Provable Guarantees for Self-Supervised Deep Learning with Spectral Contrastive Loss

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Abstract

Recent works in self-supervised learning have advanced the state-of-the-art by relying on the *contrastive learning* paradigm, which learns representations by pushing positive pairs, or similar examples from the same class, closer together while keeping negative pairs far apart. Despite the empirical successes, theoretical foundations are limited – prior analyses assume conditional independence of the positive pairs given the same class label, but recent empirical applications use heavily correlated positive pairs (i.e., data augmentations of the same image). Our work analyzes contrastive learning without assuming conditional independence of positive pairs using a novel concept of the *augmentation graph* on data. Edges in this graph connect augmentations of the same data, and ground-truth classes naturally form connected sub-graphs. We propose a loss that performs spectral decomposition on the population augmentation graph and can be succinctly written as a contrastive learning objective on neural net representations. Minimizing this objective leads to features with provable accuracy guarantees under linear probe evaluation. By standard generalization bounds, these accuracy guarantees also hold when minimizing the training contrastive loss. In all, this work provides the first provable analysis for contrastive learning where the guarantees can apply to realistic empirical settings.

1 Introduction

Recent empirical breakthroughs have demonstrated the effectiveness of self-supervised learning, which trains representations on unlabeled data with surrogate losses and self-defined supervision signals [4, 6, 10, 14, 23, 24, 35, 38, 41, 42, 50–52]. Self-supervision signals in computer vision are often defined by using data augmentation to produce multiple views of the same image. For example, the recent contrastive learning objectives [3, 12, 13, 15, 22] encourage closer representations for augmentations (views) of the same natural data than for randomly sampled pairs of data.

Despite the empirical successes, there is a limited theoretical understanding of why self-supervised losses learn representations that can be adapted to downstream tasks, for example, using linear heads. Recent mathematical analyses by Arora et al. [3], Lee et al. [28], Tosh et al. [44, 45] provide guarantees under the assumption that two views are somewhat independent conditioned on the label. However, the pair of augmented examples used in practical algorithms usually exhibit a strong correlation, even conditioned on the label. For instance, two augmentations of the same dog image share much more similarity than augmentations of two different random dog images. Thus the existing theory does not explain the practical success of self-supervised learning.

This paper presents a theoretical framework for self-supervised learning without requiring conditional independence. We design a principled, practical loss function for learning neural net representations that resembles state-of-the-art contrastive learning methods. We prove that, under a simple and

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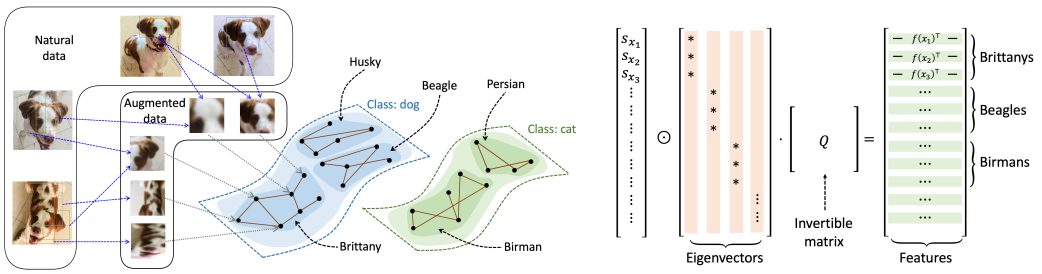

Figure 1: **Left: demonstration of the population augmentation graph.** Two augmented data are connected if they are views of the same natural data. Augmentations of data from different classes in the downstream tasks are assumed to be nearly disconnected, whereas there are more connections within the same class. We allow the existence of disconnected sub-graphs within a class corresponding to potential sub-classes. **Right: decomposition of the learned representations.** The representations (rows in the RHS) learned by minimizing the population spectral contrastive loss can be decomposed as the LHS. The scalar s_{x_i} is positive for every augmented data x_i. Columns of the matrix labeled "eigenvectors" are the top eigenvectors of the normalized adjacency matrix of the augmentation graph defined in Section 3.1. The operator $\odot$ multiplies row-wise each s_{x_i} with the x_i-th row of the eigenvector matrix. When classes (or sub-classes) are exactly disconnected in the augmentation graph, the eigenvectors are sparse and align with the sub-class structure. The invertible Q matrix does not affect the performance of the rows under the linear probe.

realistic data assumption, linear classification using representations learned on a polynomial number of unlabeled data samples can recover the ground-truth labels of the data with high accuracy.

The fundamental data property that we leverage is a notion of continuity of the population data within the same class. Though a random pair of examples from the same class can be far apart, the pair is often connected by (many) sequences of examples, where consecutive examples in the sequences are close neighbors within the same class. This property is more salient when the neighborhood of an example includes many different types of augmentations. Prior work [49] empirically demonstrates this type of connectivity property and uses it in the analysis of pseudolabeling algorithms.

More formally, we define the *population augmentation graph*, whose vertices are all the augmented data in the *population* distribution, which can be an exponentially large or infinite set. Two vertices are connected with an edge if they are augmentations of the same natural example. Our main assumption is that for some proper $m \in \mathcal{Z}^+$, the sparsest m-partition (Definition 3.4) is large. This intuitively states that we can't split the augmentation graph into too many disconnected sub-graphs by only removing a sparse set of edges. This assumption can be seen as a graph-theoretic version of the continuity assumption on population data. We also assume that there are very few edges across different ground-truth classes (Assumption 3.5). Figure 1 (left) illustrates a realistic scenario where dog and cat are the ground-truth categories, between which edges are very rare. Each breed forms a sub-graph that has sufficient inner connectivity and thus cannot be further partitioned.

Our assumption fundamentally does not require conditional independence and can allow disconnected sub-graphs within a class. The classes in the downstream task can be also somewhat flexible as long as they are disconnected in the augmentation graph. For example, when the augmentation graph consists of m disconnected sub-graphs corresponding to fine-grained classes, our assumptions allow the downstream task to have any $r \leq m$ coarse-grained classes containing these fine-grained classes as a sub-partition. Prior work [49] on pseudolabeling algorithms essentially requires an exact alignment between sub-graphs and downstream classes (i.e., $r = m$). They face this limitation because their analysis requires fitting discrete pseudolabels on the unlabeled data. We avoid this difficulty because we consider directly learning continuous representations on the unlabeled data.

We apply spectral decomposition—a classical approach for graph partitioning, also known as spectral clustering [37, 39] in machine learning—to the adjacency matrix defined on the population augmentation graph. We form a matrix where the top-k eigenvectors are the columns and interpret each row of the matrix as the representation (in $\mathbb{R}^k$) of an example. Somewhat surprisingly, we show that this feature extractor can be also recovered (up to some linear transformation) by minimizing the

following population objective which is similar to the standard contrastive loss (Section 3.2):

$$\mathcal{L}(f) = -2 \cdot \mathbb{E}_{x,x^+}\left[f(x)^\top f(x^+)\right] + \mathbb{E}_{x,x'}\left[\left(f(x)^\top f(x')\right)^2\right],$$

where (x, x^+) is a pair of augmentations of the same data, (x, x') is a pair of independently random augmented data, and f is a parameterized function from augmented data to $\mathbb{R}^k$. Figure 1 (right) illustrates the relationship between the eigenvector matrix and the learned representations. We call this loss the *population spectral contrastive loss*.

We analyze the linear classification performance of the representations learned by minimizing the population spectral contrastive loss. Our main result (Theorem 3.7) shows that when the representation dimension exceeds the maximum number of disconnected sub-graphs, linear classification with learned representations is guaranteed to have a small error. Our theorem reveals a trend that a larger representation dimension is needed when there are a larger number of disconnected sub-graphs. Our analysis relies on novel techniques tailored to linear probe performance, which have not been studied in the spectral graph theory community to the best of our knowledge.

The spectral contrastive loss also works on empirical data. Since our approach optimizes parametric loss functions, guarantees involving the population loss can be converted to finite sample results using off-the-shelf generalization bounds. The sample complexity is polynomial in the Rademacher complexity of the model family and other relevant parameters (Theorem 4.1 and Theorem 4.2).

In summary, our main theoretical contributions are: 1) we propose a simple contrastive loss motivated by spectral decomposition of the population data graph, 2) under simple and realistic assumptions, we provide downstream classification guarantees for the representation learned by minimizing this loss on population data, and 3) our analysis is easily applicable to deep networks with polynomial unlabeled samples via off-the-shelf generalization bounds.

In addition, we implement and test the proposed spectral contrastive loss on standard vision benchmark datasets. We demonstrate that the features learned by our algorithm can match or outperform several strong baselines [12, 14, 15, 21] when evaluated using a linear probe.

2 Additional related works

Empirical works on self-supervised learning. Self-supervised learning algorithms have been shown to successfully learn representations that benefit downstream tasks [4, 6, 10, 12, 13, 15, 22–24, 35, 38, 41, 42, 50–52]. Many recent self-supervised learning algorithms learn features with siamese networks [8], where two neural networks of shared weights are applied to pairs of augmented data. Introducing asymmetry to siamese networks either with a momentum encoder like BYOL [21] or by stopping gradient propagation for one branch of the siamese network like SimSiam [14] has been shown to effectively avoid collapsing. Contrastive methods [12, 15, 22] minimize the InfoNCE loss [38], where two views of the same data are attracted while views from different data are repulsed.

Theoretical works on self-supervised learning. In addition to works [3, 28, 44, 45] discussed in the introduction, several other works [5, 43, 47, 48] also theoretically study self-supervised learning. The work Tsai et al. [47] prove that self-supervised learning methods can extract task-relevant information and discard task-irrelevant information, but lacks guarantees for solving downstream tasks efficiently with simple (e.g., linear) models. Tian et al. [43] study why non-contrastive self-supervised learning methods can avoid feature collapse. Cai et al. [9] analyze domain adaptation algorithms for subpopulation shift with a similar expansion condition as [49] while also allowing disconnected parts within each class, but require access to ground-truth labels during training. In contrast, our algorithm doesn't need labels during pre-training.

3 Spectral contrastive learning on population data

In this section, we introduce our theoretical framework, the spectral contrastive loss, and the main analysis of the performance of the representations learned on population data.

We use $\overline{\mathcal{X}}$ to denote the set of all natural data (raw inputs without augmentation). We assume that each $\bar{x} \in \overline{\mathcal{X}}$ belongs to one of r classes, and let $y : \overline{\mathcal{X}} \to [r]$ denote the ground-truth (deterministic) labeling function. Let $\mathcal{P}_{\overline{\mathcal{X}}}$ be the population distribution over $\overline{\mathcal{X}}$ from which we draw training data and

test our final performance. For the ease of exposition, we assume $\overline{\mathcal{X}}$ to be a finite but exponentially large set (e.g., all real vectors in $\mathbb{R}^d$ with bounded precision).[1]

We next formulate data augmentations. Given a natural data sample $\bar{x} \in \overline{\mathcal{X}}$, we use $\mathcal{A}(\cdot|\bar{x})$ to denote the distribution of its augmentations. For instance, when $\bar{x}$ represents an image, $\mathcal{A}(\cdot|\bar{x})$ can be the distribution of common augmentations [12] that includes Gaussian blur, color distortion and random cropping. We use $\mathcal{X}$ to denote the set of all augmented data, which is the union of supports of all $\mathcal{A}(\cdot|\bar{x})$ for $\bar{x} \in \overline{\mathcal{X}}$. As with $\overline{\mathcal{X}}$, we also assume that $\mathcal{X}$ is a finite but exponentially large set, and denote $N = |\mathcal{X}|$.

We will learn an embedding function $f : \mathcal{X} \to \mathbb{R}^k$, and then evaluate its quality by the minimum error achieved with a linear probe. Concretely, a linear classifier has weights $B \in \mathbb{R}^{k \times r}$ and predicts $g_{f,B}(x) = \arg\max_{i \in [r]} (f(x)^\top B)_i$ for an augmented data x ($\arg\max$ breaks tie arbitrarily). Then, given a natural data sample $\bar{x}$, we ensemble the predictions on augmented data and predict:

$$\bar{g}_{f,B}(\bar{x}) := \arg\max_{i \in [r]} \Pr_{x \sim \mathcal{A}(\cdot|\bar{x})} [g_{f,B}(x) = i] .$$

Define the *linear probe* error as the error of the best possible linear classifier on the representations:

$$\mathcal{E}(f) := \min_{B \in \mathbb{R}^{k \times r}} \Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}} [y(\bar{x}) \neq \bar{g}_{f,B}(\bar{x})] \tag{1}$$

3.1 Augmentation graph and spectral decomposition

Our approach is based on the central concept of **population augmentation graph**, denoted by $G(\mathcal{X}, w)$, where the vertex set is all augmentation data $\mathcal{X}$ and w denotes the edge weights defined below. For any two augmented data $x, x' \in \mathcal{X}$, define the weight $w_{xx'}$ as the marginal probability of generating the pair x and x' from a random natural data $\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}$:

$$w_{xx'} := \mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}} [\mathcal{A}(x|\bar{x})\mathcal{A}(x'|\bar{x})] \tag{2}$$

Therefore, the weights sum to 1 because the total probability mass is 1: $\sum_{x,x' \in \mathcal{X}} w_{xx'} = 1$. The relative magnitude intuitively captures the closeness between x and x' with respect to the augmentation transformation. For most of the unrelated x and x', the value $w_{xx'}$ will be significantly smaller than the average value. For example, when x and x' are random croppings of a cat and a dog respectively, $w_{xx'}$ will be essentially zero because no natural data can be augmented into both x and x'. On the other hand, when x and x' are very close in ℓ_2-distance or very close in ℓ_2-distance up to color distortion, $w_{xx'}$ is nonzero because they may be augmentations of the same image with Gaussian blur and color distortion. We say that x and x' are connected with an edge if $w_{xx'} > 0$. See Figure 1 (left) for more illustrations.

Given the structure of the population augmentation graph, we apply spectral decomposition to the population graph to construct principled embeddings. The eigenvalue problems are closely related to graph partitioning as shown in spectral graph theory [17] for both worst-case graphs [11, 25, 29, 33] and random graphs [1, 30, 34]. In machine learning, spectral clustering [37, 39] is a classical algorithm that learns embeddings by eigendecomposition on an empirical distance graph and invoking k-means on the embeddings.

We will apply eigendecomposition to the *population* augmentation graph (and then later use linear probe for classification). Let $w_x = \sum_{x' \in \mathcal{X}} w_{xx'}$ be the total weights associated to x, which is often viewed as an analog of the degree of x in weighted graph. A central object in spectral graph theory is the so-called *normalized adjacency matrix*:

$$\overline{A} := D^{-1/2} A D^{-1/2} \tag{3}$$

where $A \in \mathbb{R}^{N \times N}$ is adjacency matrix with entires $A_{xx'} = w_{xx'}$ and $D \in \mathbb{R}^{N \times N}$ is a diagonal matrix with $D_{xx} = w_x$.[2]

Standard spectral graph theory approaches produce vertex embeddings as follows. Let $\gamma_1, \gamma_2, \cdots, \gamma_k$ be the k largest eigenvalues of $\overline{A}$, and $v_1, v_2, \cdots, v_k$ be the corresponding unit-norm eigenvectors.

[1] This allows us to use sums instead of integrals and avoid non-essential nuances related to calculus.

[2] We index the matrix A, D by $(x, x') \in \mathcal{X} \times \mathcal{X}$. Generally we index N-dimensional axis by $x \in \mathcal{X}$.

Let $F^\star = [v_1, v_2, \cdots, v_k] \in \mathbb{R}^{N\times k}$ be the matrix that collects these eigenvectors in columns, and we refer to it as the eigenvector matrix. Let $u_x^* \in \mathbb{R}^k$ be the x-th row of the matrix $F^\star$. It turns out that u_x^*'s can serve as desirable embeddings of x's because they exhibit clustering structure in Euclidean space that resembles the clustering structure of the graph $G(\mathcal{X}, w)$.

3.2 From spectral decomposition to spectral contrastive learning

The embeddings u_x^* obtained by eigendecomposition are nonparametric—a k-dimensional parameter is needed for every x—and therefore cannot be learned with a realistic amount of data. The embedding matrix $F^\star$ cannot be even stored efficiently. Therefore, we will instead parameterize the rows of the eigenvector matrix $F^\star$ as a neural net function, and assume embeddings u_x^* can be represented by $f(x)$ for some $f \in \mathcal{F}$, where $\mathcal{F}$ is the hypothesis class containing neural networks. As we'll show in Section 4, this allows us to leverage the extrapolation power of neural networks and learn the representation on a finite dataset.

Next, we design a proper loss function for the feature extractor f, such that minimizing this loss could recover $F^\star$ up to some linear transformation. As we will show in Section 4, the resulting population loss function on f also admits an unbiased estimator with finite training samples. Let F be an embedding matrix with u_x on the x-th row, we will first design a loss function of F that can be decomposed into parts about individual rows of F.

We employ the following matrix factorization based formulation for eigenvectors. Consider the objective

$$\min_{F\in\mathbb{R}^{N\times k}} \mathcal{L}_{\text{mf}}(F) := \left\|\overline{A} - FF^\top\right\|_F^2. \tag{4}$$

By the classical theory on low-rank approximation (Eckart–Young–Mirsky theorem [19]), any minimizer $\widehat{F}$ of $\mathcal{L}_{\text{mf}}(F)$ contains scaling of the largest eigenvectors of $\overline{A}$ up to a right transformation—for some orthonormal matrix $R \in \mathbb{R}^{k\times k}$, we have $\widehat{F} = F^\star \cdot \text{diag}([\sqrt{\gamma_1}, \ldots, \sqrt{\gamma_k}])Q$. Fortunately, multiplying the embedding matrix by any matrix on the right and any diagonal matrix on the left does not change its linear probe performance, which is formalized by the following lemma.

Lemma 3.1. *Consider an embedding matrix $F \in \mathbb{R}^{N\times k}$ and a linear classifier $B \in \mathbb{R}^{k\times r}$. Let $D \in \mathbb{R}^{N\times N}$ be a diagonal matrix with positive diagonal entries and $Q \in \mathbb{R}^{k\times k}$ be an invertible matrix. Then, for any embedding matrix $\widetilde{F} = D \cdot F \cdot Q$, the linear classifier $\tilde{B} = Q^{-1}B$ on $\widetilde{F}$ has the same prediction as B on F. As a consequence, we have*

$$\mathcal{E}(F) = \mathcal{E}(\widetilde{F}). \tag{5}$$

where $\mathcal{E}(F)$ denotes the linear probe performance when the rows of F are used as embeddings.

The proof can be found in Section C.1.

The main benefit of objective $\mathcal{L}_{\text{mf}}(F)$ is that it's based on the rows of F. Recall that vectors u_x are the rows of F. Each entry of $FF^\top$ is of the form $u_x^\top u_{x'}$, and thus $\mathcal{L}_{\text{mf}}(F)$ can be decomposed into a sum of N^2 terms involving terms $u_x^\top u_{x'}$. Interestingly, if we reparameterize each row u_x by $w_x^{1/2} f(x)$, we obtain a very similar loss function for f that resembles the contrastive learning loss used in practice [12] as shown below in Lemma 3.2. See Figure 1 (right) for an illustration of the relationship between the eigenvector matrix and the representations learned by minimizing this loss.

We formally define the positive and negative pairs to introduce the loss. Let $\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}$ be a random natural data and draw $x \sim \mathcal{A}(\cdot|\bar{x})$ and $x^+ \sim \mathcal{A}(\cdot|\bar{x})$ independently to form a positive pair (x, x^+). Draw $\bar{x}' \sim \mathcal{P}_{\overline{\mathcal{X}}}$ and $x' \sim \mathcal{A}(\cdot|\bar{x}')$ independently with $\bar{x}, x, x^+$. We call (x, x') a negative pair.[3]

Lemma 3.2 (Spectral contrastive loss)**.** *Recall that u_x is the x-th row of F. Let $u_x = w_x^{1/2} f(x)$ for some function f. Then, the loss function $\mathcal{L}_{\text{mf}}(F)$ is equivalent to the following loss function for f, called spectral contrastive loss, up to a additive constant:*

$$\begin{aligned} &\mathcal{L}_{\text{mf}}(F) = \mathcal{L}(f) + \text{const} \\ &\text{where } \mathcal{L}(f) \triangleq -2 \cdot \mathbb{E}_{x,x^+}\left[f(x)^\top f(x^+)\right] + \mathbb{E}_{x,x'}\left[\left(f(x)^\top f(x')\right)^2\right] \end{aligned} \tag{6}$$

[3]Though x and x' are simply two independent draws, we call them negative pairs following the literature [3].

The proof can be found in Section C.1.

We note that spectral contrastive loss is similar to many popular contrastive losses [12, 38, 40, 50]. For instance, the contrastive loss in SimCLR [12] can be rewritten as (with simple algebraic manipulation)

$$-f(x)^\top f(x^+) + \log\left(\exp\left(f(x)^\top f(x^+)\right) + \sum_{i=1}^{n} \exp\left(f(x)^\top f(x_i)\right)\right).$$

Here x and x^+ are a positive pair and $x_1, \cdots, x_n$ are augmentations of other data. Spectral contrastive loss can be seen as removing $f(x)^\top f(x^+)$ from the second term, and replacing the log sum of exponential terms with the average of the squares of $f(x)^\top f(x_i)$. We will show in Section 6 that our loss has a similar empirical performance as SimCLR without requiring a large batch size.

3.3 Theoretical guarantees for spectral contrastive loss on population data

In this section, we introduce the main assumptions on the data and state our main theoretical guarantee for spectral contrastive learning on population data.

To formalize the idea that G cannot be partitioned into too many disconnected sub-graphs, we introduce the notions of *Dirichlet conductance* and *sparsest m-partition*, which are standard in spectral graph theory. Dirichlet conductance represents the fraction of edges from S to its complement:

Definition 3.3 (Dirichlet conductance)**.** *For a graph $G = (\mathcal{X}, w)$ and a subset $S \subseteq \mathcal{X}$, we define the Dirichlet conductance of S as*

$$\phi_G(S) := \frac{\sum_{x \in S, x' \notin S} w_{xx'}}{\sum_{x \in S} w_x}.$$

We note that when S is a singleton, there is $\phi_G(S) = 1$ due to the definition of w_x. We introduce the sparsest m-partition to represent the number of edges between m disjoint subsets.

Definition 3.4 (Sparsest m-partition)**.** *Let $G = (\mathcal{X}, w)$ be the augmentation graph. For an integer $m \in [2, |\mathcal{X}|]$, we define the sparsest m-partition as*

$$\rho_m := \min_{S_1, \cdots, S_m} \max\{\phi_G(S_1), \ldots, \phi_G(S_m)\}$$

where $S_1, \cdots, S_m$ are non-empty sets that form a partition of $\mathcal{X}$.

When r is the number of underlying classes, we might expect $\rho_r \approx 0$ since the augmentations from different classes almost compose a disjoint r-way partition of $\mathcal{X}$. However, for $m > r$, we can expect ρ_m to be much larger. For instance, in the extreme case when $m = |\mathcal{X}| = N$, every set S_i is a singleton, which implies that $\rho_N = 1$.

Next, we formalize the assumption that very few edges cross different ground-truth classes. It turns out that it suffices to assume that the labels are recoverable from the augmentations (which is also equivalent to that two examples in different classes can rarely be augmented into the same point).

Assumption 3.5 (Labels are recoverable from augmentations)**.** *Let $\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}$ and $y(\bar{x})$ be its label. Let the augmentation $x \sim \mathcal{A}(\cdot|\bar{x})$. We assume that there exists a classifier g that can predict $y(\bar{x})$ given x with error at most α. That is, $g(x) = y(\bar{x})$ with probability at least $1 - \alpha$.*

We also introduce the following assumption which states that some universal minimizer of the population spectral contrastive loss can be realized by the hypothesis class.

Assumption 3.6 (Realizability)**.** *Let $\mathcal{F}$ be a hypothesis class containing functions from $\mathcal{X}$ to $\mathbb{R}^k$. We assume that at least one of the global minima of $\mathcal{L}(f)$ belongs to $\mathcal{F}$.*

Our main theorem bound from above the linear probe error of the features learned by minimizing the population spectral contrastive loss.

Theorem 3.7. *Assume the representation dimension $k \geq 2r$ and Assumption 3.5 holds for $\alpha > 0$. Let $\mathcal{F}$ be a hypothesis class that satisfies Assumption 3.6 and let $f^*_{\text{pop}} \in \mathcal{F}$ be a minimizer of $\mathcal{L}(f)$. Then, we have*

$$\mathcal{E}(f^*_{\text{pop}}) \leq \widetilde{O}\left(\alpha/\rho^2_{\lfloor k/2 \rfloor}\right).$$

Here we use $\widetilde{O}(\cdot)$ to hide universal constant factors and logarithmic factor in k. We note that $\alpha = 0$ when augmentations from different classes are perfectly disconnected in the augmentation graph, in which case the above theorem guarantees the exact recovery of the ground truth. Generally, we expect α to be an extremely small constant independent of k, whereas $\rho_{\lfloor k/2 \rfloor}$ increases with k and can be much larger than α when k is reasonably large. For instance, when there are t sub-graphs that have sufficient inner connections, we expect ρ_{t+1} to be on the order of a constant because any $t+1$ partition needs to break one sub-graph into two pieces and incur a large conductance. We characterize the ρ_k's growth on more concrete distributions in the next subsection.

Previous works on graph partitioning [2, 29, 31] often analyze the so-called rounding algorithms that conduct clustering based on the representations of unlabeled data and do not analyze the performance of linear probe (which has access to labeled data). These results provide guarantees on the approximation ratio—the ratio between the conductance of the obtained partition to the best partition—which may depend on graph size [2] that can be exponentially large in our setting. The approximation ratio guarantee does not lead to a guarantee on the representations' performance on downstream tasks. Our guarantees are on the linear probe accuracy on the downstream tasks and independent of the graph size. We rely on the formulation of the downstream task's labeling function (Assumption 3.5) as well as a novel analysis technique that characterizes the linear structure of the representations. In Section C, we provide the proof of Theorem 3.7 as well as its more generalized version where $k/2$ is relaxed to be any constant fraction of k.

3.4 Provable instantiation of Theorem 3.7 to mixture of manifold data

In this section, we exemplify Theorem 3.7 on an example where the natural data distribution is a mixture of manifolds, and the augmentation transformation is adding Gaussian noise.

Example 3.8 (Mixture of manifolds)**.** *Suppose $\mathcal{P}_{\overline{\mathcal{X}}}$ is mixture of $r \le d$ distributions $P_1, \cdots, P_r$, where each P_i is generated by some κ-bi-Lipschitz[4] generator $Q : \mathbb{R}^{d'} \to \mathbb{R}^d$ on some latent variable $z \in \mathbb{R}^{d'}$ with $d' \le d$ which as a mixture of Gaussian distribution:*

$$x \sim P_i \iff x = Q(z), z \sim \mathcal{N}(\mu_i, \frac{1}{d'} \cdot I_{d' \times d'}).$$

Let the data augmentation of a natural data sample $\bar{x}$ is $\bar{x} + \xi$ where $\xi \sim \mathcal{N}(0, \frac{\sigma^2}{d} \cdot I_{d \times d})$ is isotropic Gaussian noise with $0 < \sigma \lesssim \frac{1}{\sqrt{d}}$. We also assume $\min_{i \neq j} \|\mu_i - \mu_j\|_2 \gtrsim \frac{\kappa \cdot \sqrt{\log d}}{\sqrt{d'}}$.

Let the ground-truth label be the most likely mixture index i that generates x: $y(x) := \arg\max_i P_i(x)$. We note that the intra-class distance in the latent space is on the scale of $\Omega(1)$, which can be much larger than the distance between class means which is assumed to be $\gtrsim \frac{\kappa \cdot \sqrt{\log d}}{\sqrt{d'}}$. Therefore, distance-based clustering algorithms do not apply. We apply Theorem 3.7 and get the following theorem:

Theorem 3.9. *When $k \ge 2r + 2$, Example 3.8 satisfies Assumption 3.5 with $\alpha \le \frac{1}{\text{poly}(d)}$, and has $\rho_{\lfloor k/2 \rfloor} \gtrsim \frac{\sigma}{\kappa \sqrt{d}}$. As a consequence, the error bound is $\mathcal{E}(f^*_{\text{pop}}) \le \tilde{O}\left(\frac{\kappa^2}{\sigma^2 \cdot \text{poly}(d)}\right)$.*

The theorem above guarantees small error even when σ is polynomially small. In this case, the augmentation noise has a much smaller scale than the data (which is at least on the order of $1/\kappa$). This suggests that contrastive learning can non-trivially leverage the structure of the underlying data and learn good representations with relatively weak augmentation. The proof can be found in Section D.

4 Finite-sample generalization bounds

In Section 3, we provide guarantees for spectral contrastive learning on population data. In this section, we show that these guarantees can be naturally extended to the finite-sample regime with standard concentration bounds. In particular, given a training dataset $\{\bar{x}_1, \bar{x}_2, \cdots, \bar{x}_n\}$ with $\bar{x}_i \sim \mathcal{P}_{\overline{\mathcal{X}}}$, we learn a feature extractor by minimizing the following *empirical spectral contrastive loss*:

$$\widehat{\mathcal{L}}_n(f) := -\frac{2}{n} \sum_{i=1}^{n} \mathbb{E}_{\substack{x \sim \mathcal{A}(\cdot|\bar{x}_i) \\ x^+ \sim \mathcal{A}(\cdot|\bar{x}_i)}} \left[f(x)^\top f(x^+) \right] + \frac{1}{n(n-1)} \sum_{i \neq j} \mathbb{E}_{\substack{x \sim \mathcal{A}(\cdot|\bar{x}_i) \\ x' \sim \mathcal{A}(\cdot|\bar{x}_j)}} \left[\left(f(x)^\top f(x') \right)^2 \right].$$

[4] A κ bi-Lipschitz function satisfies $\frac{1}{\kappa} \|f(x) - f(y)\|_2 \le \|x - y\|_2 \le \kappa \|f(x) - f(y)\|_2$.

It is worth noting that $\widehat{\mathcal{L}}_n(f)$ is an unbiased estimator of the population spectral contrastive loss $\mathcal{L}(f)$. (See Claim E.2 for a proof.) Therefore, we can derive generalization bounds via off-the-shelf concentration inequalities. Let $\mathcal{F}$ be a hypothesis class containing feature extractors from $\mathcal{X}$ to $\mathbb{R}^k$. We extend Rademacher complexity to function classes with high-dimensional outputs and define the Rademacher complexity of $\mathcal{F}$ on n data as $\widehat{\mathcal{R}}_n(\mathcal{F}) := \max_{x_1,\cdots,x_n\in\mathcal{X}} \mathbb{E}_\sigma \left[\sup_{f\in\mathcal{F},i\in[k]} \frac{1}{n}\left(\sum_{j=1}^n \sigma_j f_i(x_j)\right)\right]$, where σ is a uniform random vector in $\{-1,1\}^n$ and $f_i(z)$ is the i-th dimension of $f(z)$.

Recall that $f^*_{\text{pop}} \in \mathcal{F}$ is a minimizer of $\mathcal{L}(f)$. The following theorem with proofs in Section E.1 bounds the population loss of a feature extractor trained with finite data:

Theorem 4.1. *For some $\kappa > 0$, assume $\|f(x)\|_\infty \le \kappa$ for all $f \in \mathcal{F}$ and $x \in \mathcal{X}$. Let $f^*_{\text{pop}} \in \mathcal{F}$ be a minimizer of the population loss $\mathcal{L}(f)$. Given a random dataset of size n, let $\hat{f}_{\text{emp}} \in \mathcal{F}$ be a minimizer of empirical loss $\widehat{\mathcal{L}}_n(f)$. Then, with probability at least $1-\delta$ over the randomness of data, we have*

$$\mathcal{L}(\hat{f}_{\text{emp}}) \le \mathcal{L}(f^*_{\text{pop}}) + c_1 \cdot \widehat{\mathcal{R}}_{n/2}(\mathcal{F}) + c_2 \cdot \left(\sqrt{\frac{\log 2/\delta}{n}} + \delta\right),$$

where constants $c_1 \lesssim k^2\kappa^2 + k\kappa$ and $c_2 \lesssim k\kappa^2 + k^2\kappa^4$.

We can apply Theorem 4.1 to any hypothesis class $\mathcal{F}$ of interest (e.g., deep neural networks) and plug in off-the-shelf Rademacher complexity bounds. For instance, in Section E.2 we give a corollary of Theorem 4.1 when $\mathcal{F}$ contains deep neural networks with ReLU activation.

The theorem above shows that we can achieve near-optimal population loss by minimizing empirical loss up to some small excess loss. The following theorem characterizes how the error propagates to the linear probe performance mildly under some spectral gap conditions.

Theorem 4.2. *Assume representation dimension $k \ge 4r+2$, Assumption 3.5 holds for $\alpha > 0$ and Assumption 3.6 holds. Recall γ_i be the i-th largest eigenvalue of the normalized adjacency matrix. Then, for any $\epsilon < \gamma_k^2$ and $\hat{f}_{\text{emp}} \in \mathcal{F}$ such that $\mathcal{L}(\hat{f}_{\text{emp}}) < \mathcal{L}(f^*_{\text{pop}}) + \epsilon$, we have:*

$$\mathcal{E}(\hat{f}_{\text{emp}}) \lesssim \frac{\alpha}{\rho^2_{\lfloor k/2\rfloor}} \cdot \log k + \frac{k\epsilon}{\Delta_\gamma^2},$$

where $\Delta_\gamma := \gamma_{\lfloor 3k/4\rfloor} - \gamma_k$ is the eigenvalue gap between the $\lfloor 3k/4\rfloor$-th and the k-th eigenvalue.

This theorem shows that the error on the downstream task only grows linearly with the error ϵ during pretraining. We can relax Assumption 3.6 to approximate realizability in the sense that $\mathcal{F}$ contains some sub-optimal feature extractor under the population spectral loss and pay an additional error term in the linear probe error bound. The proof of Theorem 4.2 can be found in Section E.3.

5 Guarantee for learning linear probe with labeled data

In this section, we provide theoretical guarantees for learning a linear probe with *labeled* data. Theorem 3.7 guarantees the existence of a linear probe that achieves a small downstream classification error. However, a priori it is unclear how large the margin of the linear classifier can be, so it is hard to apply margin theory to provide generalization bounds for 0-1 loss. We could in principle control the margin of the linear head, but using capped quadratic loss turns out to suffice and mathematically more convenient. We learn a linear head with the following *capped quadratic loss*: given a tuple $(z, y(\bar{x}))$ where $z \in \mathbb{R}^k$ is a representation of augmented data $x \sim \mathcal{A}(\cdot|\bar{x})$ and $y(\bar{x}) \in [r]$ is the label of $\bar{x}$, for a linear probe $B \in \mathbb{R}^{k\times r}$ we define loss $\ell((z, y(\bar{x})), B) := \sum_{i=1}^r \min\left\{\left(B^\top z - \vec{y}(\bar{x})\right)_i^2, 1\right\}$, where $\vec{y}(\bar{x})$ is the one-hot embedding of $y(\bar{x})$ as a r-dimensional vector (1 on the $y(\bar{x})$-th dimension, 0 on other dimensions). This is a standard modification of quadratic loss in statistical learning theory that ensures the boundedness of the loss for the ease of analysis [36].

The following Theorem 5.1 provides a generalization guarantee for the linear classifier that minimizes capped quadratic loss on a labeled dataset. The key challenge of the proof is showing the existence of a small-norm linear head B that gives small population quadratic loss, which is not obvious from Theorem 3.7 where only small 0-1 error is guaranteed. Recall γ_i is the i-th largest eigenvalue of the the normalized adjacency matrix. Given a labeled dataset $\{(\bar{x}_i, y(\bar{x}_i))\}_{i=1}^n$ where $\bar{x}_i \sim \mathcal{P}_{\overline{\mathcal{X}}}$ and $y(\bar{x}_i)$ is its label, we sample $x_i \sim \mathcal{A}(\cdot|\bar{x}_i)$ for $i \in [n]$.

Theorem 5.1. *In the setting of Theorem 3.7, assume $\gamma_k \geq C_\lambda$ for some $C_\lambda > 0$. Learn a linear probe $\widehat{B} \in \arg\min_{\|B\|_F \leq 1/C_\lambda} \sum_{i=1}^{n} \ell((f^*_{\text{pop}}(x_i), y(\bar{x}_i)), B)$ by minimizing the capped quadratic loss subject to a norm constraint. Then, with probability at least $1-\delta$ over random data, we have*

$$\Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}}}\left(\bar{g}_{f^*_{\text{pop}},\widehat{B}}(\bar{x}) \neq y(\bar{x})\right) \lesssim \frac{\alpha}{\rho^2_{\lfloor k/2 \rfloor}} \cdot \log k + \frac{r}{C_\lambda} \cdot \sqrt{\frac{k}{n}} + \sqrt{\frac{\log 1/\delta}{n}}.$$

Here the first term is the population error from Theorem 3.7. The last two terms are the generalization gap from standard concentration inequalities for linear classification and are small when the number of labeled data n is polynomial in the feature dimension k. We note that this result reveals a trade-off when choosing the feature dimension k: when n is fixed, a larger k decreases the population contrastive loss while increases the generalization gap for downstream linear classification. The proof of Theorem 5.1 is in Section F.

6 Experiments

We test spectral contrastive learning on benchmark vision datasets. We minimize the empirical spectral contrastive loss with an encoder network f and sample fresh augmentation in each iteration. The pseudo-code for the algorithm and more implementation details can be found in Section A.

Encoder / feature extractor. The encoder f contains three components: a backbone network, a projection MLP and a projection function. The backbone network is a standard ResNet architecture. The projection MLP is a fully connected network with BN applied to each layer, and ReLU activation applied to each except for the last layer. The projection function takes a vector and projects it to a sphere ball with radius $\sqrt{\mu}$, where $\mu > 0$ is a hyperparameter that we tune in experiments. We find that using a projection MLP and a projection function improves the performance.

Linear evaluation protocol. Given the pre-trained encoder network, we follow the standard linear evaluation protocol [14] and train a supervised linear classifier on frozen representations, which are from the ResNet's global average pooling layer.

Results. We report the accuracy on CIFAR-10/100 [26] and Tiny-ImageNet [27] in Table 1. Our empirical results show that spectral contrastive learning achieves better performance than two popular baseline algorithms SimCLR [12] and SimSiam [14]. In Table 2 we report results on ImageNet [18] dataset, and show that our algorithm achieves similar performance as other state-of-the-art methods. We note that our algorithm is much more principled than previous methods and doesn't rely on large batch sizes (SimCLR [12]), momentum encoders (BYOL [21] and MoCo [22]) or additional tricks such as stop-gradient (SimSiam [14]).

Datasets	CIFAR-10			CIFAR-100			Tiny-ImageNet		
Epochs	200	400	800	200	400	800	200	400	800
SimCLR (repro.)	83.73	87.72	90.60	54.74	61.05	63.88	**43.30**	**46.46**	48.12
SimSiam (repro.)	87.54	**90.31**	91.40	61.56	64.96	65.87	34.82	39.46	46.76
Ours	**88.66**	90.17	**92.07**	**62.45**	**65.82**	**66.18**	41.30	45.36	**49.86**

Table 1: Top-1 accuracy under linear evaluation protocal.

	SimCLR	BYOL	MoCo v2	SimSiam	Ours
acc. (%)	66.5	66.5	67.4	68.1	66.97

Table 2: ImageNet linear evaluation accuracy with 100-epoch pre-training. All results but ours are reported from [14]. We use batch size 384 during pre-training.

7 Conclusion

In this paper, we present a novel theoretical framework of self-supervised learning and provide provable guarantees for the learned representations on downstream linear classification. We hope our study can facilitate future theoretical analyses of self-supervised learning and inspire new practical algorithms. For instance, one interesting future direction is to test the topology of the augmentation graph on empirical data distributions and design algorithms using tools from graph theory.

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Checklist

1. For all authors...
   (a) Do the main claims made in the abstract and introduction accurately reflect the paper's contributions and scope? [Yes]
   (b) Did you describe the limitations of your work? [Yes] See Section B.
   (c) Did you discuss any potential negative societal impacts of your work? [Yes] See Section B.
   (d) Have you read the ethics review guidelines and ensured that your paper conforms to them? [Yes]
2. If you are including theoretical results...
   (a) Did you state the full set of assumptions of all theoretical results? [Yes]
   (b) Did you include complete proofs of all theoretical results? [Yes]
3. If you ran experiments...
   (a) Did you include the code, data, and instructions needed to reproduce the main experimental results (either in the supplemental material or as a URL)? [Yes] We will include our code in the supplemental matrial before the supplemental matrial deadline.
   (b) Did you specify all the training details (e.g., data splits, hyperparameters, how they were chosen)? [Yes] See Section 6 and Section A.

(c) Did you report error bars (e.g., with respect to the random seed after running experiments multiple times)? [No] Due to the constraint of time and computing resources, we couldn't run enough number of trials to compute the error bars. For instance, our experiment on ImageNet requires running on 8 GPUs for roughly 50 hours every single trial. We note that many prior works in the computer vision community also typically don't include error bars.

(d) Did you include the total amount of compute and the type of resources used (e.g., type of GPUs, internal cluster, or cloud provider)? [Yes] See Section A

4. If you are using existing assets (e.g., code, data, models) or curating/releasing new assets...

(a) If your work uses existing assets, did you cite the creators? [Yes] We use standard datasets. See Section 6.

(b) Did you mention the license of the assets? [N/A]

(c) Did you include any new assets either in the supplemental material or as a URL? [N/A]

(d) Did you discuss whether and how consent was obtained from people whose data you're using/curating? [N/A]

(e) Did you discuss whether the data you are using/curating contains personally identifiable information or offensive content? [N/A]

5. If you used crowdsourcing or conducted research with human subjects...

(a) Did you include the full text of instructions given to participants and screenshots, if applicable? [N/A]

(b) Did you describe any potential participant risks, with links to Institutional Review Board (IRB) approvals, if applicable? [N/A]

(c) Did you include the estimated hourly wage paid to participants and the total amount spent on participant compensation? [N/A]

A Experiment details

The pseudo-code for our empirical algorithm is summarized in Algorithm 1.

Algorithm 1 Spectral Contrastive Learning

Require: batch size N, structure of encoder network f

for sampled minibatch $\{\bar{x}_i\}_{i=1}^N$ **do**
  for $i \in \{1, \cdots, N\}$ **do**
    draw two augmentations $x_i = \text{aug}(\bar{x}_i)$ and $x_i' = \text{aug}(\bar{x}_i)$.
    compute $z_i = f(x_i)$ and $z_i' = f(x_i')$.
  compute loss $\mathcal{L} = -\frac{2}{N}\sum_{i=1}^N z_i^\top z_i' + \frac{1}{N(N-1)}\sum_{i\neq j}(z_i^\top z_j')^2$
  update f to minimize $\mathcal{L}$
return encoder network $f(\cdot)$

Our results with different hyperparameters on CIFAR-10/100 and Tiny-ImageNet are listed in Table 3.

Datasets	CIFAR-10			CIFAR-100			Tiny-ImageNet		
Epochs	200	400	800	200	400	800	200	400	800
SimCLR (repro.)	83.73	87.72	90.60	54.74	61.05	63.88	**43.30**	**46.46**	48.12
SimSiam (repro.)	87.54	**90.31**	91.40	61.56	64.96	65.87	34.82	39.46	46.76
Ours ($\mu = 1$)	86.47	89.90	**92.07**	59.13	63.83	65.52	28.76	33.94	40.82
Ours ($\mu = 3$)	87.72	90.09	91.84	61.05	64.79	**66.18**	40.06	42.52	**49.86**
Ours ($\mu = 10$)	**88.66**	90.17	91.01	**62.45**	**65.82**	65.16	41.30	45.36	47.84

Table 3: Top-1 accuracy under linear evaluation protocal.

Additional details about the encoder. For the backbone network, we use the CIFAR variant of ResNet18 for CIFAR-10 and CIFAR-100 experiments and use ResNet50 for Tiny-ImageNet and ImageNet experiments. For the projection MLP, we use a 2-layer MLP with hidden and output dimensions 1000 for CIFAR-10, CIFAR100, and Tiny-ImageNet experiments. We use a 3-layer MLP with hidden and output dimension 8192 for ImageNet experiments. We set $\mu = 10$ in the ImageNet experiment, and set $\mu \in \{1, 3, 10\}$ for the CIFAR-10/100 and Tiny-ImageNet experiments.

Training the encoder. We train the neural network using SGD with momentum 0.9. The learning rate starts at 0.05 and decreases to 0 with a cosine schedule. On CIFAR-10/100 and Tiny-ImageNet we use weight decay 0.0005 and train for 800 epochs with batch size 512. On ImageNet we use weight decay 0.0001 and train for 100 epochs with batch size 384. We use 1 GTX 1080 GPU for CIFAR-10/100 and Tiny-ImageNet experiments, and use 8 GTX 1080 GPUs for ImageNet experiments.

Linear evaluation protocol. We train the linear head using SGD with batch size 256 and weight decay 0 for 100 epochs, learning rate starts at 30.0 and is decayed by 10x at the 60th and 80th epochs.

Image transformation details. We use the same augmentation strategy as described in [14].

B Limitations and potential negative social impacts

Limitations. This paper provides statistical guarantees for self-supervised learning. One limitation is that we don't provide guarantees for optimization, and it is unclear whether standard optimization algorithms like SGD can provably reach the global minimum of the spectral contrastive loss. We believe this question is beyond the scope of this work but will be an interesting future direction.

Potential negative social impacts. In the long run, one possible negative impact of our research (and AI research in general) is leading to a large scale of job losses as humans are replaced by machines.

However, we believe that day is still far away given the current limitations of AI algorithms. We hope that our work can lead to more principled AI algorithms, which will overall benefit human society.

C Proofs for Section 3

We first prove a more generalized version of Theorem 3.7 in section C.2, and then prove Theorem 3.7 in Section C.3.

C.1 Proofs of Lemma 3.1 and Lemma 3.2

Proof of Lemma 3.1. Let $D = \mathrm{diag}(s)$ where $s_x > 0$ for $x \in \mathcal{X}$. Let $u_x, \tilde{u}_x \in \mathbb{R}^k$ be the x-th row of matrices F and $\widetilde{F}$, respectively. Recall that $g_{u,B}(x) = \arg\max_{i\in[r]}(u_x^\top B)_i$ is the prediction on an augmented data $x \in \overline{\mathcal{X}}$ with representation u_x and linear classifier B. Let $\widetilde{B} = Q^{-1}B$, it's easy to see that $g_{\tilde{u},\widetilde{B}}(x) = \arg\max_{i\in[r]}(s_x \cdot u_x^\top B)_i$. Notice that $s_x > 0$ doesn't change the prediction since it changes all dimensions of $u_x^\top B$ by the same scale, we have $g_{\tilde{u},\widetilde{B}}(x) = g_{u,B}(x)$ for any augmented data $x \in \mathcal{X}$. The equivalence of loss naturally follows. □

Proof of Lemma 3.2. We can expand $\mathcal{L}_{\mathrm{mf}}(F)$ and obtain

$$\begin{aligned}\mathcal{L}_{\mathrm{mf}}(F) &= \sum_{x,x'\in\mathcal{X}} \left(\frac{w_{xx'}}{\sqrt{w_x w_{x'}}} - u_x^\top u_{x'}\right)^2 \\ &= \sum_{x,x'\in\mathcal{X}} \left(\frac{w_{xx'}^2}{w_x w_{x'}} - 2\cdot w_{xx'}\cdot f(x)^\top f(x) + w_x w_{x'}\cdot \left(f(x)^\top f(x')\right)^2\right) \end{aligned} \tag{7}$$

Notice that the first term is a constant that only depends on the graph but not the variable f. By the definition of augmentation graph, $w_{xx'}$ is the probability of a random positive pair being (x, x') while w_x is the probability of a random augmented data being x. We can hence rewrite the sum of last two terms in Equation (7) as Equation (6). □

C.2 A generalized version of Theorem 3.7

For the proof we will follow the convention in literature [29] and define the *normalized Laplacian matrix* as follows:

Definition C.1. *Let $G = (\mathcal{X}, w)$ be the augmentation graph defined in Section 3.1. The normalized Laplacian matrix of the graph is defined as $L = I - D^{-1/2}AD^{-1/2}$, where A is the adjacency matrix with $A_{xx'} = w_{xx'}$ and D is a diagonal matrix with $D_{xx} = w_x$.*

It is easy to see that $L = I - \overline{A}$ where $\overline{A}$ is the normalized adjacency matrix defined in Section 3.1. Therefore, when λ_i is the i-th largest eigenvalue of L, $1 - \lambda_i$ is the i-th largest eigenvalue of $\overline{A}$.

We call a function defined on augmented data $\hat{y} : \mathcal{X} \to [r]$ an *extended labeling function*. Given an extended labeling function, we define the following quantity that describes the difference between extended labels of two augmented data of the same natural data:

$$\phi^{\hat{y}} := \sum_{x,x'\in\mathcal{X}} w_{xx'} \cdot \mathbb{1}\left[\hat{y}(x) \neq \hat{y}(x')\right]. \tag{8}$$

We also define the following quantity that describes the difference between extended label of an augmentated data and the ground truth label of the corresponding natural data:

$$\Delta(y, \hat{y}) := \Pr_{x\sim\mathcal{P}_{\overline{\mathcal{X}}},\tilde{x}\sim\mathcal{A}(\cdot|x)} \left(\hat{y}(\tilde{x}) \neq y(x)\right). \tag{9}$$

Recall the spectral contrastive loss defined in Section 3.2 is:

$$\mathcal{L}(f) := \mathbb{E}_{\substack{x_1\sim\mathcal{P}_{\overline{\mathcal{X}}},x_2\sim\mathcal{P}_{\overline{\mathcal{X}}},\\ x\sim\mathcal{A}(\cdot|x_1),x^+\sim\mathcal{A}(\cdot|x_1),x'\sim\mathcal{A}(\cdot|x_2)}} \left[-2\cdot f(x)^\top f(x^+) + \left(f(x)^\top f(x')\right)^2\right].$$

We first state a more general version of Theorem 3.7 as follows.

Theorem C.2. *Assume the set of augmented data $\mathcal{X}$ is finite. Let $f^*_{\text{pop}} \in \arg\min_{f:\mathcal{X}\to\mathbb{R}^k} \mathcal{L}(f)$ be a minimizer of the population spectral contrastive loss $\mathcal{L}(f)$ with $k \in \mathcal{Z}^+$. Let $k' \geq r$ such that $k+1 = (1+\zeta)k'$, where $\zeta \in (0,1)$ and $k' \in \mathcal{Z}^+$. Then, there exists a linear probe $B^* \in \mathbb{R}^{r\times k}$ and a universal constant c such that the linear probe predictor satisfies*

$$\mathbb{E}_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}}, x\sim\mathcal{A}(\cdot|\bar{x})}\left[\left\|\vec{y}(\bar{x}) - B^* f^*_{\text{pop}}(x)\right\|_2^2\right] \leq c\cdot\left(\text{poly}(1/\zeta)\cdot\log(k+1)\cdot\frac{\phi^{\hat{y}}}{\rho^2_{k'}} + \Delta(y,\hat{y})\right),$$

where $\vec{y}(\bar{x})$ is the one-hot embedding of $y(\bar{x})$ and $\rho_{k'}$ is the sparsest m-partition defined in Definition 3.4. Furthermore, the error of the linear probe predictor can be bounded by

$$\Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}}, x\sim\mathcal{A}(\cdot|\bar{x})}\left(g_{f^*_{\text{pop}},B^*}(x) \neq y(\bar{x})\right) \leq 2c\cdot\left(\text{poly}(1/\zeta)\cdot\log(k+1)\cdot\frac{\phi^{\hat{y}}}{\rho^2_{k'}} + \Delta(y,\hat{y})\right).$$

Also, if we let λ_i be the i-th smallest eigenvalue of the normalized Laplacian matrix of the graph of the augmented data, we can find a matrix B^* satisfying the above equations with norm bound $\|B^*\|_F \leq 1/(1-\lambda_k)$.

We provide the proof for Theorem C.2 below.

Let $\lambda_1, \lambda_2, \cdots, \lambda_k, \lambda_{k+1}$ be the $k+1$ smallest eigenvalues of the Laplacian matrix L. The following theorem gives a theoretical guarantee similar to Theorem C.2 except for that the bound depends on λ_{k+1}:

Theorem C.3. *Assume the set of augmented data $\mathcal{X}$ is finite. Let $f^*_{\text{pop}} \in \arg\min_{f:\mathcal{X}\to\mathbb{R}^k}$ be a minimizer of the population spectral contrastive loss $\mathcal{L}(f)$ with $k \in \mathcal{Z}^+$. Then, for any labeling function $\hat{y}: \mathcal{X} \to [r]$ there exists a linear probe $B^* \in \mathbb{R}^{r\times k}$ with norm $\|B^*\|_F \leq 1/(1-\lambda_k)$ such that*

$$\mathbb{E}_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}}, x\sim\mathcal{A}(\cdot|\bar{x})}\left[\left\|\vec{y}(\bar{x}) - B^* f^*_{\text{pop}}(x)\right\|_2^2\right] \leq \frac{\phi^{\hat{y}}}{\lambda_{k+1}} + 4\Delta(y,\hat{y}),$$

where $\vec{y}(\bar{x})$ is the one-hot embedding of $y(\bar{x})$. Furthermore, the error can be bounded by

$$\Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}}, x\sim\mathcal{A}(\cdot|\bar{x})}\left(g_{f^*_{\text{pop}},B^*}(x) \neq y(\bar{x})\right) \leq \frac{2\phi^{\hat{y}}}{\lambda_{k+1}} + 8\Delta(y,\hat{y}).$$

We defer the proof of Theorem C.3 to Section C.4.

To get rid of the dependency on λ_{k+1}, we use following higher-order Cheeger's inequality from [32].

Lemma C.4 (Proposition 1.2 in [32])**.** *Let $G = (V, w)$ be a weight graph with $|V| = N$. Then, for any $t \in [N]$ and $\zeta > 0$ such that $(1+\zeta)t \in [N]$, there exists a partition $S_1, S_2, \cdots, S_t$ of V with*

$$\phi_G(S_i) \lesssim \text{poly}(1/\zeta)\sqrt{\lambda_{(1+\zeta)t}\log t},$$

where $\phi_G(\cdot)$ is the Dirichlet conductance defined in Definition 3.3.

Now we prove Theorem C.2 by combining TheoremC.3 and Lemma C.4.

Proof of Theorem C.2. Let $G = (\mathcal{X}, w)$ be the augmentation graph. In Lemma C.4 let $(1+\zeta)t = k+1$ and $t = k'$ we have: there exists partition $S_1, \cdots, S_{k'} \subset \mathcal{X}$ such that $\phi_G(S_i) \lesssim \text{poly}(1/\zeta)\sqrt{\lambda_{k+1}\log(k+1)}$ for $\forall i \in [k']$. By Definition 3.4, we have $\rho_{k'} \leq \max_{i\in[k']}\phi_G(S_i) \lesssim \text{poly}(1/\zeta)\sqrt{\lambda_{k+1}\log(k+1)}$, which leads to $\frac{1}{\lambda_{k+1}} \lesssim \text{poly}(1/\zeta)\cdot log(k+1)\cdot\frac{1}{\rho^2_{k'}}$. Plugging this bound to Theorem C.3 finishes the proof. □

C.3 Proof of Theorem 3.7

We will use the following lemma which gives a connection between $\phi^{\hat{y}}$, $\Delta(y,\hat{y})$ and Assumption 3.5.

Lemma C.5. *Let $G = (\mathcal{X}, w)$ be the augmentation graph, r be the number of underlying classes. Let $S_1, S_2, \cdots, S_r$ be the partition induced by the classifier g in Assumption 3.5. Then, there exists an extended labeling function $\hat{y}$ such that*

$$\Delta(y, \hat{y}) \leq \alpha$$

and

$$\phi^{\hat{y}} = \sum_{x,x' \in \mathcal{X}} w_{xx'} \cdot \mathbb{1}\left[\hat{y}(x) \neq \hat{y}(x')\right] \leq 2\alpha.$$

Proof of Lemma C.5. We define function $\hat{y} : \mathcal{X} \to [r]$ as follows: for an augmented data $x \in \mathcal{X}$, we use function $\hat{y}(x)$ to represent the index of set that x is in, i.e., $x \in S_{\hat{y}(x)}$. By Assumption 3.5 it is easy to see $\Delta(y, \hat{y}) \leq \alpha$. On the other hand, we have

$$\begin{aligned}
\phi^{\hat{y}} =& \sum_{x,x' \in \mathcal{X}} w_{xx'} \mathbb{1}\left[\hat{y}(x) \neq \hat{y}(x')\right] \\
=& \sum_{x,x' \in \mathcal{X}} \mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}}\left[\mathcal{A}(x|\bar{x})\mathcal{A}(x'|\bar{x}) \cdot \mathbb{1}\left[\hat{y}(x) \neq \hat{y}(x')\right]\right] \\
\leq& \sum_{x,x' \in \mathcal{X}} \mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}}\left[\mathcal{A}(x|\bar{x})\mathcal{A}(x'|\bar{x}) \cdot \left(\mathbb{1}\left[\hat{y}(x) \neq y(\bar{x})\right] + \mathbb{1}\left[\hat{y}(x') \neq y(\bar{x})\right]\right)\right] \\
=& 2 \cdot \mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}}\left[\mathcal{A}(x|\bar{x}) \cdot \mathbb{1}\left[\hat{y}(x) \neq y(\bar{x})\right]\right] \\
=& 2 \cdot \Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})}\left(x \notin S_{y(\bar{x})}\right) = 2\alpha.
\end{aligned}$$

Here the inequality is because when $\hat{y}(x) \neq \hat{y}(x')$, there must be $\hat{y}(x) \neq y(\bar{x})$ or $\hat{y}(x') \neq y(\bar{x})$. □

Now we give the proof of Theorem 3.7 using Lemma C.5 and Theorem C.2.

Proof of Theorem 3.7. Let $S_1, S_2, \cdots, S_r$ be the partition of $\mathcal{X}$ induced by the classifier g given in Assumption 3.5. Define function $\hat{y} : \mathcal{X} \to [r]$ as follows: for an augmented data $x \in \mathcal{X}$, we use function $\hat{y}(x)$ to represent the index of set that x is in, i.e., $x \in S_{\hat{y}(x)}$. Let $k' = \lfloor \frac{k}{2} \rfloor$ in Theorem C.2, we have $\Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})}\left(g_{f^*_{\text{pop}}, B^*}(x) \neq y(\bar{x})\right) \lesssim \log(k) \cdot \frac{\phi^{\hat{y}}}{\rho^2_{\lfloor k/2 \rfloor}} + \Delta(y, \hat{y})$. By Lemma C.5 we have $\phi^{\hat{y}} \leq 2\alpha$ and $\Delta(y, \hat{y}) \leq \alpha$, so we have $\Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})}\left(g_{f^*_{\text{pop}}, B^*}(x) \neq y(\bar{x})\right) \lesssim \frac{\alpha}{\rho^2_{\lfloor k/2 \rfloor}} \cdot \log(k)$. Notice that by definition of ensembled linear probe predictor, $\bar{g}_{f^*_{\text{pop}}, B^*}(\bar{x}) \neq y(\bar{x})$ happens only if more than half of the augmentations of $\bar{x}$ predicts differently from $y(\bar{x})$, so we have $\Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}}\left(\bar{g}_{f^*_{\text{pop}}, B^*}(\bar{x}) \neq y(\bar{x})\right) \leq 2 \Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})}\left(g_{f^*_{\text{pop}}, B^*}(x) \neq y(\bar{x})\right) \lesssim \frac{\alpha}{\rho^2_{\lfloor k/2 \rfloor}} \cdot \log(k)$. □

C.4 Proof of Theorem C.3

The proof of Theorem C.3 contains two steps. First, we show that when the feature extractor is composed of the minimal eigenvectors of the normalized Laplacian matrix L, we can achieve good linear probe accuracy. Then we show that minimizing $\mathcal{L}(f)$ gives us a feature extractor equally good as the eigenvectors.

For the first step, we use the following lemma which shows that the smallest eigenvectors of L can approximate any function on $\mathcal{X}$ up to an error proportional to the Rayleigh quotient of the function.

Lemma C.6. *Let L be the normalized Laplacian matrix of some graph G. Let $N = |\mathcal{X}|$ be total number of augmented data, v_i be the i-th smallest unit-norm eigenvector of L with eigenvalue λ_i (make them orthogonal in case of repeated eignevalues). Let $R(u) := \frac{u^\top L u}{u^\top u}$ be the Rayleigh quotient of a vector $u \in \mathbb{R}^N$. Then, for any $k \in \mathcal{Z}^+$ such that $k < N$ and $\lambda_{k+1} > 0$, there exists a vector $b \in \mathbb{R}^k$ with norm $\|b\|_2 \leq \|u\|_2$ such that*

$$\left\| u - \sum_{i=1}^{k} b_i v_i \right\|_2^2 \leq \frac{R(u)}{\lambda_{k+1}} \|u\|_2^2 .$$

Proof of Lemma C.6. We can decompose the vector u in the eigenvector basis as:

$$u = \sum_{i=1}^{N} \zeta_i v_i.$$

We have

$$R(u) = \frac{\sum_{i=1}^{N} \lambda_i \zeta_i^2}{\|u\|_2^2}.$$

Let $b \in \mathbb{R}^k$ be the vector such that $b_i = \zeta_i$. Obviously we have $\|b\|_2^2 \le \|u\|_2^2$. Noticing that

$$\left\| u - \sum_{i=1}^{k} b_i v_i \right\|_2^2 = \sum_{i=k+1}^{N} \zeta_i^2 \le \frac{R(u)}{\lambda_{k+1}} \|u\|_2^2,$$

which finishes the proof. □

We also need the following claim about the Rayleigh quotient $R(u)$ when u is a vector defined by an extended labeling function $\hat{y}$.

Claim C.7. *In the setting of Lemma C.6, let $\hat{y}$ be an extended labeling function. Fix $i \in [r]$. Define function $u_i^{\hat{y}}(x) := \sqrt{w_x} \cdot \mathbb{1}[\hat{y}(x) = i]$ and $u_i^{\hat{y}}$ is the corresponding vector in $\mathbb{R}^N$. Also define the following quantity:*

$$\phi_i^{\hat{y}} := \frac{\sum_{x,x' \in \mathcal{X}} w_{xx'} \cdot \mathbb{1}\left[(\hat{y}(x) = i \wedge \hat{y}(x') \neq i) \text{ or } (\hat{y}(x) \neq i \wedge \hat{y}(x') = i)\right]}{\sum_{x \in \mathcal{X}} w_x \cdot \mathbb{1}[\hat{y}(x) = i]}.$$

Then, we have

$$R(u_i^{\hat{y}}) = \frac{1}{2}\phi_i^{\hat{y}}.$$

Proof of Claim C.7. Let f be any function $\mathcal{X} \to \mathbb{R}$, define function $u(x) := \sqrt{w_x} \cdot f(x)$. Let $u \in \mathbb{R}^N$ be the vector corresponding to u. Let A be the adjacency matrix with $A_{xx'} = w_{xx'}$ and D be the diagonal matrix with $D_{xx} = w_x$. By definition of Laplacian matrix, we have

$$\begin{aligned} u^\top L u &= \|u\|_2^2 - u^\top D^{-1/2} A D^{-1/2} u \\ &= \sum_{x \in \mathcal{X}} w_x f(x)^2 - \sum_{x,x' \in \mathcal{X}} w_{xx'} f(x) f(x') \\ &= \frac{1}{2} \sum_{x,x' \in \mathcal{X}} w_{xx'} \cdot (f(x) - f(x'))^2. \end{aligned}$$

Therefore we have

$$\begin{aligned} R(u) &= \frac{u^\top L u}{u^\top u} \\ &= \frac{1}{2} \cdot \frac{\sum_{x,x' \in \mathcal{X}} w_{xx'} \cdot (f(x) - f(x'))^2}{\sum_{x \in \mathcal{X}} w_x \cdot f(x)^2}. \end{aligned}$$

Setting $f(x) = \mathbb{1}[\hat{y}(x) = i]$ finishes the proof. □

To see the connection between the feature extractor minimizing the population spectral contrastive loss $\mathcal{L}(f)$ and the feature extractor corresponding to eigenvectors of the Laplacian matrix, we use the following lemma which states that the minimizer of the matrix approximation loss defined in Section 3.2 is equivalent to the minimizer of population spectral contrastive loss up to a data-wise scaling.

Lemma C.8. *Let $\hat{F}_{\text{mf}}$ be the matrix form of a feature extractor $\hat{f}_{\text{mf}} : \mathcal{X} \to \mathbb{R}^k$. Then, $\hat{F}_{\text{mf}}$ is a minimizer of $\mathcal{L}_{\text{mf}}(F)$ if and only if*

$$\hat{f}(x) := \frac{1}{\sqrt{w_x}} \cdot \hat{f}_{\text{mf}}(x)$$

is a minimizer of the population spectral contrastive loss $\mathcal{L}(f)$.

Proof of Lemma C.8. Notice that

$$\begin{aligned}\mathcal{L}_{\text{mf}}(F) &= \left\|(I-L)-FF^\top\right\|_F^2 \\ &= \sum_{x,x'\in\mathcal{X}}\left(\frac{w_{xx'}}{\sqrt{w_x w_{x'}}}-f(x)^\top f(x')\right)^2 \\ &= \sum_{x,x'\in\mathcal{X}}\left(f(x)^\top f(x')\right)^2 - 2\sum_{x,x'\in\mathcal{X}}\frac{w_{xx'}}{\sqrt{w_x w_{x'}}}f(x)^\top f(x') + \|I-L\|_F^2 . \end{aligned} \tag{10}$$

Recall that the definition of spectral contrastive loss is

$$\mathcal{L}(f) := -2\cdot\mathbb{E}_{x,x^+}\left[f(x)^\top f(x^+)\right] + \mathbb{E}_{x,x'}\left[\left(f(x)^\top f(x')\right)^2\right],$$

where (x, x^+) is a random positive pair, (x, x') is a random negative pair. We can rewrite the spectral contrastive loss as

$$\begin{aligned}\mathcal{L}(f) &= -2\sum_{x,x'\in\mathcal{X}} w_{xx'}\cdot f(x)^\top f(x') + \sum_{x,x'\in\mathcal{X}} w_x w_{x'}\cdot\left(f(x)^\top f(x')\right)^2 \\ &= \sum_{x,x'\in\mathcal{X}}\left(\left(\frac{f(x)}{\sqrt{w_x}}\right)^\top\left(\frac{f(x')}{\sqrt{w_{x'}}}\right)\right)^2 - 2\sum_{x,x'\in\mathcal{X}}\frac{w_{xx'}}{\sqrt{w_x w_{x'}}}\left(\frac{f(x)}{\sqrt{w_x}}\right)^\top\left(\frac{f(x')}{\sqrt{w_{x'}}}\right). \end{aligned} \tag{11}$$

Compare Equation (10) and Equation (11), we see their minimizers only differ by a term $\sqrt{w_x}$, which finishes the proof. □

Note that the minimizer of matrix approximation loss is exactly the largest eigenvectors of $I - L$ (also the smallest eigenvectors of L) due to Eckart–Young–Mirsky theorem, Lemma C.8 indicates that the minimizer of $\mathcal{L}(f)$ is equivalent to the smallest eigenvectors of L up to data-wise scaling. Now we are ready to prove Theorem C.3 by combining Lemma C.6, Claim C.7 and Lemma C.8.

Proof of Theorem C.3. Let $F_{\text{sc}} = [v_1, v_2, \cdots, v_k]$ be the matrix that contains the smallest k eigenvectors of L as columns, and $f_{\text{sc}} : \mathcal{X} \to \mathbb{R}^k$ is the corresponding feature extractor. For each $i \in [r]$, we define function $u_i^{\hat{y}}(x) := \sqrt{w_x}\cdot\mathbb{1}\left[\hat{y}(x) = i\right]$ and $u_i^{\hat{y}}$ be the corresponding vector in $\mathbb{R}^N$. By Lemma C.6, there exists a vector $b_i \in \mathbb{R}^k$ with norm bound $\|b_i\|_2 \le \left\|u_i^{\hat{y}}\right\|_2$ such that

$$\left\|u_i^{\hat{y}} - F_{\text{sc}} b_i\right\|_2^2 \le \frac{R(u_i^{\hat{y}})}{\lambda_{k+1}}\left\|u_i^{\hat{y}}\right\|_2^2 . \tag{12}$$

By Claim C.7, we have

$$R(u_i^{\hat{y}}) = \frac{1}{2}\phi_i^{\hat{y}} = \frac{1}{2}\cdot\frac{\sum_{x,x'\in\mathcal{X}} w_{xx'}\cdot\mathbb{1}\left[(\hat{y}(x) = i \wedge \hat{y}(x') \neq i) \text{ or } (\hat{y}(x) \neq i \wedge \hat{y}(x') = i)\right]}{\sum_{x\in\mathcal{X}} w_x\cdot\mathbb{1}\left[\hat{y}(x) = i\right]}.$$

So we can rewrite Equation (12) as:

$$\begin{aligned}\left\|u_i^{\hat{y}} - F_{\text{sc}} b_i\right\|_2^2 &\le \frac{\phi_i^{\hat{y}}}{2\lambda_{k+1}}\cdot\sum_{x\in\mathcal{X}} w_x\cdot\mathbb{1}\left[\hat{y}(x) = i\right] \\ &= \frac{1}{2\lambda_{k+1}}\sum_{x,x'\in\mathcal{X}} w_{xx'}\cdot\mathbb{1}\left[(\hat{y}(x) = i \wedge \hat{y}(x') \neq i) \text{ or } (\hat{y}(x) \neq i \wedge \hat{y}(x') = i)\right]. \end{aligned} \tag{13}$$

Let matrix $U = [u_1^{\hat{y}}, \cdots, u_r^{\hat{y}}]$ contains all $u_i^{\hat{y}}$ as columns, and let $u : \mathcal{X} \to \mathbb{R}^r$ be the corresponding feature extractor. Define matrix $B \in \mathbb{R}^{N\times k}$ such that $B^\top = [b_1, \cdots, b_r]$. Summing Equation (13) over all $i \in [r]$ and by the definition of $\phi^{\hat{y}}$ we have

$$\left\|U - F_{\text{sc}} B^\top\right\|_F^2 \le \frac{1}{2\lambda_{k+1}}\sum_{x,x'\in\mathcal{X}} w_{xx'}\cdot\mathbb{1}\left[\hat{y}(x) \neq \hat{y}(x')\right] = \frac{\phi^{\hat{y}}}{2\lambda_{k+1}}, \tag{14}$$

where

$$\|B\|_F^2 = \sum_{i=1}^{r} \|b_i\|_2^2 \le \sum_{i=1}^{r} \left\| u_i^{\hat{y}} \right\|_2^2 = \sum_{x \in \mathcal{X}} w_x = 1.$$

Now we come back to the feature extractor f_{pop}^* that minimizes the spectral contrastive loss function $\mathcal{L}(f)$. By Lemma C.8, function $f_{\text{mf}}^*(x) := \sqrt{w_x} \cdot f_{\text{pop}}^*(x)$ is a minimizer of matrix approximation loss. Let F_{mf}^* be the corresponding matrix. By Eckard-Young-Mirsky theorem, we have

$$F_{\text{mf}}^* = F_{\text{sc}} D_\lambda Q,$$

where Q is an orthonomal matrix and

$$D_\lambda = \begin{bmatrix} \sqrt{1-\lambda_1} & & & \\ & \sqrt{1-\lambda_2} & & \\ & & \cdots & \\ & & & \sqrt{1-\lambda_k} \end{bmatrix}.$$

Let

$$B^* = B D_\lambda^{-1} Q^{-1},$$

and let $\bar{y}(\bar{x})$ be the one-hot embedding of $y(\bar{x})$, $\vec{\hat{y}}(x)$ be the one-hot embedding of $\hat{y}(x)$, we have

$$\begin{aligned}
&\mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left[\left\| \vec{y}(\bar{x}) - B^* f_{\text{pop}}^*(x) \right\|_2^2 \right] \\
\le & 2\mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left[\left\| \vec{\hat{y}}(x) - B^* f_{\text{pop}}^*(x) \right\|_2^2 \right] + 2\mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left[\left\| \vec{\hat{y}}(x) - \vec{y}(\bar{x}) \right\|_2^2 \right] \\
= & 2 \sum_{x \in \mathcal{X}} w_x \cdot \left\| \vec{\hat{y}}(x) - B^* f_{\text{pop}}^*(x) \right\|_2^2 + 4\Delta(y, \hat{y}) && \text{(because } w_x \text{ is the probability of } x\text{)} \\
= & 2 \sum_{x \in \mathcal{X}} \|u(x) - B^* f_{\text{mf}}^*(x)\|_2^2 + 4\Delta(y, \hat{y}) && \text{(because } f_{\text{mf}}^*(x) = \sqrt{w_x} \cdot f_{\text{pop}}^*(x)\text{)} \\
= & 2 \left\| U - F_{\text{mf}}^* B^{*\top} \right\|_F^2 + 4\Delta(y, \hat{y}) && \text{(rewrite in matrix form)} \\
= & 2 \left\| U - F_{\text{sc}} B^\top \right\|_F^2 + 4\Delta(y, \hat{y}) && \text{(by definition of } B^*\text{)} \\
\le & \frac{\phi^{\hat{y}}}{\lambda_{k+1}} + 4\Delta(y, \hat{y}). && \text{(by Equation (14))}
\end{aligned}$$

To bound the error rate, we first notice that $g_{F_{\text{mf}}^*, B^*}(x) \neq y(\bar{x})$ happens only if $\|\vec{y}(\bar{x}) - B^* f_{\text{sc}}(x)\|_2^2 \ge \frac{1}{2}$, we have for any $x \in \mathcal{X}$,

$$\left\| \vec{y}(\bar{x}) - B^* \hat{f}_{\text{ma}}(x) \right\|_2^2 \ge \frac{1}{2} \cdot \mathbb{1}\left[g_{F_{\text{mf}}^*, B^*}(x) \neq y(\bar{x}) \right]. \tag{15}$$

Now we bound the error rate on $\mathcal{X}$ as follows:

$$\begin{aligned}
&\Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left(g_{f_{\text{pop}}^*, B^*}(x) \neq y(\bar{x}) \right) \\
\le & 2\mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left[\left\| \vec{y}(\bar{x}) - B^* f_{\text{pop}}^*(x) \right\|_2^2 \right] && \text{(by Equation (15))} \\
\le & \frac{2\phi^{\hat{y}}}{\lambda_{k+1}} + 8\Delta(y, \hat{y}).
\end{aligned}$$

Finally we bound the norm of B^* as

$$\|B^*\|_F^2 = Tr\left(B^* B^{*\top} \right) = Tr\left(B D_\lambda^{-2} B^\top \right) \le \frac{1}{1-\lambda_k} \|B\|_F^2 = \frac{1}{1-\lambda_k}.$$

□

D Proofs for Section 3.4

In this section, we give a proof of Theorem 3.9.

The following lemma shows that the augmented graph for Example 3.8 satisfies Assumption 3.5 with some bounded α.

Lemma D.1. *In the setting of Theorem 3.9, the data distribution satisfies Assumption 3.5 with* $\alpha \le \frac{1}{\text{poly}(d')}$.

Proof of Lemma D.1. For any $z \sim \mathcal{N}(\mu_i, \frac{1}{d'} \cdot I_{d' \times d'})$ and any $j \neq i$, by the tail bound of gaussian distribution we have

$$\Pr_{z \sim \mathcal{N}(\mu_i, \frac{1}{d'} \cdot I_{d' \times d'})} \left((z - \mu_i)^\top \left(\frac{\mu_j - \mu_i}{\|\mu_j - \mu_i\|_2} \right) \lesssim \frac{\sqrt{\log d}}{\sqrt{d'}} \right) \ge 1 - \frac{1}{\text{poly}(d)}.$$

Also, for $\xi \sim \mathcal{N}(0, \frac{1}{d} \cdot I_{d \times d})$, when $\sigma \le \frac{1}{\sqrt{d}}$ we have

$$\Pr_{\xi \sim \mathcal{N}(0, \frac{\sigma^2}{d} \cdot I_{d' \times d'})} \left(\|\xi\|_2 \lesssim \frac{\sqrt{\log d}}{\sqrt{d}} \right) \ge 1 - \frac{1}{\text{poly}(d)}.$$

Notice that $\left\|Q^{-1}(Q(z) + \xi) - z\right\|_2 \le \kappa \|\xi\|$, we can set $\|\mu_i - \mu_j\| \gtrsim \kappa \frac{\sqrt{\log d}}{\sqrt{d}}$. Therefore, when $\|\mu_i - \mu_j\| \gtrsim \kappa \frac{\sqrt{\log d}}{\sqrt{d'}}$ we can combine the above two cases and have

$$\Pr_{z \sim \mathcal{N}(\mu_i, \frac{1}{d'} \cdot I_{d' \times d'}), \xi \sim \mathcal{N}(0, \frac{\sigma^2}{d} \cdot I_{d' \times d'})} \left(P_i(z) > P_j(Q^{-1}(Q(z) + \xi)) \right) \ge 1 - \frac{1}{\text{poly}(d)}.$$

Since $r \le d$, we have

$$\Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} (y(x) \neq y(\bar{x})) \ge 1 - \frac{1}{\text{poly}(d)}.$$

□

We use the following lemma to give a lower bound for the sparest m-partition of the augmentation graph in Example 3.8.

Lemma D.2. *In the setting of Theorem 3.9, for any $k' > r$ and $\tau > 0$, we have*

$$\rho_{k'} \ge \frac{c_{\tau/\kappa}}{18} \cdot \exp\left(-\frac{2c_\sigma \tau + \tau^2}{2\sigma^2/d} \right),$$

where

$$c_\sigma := \sigma \cdot \Phi_d^{-1}(\frac{2}{3})$$

with $\Phi_d(z) := \Pr_{\xi \sim \mathcal{N}(0, \frac{1}{d} I_{d \times d})}(\|\xi\|_2 \le z)$, and

$$c_{\tau/\kappa} := \min_{p \in [0, \frac{3}{4}]} \frac{\Phi(\Phi^{-1}(p) + \tau\sqrt{d}/\kappa)}{p} - 1$$

with $\Phi(z) := \int_{-\infty}^{z} \frac{e^{-u^2/2}}{\sqrt{2\pi}} du$.

The proof of Lemma D.2 can be found in Section D.1. Now we give the proof of Example 3.9.

Proof of Theorem 3.9. The result on α is directly from Lemma D.1. By concentration inequality, there must exists some universal constant $C > 0$ such that for any $d \ge C$, we have $1 - \Phi_d(\sqrt{\frac{3}{2}}) \le \frac{1}{3}$. When this happens, we have $\Phi_d^{-1}(\frac{2}{3}) \le \sqrt{\frac{3}{2}}$. Since for $d \le C$ we can just treat d as constant, we have $\Phi_d^{-1}(\frac{2}{3}) \lesssim 1$. Set $\tau = \sigma/d$ in Lemma D.2, we have $\rho_{k'} \gtrsim \frac{\sigma}{\kappa\sqrt{d}}$. Set $k' = \lfloor k/2 \rfloor$, we apply Theorem 3.7 and get the bound we need. □

D.1 Proof of Lemma D.2

In this section we give a proof for Lemma D.2. We first introduce the following claim which states that for a given subset of augmented data, any two data close in L_2 norm cannot have a very different chance of being augmented into this set.

Claim D.3. *In the setting of Theorem 3.9, given a set $S \subseteq \mathbb{R}^d$. If $x \in \mathbb{R}^d$ satisfies $\Pr(S|x) := \Pr_{\tilde{x}\sim\mathcal{A}(\cdot|x)}(\tilde{x} \in S) \geq \frac{2}{3}$. Then, for any x' such that $\|x - x'\|_2 \leq \tau$, we have*

$$\Pr(S|x') \geq \frac{1}{3} \cdot \exp\left(-\frac{2c_\sigma\tau + \tau^2}{2\sigma^2}\right),$$

where

$$c_\sigma := \sigma \cdot \Phi_d^{-1}(\frac{2}{3}),$$

with $\Phi_d(z) := \Pr_{\xi\sim\mathcal{N}(0,\frac{1}{d}\cdot I_{d\times d})}(\|\xi\|_2 \leq z)$.

Proof of Claim D.3. By the definition of augmentation, we know

$$\Pr(S|x) = \mathbb{E}_{\xi\sim\mathcal{N}(0,\frac{\sigma^2}{d}\cdot I_{d\times d})}\left[\mathbb{1}\left[x + \xi \in S\right]\right].$$

By the definition of c_σ, we have

$$\Pr_{\xi\sim\mathcal{N}(0,\frac{\sigma^2}{d}\cdot I_{d\times d})}(\|\xi\|_2 \leq c_\sigma) = \frac{2}{3}.$$

Since $\Pr(S|x) \geq \frac{2}{3}$ by assumption, we have

$$\mathbb{E}_{\xi\sim\mathcal{N}(0,\frac{\sigma^2}{d}\cdot I_{d\times d})}\left[P(S|x+\xi) \cdot \mathbb{1}\left[\|\xi\|_2 \leq c_\sigma\right]\right] \geq \frac{1}{3}.$$

Now we can bound the quanity of our interest:

$$\begin{aligned}
\Pr(S|x') &= \frac{1}{(2\pi\sigma^2/d)^{d/2}} \int_\xi e^{-\frac{\|\xi\|_2^2}{2\sigma^2/d}} P(S|x'+\xi)d\xi \\
&= \frac{1}{(2\pi\sigma^2/d)^{d/2}} \int_\xi e^{-\frac{\|\xi+x-x'\|_2^2}{2\sigma^2/d}} P(S|x+\xi)d\xi \\
&\geq \frac{1}{(2\pi\sigma^2/d)^{d/2}} \int_\xi e^{-\frac{\|\xi+x-x'\|_2^2}{2\sigma^2/d}} P(S|x+\xi) \cdot \mathbb{1}\left[\|\xi\|_2 \leq c_\sigma\right] d\xi \\
&\geq \frac{1}{(2\pi\sigma^2/d)^{d/2}} \int_\xi e^{-\frac{2c_\sigma\tau+\tau^2+\|\xi\|_2^2}{2\sigma^2/d}} P(S|x+\xi) \cdot \mathbb{1}\left[\|\xi\|_2 \leq c_\sigma\right] d\xi \\
&= e^{-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}} \cdot \mathbb{E}_{\xi\sim\mathcal{N}(0,\frac{\sigma^2}{d}I_{d\times d})}\left[P(S|x+\xi) \cdot \mathbb{1}\left[\|\xi\|_2 \leq c_\sigma\right]\right] \\
&\geq \frac{1}{3} \cdot \exp\left(-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}\right).
\end{aligned}$$

□

We now give the proof of Lemma D.2.

Proof of Lemma D.2. Let $S_1, \cdots, S_{k'}$ be the disjoint sets that gives $\rho_{k'}$ in Definition 3.4. First we notice that when $k' > r$, there must exist $t \in [k']$ such that for all $i \in [r]$, we have

$$\Pr_{x\sim P_i, \tilde{x}\sim\mathcal{A}(\cdot|x)}(\tilde{x} \in S_t) \leq \frac{1}{2}. \tag{16}$$

WLOG, we assume $t = 1$. So we know that

$$\rho_{k'} = \max_{i\in[k']} \phi_G(S_i) \geq \phi_G(S_1) \geq \min_{j\in[r]} \frac{\mathbb{E}_{x\sim P_j}\left[\Pr(S_1|x)(1 - \Pr(S_1|x))\right]}{\mathbb{E}_{x\sim P_j}\left[\Pr(S_1|x)\right]}, \tag{17}$$

where

$$\Pr(S|x) := \Pr_{\tilde{x}\sim\mathcal{A}(\cdot|x)}(\tilde{x} \in S).$$

WLOG, we assume $j = 1$ minimizes the RHS of Equation (17), so we only need to prove

$$\frac{\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)(1-\Pr(S_1|x))\right]}{\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)\right]} \geq \frac{c_{\tau/\kappa}}{18} \cdot \exp\left(-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}\right).$$

We define the following set

$$R := \left\{x \middle| \Pr(S_1|x) \geq \frac{2}{3}\right\}.$$

Notice that

$$\begin{aligned}\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)\right] &= \int_x P_1(x)\Pr(S_1|x)dx \\ &= \int_{x\in R} P_1(x)\Pr(S_1|x)dx + \int_{x\notin R} P_1(x)\Pr(S_1|x)dx. \end{aligned} \tag{18}$$

We can consider the following two cases.

Case 1: $\int_{x\notin R} P_1(x)\Pr(S_1|x)dx \geq \frac{1}{2}\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)\right]$.

This is the easy case because we have

$$\begin{aligned}\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)(1-\Pr(S_1|x))\right] &\geq \int_{x\notin R} P_1(x)\Pr(S_1|x)(1-\Pr(S_1|x))dx \\ &\geq \frac{1}{3}\int_{x\notin R} P_1(x)\Pr(S_1|x)dx \\ &\geq \frac{1}{6}\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)\right].\end{aligned}$$

Case 2: $\int_{x\in R} P_1(x)\Pr(S_1|x)dx \geq \frac{1}{2}\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)\right]$.

Define neighbourhood of R as

$$N(R) := \left\{x \middle| \|x-a\|_2 \leq \tau \text{ for some } a \in R\right\}.$$

We have

$$\begin{aligned}\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)(1-\Pr(S_1|x))\right] &\geq \int_{x\in N(R)\setminus R} P_1(x)\Pr(S_1|x)(1-\Pr(S_1|x))dx \\ &\geq \frac{1}{9}\cdot\exp\left(-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}\right)\cdot\int_{x\in N(R)\setminus R} P_1(x)dx,\end{aligned}$$

where the second inequality is by Claim D.3. Notice that

$$\int_{x\in R} P_1(x)dx \leq \frac{3}{2}\int_{x\in R} P_1(x)\Pr(S_1|x)dx \leq \frac{3}{2}\int_x P_1(x)\Pr(S_1|x)dx \leq \frac{3}{4},$$

where we use Equation (16). Define set $\widetilde{R} := Q^{-1}(R)$ be the set in the ambient space corresponding to R. Define

$$\widetilde{N}(\widetilde{R}) := \left\{x' \in \mathbb{R}^{d'} \middle| \|x'-a\|_2 \leq \frac{\tau}{\kappa} \text{ for some } a \in \widetilde{R}\right\}$$

Due to Q being κ-bi-lipschitz, it is easy to see $\widetilde{N}(\widetilde{R}) \subseteq Q^{-1}\left(N(R)\right)$. According to the Gaussian isoperimetric inequality [7], we have

$$\int_{x\in N(R)\setminus R} P_1(x)dx \geq c_{\tau/\kappa}\int_{x\in R} P_1(x)dx,$$

where

$$c_{\tau/\kappa} := \min_{0\le p\le 3/4} \frac{\Phi(\Phi^{-1}(p)+\tau\sqrt{d}/\kappa)}{p} - 1,$$

with $\Phi(\cdot)$ is the Gaussian CDF function defined as

$$\Phi(z) := \int_{-\infty}^{z} \frac{e^{-u^2/2}}{\sqrt{2\pi}} du.$$

So we have

$$\begin{aligned}
\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)(1-\Pr(S_1|x))\right] &\ge \frac{c_{\tau/\kappa}}{9}\cdot\exp\left(-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}\right)\cdot\int_{x\in R} P_1(x)dx \\
&\ge \frac{c_{\tau/\kappa}}{9}\cdot\exp\left(-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}\right)\cdot\int_{x\in R} P_1(x)\Pr(S_1|x)dx \\
&\ge \frac{c_{\tau/\kappa}}{18}\cdot\exp\left(-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}\right)\cdot\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)\right].
\end{aligned}$$

By Equation (18), either case 1 or case 2 holds. Combining case 1 and case 2, we have

$$\begin{aligned}
\frac{\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)(1-\Pr(S_1|x))\right]}{\mathbb{E}_{x\sim P_1}\left[\Pr(S_1|x)\right]} &\ge \min\left\{\frac{1}{6}, \frac{c_{\tau/\kappa}}{18}\cdot\exp\left(-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}\right)\right\} \\
&= \frac{c_{\tau/\kappa}}{18}\cdot\exp\left(-\frac{2c_\sigma\tau+\tau^2}{2\sigma^2/d}\right).
\end{aligned}$$

□

E Proofs for Section 4

E.1 Proof of Theorem 4.1

We restate the empirical spectral contrastive loss defined in Section 4 as follows:

Definition E.1 (Empirical spectral contrastive loss)**.** *Consider a dataset $\widehat{\mathcal{X}} = \{\bar{x}_1, \bar{x}_2, \cdots, \bar{x}_n\}$ containing n data points i.i.d. sampled from $\mathcal{P}_{\overline{\mathcal{X}}}$. Let $\hat{\mathcal{P}}_{\mathcal{X}}$ be the uniform distribution over $\widehat{\mathcal{X}}$. Let $\hat{P}_{\bar{x},\bar{x}'}$ be the uniform distribution over data pairs $(\bar{x}_i, \bar{x}_j)$ where $i \ne j$. We define the empirical spectral contrastive loss of a feature extractor f as*

$$\widehat{\mathcal{L}}_f(:) = -2\mathbb{E}_{\substack{\bar{x}\sim\hat{\mathcal{P}}_{\mathcal{X}},\\ x\sim\mathcal{A}(\cdot|\bar{x}),x'\sim\mathcal{A}(\cdot|\bar{x})}}\left[f(x)^\top f(x')\right] + \mathbb{E}_{\substack{(\bar{x},\bar{x}')\sim\hat{P}_{\bar{x},\bar{x}'},\\ x\sim\mathcal{A}(\cdot|\bar{x}),x'\sim\mathcal{A}(\cdot|\bar{x}')}}\left[\left(f(x)^\top f(x')\right)^2\right].$$

The following claim shows that $\widehat{\mathcal{L}}_n(f)$ is an unbiased estimator of population spectral contrastive loss.

Claim E.2. *$\widehat{\mathcal{L}}_n(f)$ is an unbiased estimator of $\mathcal{L}(f)$, i.e.,*

$$\mathbb{E}_{\widehat{\mathcal{X}}}\left[\widehat{\mathcal{L}}_n(f)\right] = \mathcal{L}(f).$$

Proof. This is because

$$\begin{aligned}
\mathbb{E}_{\widehat{\mathcal{X}}}\left[\widehat{\mathcal{L}}_n(f)\right] &= -2\cdot\mathbb{E}_{\widehat{\mathcal{X}}}\left[\mathbb{E}_{\substack{\bar{x}\sim\hat{\mathcal{P}}_{\mathcal{X}},\\ x\sim\mathcal{A}(\cdot|\bar{x}),x'\sim\mathcal{A}(\cdot|\bar{x})}}\left[f(x)^\top f(x')\right]\right] + \mathbb{E}_{\widehat{\mathcal{X}}}\left[\mathbb{E}_{\substack{(\bar{x},\bar{x}')\sim\hat{P}_{\bar{x},\bar{x}'},\\ x\sim\mathcal{A}(\cdot|\bar{x}),x'\sim\mathcal{A}(\cdot|\bar{x}')}}\left[\left(f(x)^\top f(x')\right)^2\right]\right] \\
&= -2\mathbb{E}_{\substack{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},\\ x\sim\mathcal{A}(\cdot|\bar{x}),x'\sim\mathcal{A}(\cdot|\bar{x})}}\left[f(x)^\top f(x')\right] + \mathbb{E}_{\substack{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},\bar{x}'\sim\mathcal{P}_{\overline{\mathcal{X}}},\\ x\sim\mathcal{A}(\cdot|\bar{x}),x'\sim\mathcal{A}(\cdot|\bar{x}')}}\left[\left(f(x)^\top f(x')\right)^2\right] = \mathcal{L}(f).
\end{aligned}$$

□

To make use of the Radmacher complexity theory, we need to write the empirical loss as the sum of i.i.d. terms, which is achieved by the following sub-sampling scheme:

Definition E.3. *Given dataset $\widehat{\mathcal{X}}$, we sample a subset of tuples as follows: first sample a permutation $\pi : [n] \to [n]$, then we sample tuples $S = \{(z_i, z_i^+, z_i')\}_{i=1}^{n/2}$ as follows:*

$$\begin{aligned} z_i &\sim \mathcal{A}(\cdot|\bar{x}_{\pi(2i-1)}), \\ z_i^+ &\sim \mathcal{A}(\cdot|\bar{x}_{\pi(2i-1)}), \\ z_i' &\sim \mathcal{A}(\cdot|\bar{x}_{\pi(2i)}). \end{aligned}$$

We define the following loss on S:

$$\widehat{\mathcal{L}}_S(f) := \frac{1}{n/2} \sum_{i=1}^{n/2} \left[\left(f(z_i)^\top f(z_i') \right)^2 - 2 f(z_i)^\top f(z_i^+) \right].$$

It is easy to see that $\widehat{\mathcal{L}}_S(f)$ is an unbiased estimator of $\widehat{\mathcal{L}}_n(f)$:

Claim E.4. *For given $\widehat{\mathcal{X}}$, if we sample S as above, we have:*

$$\mathbb{E}_S \left[\widehat{\mathcal{L}}_S(f) \right] = \widehat{\mathcal{L}}_f(.)$$

Proof. This is obvious by the definition of $\widehat{\mathcal{L}}_S(f)$ and $\widehat{\mathcal{L}}_n(f)$. □

The following lemma reveals the relationship between the Rademacher complexity of feature extractors and the Rademacher complexity of the loss defined on tuples:

Lemma E.5. *Let $\mathcal{F}$ be a hypothesis class of feature extractors from $\mathcal{X}$ to $\mathbb{R}^k$. Assume $\|f(x)\|_\infty \le \kappa$ for all $x \in \mathcal{X}$. For $i \in [k]$, define $f_i : \mathcal{X} \to \mathbb{R}$ be the function such that $f_i(x)$ is the i-th dimension of $f(x)$. Let $\mathcal{F}_i$ be the hypothesis containing f_i for all $f \in \mathcal{F}$. For $m \in \mathcal{Z}^+$, let $\widehat{\mathcal{R}}_m(\mathcal{F}_i)$ be the maximal possible empirical Rademacher complexity of $\mathcal{F}_i$ over m data:*

$$\widehat{\mathcal{R}}_m(\mathcal{F}_i) := \max_{\{x_1, x_2, \cdots, x_m\}} \mathbb{E}_\sigma \left[\sup_{f_i \in \mathcal{F}_i} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f_i(x_j) \right) \right],$$

where $x_1, x_2, \cdots, x_m$ are in $\mathcal{X}$, and σ is a uniform random vector in $\{-1, 1\}^m$. Then, the empirical Rademacher complexity on any m tuples $\{(z_i, z_i^+, z_i')\}_{i=1}^m$ can be bounded by

$$\mathbb{E}_\sigma \left[\sup_{f \in \mathcal{F}} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j \left(\left(f(z_j)^\top f(z_j') \right)^2 - 2 f(z_j)^\top f(z_j^+) \right) \right) \right] \le (16k^2\kappa^2 + 16k\kappa) \cdot \max_{i \in [k]} \widehat{\mathcal{R}}_m(\mathcal{F}_i).$$

Proof.

$$\begin{aligned} & \mathbb{E}_\sigma \left[\sup_{f \in \mathcal{F}} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j \left(\left(f(z_j)^\top f(z_j') \right)^2 - 2 f(z_j)^\top f(z_j^+) \right) \right) \right] \\ \le & \mathbb{E}_\sigma \left[\sup_{f \in \mathcal{F}} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j \left(f(z_j)^\top f(z_j') \right)^2 \right) \right] + 2\mathbb{E}_\sigma \left[\sup_{f \in \mathcal{F}} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f(z_j)^\top f(z_j^+) \right) \right] \\ \le & 2k\kappa \mathbb{E}_\sigma \left[\sup_{f \in \mathcal{F}} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f(z_j)^\top f(z_j') \right) \right] + 2\mathbb{E}_\sigma \left[\sup_{f \in \mathcal{F}} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f(z_j)^\top f(z_j^+) \right) \right] \\ \le & (2k^2\kappa + 2k) \max_{\substack{z_1, z_2, \cdots, z_m \\ z_1', z_2', \cdots, z_m'}} \max_{i \in [k]} \mathbb{E}_\sigma \left[\sup_{f_i \in \mathcal{F}_i} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f_i(z_j) f_i(z_j') \right) \right], \end{aligned}$$

here the second inequality is by Talagrand's lemma. Notice that for any $z_1, z_2 \cdots z_m$ and $z_1', z_2', \cdots, z_m'$ in $\mathcal{X}$ and any $i \in [k]$ we have

$$\begin{aligned}
&\mathbb{E}_\sigma \left[\sup_{f_i \in \mathcal{F}_i} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f_i(z_j) f_i(z_j') \right) \right] \\
\leq &\frac{1}{2} \mathbb{E}_\sigma \left[\sup_{f_i \in \mathcal{F}_i} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j \left(f_i(z_j) + f_i(z_j') \right)^2 \right) \right] + \frac{1}{2} \mathbb{E}_\sigma \left[\sup_{f_i \in \mathcal{F}_i} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j \left(f_i(z_j) - f_i(z_j') \right)^2 \right) \right] \\
\leq &4\kappa \mathbb{E}_\sigma \left[\sup_{f_i \in \mathcal{F}_i} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f_i(z_j) \right) \right] + 4\kappa \mathbb{E}_\sigma \left[\sup_{f_i \in \mathcal{F}_i} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f_i(z_j') \right) \right],
\end{aligned}$$

where the first inequaltiy is by Talagrand's lemma. Combine these two equations and we get:

$$\begin{aligned}
&\mathbb{E}_\sigma \left[\sup_{f \in \mathcal{F}} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j \left(\left(f(z_j)^\top f(z_j') \right)^2 - 2 f(z_j)^\top f(z_j^+) \right) \right) \right] \\
\leq &(16k^2\kappa^2 + 16k\kappa) \max_{z_1, z_2, \cdots, z_m} \max_{i \in [k]} \mathbb{E}_\sigma \left[\sup_{f_i \in \mathcal{F}_i} \left(\frac{1}{m} \sum_{j=1}^m \sigma_j f_i(z_j) \right) \right].
\end{aligned}$$

□

Proof of Theorem 4.1. By Claim E.2 and Claim E.4, we know that $\mathbb{E}_S[\widehat{\mathcal{L}}_S(f)] = \mathcal{L}(f)$, where S is sampled by first sampling $\widehat{\mathcal{X}}$ then sample S according to Definition E.3. Notice that when $\widehat{\mathcal{X}}$ contains n i.i.d. samples natural data, the set of random tuples S contains n i.i.d tuples. Therefore, we can apply generalization bound with Rademacher complexity to get a uniform convergence bound. In particular, by Lemma E.5 and notice the fact that $\left(f(z_j)^\top f(z_j') \right)^2 - 2 f(z_j)^\top f(z_j^+)$ always take values in range $[-2k\kappa^2, 2k\kappa^2 + k^2\kappa^4]$, we apply standard generalization analysis based on Rademacher complexity and get: with probability at least $1 - \delta^2/4$ over the randomness of $\widehat{\mathcal{X}}$ and S, we have for any $f \in \mathcal{F}$,

$$\mathcal{L}(f) \leq \widehat{\mathcal{L}}_S(f) + (32k^2\kappa^2 + 32k\kappa) \max_{i \in [k]} \widehat{\mathcal{R}}_{n/2}(\mathcal{F}_i) + (4k\kappa^2 + k^2\kappa^4) \cdot \sqrt{\frac{4 \log 2/\delta}{n}}. \tag{19}$$

This means with probability at least $1 - \delta/2$ over random $\widehat{\mathcal{X}}$, we have: with probability at least $1 - \delta/2$ over random tuples S conditioned on $\widehat{\mathcal{X}}$, Equation (19) holds. Since both $\mathcal{L}(f)$ and $\widehat{\mathcal{L}}_n(f)$ take value in range $[-2k\kappa^2, 2k\kappa^2 + k^2\kappa^4]$, we have: with probability at least $1 - \delta/2$ over random $\widehat{\mathcal{X}}$, we have for any $f \in \mathcal{F}$,

$$\mathcal{L}(f) \leq \widehat{\mathcal{L}}_n(f) + (32k^2\kappa^2 + 32k\kappa) \cdot \max_{i \in [k]} \widehat{\mathcal{R}}_{n/2}(\mathcal{F}_i) + (4k\kappa^2 + k^2\kappa^4) \cdot \left(\sqrt{\frac{4 \log 2/\delta}{n}} + \frac{\delta}{2} \right).$$

Since negating the functions in a function class doesn't change its Rademacher complexity, we also have the other direction: with probability at least $1 - \delta/2$ over random $\widehat{\mathcal{X}}$, we have for any $f \in \mathcal{F}$,

$$\mathcal{L}(f) \geq \widehat{\mathcal{L}}_n(f) - (32k^2\kappa^2 + 32k\kappa) \cdot \max_{i \in [k]} \widehat{\mathcal{R}}_{n/2}(\mathcal{F}_i) + (4k\kappa^2 + k^2\kappa^4) \cdot \left(\sqrt{\frac{4 \log 2/\delta}{n}} + \frac{\delta}{2} \right).$$

Combine them together we get the excess risk bound: with probability at least $1 - \delta$, we have

$$\mathcal{L}(\hat{f}) \leq \mathcal{L}(f_{\mathcal{F}}^*) + (64k^2\kappa^2 + 64k\kappa) \cdot \max_{i \in [k]} \widehat{\mathcal{R}}_{n/2}(\mathcal{F}_i) + (8k\kappa^2 + 2k^2\kappa^4) \cdot \left(\sqrt{\frac{4 \log 2/\delta}{n}} + \frac{\delta}{2} \right),$$

where $\hat{f}$ is minimizer of $\widehat{\mathcal{L}}_n(f)$ in $\mathcal{F}$ and $f_{\mathcal{F}}^*$ is minimizer of $\mathcal{L}(f)$ in $\mathcal{F}$. Set $c_1 = 64k^2\kappa^2 + 64k\kappa$ and $c_2 = 16k\kappa^2 + 4k^2\kappa^4$ and notice that $\max_{i \in [k]} \widehat{\mathcal{R}}_{n/2}(\mathcal{F}_i) = \widehat{\mathcal{R}}_{n/2}(\mathcal{F})$ finishes the proof. □

E.2 Generalization bound for spectral contrastive learning with deep neural networks

In this section, we examplify Theorem 4.1 with the norm-contralled Rademacher complexity bound introduced in [20], which gives the following theorem.

Theorem E.6. *Assume $\mathcal{X}$ is a subset of Euclidean space $\mathbb{R}^d$ and $\|x\|_2 \le C_x$ for any $x \in \mathcal{X}$. Let $\mathcal{F}$ be a hypothesis class of norm-contralled l-layer deep neural networks defined as*

$$\{x \to P_\kappa(W_l\sigma(W_{l-1}\sigma(\cdots\sigma(W_1x)))) : \|W_i\|_F \le C_{w,i}\}$$

where $\sigma(\cdot)$ is element-wise ReLU activation, $P_\kappa(\cdot)$ is element-wise projection to interval $[-\kappa, \kappa]$ for some $\kappa > 0$, $C_{w,i}$ is the norm bound of the i-th layer, W_l has k rows and W_1 has d columns. Then, with probability at least $1-\delta$ over randomness of a dataset with size $2n$, we have

$$\mathcal{L}(\hat{f}) \le \mathcal{L}^*_{\mathcal{F}} + c_1 \cdot \frac{C_x C_w \sqrt{l}}{\sqrt{n}} + c_2 \cdot \left(\sqrt{\frac{\log 1/\delta}{n}} + \delta\right),$$

where $\hat{f}$ is the minimizer of $\widehat{\mathcal{L}}_{2n}(f)$ in $\mathcal{F}$, $\mathcal{L}^*_{\mathcal{F}}$ is the minimal $\mathcal{L}(f)$ achievable by any function $f \in \mathcal{F}$, $C_w := \prod_{i=1}^l C_{w,i}$, constants $c_1 \lesssim k^2\kappa^2 + k\kappa$ and $c_2 \lesssim k\kappa^2 + k^2\kappa^4$.

Proof of Theorem E.6. Consider the following hypothesis class of real-valued neural networks:

$$\mathcal{F}_{\text{real}} \triangleq \left\{x \to \widehat{W}_l\sigma(W_{l-1}\sigma(\cdots\sigma(W_1x))) : \|W_i\|_F \le C_{w,i}\right\}$$

where $\sigma(\cdot)$ is element-wise ReLU activation and $C_{w,i}$ is the norm bound of the i-th layer defined in the theorem, W_l has k rows and $\widehat{W}_1$ is a vector. By Theorem 1 of [20], we have

$$\widehat{\mathcal{R}}_n\left(\mathcal{F}_{\text{real}}\right) \le \frac{C_x(\sqrt{2\log(2)l}+1)C_w}{\sqrt{n}}.$$

Let the projection version of this hyposis class be:

$$\mathcal{F}_{\text{real+proj}} \triangleq \left\{x \to P_\kappa(\widehat{W}_l\sigma(W_{l-1}\sigma(\cdots\sigma(W_1x)))) : \|W_i\|_F \le C_{w,i}\right\},$$

where $P_\kappa(\cdot)$ projects a real number into interval $[-C_w, C_w]$. Notice that $P_\kappa(\cdot)$ is 1-Lipschitz, by Telegrand's lemma we have

$$\widehat{\mathcal{R}}_n\left(\mathcal{F}_{\text{real+proj}}\right) \le \frac{C_x(\sqrt{2\log(2)l}+1)C_w}{\sqrt{n}}.$$

For each $i \in [k]$, define function $f_i : \mathcal{X} \to \mathbb{R}$ such that $f_i(x)$ is the i-th dimension of $f(x)$, define $\mathcal{F}_i$ be the hypothesis class including all f_i for $f \in \mathcal{F}$. Then when $\mathcal{F}$ is the composition of deep neural networks and projection function as defined in the theorem, it is obvious to see that $\mathcal{F}_i = \mathcal{F}_{\text{real+proj}}$ for all $i \in [k]$. Therefore, by Theorem 4.1 we have

$$\mathcal{L}(\hat{f}) \le \mathcal{L}^*_{\mathcal{F}} + c_1 \cdot \frac{C_x(\sqrt{2\log(2)l}+1)C_w}{\sqrt{n}} + c_2 \cdot \left(\sqrt{\frac{\log 2/\delta}{n}} + \delta\right),$$

and absorbing the constants into c_1 finishes the proof. □

E.3 Proof of Theorem 4.2

In this section we give the proof of Theorem 4.2. We first introduce the following definitions of ϵ-optimal minimizers of matrix approximation loss and population spectral contrastive loss:

Definition E.7. *We say a function $\hat{f}_{\text{mf}}$ is ϵ-optimal minimizer of matrix approximation loss $\mathcal{L}_{\text{mf}}$ if*

$$\mathcal{L}_{\text{mf}}(\widehat{F}_{\text{mf}}) \le \min_F \mathcal{L}_{\text{mf}}(F) + \epsilon,$$

where $\widehat{F}_{\text{mf}}$ is $\hat{f}_{\text{mf}}$ written in the matrix form. We say a function $\hat{f}$ is ϵ-optimal minimizer of spectral contrastive loss $\mathcal{L}$ if

$$\mathcal{L}(\hat{f}) \le \min_f \mathcal{L}(f) + \epsilon.$$

We introduce the following generalized version of Theorem C.3, which captures the main effects of error in the representation.

Theorem E.8. *[Generalization of Theorem C.3] Assume the set of augmented data $\mathcal{X}$ is finite. Let λ_i be the i-th smallest eigenvalue of the normalize laplacian matrix. Let $\hat{f} \in \arg\min_{f:\mathcal{X}\to\mathbb{R}^k}$ be a ϵ-optimal minimizer of the spectral contrastive loss function $\mathcal{L}(f)$ with $k \in \mathcal{Z}^+$. Then, for any labeling function $\hat{y} : \mathcal{X} \to [r]$ there exists a linear probe $\widehat{B} \in \mathbb{R}^{r\times k}$ such that*

$$\Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},x\sim\mathcal{A}(\cdot|\bar{x})}\left(g_{\hat{f},\widehat{B}}(x) \neq y(\bar{x})\right) \leq \min_{1\leq k'\leq k}\left(\frac{2\phi^{\hat{y}}}{\lambda_{k'+1}} + \frac{4k'\epsilon}{(\lambda_{k+1}-\lambda_{k'})^2}\right) + \Delta(y,\hat{y}),$$

where $\phi^{\hat{y}}$ and $\Delta(y,\hat{y})$ are defined in Equations 8 and 9 respectively.

The proof of lemma E.8 is deferred to Section E.4.

Now we are ready to prove Theorem 4.2 using Theorem E.8.

Proof of Theorem 4.2. In Theorem E.8 we let $k' = \lfloor \frac{3}{4}k \rfloor$ on the RHS of the bound and get: for any $\hat{y} : \mathcal{X} \to [r]$ there exists $\widehat{B} \in \mathbb{R}^{r\times k}$ such that

$$\Pr_{x\sim\mathcal{P}_{\overline{\mathcal{X}}},\tilde{x}\sim\mathcal{A}(\cdot|x)}\left(g_{\hat{f},\widehat{B}}(\tilde{x}) \neq y(x)\right) \leq \frac{2\phi^{\hat{y}}}{\lambda_{\lfloor\frac{3}{4}k\rfloor+1}} + \frac{3k\epsilon}{(\lambda_{k+1}-\lambda_{\lfloor\frac{3}{4}k\rfloor})^2} + \Delta(y,\hat{y}).$$

Let $S_1, S_2, \cdots, S_r$ be the partition of $\mathcal{X}$ induced by the classifier g in Assumption 3.5. Define function $\hat{y} : \mathcal{X} \to [r]$ as follows: for an augmented data $x \in \mathcal{X}$, we use function $\hat{y}(x)$ to represent the index of set that x is in, i.e., $x \in S_{\hat{y}(x)}$. Then by Lemma C.5 we have $\phi^{\hat{y}} \leq 2\alpha$ and $\Delta(y,\hat{y}) \leq \alpha$. In Lemma C.4 let $(1+\zeta)t = \lfloor \frac{3}{4}k \rfloor + 1$ and $t = \lfloor \frac{k}{2} \rfloor$, then there is $\zeta \geq 0.5$, so we have: there exists a partition $S_1, \cdots, S_{\lfloor\frac{k}{2}\rfloor} \subset \mathcal{X}$ such that $\phi_G(S_i) \lesssim \sqrt{\lambda_{\lfloor\frac{3}{4}k\rfloor+1}\log(k)}$ for $\forall i \in [\lfloor\frac{k}{2}\rfloor]$. By Definition 3.4, we have $\rho_{\lfloor\frac{k}{2}\rfloor} \lesssim \sqrt{\lambda_{\lfloor\frac{3}{4}k\rfloor+1}\log(k)}$, which leads to $\frac{1}{\lambda_{\lfloor\frac{3}{4}k\rfloor+1}} \lesssim \frac{log(k)}{\rho^2_{\lfloor\frac{k}{2}\rfloor}}$. So we have

$$\begin{aligned}\Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},x\sim\mathcal{A}(\cdot|\bar{x})}\left(g_{\hat{f},\widehat{B}}(x) \neq y(\bar{x})\right) &\lesssim \frac{\alpha}{\rho^2_{\lfloor\frac{k}{2}\rfloor}}\cdot\log(k) + \frac{k\epsilon}{(\lambda_{k+1}-\lambda_{\lfloor\frac{3}{4}k\rfloor})^2} \\ &\lesssim \frac{\alpha}{\rho^2_{\lfloor\frac{k}{2}\rfloor}}\cdot\log(k) + \frac{k\epsilon}{(\lambda_{k}-\lambda_{\lfloor\frac{3}{4}k\rfloor})^2}.\end{aligned}$$

Notice that by the definition of ensembled linear probe predictor, $\bar{g}_{\hat{f},\widehat{B}}(\bar{x}) \neq y(\bar{x})$ happens only if more than half of the augmentations of $\bar{x}$ predicts differently from $y(\bar{x})$, so we have $\Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}}}\left(\bar{g}_{\hat{f},\widehat{B}} \neq y(\bar{x})\right) \leq 2\Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},x\sim\mathcal{A}(\cdot|\bar{x})}\left(g_{\hat{f},\widehat{B}}(x) \neq y(\bar{x})\right)$ which finishes the proof. □

E.4 Proof of Theorem E.8

In this section, we give the proof for Theorem E.8.

Lemma E.9 (Generalization of Lemma C.8)**.** *Let $\hat{F}_{\text{mf}}$ be the matrix form of a feature extractor $\hat{f}_{\text{mf}} : \mathcal{X} \to \mathbb{R}^k$. Then, $\hat{F}_{\text{mf}}$ is a ϵ-optimal minimizer of $\mathcal{L}_{\text{mf}}(F)$ if and only if*

$$\hat{f}(x) := \frac{1}{\sqrt{w_x}}\cdot\hat{f}_{\text{mf}}(x)$$

is a ϵ-optimal minimizer of spectral contrastive loss $\mathcal{L}(f)$.

Proof of Lemma E.9. The proof follows the proof of Lemma C.8. □

We will use the following important lemma about ϵ-optimal minimizer of $\mathcal{L}_{\text{mf}}$:

Lemma E.10. *Let λ_i be the i-th minimal eigenvalue of the normalized Laplacian matrix L with corrsponding unit-norm eigenvector v_i. Let $f : \mathcal{X} \to \mathbb{R}^k$ be ϵ-optimal minimizer of $\mathcal{L}_{\text{mf}}$ where $\epsilon < (1-\lambda_k)^2$. Let $F \in \mathbb{R}^{N\times k}$ be the matrix form of f, where $N = |\mathcal{X}|$. Let $\Pi_f^{\perp} v_i$ be the projection of v_i onto the subspace orthogonal to the column span of F. Then, for $i \le k$ we have*

$$\left\|\Pi_f^{\perp} v_i\right\|_2^2 \le \frac{\epsilon}{(\lambda_{k+1}-\lambda_i)^2}.$$

Proof. For function f with matrix form F, we overload notation $\mathcal{L}_{\text{mf}}(\cdot)$ and use $\mathcal{L}_{\text{mf}}(f)$ to represent $\mathcal{L}_{\text{mf}}(F)$.

We first prove that the column rank of f is k. If the column rank of f is less than k, then there must exists some function $f' : \mathcal{X} \to \mathbb{R}^{k-1}$ such that $\mathcal{L}_{\text{mf}}(f') = \mathcal{L}_{\text{mf}}(f)$. According to the Eckart–Young–Mirsky Theorem, we have $\min_{f:\mathcal{X}\to\mathbb{R}^k} \mathcal{L}_{\text{mf}}(f) = \sum_{j=k+1}^{N}(1-\lambda_j)^2$ and $\min_{f:\mathcal{X}\to\mathbb{R}^{k-1}} \mathcal{L}_{\text{mf}}(f) = \sum_{j=k}^{N}(1-\lambda_j)^2$. Therefore, $\mathcal{L}_{\text{mf}}(f) = \mathcal{L}_{\text{mf}}(f') \ge (1-\lambda_k)^2 + \min_{f:\mathcal{X}\to\mathbb{R}^k} \mathcal{L}_{\text{mf}}(f)$, contradicting with $\epsilon < (1-\lambda_k)^2$. As a result, the column rank of f has to be k.

Recall normalized adjacency matrix $\overline{A} = I - L$. We use $\overline{A}_i$ to denote the i-th column of $\overline{A}$. We use $\widehat{A}$ to denote matrix $FF^\top$ and $\widehat{A}_i$ to denote the i-th column of $\widehat{A}$. Let $z_1, \cdots, z_k$ be unit-norm orthogonal vectors in the column span of F. Since the column span of $\widehat{A}$ is the same as the column span of F, we know columns of $\widehat{A}$ are in $span\{z_1, \cdots, z_k\}$. Let $z_{k+1}, \cdots, z_N$ be unit-norm orthogonal vectors such that together with $z_1, \cdots, z_k$ they form an orthonormal basis of $\mathbb{R}^N$. We use Π_f and $\Pi_f^{\perp}$ to denote matrices $\sum_{j=1}^{k} z_j z_j^\top$ and $\sum_{j=k+1}^{N} z_j z_j^\top$ respectively, then for any vector $v \in \mathbb{R}^N$, vectors $\Pi_f v$ and $\Pi_f^{\perp} v$ are the projections of v onto the column span of f and its orthogonal space respectively.

We first give a lower bound of $\mathcal{L}_{\text{mf}}(f)$ as follows:

$$\begin{aligned}
\mathcal{L}_{\text{mf}}(f) &= \left\|\overline{A} - \widehat{A}\right\|_F^2 = \sum_{j=1}^{N}\left\|\overline{A}_j - \widehat{A}_j\right\|_2^2 \ge \sum_{j=1}^{N}\left\|\overline{A}_j - \Pi_f \overline{A}_j\right\|_2^2 \\
&= \sum_{j=1}^{N}\left\|\overline{A}_j - \left(\sum_{t=1}^{k} z_t z_t^\top\right)\overline{A}_j\right\|_2^2 = \sum_{j=1}^{N}\left\|\left(\sum_{t=k+1}^{N} z_t z_t^\top\right)\overline{A}_j\right\|_2^2 \\
&= \left\|\left(\sum_{t=k+1}^{N} z_t z_t^\top\right)\overline{A}\right\|_F^2 = \left\|\Pi_f^{\perp}\overline{A}\right\|_F^2.
\end{aligned}$$

where the first equality is by definition of $\mathcal{L}_{\text{mf}}(f)$, the second equality is by writing the Frobenius norm square as the sum of column norm square, the inequality is because $\widehat{A}_j$ must be in the span of $z_1, \cdots, z_k$ while $\Pi_f \overline{A}_j$ is the vector in this span that is closest to $\overline{A}_j$, the third equality is writing the projection function in the matrix form, the fourth equality is because $z_1, \cdots z_d$ are an orthonormal basis, the fifth equality is rewriting to Frobenius norm, and the last equality is by definition of $\Pi_f^{\perp}$.

Notice that

$$\left\|\Pi_f^{\perp}\overline{A}\right\|_F^2 = Tr\left(\overline{A}^\top {\Pi_f^{\perp}}^\top \Pi_f^{\perp}\overline{A}\right) = Tr\left(\overline{A}^\top \Pi_f^{\perp}\overline{A}\right) = Tr\left(\overline{A}\overline{A}^\top \Pi_f^{\perp}\right).$$

We can rewrite the above lower bound as

$$\mathcal{L}_{\text{mf}}(f) \ge Tr\left(\overline{A}\overline{A}^\top \Pi_f^{\perp}\right) = Tr\left(\sum_{j=1}^{N}(1-\lambda_j)^2 v_j v_j^\top \sum_{t=k+1}^{N} z_t z_t^\top\right) = \sum_{j=1}^{N}\sum_{t=k+1}^{N}(1-\lambda_j)^2\langle v_j, z_t\rangle^2.$$

We define variable $S_j \triangleq \sum_{t=1}^{j}\sum_{l=k+1}^{d}\langle v_t, z_l\rangle^2$ for any $j \in [N]$. Also denote $\lambda_{d+1} = 1$. We have the following equality:

$$\sum_{j=1}^{N}\sum_{t=k+1}^{N}(1-\lambda_j)^2\langle v_j, z_t\rangle^2 = \sum_{j=1}^{N}\left((1-\lambda_j)^2 - (1-\lambda_{j+1})^2\right)S_j.$$

Notice that $S_j \geq 0$ and also when $i \leq j \leq k$, we have $S_j \geq \left\|\Pi_f^\perp v_i\right\|_2^2$, we have

$$\sum_{j=1}^{N}\sum_{t=k+1}^{N}(1-\lambda_j)^2\langle v_j, z_t\rangle^2 \geq \left((1-\lambda_i)^2-(1-\lambda_{k+1})^2\right)\left\|\Pi_f^\perp v_i\right\|_2^2 + \sum_{j=k+1}^{N}\left((1-\lambda_j)^2-(1-\lambda_{j+1})^2\right)S_j,$$

where we replace every S_j with 0 when $j < k$, replace S_j with $\left\|\Pi_f^\perp v_i\right\|_2^2$ when $i \leq j \leq k$, and keep S_j when $j \geq k+1$. Now notice that

$$S_N = \sum_{t=1}^{N}\sum_{l=k+1}^{N}\langle v_t, z_l\rangle^2 = \sum_{l=k+1}^{N}\sum_{t=1}^{N}\langle v_t, z_l\rangle^2 = \sum_{l=k+1}^{N}\|z_l\|_2^2 = N-k,$$

and also

$$S_{j+1} - S_j = \sum_{l=k+1}^{N}\langle v_{j+1}, z_l\rangle^2 \leq \sum_{l=1}^{N}\langle v_{j+1}, z_l\rangle^2 = 1,$$

there must be $S_j \geq j-k$ when $j \geq k+1$. So we have

$$\begin{aligned}
&\sum_{j=1}^{N}\sum_{t=k+1}^{N}(1-\lambda_j)^2\langle v_j, z_t\rangle^2 \\
&\geq \left((1-\lambda_i)^2-(1-\lambda_{k+1})^2\right)\left\|\Pi_f^\perp v_i\right\|_2^2 + \sum_{j=k+1}^{N}\left((1-\lambda_j)^2-(1-\lambda_{j+1})^2\right)(j-k) \\
&= \left((1-\lambda_i)^2-(1-\lambda_{k+1})^2\right)\left\|\Pi_f^\perp v_i\right\|_2^2 + \sum_{j=k+1}^{N}(1-\lambda_j)^2 \\
&= \left((1-\lambda_i)^2-(1-\lambda_{k+1})^2\right)\left\|\Pi_f^\perp v_i\right\|_2^2 + \mathcal{L}_{\text{mf}}(f^*_{\text{pop}}),
\end{aligned}$$

where f^*_{pop} is the minimizer of $\mathcal{L}_{\text{mf}}$, and the last equality is by Eckart–Young–Mirsky Theorem. So we know $\mathcal{L}_{\text{mf}}(f) \geq \left((1-\lambda_i)^2-(1-\lambda_{k+1})^2\right)\left\|\Pi_f^\perp v_i\right\|_2^2 + \mathcal{L}_{\text{mf}}(f^*_{\text{pop}})$, which implies that $\left\|\Pi_f^\perp v_i\right\|_2^2 \leq \frac{\epsilon}{(1-\lambda_i)^2-(1-\lambda_{k+1})^2} \leq \frac{\epsilon}{(\lambda_{k+1}-\lambda_i)^2}$. □

The following lemma generalizes Lemma C.6.

Lemma E.11 (Generalization of Lemma C.6)**.** *Let L be the normalized Laplacian matrix of graph $G = (\mathcal{X}, w)$, where $|\mathcal{X}| = N$. Let $f : \mathcal{X} \to \mathbb{R}^k$ be an ϵ-optimal minimizer of $\mathcal{L}_{\text{mf}}(f)$ where $\epsilon < (1-\lambda_k)^2$. Let F be the matrix form of f and F_i is the i-th column of F. Let $R(u) := \frac{u^\top L u}{u^\top u}$ be the Rayleigh quotient of a vector $u \in \mathbb{R}^N$. Then, for any $k \in \mathcal{Z}^+$ such that $k < N$, there exists a vector $b \in \mathbb{R}^k$ such that*

$$\left\|u - \sum_{i=1}^{k} b_i F_i\right\|_2^2 \leq \min_{1\leq k'\leq k}\left(\frac{2R(u)}{\lambda_{k'+1}} + \frac{2k'\epsilon}{(\lambda_{k+1}-\lambda_{k'})^2}\right)\|u\|_2^2 .$$

Proof. Let k' be the choice that minimizes the right hand side. We use $p_v(u)$ to denote the projection of u onto the span of $v_1, \cdots, v_{k'}$. We use $p_{v,f}(u)$ to denote the projection of $p_v(u)$ onto the span of $f_1, \cdots, f_k$. Then we know that

$$\|u - p_{v,f}(u)\|_2^2 \leq 2\|u - p_v(u)\|_2^2 + 2\|p_v(u) - p_{v,f}(u)\|_2^2 . \tag{20}$$

By the proof of Lemma C.6, we know that

$$\|u - p_v(u)\|_2^2 \leq \frac{R(u)}{\lambda_{k'+1}}\|u\|_2^2 . \tag{21}$$

On the other hand, we have

$$\begin{aligned}\|p_v(u) - p_{vf}(u)\|_2^2 &= \left\|\Pi_f^{\perp} p_v(u)\right\|_2^2 \\ &= \left\|\sum_{i=1}^{k'} \Pi_f^{\perp} v_i v_i^{\top} u\right\|_2^2 \\ &\le \left(\sum_{i=1}^{k'} \left\|\Pi_f^{\perp} v_i\right\|_2^2\right) \cdot \left(\sum_{i=1}^{k'} (v_i^{\top} u)^2\right) \\ &\le \frac{k'\epsilon}{(\lambda_{k+1} - \lambda_{k'})^2} \|u\|_2^2, \end{aligned} \tag{22}$$

where the first inequality if by Cauchy–Schwarz inequality and the second inequality if by Lemma E.10. Plugging Equation (21) and Equation (22) into Equation (20) finishes the proof. □

Now we prove Theorem E.8 using the above lemmas.

Proof of Theorem E.8. By Lemma E.9, the ϵ-optimal minimizer of $\mathcal{L}(f)$ is only different from ϵ-optimal minimizer of $\mathcal{L}_{\text{mf}}(f)$ by a positive constant for each x. Since this difference won't influence the perdiction accuracy, we only need to prove this theorem assuming $\hat{f}$ is ϵ-optimal minimizer of $\mathcal{L}_{\text{mf}}(f)$.

For each $i \in [r]$, we define the function $u_i(x) = \mathbb{1}[\hat{y}(x) = i] \cdot \sqrt{w_x}$. Let $u : \mathcal{X} \to \mathbb{R}^k$ be the function such that $u(x)$ has u_i at the i-th dimension. By Lemma E.11, there exists a vector $b_i \in \mathbb{R}^k$ such that

$$\left\|u_i - \widehat{F} b_i\right\|_2^2 \le \min_{1 \le k' \le k} \left(\frac{2R(u_i)}{\lambda_{k'+1}} + \frac{2k'\epsilon}{(\lambda_{k+1} - \lambda_{k'})^2}\right) \|u_i\|_2^2$$

Let matrices $U = [u_1, \cdots, u_r]$ and $\widehat{B}^{\top} = [b_1, \cdots, b_r]$. We sum the above equation over all $i \in [r]$ and get

$$\begin{aligned}\left\|U - \widehat{F}\widehat{B}^{\top}\right\|_F^2 &\le \sum_{i=1}^{r} \min_{1 \le k' \le k} \left(\frac{2R(u_i)}{\lambda_{k'+1}} + \frac{2k'\epsilon}{(\lambda_{k+1} - \lambda_{k'})^2}\right) \|u_i\|_2^2 \\ &\le \min_{1 \le k' \le k} \sum_{i=1}^{r} \left(\frac{2R(u_i)}{\lambda_{k'+1}} \|u_i\|_2^2 + \frac{2k'\epsilon}{(\lambda_{k+1} - \lambda_{k'})^2} \|u_i\|_2^2\right). \end{aligned} \tag{23}$$

Notice that

$$\begin{aligned}\sum_{i=1}^{r} R(u_i) \|u_i\|_2^2 &= \sum_{i=1}^{r} \frac{1}{2} \phi_i^{\hat{y}} \sum_{x \in \mathcal{X}} w_x \cdot \mathbb{1}[\hat{y}(x) = i] \\ &= \frac{1}{2} \sum_{i=1}^{r} \sum_{x, x' \in \mathcal{X}} w_{xx'} \cdot \mathbb{1}[(\hat{y}(x) = i \wedge \hat{y}(x') \ne i) \text{ or } (\hat{y}(x) \ne i \wedge \hat{y}(x') = i)] \\ &= \frac{1}{2} \sum_{x, x' \in \mathcal{X}} w_{xx'} \cdot \mathbb{1}[\hat{y}(x) \ne \hat{y}(x')] = \frac{1}{2} \phi^{\hat{y}}, \end{aligned} \tag{24}$$

where the first equality is by Claim C.7. On the other hand, we have

$$\sum_{i=1}^{r} \|u_i\|_2^2 = \sum_{i=1}^{r} \sum_{x \in \mathcal{X}} w_x \cdot \mathbb{1}[\hat{y}(x) = i] = \sum_{x \in \mathcal{X}} w_x = 1. \tag{25}$$

Plugging Equation (24) and Equation (25) into Equation (23) gives us

$$\left\|U - \widehat{F}\widehat{B}^{\top}\right\|_F^2 \le \min_{1 \le k' \le k} \left(\frac{\phi^{\hat{y}}}{\lambda_{k'+1}} + \frac{2k'\epsilon}{(\lambda_{k+1} - \lambda_{k'})^2}\right).$$

Notice that by definition of $u(x)$, we know that prediction $g_{\hat{f},\widehat{B}}(x) \neq \hat{y}(x)$ only happens if $\left\|u(x) - \widehat{B}\hat{f}(x)\right\|_2^2 \geq \frac{w_x}{2}$. Hence we have

$$\sum_{x\in\mathcal{X}} \frac{1}{2} w_x \cdot \mathbb{1}\left[g_{\hat{f},\widehat{B}}(x) \neq \hat{y}(x)\right] \leq \sum_{x\in\mathcal{X}} \left\|u(x) - \widehat{B}\hat{f}(x)\right\|_2^2 = \left\|U - \widehat{F}\widehat{B}^\top\right\|_F^2.$$

Now we are ready to bound the error rate on $\mathcal{X}$:

$$\begin{aligned}
&\Pr_{x\sim\mathcal{X}}(g_{\hat{f},\widehat{B}}(x) \neq \hat{y}(x)) = \sum_{x\in\mathcal{X}} w_x \cdot \mathbb{1}\left[g_{\hat{f},\widehat{B}}(x) \neq \hat{y}(x)\right] \\
&\leq 2 \cdot \left\|U - \hat{f}\widehat{B}^\top\right\|_F^2 \leq \min_{1\leq k'\leq k}\left(\frac{2\phi^{\hat{y}}}{\lambda_{k'+1}} + \frac{4k'\epsilon}{(\lambda_{k+1}-\lambda_{k'})^2}\right).
\end{aligned}$$

Here for the equality we are using the fact that $\Pr(x) = w_x$. We finish the proof by noticing that by the definition of $\Delta(y,\hat{y})$:

$$\begin{aligned}
\Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},x\sim\mathcal{A}(\cdot|\bar{x})}\left(g_{\hat{f},\widehat{B}}(x) \neq y(\bar{x})\right) &\leq \Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},x\sim\mathcal{A}(\cdot|\bar{x})}\left(g_{\hat{f},\widehat{B}}(x) \neq \hat{y}(x)\right) + \Pr_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},x\sim\mathcal{A}(\cdot|\bar{x})}(y(\bar{x}) \neq \hat{y}(x)) \\
&\leq \min_{1\leq k'\leq k}\left(\frac{2\phi^{\hat{y}}}{\lambda_{k'+1}} + \frac{4k'\epsilon}{(\lambda_{k+1}-\lambda_{k'})^2}\right) + \Delta(y,\hat{y}).
\end{aligned}$$

□

F Proofs for Section 5

In this section we give the proof of Theorem 5.1. We first introduce the following lemma, which states the expected norm of representations:

Lemma F.1. *Let $f^*_{\text{pop}} : \mathcal{X} \to \mathbb{R}^k$ be a minimizer of population spectral contrastive loss $\mathcal{L}(f)$. Then, we have*

$$\mathbb{E}_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},x\sim\mathcal{A}(\cdot|\bar{x})}\left[\left\|f^*_{\text{pop}}(x)\right\|_2^2\right] \leq k. \tag{26}$$

Proof of Lemma F.1. By Lemma C.8 and the definition of w, we have

$$\mathbb{E}_{\bar{x}\sim\mathcal{P}_{\overline{\mathcal{X}}},x\sim\mathcal{A}(\cdot|\bar{x})}\left[\left\|f^*_{\text{pop}}(x)\right\|_2^2\right] = \sum_{x\in\mathcal{X}} w_x \left\|f^*_{\text{pop}}(x)\right\|_2^2 = \sum_{x\in\mathcal{X}} \left\|\hat{f}_{\text{ma}}(x)\right\|_2^2 = \left\|\widehat{F}_{\text{mf}}\right\|_F^2, \tag{27}$$

where $\widehat{F}_{\text{mf}}$ is a minimizer of the matrix approximation loss defined in Section 3.2. By Eckard-Young-Mirsky theorem, $\widehat{F}_{\text{mf}}$ looks like

$$\widehat{F}_{\text{mf}} = F_{\text{sc}} D_\lambda Q,$$

where $F_{\text{sc}} = [v_1, v_2, \cdots, v_k]$ contains the k smallest eigenvectors of the laplacian matrix L as columns, Q is an orthonomal matrix and

$$D_\lambda = \begin{bmatrix} \sqrt{1-\lambda_1} & & & \\ & \sqrt{1-\lambda_2} & & \\ & & \cdots & \\ & & & \sqrt{1-\lambda_k} \end{bmatrix}.$$

So we have

$$\left\|\widehat{F}_{\text{mf}}\right\|_F^2 = Tr\left(F_{\text{sc}} D_\lambda^2 F_{\text{sc}}^\top\right) \leq Tr\left(F_{\text{sc}} F_{\text{sc}}^\top\right) = k,$$

where we use the fact that D_λ has diagonal values less than 1 and v_i is unit-norm. Pluggin this into Equation (27) finishes the proof. □

Now we give the proof of Theorem 5.1:

Proof of Theorem 5.1. Let f^*_{pop} be the minimizer of population spectral contrastive loss $\mathcal{L}(f)$. We abuse notation and use y_i to denote $y(\bar{x}_i)$, and let $z_i = f^*_{\text{pop}}(x_i)$. We first study the average empirical Rademacher complexity of the capped quadratic loss on a dataset $\{(z_i, y_i)\}_{i=1}^n$, where (z_i, y_i) is sampled as in Section 5:

$$\begin{aligned}
\widehat{\mathcal{R}}_n(\ell) :=& \mathbb{E}_{\{(z_i,y_i)\}_{i=1}^n} \mathbb{E}_\sigma \left[\sup_{\|B\|_F \le 1/C_\lambda} \frac{1}{n} \left[\sum_{i=1}^n \sigma_i \ell((z_i, y_i), B) \right] \right] \\
\le& 2r \mathbb{E}_{\{(z_i,y_i)\}_{i=1}^n} \mathbb{E}_\sigma \left[\sup_{\|b\|_2 \le 1/C_\lambda} \frac{1}{n} \left[\sum_{i=1}^n \sigma_i w^\top z_i \right] \right] \\
\le& \frac{2r}{C_\lambda} \sqrt{\frac{\mathbb{E}[\|z_i\|^2]}{n}} \le \frac{2r\sqrt{k}}{C_\lambda \sqrt{n}},
\end{aligned}$$

where the first inequality uses Talagrand's lemma and the fact that ℓ_σ is 2-Lipschitz, the second inequality is by standard Rademacher complexity of linear models, the third inequality is by Lemma F.1.

By Theorem C.2, there exists a linear probe B^* with norm bound $\|B^*\|_F \le 1/(1-\lambda_k) \le 1/C_\lambda$ such that

$$\mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left[\ell \left((f^*_{\text{pop}}(x), y(\bar{x})), B^* \right) \right] \lesssim \text{poly}(1/\zeta) \log(k+1) \cdot \frac{\phi^{\hat{y}}}{\rho^2_{k'}} + \Delta(y, \hat{y}),$$

where $(1+\zeta)k' = k+1$. Let $\widehat{B}$ be the minimizer of $\sum_{i=1}^n \ell\left((z_i, y_i), B\right)$ subject to $\|B\|_F \le 1/C_\lambda$, then by standard generalization bound, we have: with probability at least $1-\delta$, we have

$$\mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left[\ell \left((f^*_{\text{pop}}(x), y(\bar{x})), \widehat{B} \right) \right] \lesssim \text{poly}(1/\zeta) \log(k+1) \cdot \frac{\phi^{\hat{y}}}{\rho^2_{k'}} + \Delta(y, \hat{y}) + \frac{r\sqrt{k}}{C_\lambda \sqrt{n}} + \sqrt{\frac{\log 1/\delta}{n}}.$$

Follow the same steps as in the proof of Theorem 3.7, we can get a genalization bound of

$$\mathbb{E}_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left[\ell \left((f^*_{\text{pop}}(x), y(\bar{x})), \widehat{B} \right) \right] \lesssim \frac{\alpha}{\rho^2_{\lfloor k/2 \rfloor}} \cdot \log k + \frac{r}{C_\lambda} \cdot \sqrt{\frac{k}{n}} + \sqrt{\frac{\log 1/\delta}{n}}.$$

Notice that $y(\bar{x}) \neq g_{f^*_{\text{pop}}, \widehat{B}}(x)$ only if $\ell \left((f^*_{\text{pop}}(x), y(\bar{x})), \widehat{B} \right) \ge \frac{1}{2}$, we have the error bound

$$\Pr_{\bar{x} \sim \mathcal{P}_{\overline{\mathcal{X}}}, x \sim \mathcal{A}(\cdot|\bar{x})} \left(g_{f^*_{\text{pop}}, \widehat{B}}(x) \neq y(\bar{x}) \right) \lesssim \frac{\alpha}{\rho^2_{\lfloor k/2 \rfloor}} \cdot \log k + \frac{r}{C_\lambda} \cdot \sqrt{\frac{k}{n}} + \sqrt{\frac{\log 1/\delta}{n}}.$$

The result on $\bar{g}_{f^*_{\text{pop}}, \widehat{B}}$ naturally follows by the definition of $\bar{g}$. □