v}}^{\text{ML}}; \mathcal{C}_n) \leq 4(\sqrt{\delta(\mathcal{C}_n)} + \varepsilon_n)/\beta$ follows simply by putting together the above inequalities along with the previous lemma for $\lambda_{\max}(\mathbf{Z}; \mathcal{C}_n)$.

For the other claim we have as above:

$$\begin{aligned} \langle \hat{\mathbf{v}}^{\text{ML}}, \mathbf{X}\hat{\mathbf{v}}^{\text{ML}} \rangle &\geq \beta + \langle \mathbf{v}_0, \mathbf{Z}\mathbf{v}_0 \rangle, \\ \langle \hat{\mathbf{v}}^{\text{ML}}, \mathbf{X}\hat{\mathbf{v}}^{\text{ML}} \rangle &= \beta \langle \mathbf{v}_0, \hat{\mathbf{v}}^{\text{ML}} \rangle^2 + \langle \hat{\mathbf{v}}^{\text{ML}}, \mathbf{Z}\hat{\mathbf{v}}^{\text{ML}} \rangle. \end{aligned}$$

Together, this implies:

$$\begin{aligned} \beta(1 - \langle \mathbf{v}_0, \hat{\mathbf{v}}^{\text{ML}} \rangle^2) &\leq \langle \mathbf{Z}, -\mathbf{v}_0\mathbf{v}_0^\top + \hat{\mathbf{v}}^{\text{ML}}(\hat{\mathbf{v}}^{\text{ML}})^\top \rangle \\ &\leq \|\mathbf{Z}\|_2 \|\mathbf{v}_0\mathbf{v}_0^\top - \hat{\mathbf{v}}^{\text{ML}}(\hat{\mathbf{v}}^{\text{ML}})^\top\|_*, \end{aligned}$$

where the last line follows from Holder's inequality and $\|\cdot\|_*$ denotes the nuclear norm (or sum of singular values). Since $\mathbf{v}_0\mathbf{v}_0^\top - \hat{\mathbf{v}}^{\text{ML}}(\hat{\mathbf{v}}^{\text{ML}})^\top$ has rank at most two:

$$\begin{aligned} \|\mathbf{v}_0\mathbf{v}_0^\top - \hat{\mathbf{v}}^{\text{ML}}(\hat{\mathbf{v}}^{\text{ML}})^\top\|_* &\leq \sqrt{2} \|\mathbf{v}_0\mathbf{v}_0^\top - \hat{\mathbf{v}}^{\text{ML}}(\hat{\mathbf{v}}^{\text{ML}})^\top\|_F \\ &= 2\sqrt{1 - \langle \mathbf{v}_0, \hat{\mathbf{v}}^{\text{ML}} \rangle^2}. \end{aligned}$$

Thus we have:

$$1 - \langle \mathbf{v}_0, \hat{\mathbf{v}}^{\text{ML}} \rangle^2 \leq \frac{2\|\mathbf{Z}\|_2^2}{\beta^2}.$$

Using the fact that $1 - a^2 \geq 1 - |a|$ when $a \in [-1, 1]$, we obtain the desired result for the risk $\mathcal{R}(\hat{\mathbf{v}}^{\text{ML}}; \mathcal{C}_n)$. \square

4 Proof of Theorem 3

Note that the problem (1) is unchanged (except for an additive constant in the objective function) if we replace \mathbf{X} with $\mathbf{X}_\rho = \mathbf{X} + \rho \mathbf{I}$. We will take advantage of this freedom and consider the modified iteration

$$\hat{\mathbf{v}}^{t+1} = \frac{\mathcal{P}_{\mathcal{C}}(\mathbf{u}^t)}{\|\mathcal{P}_{\mathcal{C}}(\mathbf{u}^t)\|_2}, \quad (14)$$

$$\mathbf{u}^t = \mathbf{X}_\rho \hat{\mathbf{v}}^t. \quad (15)$$

It is convenient to define

$$\boldsymbol{\mu}^{t+1} \equiv -\mathcal{P}_{\mathcal{C}_n^\circ}(\mathbf{X}_\rho \hat{\mathbf{v}}^t), \quad (16)$$

$$\lambda_{t+1} \equiv \|\mathcal{P}_{\mathcal{C}_n}(\mathbf{X}_\rho \hat{\mathbf{v}}^t)\|_2. \quad (17)$$

Then we have the identity

$$\mathbf{X}_\rho \widehat{\mathbf{v}}^{t-1} = \lambda_t \widehat{\mathbf{v}}^t - \boldsymbol{\mu}^t, \quad (18)$$

which is the Moreau's decomposition of $\mathbf{X}_\rho \widehat{\mathbf{v}}^{t-1}$. In particular

$$\langle \widehat{\mathbf{v}}^t, \boldsymbol{\mu}^t \rangle = 0. \quad (19)$$

Lemma 4.1. *With the above definitions we have*

$$\lambda_t = \lambda_{t+1} \langle \widehat{\mathbf{v}}^{t+1}, \widehat{\mathbf{v}}^{t-1} \rangle - \langle \boldsymbol{\mu}^{t+1}, \widehat{\mathbf{v}}^{t-1} \rangle \quad (20)$$

$$\leq \lambda_{t+1} \langle \widehat{\mathbf{v}}^{t+1}, \widehat{\mathbf{v}}^{t-1} \rangle. \quad (21)$$

Proof. Taking the scalar product of both sides of Eq. (18) by $\widehat{\mathbf{v}}^t$ and using Eq. (19), together with the fact that $\|\widehat{\mathbf{v}}^t\|_2 = 1$ by construction, we get

$$\lambda_t = \langle \widehat{\mathbf{v}}^t, \mathbf{X} \widehat{\mathbf{v}}^{t-1} \rangle = \langle \widehat{\mathbf{v}}^{t-1}, \mathbf{X} \widehat{\mathbf{v}}^t \rangle \quad (22)$$

$$= \langle \widehat{\mathbf{v}}^{t-1}, \lambda_{t+1} \widehat{\mathbf{v}}^{t+1} - \boldsymbol{\mu}^{t+1} \rangle. \quad (23)$$

This proves Eq. (20). Equation (21) follows from $\widehat{\mathbf{v}}^{t-1} \in \mathcal{C}_n$, $\boldsymbol{\mu}^{t+1} \in \mathcal{C}_n^*$, which imply $\langle \widehat{\mathbf{v}}^{t-1}, \boldsymbol{\mu}^{t+1} \rangle \geq 0$. \square

Lemma 4.2. *For any $t \geq 1$, we have*

$$\rho + \lambda_{\min}(\mathbf{X}) \leq \lambda_t \leq \rho + \lambda_{\max}(\mathbf{X}).$$

Proof. For $\mathbf{y} \in \mathcal{C}_n^o$, $\max \{ \langle \mathbf{y}, \mathbf{z} \rangle : \mathbf{z} \in \mathcal{C}_n \cap \mathbb{S}^{n-1} \} = 0$ and $\text{P}_{\mathcal{C}_n}(\mathbf{y}) = 0$. For any $\mathbf{y} \notin \mathcal{C}_n^o$, $\text{P}_{\mathcal{C}_n}(\mathbf{y}) \neq 0$, so by using Cauchy-Schwarz inequality, and the fact that $\text{P}_{\mathcal{C}_n}(\mathbf{y}) / \|\text{P}_{\mathcal{C}_n}(\mathbf{y})\|_2 \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$,

$$\|\text{P}_{\mathcal{C}_n}(\mathbf{y})\|_2 = \max \{ \langle \mathbf{y}, \mathbf{z} \rangle : \mathbf{z} \in \mathcal{C}_n \cap \mathbb{S}^{n-1} \}.$$

Then, since $\widehat{\mathbf{v}}^{t-1} \in \mathcal{C}_n \cap \mathbb{S}^{n-1}$, we have

$$\lambda_t = \|\text{P}_{\mathcal{C}_n}(\mathbf{X}_\rho \widehat{\mathbf{v}}^{t-1})\|_2 \geq \langle \widehat{\mathbf{v}}^{t-1}, \mathbf{X}_\rho \widehat{\mathbf{v}}^{t-1} \rangle,$$

which yields the desired lower bound by definition. The upper bound follows since $\lambda_t \leq \|\mathbf{X}_\rho \widehat{\mathbf{v}}^{t-1}\|_2$. \square

Proposition 4.3. *Assume $\rho + \lambda_{\min}(\mathbf{X}) > 0$. Then the sequence $\{\lambda_t\}_{t \geq 0}$ is bounded and non-decreasing and therefore it has a limit*

$$\lambda_* = \lim_{t \rightarrow \infty} \lambda_t = \lim_{t \rightarrow \infty} \langle \widehat{\mathbf{v}}^t, \mathbf{X}_\rho \widehat{\mathbf{v}}^t \rangle. \quad (24)$$

Further, if $\widehat{\mathbf{v}}_$ is any sub-sequential limit of $\{\widehat{\mathbf{v}}^t\}_{t \geq 0}$ (i.e. $\lim_{k \rightarrow \infty} \widehat{\mathbf{v}}^{t(k)} = \widehat{\mathbf{v}}_*$ for some sequence $\{t(k)\}$) then it satisfies the stationarity condition*

$$\mathbf{X} \widehat{\mathbf{v}}_* = (\lambda_* - \rho) \widehat{\mathbf{v}}_* - \boldsymbol{\mu}_*. \quad (25)$$

Proof. The existence of the limit λ_* follows immediately from Lemma 4.1 and 4.2. Next multiplying the identity (18) by $\widehat{\mathbf{v}}^t$, we get, using Cauchy-Schwartz,

$$\lambda_t = \langle \widehat{\mathbf{v}}^t, \mathbf{X}_\rho \widehat{\mathbf{v}}^{t-1} \rangle \leq \sqrt{\langle \widehat{\mathbf{v}}^t, \mathbf{X}_\rho \widehat{\mathbf{v}}^t \rangle \langle \widehat{\mathbf{v}}^{t-1}, \mathbf{X}_\rho \widehat{\mathbf{v}}^{t-1} \rangle}.$$

Multiplying Eq. (18) by $\widehat{\mathbf{v}}^{t-1}$, we get $\langle \widehat{\mathbf{v}}^{t-1}, \mathbf{X} \widehat{\mathbf{v}}^{t-1} \rangle \leq \lambda_t \langle \widehat{\mathbf{v}}^t, \widehat{\mathbf{v}}^{t-1} \rangle$ and, changing the iteration number, $\langle \widehat{\mathbf{v}}^t, \mathbf{X} \widehat{\mathbf{v}}^t \rangle \leq \lambda_{t+1} \langle \widehat{\mathbf{v}}^t, \widehat{\mathbf{v}}^{t+1} \rangle$. Substituting in the above, we obtain

$$\lambda_t \leq \lambda_{t+1} \langle \widehat{\mathbf{v}}^t, \widehat{\mathbf{v}}^{t-1} \rangle \langle \widehat{\mathbf{v}}^t, \widehat{\mathbf{v}}^{t+1} \rangle.$$

Since $\lambda_t \rightarrow \lambda_*$, we conclude that $\langle \widehat{\mathbf{v}}^t, \widehat{\mathbf{v}}^{t-1} \rangle \rightarrow 1$ or

$$\lim_{t \rightarrow \infty} \|\widehat{\mathbf{v}}^t - \widehat{\mathbf{v}}^{t+1}\|_2 = 0.$$

Using this in the identity $\lambda_t = \langle \widehat{\mathbf{v}}^t, \mathbf{X}_\rho \widehat{\mathbf{v}}^{t-1} \rangle$ derived above, we deduce that $\langle \widehat{\mathbf{v}}^t, \mathbf{X}_\rho \widehat{\mathbf{v}}^t \rangle \rightarrow \lambda_*$. Using it in Eq. (18) (with $t = t(k)$ as per the statement) we get the stationarity condition (25). \square

Remark 4.4. Corollary 4.1 follows directly from the proof of Theorem 3 and the Gaussian comparison lemma 3.1.

References

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